Question 1

Show that in general, $MSE(\hat{\theta}) = bias(\hat{\theta})^2 + var(\hat{\theta})$, where var is the variance, and bias is given by

$$bias(\hat{\theta}) = \theta - \mathbb{E}\left[\hat{\theta}\right]$$

Answer 1

$$var(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$
 (known) or,
$$var(\hat{\theta}) = \mathbb{E}[\hat{\theta}^2] - \mathbb{E}[\hat{\theta}]^2$$
 (1)

$$bias(\hat{\theta}) = \theta - \mathbb{E}[\hat{\theta}]$$
 (given)
or,
$$bias(\hat{\theta})^2 = (\theta - \mathbb{E}[\hat{\theta}])^2$$
 (squaring both sides)
or,
$$bias(\hat{\theta})^2 = \theta^2 - 2\theta \mathbb{E}[\hat{\theta}] + \mathbb{E}[\hat{\theta}]^2$$
 (2)

$$MSE(\hat{L}) = \mathbb{E}\left[(\hat{L} - L)^2\right]$$
 (given)
or,
$$MSE(\hat{\theta}) = \mathbb{E}\left[(\hat{\theta} - \theta)^2\right]$$
or,
$$MSE(\hat{\theta}) = \mathbb{E}\left[\hat{\theta}^2 - 2\hat{\theta}\theta + \theta^2\right]$$
or,
$$MSE(\hat{\theta}) = \mathbb{E}\left[\hat{\theta}^2\right] - 2\theta\mathbb{E}\left[\hat{\theta}\right] + \theta^2$$
or,
$$MSE(\hat{\theta}) = \theta^2 - 2\theta\mathbb{E}\left[\hat{\theta}\right] + \mathbb{E}\left[\hat{\theta}^2\right]$$
 (rearranging)
or,
$$MSE(\hat{\theta}) = (\theta^2 - 2\theta\mathbb{E}\left[\hat{\theta}\right] + \mathbb{E}\left[\hat{\theta}^2\right] + (\mathbb{E}\left[\hat{\theta}^2\right] - \mathbb{E}\left[\hat{\theta}\right]^2)$$
 (adding and subtracting $\mathbb{E}\left[\hat{\theta}\right]^2$)
or,
$$MSE(\hat{\theta}) = bias(\hat{\theta})^2 + var(\hat{\theta})$$
 (from 2 and 1)

Question 2

or,

Compute the bias of \hat{L}_{MOM} and \hat{L}_{MLE} . In general, \hat{L}_{MLE} consistently underestimates L - why? Hint: What is the pdf for \hat{L}_{MLE} ?

Answer 2

$$\hat{L}_{MOM} = 2\bar{X}_n \qquad \qquad \text{(given)}$$
 and,
$$bias(\hat{\theta}) = \theta - \mathbb{E}[\hat{\theta}] \qquad \qquad \text{(given)}$$
 Therefore,
$$bias(\hat{L}_{MOM}) = L - \mathbb{E}[2\hat{L}]$$

$$bias(\hat{L}_{MOM}) = L - 2\mathbb{E}[\hat{L}]$$

$$bias(\hat{L}_{MOM}) = L - 2\frac{L}{2} \qquad \qquad \text{(from Estimation notes 13)}$$

$$bias(\hat{L}_{MOM}) = 0$$

Therefore, \hat{L}_{MOM} is unbiased.

$$\hat{L}_{MLE} = \max_{i=1,\dots,n} X_i \tag{given}$$

The c.d.f. is:

The c.d.f. is:
$$P(\hat{L}_{MLE} \le x) = P\left(\max_{i=1,...,n} X_i\right)$$
or,
$$F(x) = P(X_1 \le x, X_2 \le x, ..., X_n \le x)$$
or,
$$F(x) = P(X_1 \le x, X_2 \le x, ..., X_n \le x)$$
or,
$$F(x) = P(X_1 \le x, X_2 \le x, ..., X_n \le x)$$
or,
$$F(x) = P(X_1 \le x) P(X_2 \le x) ... P(X_n \le x)$$
or,
$$F(x) = P(X \le x)^n$$
or,
$$F(x) = \left(\frac{x}{L}\right)^n$$
(1)

