A Methodological Evaluation of Meta-Analyses in tDCS - Motor Learning Research

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Abstract

With transcranial direct-current stimulation’s (tDCS) rising popularity both in motor learning research and as a commercial product, it is becoming increasingly important that the quality of evidence on its effectiveness be evaluated. Special attention should be paid to meta-analyses, as they usually have a larger impact than other types of studies. There is evidence for several methodological issues in the tDCS motor learning literature such as limited reproducibility, largely untested replicability, and marked hetereogeneity in the technical paramaters employed. These issues inevitably impact the quality of the meta-analyses synthesising these studies, which display methodological problems of their own. The aim of this project was to evaluate the methodological quality of meta-analyses estimating the effect of tDCS on motor learning with respect to three main aspects: transparency, reproducibility, and publication bias control. Akin to previous reviews with similar aims, we show that the methods and results sections of meta-analyses are severely underreported, which compromises the ability to judge the soundness of the methodogloical procedure adopted as well as its reproducibility. Furthermore, only one of the three meta-analyses reported having taken non-statistical approaches to control for publication bias and all three used older statistical methods which are known to produce results of limited validity. These results reemphasise the need to evaluate the quality of meta-analyses as they have the largest impact on research, clinical practice, and even consumer behaviour.

A Methodological Evaluation of Meta-Analyses in tDCS - Motor Learning Research

Transcranial direct-current stimulation (tDCS) refers to a non-invasive brain stimulation technique which involves delivering constant, low current to the brain via electrodes fixed on the scalp (Gazzaniga, Ivry, & Mangun, 2018). tDCS is believed to be capable of both temporarily and permanently altering brain function by increasing or decreasing excitability in targeted cortical regions depending on the type of stimulation implemented (Stagg & Nitsche, 2011). Positive current delivered via the anodal electrode causes neurons just beneath the region stimulated to depolarise, that is, to increase their excitability, rendering them more likely to start off an action potential. Neurons targeted by negative stimulation, on the other hand, become hyperpolerised, i.e., less likely to “fire” (Gazzaniga et al., 2018). An important third type of stimulation is sham stimulation, which involves delivering an initial brief current that quickly fades off, leaving an inactive electrode for the remainder of the experiment. Sham stimulation is used as the control condition in experimental settings. tDCS is known to be safe and to cause negligible side-effects, if any (Gianni et al., 2021; Nitsche et al., 2008).

The typical set-up of a tDCS protocol involves no more than a handful of inexpensive components and relatively simple steps (Gebodh et al., 2019; Woods et al., 2016). A standard tDCS device consists of an anodal (positively charged) and a cathodal (negatively charged) electrode attached to a battery-powered device which controls the type and intensity of the current as well as the duration of the stimulation. To ensure that the administered current actually reaches the target area, a few measures are usually taken to minimise resistance between the skin and the active electrodes. This includes using a sponge and an electrolyte-based contact medium such as gel and any materials necessary to make sure the electrodes remain fixed in their positions. Electroencephalography (EEG) or magnetic resonance imaging (MRI) is sometimes used to more accurately locate the region of interest or to monitor the physiological effects of tDCS in real time (Nitsche et al., 2008; Woods et al., 2016). More recent variations of tDCS used for at-home clinical interventions are even simpler (Riggs et al., 2018).

When considering these advantages, tDCS’s rapidly increasing popularity over the last two decades is not surprising (Buch et al., 2017). In clinical research, its efficacy for treating or attenuating depression (Brunoni et al., 2016), memory deficits in Alzheimer’s patients (Bennabi et al., 2014), pain (Luedtke et al., 2012), schizophrenia (Liu et al., 2021), and others, has been investigated. Interest in tDCS was not restricted to clinical settings: studies on healthy subjects have been conducted to test its effects on cognitive abilities such as language and memory (Horvath, Forte, & Carter, 2015), affective states (Austin et al., 2016), and motor skills, such as surgery (Hung et al., 2021a) and musical performance (Rosen et al., 2016).

Indeed, tDCS has become so established as a research tool that it is already approved for clinical use (with some restrictions) in several countries around the world (Fregni et al., 2015). Furthermore, an ever-richer variety of commercial tDCS products has become available on the market (Davis, 2016; Wexler, 2018; Zettler, 2017), prompting experts (Wurzman, Hamilton, Pascual-Leone, & Fox, 2016) to voice concerns over the increasing prevalence of this “do-it-yourself” use of tDCS. Out of the 449 such at-home tDCS consumers surveyed by Wexler (2018), 52% (237) did so for enhancement purposes. It appears likely that this *hype* surrounding tDCS products is partly due to considerable media coverage (Dubljević, Saigle, & Racine, 2014; Steenbergen et al., 2016).

One increasingly prevalent implementation of tDCS is to improve motor learning, that is, the acquisition of new motor skills or the improvement of existing ones. Here, complexity is compounded by the multifacetedness of motor skills and the variety of ways to measure them (Buch et al., 2017). They can range from the most mundane daily activities like walking or doing the dishes to finer skills that can require years to master such as playing a musical instrument or performing surgery. tDCS is most often used as a supplement to specialised training (Buch et al., 2017). The two main types of improvement that can occur are referred to as “online learning” (improvement over short time periods such as within a single training session or a day) and consolidated learning (over several hours, days, or training sessions).

Experimental paradigms aiming to evaluate such patterns of improvement typically may take the following form (Buch et al., 2017; Reis et al., 2009): following baseline measurements of different outcomes on the first day, study participants undergo a series of training sessions in a specific motor task (e.g., squeezing a hand-held force transducer to move the cursor on the computer screen quickly and accurately between a starting and different target positions). The participants receive tDCS (or sham, depending group and time point) during this training. Performance is measured subsequently (online) and/or 1 to several weeks or months later (consolidation/retention). Measuring motor learning is a complex task, too, as motor skills require an optimisation of a speed-accuracy trade-off. Different tasks and measurement instruments take this into account in different ways depending on context. For example, some tasks involve instructing participants to focus on either speed or accuracy. Another approach is to stick to neutral instructions and instead explicitly model this trade-off, as done by Reis et al. (2009):

Where is a fixed constant.

Although tDCS might hold promise for those and other applications, treating tDCS’s positive effects as an established fact is deemed by some as premature, at best (Buch et al., 2017). The technique’s apparent simplicity might lead researchers to disregard aspects which are crucial to obtaining reliable results (Gebodh et al., 2019; Woods et al., 2016). When conducting a tDCS experiment, the researcher must select the type of stimulation, electrode size, shape and material; the duration of the stimulation; the location of the electrodes; and the current intensity, all of which are factors that have a substantial impact on stimulation efficacy.

The wide range of different possible tDCS settings appears to be well utilised by researchers as there is evidence of considerable heterogeneity in the literature with respect to tDCS-related parameters, tasks, and outcomes (Buch et al., 2017). Unsurprisingly, tDCS research also appears to suffer from the same methodological problems as neighbouring fields in the behavioural and cognitive (neuro)sciences, such as suboptimal reproducibility, untested replicability, underpowered studies (Minarik et al., 2016), and publication bias (Button et al., 2013). These issue make it difficult to evaluate the evidence or reach conclusions based on it. Furthermore, they hamper cumulative science as it is difficult to synthesise the results of studies for the purpose of a meta-analysis when they vary too much in quality and methodological approach (Borenstein, Hedges, Higgins, & Rothstein, 2009b).

Meta-analysis refers to a method of quantitatively synthesising the results of a set of studies while weighing individual studies differently depending on certain criteria, most commonly sample sizes. Meta-analyses are usually cited more frequently than primary studies about the same topics, are commonly assumed to provide the most accurate estimate of an effect, and have a large impact on theory development as well as policy and clinical practice (Gopalakrishnan & Ganeshkumar, 2013; Gøtzsche, Hróbjartsson, Marić, & Tendal, 2007; Ioannidis, 2016; Lakens et al., 2017; Morganti, 2007). Regrettably, meta-analyses can and very often *do* suffer from methodological problems that go beyond those of the primary studies included (Ioannidis, 2016).

One crucial issue is reporting transparency. Researchers must make a great number of decisions at multiple stages of a study’s timeline. These decisions impact the outcome of the study to various degrees and only by providing a full, justified account of which path was taken at each fork can other researchers evaluate the soundness of the methodological approach (Botvinik-Nezer et al., 2020; Gelman & Loken, 2013; Hyatt et al., 2020; Simmons, Nelson, & Simonsohn, 2011). Similarly, meta-analysts are faced with a plethora of different decisions with regards to which databases to search and using which strings; primary study selection and exclusion; data extraction; statistical methods, among others (Ada, Sharman, & Balkundi, 2012; Geyskens, Krishnan, Steenkamp, & Cunha, 2009; Guzzo, Jackson, & Katzell, 1987; Nieminen, Nicklin, McClure, & Chakrabarti, 2011; Schalken & Rietbergen, 2017; Valentine, Pigott, & Rothstein, 2010; Voracek, Kossmeier, & Tran, 2019). Although the extent to which these “researcher degrees of freedom” (Simmons et al., 2011) impact the conclusions of the meta-analysis is not a completely uncontroversial issue[[1]](#footnote-20), there is consensus regarding the importance of transparently reporting these decisions (Aguinis, Pierce, Bosco, Dalton, & Dalton, 2011), as it is difficult to assess the quality and trustworthiness of that which one does not have access to (Page, McKenzie, et al., 2021).

Hundreds of different sets of reporting guidelines have been developed for different types of studies different fields (“Reporting guidelines | The EQUATOR Network,” n.d.) including several for meta-analyses, e.g., the Cochrane Handbook (Higgins et al., 2019), Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA, Moher et al., 2000; Moher, Liberati, Tetzlaff, Altman, & Group, 2009; Page, McKenzie, et al., 2021), Meta-Analysis Reporting Standards (MARS, American Psychological Association, n.d.), with PRISMA being the mostly widely known and adopted (Page & Moher, 2017). Although PRISMA’s predecessor has been available since 2000, meta-scientific evaluations of adherence to PRISMA and other guidelines have suggested that reporting standards of meta-analyses are generally poor and that these guidelines are rarely fully adhered to (see Table 1 for an overview of such studies).

A related issue is that the less information a meta-analyst provides about their analytical procedures, the harder it is to reproduce their meta-analysis (Aguinis et al., 2011; Gøtzsche et al., 2007; Lakens et al., 2017; Maassen, Assen, Nuijten, Olsson-Collentine, & Wicherts, 2020). Reproducibility (and relatedly, replicability) remain a hot topic in the behavioural and biomedical sciences, with various fuzzy or inconsistent definitions still floating around (Goodman, Fanelli, & Ioannidis, 2016; Plesser, 2018). Both terms have been somewhat indiscriminately used to refer to the ability to obtain the same results of a given study by repeating their procedure. One widely adopted way to distinguish them is by defining reproducibility as the ability to obtain the numerical findings of the original work using their methods (and if applicable, data pre-processing and analysis code) and the original data. (Direct) replication, on the other hand, refers to repeating the entire study using new data (Broman et al., 2017).