Calculating the p.d.f.:

$$f(x) = \frac{d}{dx}F(x)$$
or,
$$f(x) = \frac{d}{dx}\left(\frac{x}{L}\right)^{n}$$
or,
$$f(x) = \frac{n}{L}\left(\frac{x}{L}\right)^{n-1}$$
or,
$$f(x) = \frac{n}{L^{n}}x^{n-1}$$
(2)

Calculating the expectation:

or,
$$\mathbb{E}[\hat{L}_{MLE}] = \int_{0}^{L} x f(x) dx$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \int_{0}^{L} x \frac{n}{L^{n}} x^{n-1} dx$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \int_{0}^{L} \frac{n}{L^{n}} x^{n} dx$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \frac{n}{L^{n}} \int_{0}^{L} x^{n} dx$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \frac{n}{L^{n}} |\frac{x^{n+1}}{n+1}|_{0}^{L}$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \frac{n}{L^{n}} \frac{L^{n+1}}{n+1}$$
or,
$$\mathbb{E}[\hat{L}_{MLE}] = \frac{nL}{n+1}$$
(3)

$$bias(\hat{\theta}) = \theta - \mathbb{E}[\hat{\theta}]$$
 (given)
or,
$$bias(\hat{L}_{MLE}) = L - \mathbb{E}[\hat{L}_{MLE}]$$

or,
$$bias(\hat{L}_{MLE}) = L - \frac{nL}{n+1}$$

or,
$$bias(\hat{L}_{MLE}) = \frac{nL + L - nL}{n+1}$$

or,
$$bias(\hat{L}_{MLE}) = \frac{L}{n+1}$$

Therefore, \hat{L}_{MLE} is biased.

The denominator is always more than 1 (actually more than 2) because $n \ge 1$, \hat{L}_{MLE} will always be less than L. Therefore \hat{L}_{MLE} consistently underestimates L.

Question 3

Compute the variance of \hat{L}_{MOM} and \hat{L}_{MLE} .

Answer 3

or,

or,
$$var(L_{MOM}) = var(2X_n)$$
or, $var(L_{MOM}) = var(2X_n)$
or, $var(L_{MOM}) = var(X_n)$
or, $var(L_{MOM}) = 4var(X_n)$
or, $var(L_{MOM}) = 4var(X_n)$
or, $var(L_{MOM}) = 4var(X_{n-1})$
or, $var(L_{MOM}) = \frac{1}{n^2}var(\sum_{i=1}^n X_i)$
or, $var(L_{MOM}) = \frac{1}{n^2}var(\sum_{i=1}^n X_i)$
or, $var(L_{MOM}) = \left(\frac{1}{n^2}\right)\left(\frac{n^{(1-o)^2}}{12}\right)$ (uniform distribution on the interval $[0, L]$)
or, $var(L_{MOM}) = \frac{L}{n}(L_3)$
or, $var(L_{MOM}) = \frac{L}{n}$
or, $var(L_{MOM}) = \frac{L}{n}$
or, $var(L_{MOM}) = \frac{L}{n^2}$
or, $var(L_{MOM}) = \frac{L}{n}$

$$\int_0^L x^2 f(x) dx$$
or, $\mathbb{E}\left[L_{MLE}^2\right] = \int_0^L x^2 f(x) dx$
or, $\mathbb{E}\left[L_{MLE}^2\right] = \frac{1}{n^2} \int_0^L x^{n+1} dx$
or, $\mathbb{E}\left[L_{MLE}\right] = \frac{1}{n^2} \int_0^L x^{n+2} dx$
or, $\mathbb{E}\left[L_{MLE}\right] = \frac{1}{n^2} \int_0^L x^{n+1} dx$
or, $\mathbb{E}\left[L_{MLE}\right] = \frac{1}{n^2} \int_$

(3)

Question 4

Which one is the better estimator, i.e., which one has the smaller mean squared error?