This definition, however, does not do full justice to the nuances of the concept of reproducibility. Importantly, this definition fails to account for the different types of reproducibility and emphasises one type in particular: *results reproducibility* (Goodman et al., 2016). Two further types can be differentiated: *methods reproducibility*, which refers to “the provision of enough detail about study procedures and data so the same procedures could, in theory or in actuality, be exactly repeated” (Goodman et al., 2016, p. 2), and *inferential reproducibility*, which describes the ability to draw “qualitatively similar conclusions from either an independent replication of a study or a reanalysis of the original study” (Goodman et al., 2016, p. 4). For example, a study might report its methods extensively enough for it to be easily reproducible methodologically, but for which one obtains different results from the original upon a reproduction attempt, which could happen due to erroneous descriptions of methods or mistakes in the implementation of the described methods.

Similarly, inferential reproducibility can depend on the magnitude of the effect reported in a given study as two interpretors of the same results might be in stark disagreement regarding the inferences which can be drawn from the them. A -value of 0.04, for example, might have an entirely different “significance” for a hard frequentist than for a more Bayes-inclined researcher. Another limitation to the definition above is its implication of a dichotomous nature of reproducibility. Especially with regards to methods and inferential reproducibility, it is somewhat imprudent to think of reproducibility as a binary feature since there are so many factors (think methodological decisions discussed above) that might play a role in deeming a study reproducible or not (Broman et al., 2017). Concretely, although it is possible to construct a framework which allows one to make an unambiguous yes or no decision when the results of a study are reproducible (e.g., see Steiner, Wong, & Anglin, 2019), methods reproducibility is better served by considering it as a function of how much information is provided about the methodological procedure.

Applied to meta-analysis, evaluating all types of reproducibility involves attempting to repeat 1. study search, 2. study screening and selection, 3. data extraction, 4. computing effect size estimates for each primary study included (primary ES), 5. computing the main pooled ES estimate (pooled ES), and often 6. computing “corrected” pooled ESs, 7. computing multiple pooled ESs based on study characteristics and/or conducting subgroup analyses (Cooper, Hedges, & Valentine, 2009). Here, the issue of reproducibility acquires yet another layer of complexity: reproducibility with regards to elements 3 to 7 depends not only on what is reported in the meta-analysis, but what is reported in the included primary studies. How hard it is to reproduce a primary ES, for example, is a function of the amount and precision of the information the meta-analysts reported about how they computed the ES *and* the degree to which this information corresponds to data that is accessible in the primary study.

Three reviews[[2]](#footnote-21) testing the reproducibility of meta-analyses exist as of 2021 (Gøtzsche et al., 2007; Lakens et al., 2017; Maassen et al., 2020, see Table 2). They emphasised somewhat different methodological aspects (e.g., complete reporting vs. computational correctness) but were uniform in their focus on reproducing data extraction and ES computation (both primary and pooled). Their conclusions about the reproducibility of meta-analyses were, although of varyingly grave consequences, also similar: the reproducibility of meta-analyses was severely limited due to under-reporting and errors. Despite the growing popularity of tDCS and the ensuing increase in number of meta-analyses estimating its effect on motor learning, no evaluation of the transparency or reproducibility of meta-analyses in this field has been conducted as far as we are aware.

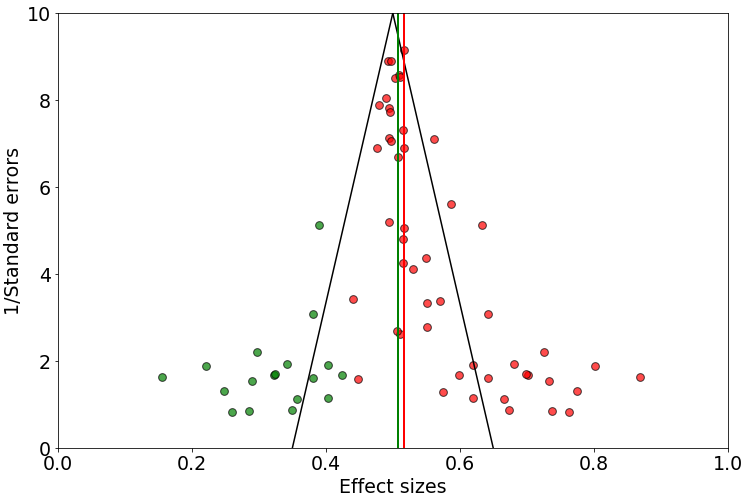
Beyond transparency in reporting and any effect it may have on reproducibility, a further issue that merits consideration when evaluating the methodology of meta-analyses is how the meta-analysts dealt with publication bias. Publication bias, that is, the tendency of researchers to suppress findings that do not go their way (“negative” results) and journals to selectively publish studies that report significant results, has been known to distort the scientific literature for over six decades, having been first discussed in 1959 by Sterling. It remains a major concern, especially in the “softer” social sciences (Fanelli, 2010), where the proportion of studies reporting “positive” results is markedly higher than in the physical sciences. This remarkably elevated rate of positive results combined with the meagre average power of studies in the behavioural and neurosciences give a clear indication of the presence of publication bias (Button et al., 2013; Szucs & Ioannidis, 2017; Szucs & Ioannidis, 2020). Since this holds true for the neighbouring field of transcranial *magnetic* stimulation, another widely used non-invasive brain stimulation technique (Amad et al., 2019), there is no reason to believe tDCS research to be spared.