Answer 4

$$MSE(\hat{\theta}) = bias(\hat{\theta})^2 + var(\hat{\theta})$$
 (proven in Answer 1) or,
$$MSE(\hat{L}_{MOM}) = bias(\hat{L}_{MOM})^2 + var(\hat{L}_{MOM})$$
 (proven in Answer 2 and 3) or,
$$MSE(\hat{L}_{MOM}) = \frac{l^2}{3n}$$
 (proven in Answer 2 and 3) or,
$$MSE(\hat{L}_{MOM}) = \frac{l^2}{3n}$$
 (proven in Answer 2 and 3) or,
$$MSE(\hat{L}_{MOM}) = \frac{l^2}{3n}$$
 (proven in Answer 1) or,
$$MSE(\hat{L}_{MLE}) = bias(\hat{L}_{MLE})^2 + var(\hat{L}_{MLE})$$
 (proven in Answer 1) or,
$$MSE(\hat{L}_{MLE}) = \frac{l^2}{(n+2)} + \frac{nl^2}{(n+2)(n+1)^2}$$
 (proven in Answer 2 and 3) or,
$$MSE(\hat{L}_{MLE}) = \frac{l^2(n+2)+nl^2}{(n+2)(n+1)^2}$$
 or,
$$MSE(\hat{L}_{MLE}) = \frac{nl^2+2l^2+nl^2}{(n+2)(n+1)^2}$$
 or,
$$MSE(\hat{L}_{MLE}) = \frac{nl^2+2l^2+nl^2}{(n+2)(n+1)^2}$$
 or,
$$MSE(\hat{L}_{MLE}) = \frac{2l^2(n+1)}{(n+2)(n+1)^2}$$
 (2)
$$MSE(\hat{L}_{MLE}) = \frac{l^2}{(n+2)(n+1)^2}$$
 or,
$$MSE(\hat{L}_{MLE}) = \frac{l^2}{(n+2)(n+1)^2}$$
 (2)
$$MSE_{diff} = MSE(\hat{L}_{MOM}) - MSE(\hat{L}_{MLE})$$
 or,
$$MSE_{diff} = \frac{l^2}{3n} - \frac{2l^2}{(n+2)(n+1)}$$
 (2)
$$MSE_{diff} = \frac{l^2(n+2)(n+1)}{(3n)(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{l^2(n+2)(n+1)}{(3n)(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{l^2(n+2)(n+1)}{(3n)(n+2)(n+1)}$$
 or,
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 or,
$$MSE_{diff} = \frac{l^2(n+2)(n+2)(n+2)(n+2)}{(3n)(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{l^2(n+2)(n+2)(n+2)(n+2)}{(3n)(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{l^2(n+2)(n+2)(n+2)(n+2)(n$$

 MSE_{diff} is positive for n < 1 and n > 2 MSE_{diff} is negative for n between 1 and 2 (exclusive). MSE_{diff} is 0 for n=1 and n=2

But $n \ge 1$, so MSE_{diff} is negative for n < 2 and positive for n > 2. In general n >>> 1, so MSE_{diff} is positive.

Considering MSE_{diff} as positive, we can say $MSE(\hat{L}_{MLE})$ is less than $MSE(\hat{L}_{MOM})$. Therefore, \hat{L}_{MLE} is a better estimator.

Ouestion 5

```
Experimentally verify your computations in the following way: Taking n = 100 and L = 10, - For j = 1, \ldots, 1000:

*Simulate X_1^j, \ldots, X_n^j and compute values for \hat{L}_{MOM}^j and \hat{L}_{MLE}^j
- Estimate the mean squared error for each population of estimator values.
```

- How do these estimated MSEs compare to your theoretical MSEs?

Answer 5

Code:

```
print(chr(27) + "[2J")
import numpy as np
n = 100 #features
 L = 10 #dist. between 0 to L
MAX ITER = 1000
                               #keeping results same
np.random.seed(0)
#Generating the (pseudo) random data (uniform distribution)
for i in range(MAX_ITER):
    data.append(np.random.uniform(0, L, n))
estd_l_mom = []
estd l mle = []
                               #estimated L M.L.E. vector
for i in range(len(data)):
    estd_l_mom.append(2*np.mean(data[i]))
    estd l mle.append(max(data[i]))
estd mse mom = 0
for val in estd 1 mom:
       estd_mse_mom += pow((val - L),2)
estd mse mom = estd mse mom/MAX ITER
estd mse mle = 0
for val in estd l mle:
       estd mse mle += pow((val - L),2)
estd mse mle = estd mse mle/MAX ITER
print("Estimated MSE MOM: ",estd mse mom)
print("Estimated MSE MLE: ", estd mse mle)
# Calculating theoretical MSEs
theoretical mse mom = pow(L, 2)/(3*n)
theoretical mse mle = (2*pow(L,2))/((n+2)*(n+1))
print("Theoretical MSE MOM: ", theoretical mse mom)
print("Theoretical MSE MLE: ", theoretical mse mle)
# Difference between estimated and theoretical MSEs
print("Absolute difference between theoretical and estimated MSE w.r.t.
MOM: ", abs (theoretical mse mom - estd_mse_mom))
```