Since this apparent overrepresentation of positive results can (and often does) inflate the effect size estimates of meta-analyses in a certain direction (Borenstein, Hedges, Higgins, & Rothstein, 2009a; Friese & Frankenbach, 2020; Vevea, Coburn, & Sutton, 2019), meta-analysis reporting guidelines such as PRISMA have consistently recommended reporting any attempts to account for publication bias (Moher et al., 2000, 2015; Page, McKenzie, et al., 2021). Assuming that the meta-analyst cannot influence which studies get published in the body of evidence they want to synthesise, they have at their disposal several measures they can take to mitigate the ubiquitous effects of publication bias. Locating as many unpublished studies as possible is likely to be the most effective method and searching repositories of potentially unpublished results are mandatory (e.g., ClinicalTrials.gov) when conducting a Cochrane systematic review (Higgins et al., 2019). Other such non-statistical approaches to minimising the effect of publication bias include searching pre-print and theses repositories, contacting authors of relevant studies to inquire about potentially file-drawered studies, not restricting the study search to articles written in English, etc..

This endeavour proving fruitless, the meta-analyst may next (or in addition) attempt to detect the presence of publication bias and/or adjust the meta-analytic ES estimate for it statistically. Many procedures for this purpose have been developed in the last three decades displaying varying performance in general and under certain conditions (for extensive reviews/comparisons of these methods see Banks, Kepes, & McDaniel, 2012; Borenstein et al., 2009a; Carter, Schönbrodt, Gervais, & Hilgard, 2019; Harrer, Cuijpers, Furukawa, & Ebert, 2021; Ioannidis, 2008; Jin, Zhou, & He, 2015; Marks-Anglin & Chen, 2020; McShane, Böckenholt, & Hansen, 2016; Moreno et al., 2009a; Peters et al., 2010; Renkewitz & Keiner, 2019; Rothstein, Sutton, & Borenstein, 2005b; Rücker, Carpenter, & Schwarzer, 2011; Schwarzer, Carpenter, & Rücker, 2015; T. D. Stanley, Doucouliagos, Ioannidis, & Carter, 2021; van Aert, Wicherts, & van Assen, 2016; Vevea et al., 2019).

These statistical methods can be divided into two main categories: small-study effects-based and -values-based. Small-study effect methods have been available for more than two decades and are very widely used (Borenstein et al., 2009a; Harrer et al., 2021; Vevea et al., 2019). In essence, they are based on the often observed positive correlation between the ES derived from a primary study and the ES’s corresponding standard error (SE, which can be seen as the inverse of the study’s sample size, ). This association between and ES should not exist, theoretically, but does due to many potential reasons, one of which publication bias. The main assumption here is that large studies mostly get published (and thus become easily findable by meta-analysts) regardless of their outcome, whereas small studies only get published when they are significant, which can (due to the small ) only happen when the ES is large, ergo a negative correlation between and ES. Some popular examples of methods in this category include:

* Trim-and-fill (Duval & Tweedie, 2000b): a simple non-parametric method which is based on an earlier, purely graphical, procedure, the funnel plot (Light & Pillemer, 1984). The trim-and-fill method attempts to correct for bias-induced asymmetry in the funnel plot by 1. *trimming* an arbitrary number of small studies with large ESs, usually the rightmost studies on the funnel plot, 2. computing the pooled ES based on this new set of studies, 3. imputing for each trimmed study a *filled* study which mirrors it on the other side of the pooled ES computed in the previous step (i.e., identical SE, filled ES trimmed ES pooled ES from step 2), 4. computing the pooled ES based on both the original studies and the trimmed ones. Example results of this procedure are depicted in Figure 1.



*Figure* *1.*  Funnel plot of simulated data illustrating asymmetry and the trim-and-fill method. The red dots represent the observed studies, the green dots the filled studies. The red and green lines are the fixed-effect pooled ES estimates of the observed studies only and the observed plus the filled studies, respectively.

* Egger’s test (Egger, Smith, Schneider, & Minder, 1997) parametrically checks for funnel plot asymmetry by regressing the ratio of ES to SE on the inverse of SE:
* PET-PEESE (T. Stanley, 2008; T. D. Stanley & Doucouliagos, 2014) is a more recent method that is gaining in popularity. Similarly to Egger’s test, both the precision-effect test (PET) and precision-effect estimate with standard error (PEESE) regress the ES on a proxy of its precision, namely the standard error in the case of PET,
* and the variance in the case of PEESE.
* However, the theoretical idea motivating this manoeuvre is quite different: here it is assumed that, because the intercept resulting from either one of these two regression analyses represents the ES when the sample variance (or SE) equals nought, this intercept should estimate the true ES.

The most prominent -value based method is Rosenthal’s (1979) Fail-Safe method. Using it, one can compute the number of additional studies reporting non-significant results needed to make the -value corresponding to the pooled ES no longer significant. This “traditional” procedure remains one of the most widely employed for detecting publication bias despite several explicit recommendations against its use (e.g., Borenstein et al., 2009a; Higgins et al., 2019; Rothstein, Sutton, & Borenstein, 2005a). Due its many limitations of both mathematical and theoretical nature, the Fail-Safe is “[…] now generally regarded as valueless” (Vevea et al., 2019, p. 390).

Two more recent -value based methods are the so called -curve (Simonsohn, Nelson, & Simmons, 2014b, 2014a; Simonsohn, Simmons, & Nelson, 2015) and -uniform (van Assen, van Aert, & Wicherts, 2015). -uniform is based on the notion that conditional -values are uniformly distributed given a fixed true effect size. -curve analysis, on the other hand, derives its logic from the fact that -values are uniformly distributed when the true effect equals nought (the null hypothesis is true), and right skewed when there is a true effect (the null hypothesis is false). Importantly, this also applies to the range of -values commonly defined as the region of significance in a given research field, e.g., in the behavioural, social, and biomedical sciences. Significant -values stemming from a biased pool of studies should, on the contrary, exhibit a left-skewed distribution.

This left skewness is commonly attributed to selective reporting and publishing of effects whose -values are over the significance threshold regardless of their magnitude. Such bodies of research are considered to possess little or no *evidential value* (Simonsohn et al., 2014a) as one assumes most of their significant effects to be the product of a combination of -hacking[[3]](#footnote-23) and selective reporting/publication bias. -curve analysis therefore involves testing for evidential value in a set of ESs by conducting 6 statistical tests on the subset containing exclusively significant ESs: one binomial test of right skewness, two parametric tests of right skewness, one binomial tests of flatness, and two parametric tests of flatness. There are two parametric tests each as the second one is conducted on half the curve only, i.e., the interval , in order to account for “ambitious -hacking.” The adjusted ES estimate produced by the analysis is, however, based on the full curve.

The last class of publication bias assessment tools, selections models, cannot be exclusively assigned to either category as they are a very versatile methods which allow their user to model *any* process assumed to generate bias in the meta-analytic estimates (Harrer et al., 2021; Vevea et al., 2019). The main purpose of these models is to adjust the observed pooled ES estimate by merging two functions: one describing the distribution of ESs in the absence of bias and the other describing the mechanism by which ESs are assumed to be selected for reporting or publication. Selection models have also been around for decades, although they have enjoyed much less use as they are considerably more complex conceptually and have been barely implemented in commercial point and click software (Vevea et al., 2019).

One relatively simple variety is the three-parameter selection model (McShane et al., 2016), which, as the name suggests, only estimates three parameters: the true average effect , the between-study variance , and the likelihood of obtaining non-significant ES in relation to obtaining a significant one, . What is fixed is the cut-off value , which is usually the significance threshold for a one-sided -value, 0.025. The relative likelihood of obtaining a -value in this interval, , is set to one. Hence, if the estimated is close to one, the adjusted pooled ES estimate will deviate little from the one based on a standard random effects model as this would indicate a similar likelihood of obtaining a non-significant ES to obtaining a significant one. Otherwise, the significant ESs will be downweighted if and upweighted if , which diminishes or augments the pooled ES estimate, respectively.

Fail-Safe ’s shortcomings were already mentioned. However, this should not be taken to mean that it is the only method with limitations. In fact, Fail-Safe ’s meaninglessness is one of the very few things the publication bias assessment literature appears to be in agreement about (Carter et al., 2019; Harrer et al., 2021). The methods vary considerably in their assumptions and hence under what conditions they perform well. Different developers of such methods disagree *strongly* about which conditions are realistic in which disciplines and to what degree this should inform the methods’ construction. For example, Simonsohn (2017), one of the developers of the -curve, deems PET-PEESE to be worse than homoeopathy as it (PET-PEESE) can be actively harmful whereas homoeopathy is simply ineffective. Vevea et al. (2019)[[4]](#footnote-24), on the other hand, see the -curve and -uniform methods as reinventions of the wheel and as “[…] modified versions of simplistic early weight-function models” (p.392). Duval and Tweedie (2000a)[[5]](#footnote-25) quote DuMouchel and Harris (1997) as writing in reference to selection models: “attempts to assess publication bias beyond simple graphs like the funnel plot seem to involve a tour de force of modeling, and as such are bound to run up against resistance from those who are not statistical modeling wonks” (p. 95).

Given this intricate state of affairs when it comes to statistical methods of publication bias assessment and correction, it should have become clear that as of yet, there is neither a single method nor a single set of methods which can be recommended as an all-purpose tool (Harrer et al., 2021; Vevea et al., 2019). Simulations studies have shown that no single method outperforms all the other in all conditions (Carter et al., 2019). Another point on which a consensus seems to prevail is that these methods should be used as sensitivity analyses, both with respect to a single method (e.g., by varying the assumptions used) and with respect to what and how many tests one uses (Carter et al., 2019; McShane et al., 2016; Vevea et al., 2019). We are not aware of studies investigating how publication bias is accounted for in tDCS-motor learning research.

In sum, 4 principal premises motivated our work: 1. tDCS appears to be remarkably popular as a research tool in basic and clinical research as well as in form of commercial gadgets, 2. meta-analyses of tDCS’s effect on motor learning might be substantially impacting research and clinical practice and/or tDCS’s uptake as a commercial product, 3. meta-analyses are often of suboptimal methodological quality, which compromises their credibility and informativeness, 4. no methodological evaluation of meta-analyses in tDCS-motor learning research has been conducted as of yet. Our aim was thus to provide such an evaluation while partitioning the concept of methodological quality of a meta-analysis into 3 main aspects:

1. Reporting quality/transparency: did the meta-analysis adhere to the PRISMA reporting guidelines? Which items were neglected, if any?
2. Reproducibility: how hard is it to reproduce the pooled ES estimate reported in the meta-analysis, if at all, based on the information provided therein? What necessary pieces of information were missing, if any? If enough information was provided to reproduce the methods, does the reproduced pooled ES equal the reported pooled ES?
3. Consideration of publication bias: Did the meta-analysis report attempts to minimise the effects of publication bias? Which statistical or non-statistical methods were used? By conducting a publication bias analysis of our own, do we arrive at the same conclusion regarding the presence of publication bias as the meta-analysts?