print("Absolute difference between theoretical and estimated MSE w.r.t.
MLE: ",abs(theoretical_mse mle - estd_mse_mle))

Output:

The difference between the estimated and theoretical MSEs are quite small.

Also, we notice that $MSE(\hat{L}_{MLE}) < MSE(\hat{L}_{MOM})$ for both estimated and theoretical values and hence we can confirm our result from Answer 4 that $MSE(\hat{L}_{MLE})$ is a better estimator.

Question 6

You should have shown that \hat{L}_{MLE} , while biased, has a smaller error over all. Why? The mathematical justification for it is above, but is there an explanation for this?

Answer 6

We have proved in Answer 2 that \hat{L}_{MLE} is biased.

We have proved in Answer 4 that $MSE(\hat{L}_{MLE}) < MSE(\hat{L}_{MOM})$.

From the following: $MSE(\widehat{\theta}) = bias(\widehat{\theta})^2 + var(\widehat{\theta})$, we notice that the weightage on bias is more compared to the variance, and yet $MSE(\widehat{L}_{MLE}) < MSE(\widehat{L}_{MOM})$. This implies that $var(\widehat{L}_{MOM})$ is very high compared to $var(\widehat{L}_{MLE})$ and the same was proven in Answer 3. As n increases $var(\widehat{L}_{MOM})$ increases at a higher rate compared to $var(\widehat{L}_{MLE})$, although $bias(\widehat{L}_{MOM}) = 0$ and $bias(\widehat{L}_{MLE}) > 0$. \widehat{L}_{MLE} has a better bias vs variance trade off and thus we get lower value of $MSE(\widehat{L}_{MLE})$ compared to $MSE(\widehat{L}_{MOM})$.

Question 7

Find $P(\hat{L}_{MLE} < L - \varepsilon)$ as a function of L, ε , n. Estimate how many samples I would need to be sure that my estimate was within ε with probability at least δ .

Answer 7

$$\hat{L}_{MLE} = \max_{i=1,\dots,n} X_i$$
 (given)

$$F(x) = \left(\frac{x}{L}\right)^n$$
 (from Answer 2 ... 1)

Therefore,
$$P(\hat{L}_{MLE} \leq L - \varepsilon) = \left(\frac{L - \varepsilon}{L}\right)^n$$

or, $P(\hat{L}_{MLE} \leq L - \varepsilon) = \left(1 - \frac{\varepsilon}{L}\right)^n$ (1)

$$P(L - \hat{L}_{MLE} < \varepsilon) > \delta \qquad \qquad \text{(question requirement)}$$
 or,
$$1 - P(L - \hat{L}_{MLE} > \varepsilon) > \delta$$
 or,
$$1 - \delta > P(L - \hat{L}_{MLE} > \varepsilon)$$
 or,
$$P(L - \hat{L}_{MLE}) < 1 - \delta \qquad \qquad (2)$$

$$\left(1 - \frac{\varepsilon}{L}\right)^n < 1 - \delta \qquad \qquad \text{(from 1 and 2)}$$
 or,
$$n \ln\left(1 - \frac{\varepsilon}{L}\right) < \ln(1 - \delta) \qquad \qquad \text{(taking log on both side)}$$
 or,
$$n < \frac{\ln(1 - \delta)}{\ln\left(1 - \frac{\varepsilon}{L}\right)}$$

Since δ lies between 0 and 1, this makes $\ln(1-\delta)$ negative and the inequality is reversed:

Since
$$\delta$$
 lies between 0 and 1, or,
$$n > -\frac{\ln(1-\delta)}{\ln(1-\frac{\varepsilon}{L})}$$
 or,
$$n > \frac{\ln(1-\delta)^{-1}}{\ln(1-\frac{\varepsilon}{L})}$$
 or,
$$n > \frac{\ln(\frac{1}{1-\delta})}{\ln(1-\frac{\varepsilon}{L})}$$

Thus, we need at least $\frac{\ln(\frac{1}{1-\delta})}{\ln(1-\frac{\epsilon}{\epsilon})} + 1$ samples to be sure that our estimate was within ϵ with probability at least δ .