Secondary methodological aspects we evaluated included:

1. Had the meta-analysis been pre-registered? If yes, did the pre-registration protocol adhere to meta-analysis pre-registration protocol reporting guidelines PRISMA-P (Moher et al., 2015)? Pre-registering the hypotheses and data analysis plan is essential for distinguishing between confirmatory and exploratory findings (Nosek, Ebersole, DeHaven, & Mellor, 2018). In the context of reviews and meta-analyses, pre-registration protocols “[…] act as a guard against arbitrary decision making during review conduct, enable readers to assess for the presence of selective reporting against completed reviews, and, when made publicly available, reduce duplication of efforts and potentially prompt collaboration” (Shamseer et al., 2015, p. 1).
2. Did the meta-analysis discuss outliers? How was the presence of outliers addressed? Extreme effect size values may affect the validity and robustness of meta-analytic results and there is a general consensus that meta-analyses should examine the presence of outliers and to what extent they influence conclusions (Viechtbauer & Cheung, 2010).

# 1 Methods

Although our methodological approach mostly followed the plan pre-defined in the thesis proposal (accessible on the thesis’ Open Science Framework [OSF] project [osf.io/uagvf/](https://osf.io/uagvf/)), there were important deviations from the plan, especially with respect to reproducibility testing. A document listing these deviations can also be found on the OSF project. For data wrangling, analysis and visualisation, we used R (Version 4.1.2, R Core Team, 2021) and the packages dmetar (Version 0.0.9000, Harrer, Cuijpers, Furukawa, & Ebert, 2019), dplyr (Version 1.0.7, Wickham, François, Henry, & Müller, 2021), forcats (Version 0.5.1, Wickham, 2021a), ggplot2 (Version 3.3.5, Wickham, 2016), MAd (Version 0.8.2.1, Hoyt, 2014), Matrix (Version 1.3.4, Bates & Maechler, 2021), meta (Version 5.1.0, Balduzzi, Rücker, & Schwarzer, 2019), metafor (Version 3.0.2, Viechtbauer, 2010), purrr (Version 0.3.4, Henry & Wickham, 2020), readr (Version 2.1.0, Wickham & Hester, 2021), readxl (Version 1.3.1, Wickham & Bryan, 2019), stringr (Version 1.4.0, Wickham, 2019), tibble (Version 3.1.6, Müller & Wickham, 2021), and tidyr (Version 1.1.4, Wickham, 2021b).

The package renv (Version 0.14.0, Ushey, RStudio, & PBC, 2021) was used to ensure long-term reproducibility of our analyses. This thesis was written using RMarkdown and the associated packages papaja (Version 0.1.0.9997, Aust & Barth, 2020), kableExtra (Version 1.3.4, Zhu, 2021), knitr (Version 1.36, Xie, 2014, 2015, 2021). Both writing and analysis were done using RStudio (Version 2021.9.1.372, RStudio Team, 2021). Data were extracted from figures using WebPlotDigitizer ([apps.automeris.io/wpd/](https://apps.automeris.io/wpd/), Rohatgi, 2021). A video demonstration of how we extracted data from figures is available on the OSF project.

## 1.1 Sample of meta-analyses

Three meta-analyses (Hung et al., 2021b; Kang, Summers, & Cauraugh, 2016; Kang, Weingart, & Cauraugh, 2018) were selected based on the following eligibility criteria:

* Meta-analysis studies which quantitatively synthesise multiple (at least 3) primary studies on the effects of tDCS on motor learning.
* No restriction on primary outcomes (e.g., how speed or accuracy were measured), designs of primary studies (e.g., randomised vs. crossover designs), or participants (e.g., clinical or healthy subjects) in the primary studies were imposed.

Exclusion criteria:

* Reviews of any type without a quantitative synthesis
* Primary studies
* Bayesian meta-analyses
* Reviews which did not report a “main” meta-analysis, but rather multiple meta-analyses of subgroups of studies.

All three meta-analyses were found via quick, non-systematic Google Scholar or Web of Science searches.

## 1.2 Data extraction

### 1.2.1 Elements extracted.

There were two types of elements extracted:

1. Primary study level variables (e.g., sample sizes, means and SDs), were extracted from both the primary studies and the meta-analyses.
2. Meta-analysis level variables (e.g., pooled SMD, publication bias control related variables).

Both data sheets, along with older/empty versions, are also available on the OSF project page. They can be found by navigating to the GitHub thesis branch in the folder “data\_thesis” under the names “Data\_ps\_raw\_updated” and “Data\_ma\_raw,” respectively. The codebook explaining column headers in the data sheets is also accessible on the OSF project (folder “Post\_data\_extraction”).

### 1.2.2 Data extraction procedure.

The data extraction procedure for the purpose of testing reproducibility (see below) differed substantially from the rest. Data for all non-reproducibility-related variables were extracted directly to Google Sheets.

## 1.3 Procedure, coding, and data analysis

### 1.3.1 Reporting quality.

The adherence of each meta-analysis to the PRISMA reporting guidelines (Liberati et al., 2009; Moher et al., 2009) were checked. Concretely, for each of the 27 PRISMA items, we coded whether the meta-analysis reported the relevant information as recommended, regardless of whether the meta-analysis reported having adhered to any reporting guidelines. Besides coding whether the meta-analysis had been pre-registered, we checked whether the pre-registration protocol adhered to the PRISMA-P reporting guidelines (Moher et al., 2015; Shamseer et al., 2015) for such protocols, provided that the meta-analysis had been pre-registered and published later than 2015.