Question 8

Show that

$$\hat{L} = \left(\frac{n+1}{n}\right) \max_{i=1,\dots,n} X_i$$

is an unbiased estimator, and has a smaller MSE still.

Answer 8

$$\hat{L} = \left(\frac{n+1}{n}\right) \max_{i=1,\dots,n} X_i \qquad \text{(given...1)}$$

$$bias(\hat{\theta}) = \theta - \mathbb{E}[\hat{\theta}] \qquad \text{(given)}$$
or,
$$bias(\hat{L}) = L - \mathbb{E}[\hat{L}]$$
or,
$$bias(\hat{L}) = L - \mathbb{E}\left[\left(\frac{n+1}{n}\right) \max_{i=1,\dots,n} X_i\right] \qquad \text{(from 1)}$$
or,
$$bias(\hat{L}) = L - \left(\frac{n+1}{n}\right) \mathbb{E}\left[\max_{i=1,\dots,n} X_i\right]$$
or,
$$bias(\hat{L}) = L - \left(\frac{n+1}{n}\right) \left(\frac{nL}{n+1}\right) \qquad \text{(from Answer 2 ... 3)}$$
or,
$$bias(\hat{L}) = 0 \qquad (2)$$

Thus, $\hat{L} = \left(\frac{n+1}{n}\right) \max_{i=1,\dots,n} X_i$ is an unbiased estimator.

$$var(\hat{L}) = var\left(\left(\frac{n+1}{n}\right) \max_{i=1,\dots,n} X_i\right)$$
or,
$$var(\hat{L}) = \left(\frac{n+1}{n}\right)^2 var\left(\max_{i=1,\dots,n} X_i\right)$$
or,
$$var(\hat{L}) = \left(\frac{n+1}{n}\right)^2 \left(\frac{nL^2}{(n+2)(n+1)^2}\right)$$
or,
$$var(\hat{L}) = \frac{L^2}{n(n+2)}$$
(from Answer 3 ... 3)
(3)

$$MSE(\hat{\theta}) = bias(\hat{\theta})^2 + var(\hat{\theta})$$
 (proven in Answer 1) or,
$$MSE(\hat{L}) = bias(\hat{L})^2 + var(\hat{L})$$
 or,
$$MSE(\hat{L}) = 0^2 + \frac{L^2}{n(n+2)}$$
 (from 2 and 3) or,
$$MSE(\hat{L}) = \frac{L^2}{n(n+2)}$$

$$MSE(\hat{L}_{MLE}) = \frac{2L^2}{(n+2)(n+1)}$$
 (from Answer 2...2)

We need to prove $MSE(\hat{L}) < MSE(\hat{L}_{MLE})$ or $MSE(\hat{L}) - MSE(\hat{L}_{MLE}) < 0$

$$MSE_{diff} = MSE(\hat{L}) - MSE(\hat{L}_{MLE})$$
 or,
$$MSE_{diff} = \frac{L^2}{n(n+2)} - \frac{2L^2}{(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{L^2(n+1) - 2L^2(n)}{n(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{nL^2 + L^2 - 2nL^2}{n(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{L^2 - nL^2}{n(n+2)(n+1)}$$
 or,
$$MSE_{diff} = \frac{(1-n)L^2}{n(n+2)(n+1)}$$

Since $n \ge 1$, $MSE_{diff} \le 0$, therefore $MSE(\hat{L}) \le MSE(\hat{L}_{MLE})$.

In general n has larger values than 1, hence $MSE(\hat{L}) < MSE(\hat{L}_{MLE})$ and the difference is more for larger values of n.