### 1.3.2 Reproducibility.

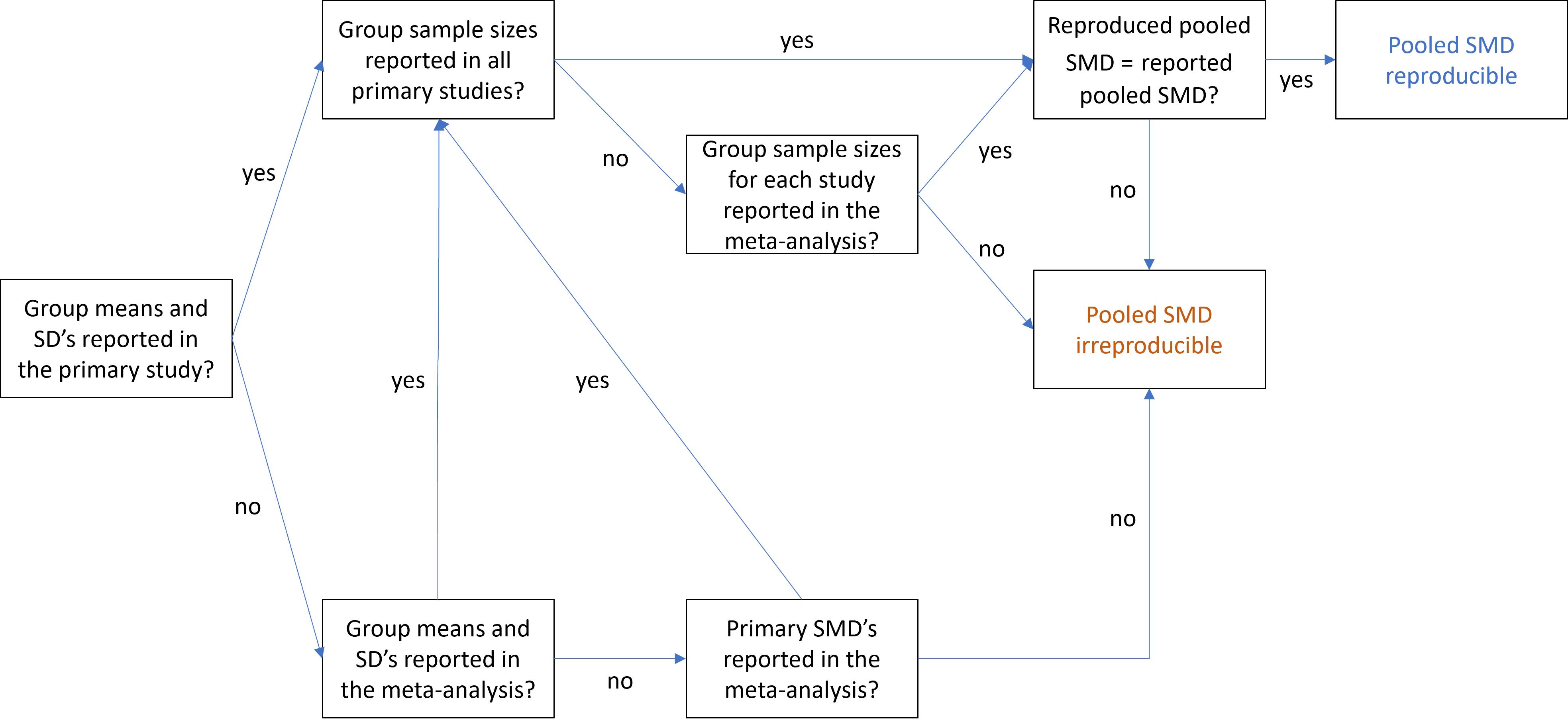
We based our treatment of meta-analysis reproducibility on the definition and principles of reproducibility put forward by the American Statistical Association (Broman et al., 2017): a meta-analysis is reproducible if its authors provided enough information to go through all the necessary procedures (search, screening, data extraction, calculation of primary ESs, calculation of the pooled ES…) to arrive at the same results. However, although we acknowledge the importance of all these steps, we, like Gøtzsche et al. (2007), Lakens et al. (2017), and Maassen et al. (2020), focused on data extraction and calculation of ES estimates.

Since reproducing the primary ESs is a necessary step towards reproducing the pooled SMD, we constructed a scheme for classifying their reproducibility status. The two variables in this scheme are A. whether the primary ES could be successfully reproduced or approximated numerically[[6]](#footnote-33) (results reproducibility) and B. whether the procedure we followed in reproducing the ES strictly corresponded to the information given in the meta-analysis or to the procedure apparently adopted for at least two other ESs included in the meta-analysis (methods reproducibility).

Table 3 depicts the classification system in which we use SMD (for standardised mean difference) instead of ES since SMD was the only measure of ES employed in the meta-analyses we evaluated. “Procedure” here includes such analytic decisions as using -values or test statistics in combination with sample sizes to estimate the SMD for a given primary study, using the raw means and SDs instead of means and SDs of changes in the outcome from baseline or vice versa, using follow-up means and SDs instead of post means and SDs, etc.. The second variable in the classification system was thus mainly adopted to capture cases where there was a discrepancy between how and using which values from the primary study the meta-analysts reported having computed a primary ES and how they actually computed it.

Reproducibility testing was an iterative process which involved several rounds of data extraction. The initial round of data extraction and testing reproducibility (depicted in Figure 2) did not yield a lot of reproducible SMDs as most primary studies did not report the values necessary to directly compute an SMD and/or its sampling variance (e.g., for a between groups Cohen’s this would be the group means, SDs and sample sizes [s]). Therefore, a less strict data extraction procedure was adopted, which involved the following steps:

1. For each primary SMD, we first looked for the raw means and SDs of the outcome reported as having been used by the meta-analysts. If the primary study reported multiple sets of means and SDs which can be seen as corresponding to the outcome described in the meta-analysts (e.g., the outcome in the meta-analysis for a given primary study is “Fugl-Myer Test” but the primary study reports values for “Fugl-Myer Test - upper limbs” and “Fugl-Myer Test - full”), all sets were extracted.
2. If no means and SDs for the relevant outcome were reported in the primary study, means and SDs were extracted from figures. If no figures were reported which contained means and SDs (or SEs or confidence intervals [CIs]), and/or -values for tests on the relevant outcome were extracted, which in combination with s can be converted to SMDs.
3. Based on all extracted values, we computed each primary SMD using the estimator (Cohen’s or Hedges’ ) reported as having been used by the meta-analysts. If this information was not given in the meta-analysis, we tried both formulas and for further analysis used the one which consistently approximated the reported SMDs better.
4. If none of the values extracted reproduced a given SMD, we double checked the correctness of the data extracted and, in some cases, extracted more values from the primary study and computed the SMD based on those.
5. We computed the sampling variances based on the SMDs and the corresponding s.
6. For each meta-analysis, we fit three meta-analytic models: 1. Based on the faithfully reproduced (but not necessarily reproducible) SMDs. 2. Based on the faithfully reproduced SMDs plus brute-force reproduced. 3. Based on the SMDs and sampling variances reported in the meta-analysis[[7]](#footnote-34). This was done to test for analytical reproducibility of the pooled SMD and to find out which between-study heterogeneity estimator was used.



*Figure* *2.*  Initial procedure for reproducing the pooled SMD.

### 1.3.3 Publication bias control.

The non-statistical approaches to accounting for publication bias that we coded were:

1. Searched clinical trial registries (e.g., ClinicalTrials.org)
2. Searched thesis and dissertation repositories (e.g., ProQuest)
3. Did not restrict their search to studies written in English
4. Contacted known researchers in the field to inquire about unpublished results
5. Contacted authors of included studies to ask for raw data or unpublished results

Besides coding whether any statistical methods were used at all and which, we tested for publication bias in each meta-analysis using 3 different methods: PET-PEESE (T. Stanley, 2008; T. D. Stanley & Doucouliagos, 2014), -curve (Simonsohn et al., 2014b, 2014a), and the three-parameter selection model (McShane et al., 2016). For each meta-analysis, we concluded that publication bias is a concern in the studies synthesised if at least two out of the three methods used yielded results indicating the presence of publication bias. Although such a “majority vote” approach is not recommended for actual publication bias assessment (Carter et al., 2019, p. 140), we adopted it so as to have an unambiguous decision rule whether our analysis agrees with that of the meta-analysts or not. For these analyses, we used the SMDs and the sample sizes reported in the meta-analyses along with the sampling variances extracted from the funnel plots in the case of the first two meta-analyses and calculated based on the CIs in the case of the third meta-analysis.

### 1.3.4 Outlier/influential study analysis.

Besides coding whether each meta-analysis checked for the existence of outliers among the included studies and how this was done, we tested this ourselves for each meta-analysis using the leave-one-out method (Harrer et al., 2021; Tobias, 1999).

# 2 Results

## 2.1 General description of the meta-analyses

Table 4 gives an overview of the three meta-analyses we reviewed. Meta-analyses 1 and 2 aimed to estimate the effectiveness of tDCS for improving motor function in post-stroke patients, although meta-analysis 2 focused exclusively on the effects of cathodal tDCS. Meta-analysis 3 investigated effectiveness of tDCS for improving surgical performance of surgery trainees. Meta-analyses 2 and 3 reported more primary SMDs than primary studies because they drew two comparisons from some primary studies. In meta-analysis 1, the double comparisons represented cathodal vs. sham and anodal vs. sham pairs, whereas in meta-analysis comparison pairs were based on two different outcomes. The first meta-analysis synthesised the results of 13 randomised controlled trials (RCTs) and 4 crossover trials, the second 6 RCTs and 9 crossover trials, the third 5 RCTs and one crossover trial. None of three meta-analyses had been pre-registered. According to Google Scholar, the meta-analyses were cited 197, 12, and two times, respectively, as of 16.11.2021. The average group sample sizes of the primary studies included in the three meta-analyses, i.e., mean of , were ~ 13, ~ 12, and ~ 16, respectively.

## 2.2 Adherence to guidelines

Meta-analyses 1 and 2 did not report having adhered to any reporting guidelines. Despite this, they can be seen as having reported the content of 19 and 22 items, respectively, out of the 27 PRISMA (Moher et al., 2009) items. Meta-analysis 3 reported having adhered to the PRISMA (Page, McKenzie, et al., 2021) guidelines and reported the content of 26 out of the 27 items. Table 5 summarises the three meta-analyses’ adherence to PRISMA items. Note that meta-analysis 3 indicated that their work had *not* been pre-registered.

## 2.3 Reproducibility

We faced enormous difficulties in reproducing the primary SMDs and their sampling variances for all three meta-analyses, not least due to scanty reporting of the methods sections. None of the three meta-analyses shared their data or provided data analysis code. The process necessitated several very time-consuming rounds of data extraction and testing. Which relevant pieces of information were missing and how many primary SMDs could be successfully reproduced is reported below for each meta-analysis separately. Naturally, the effect of irreproducible primary SMDs on the pooled SMDs is also reported. Since all three meta-analyses reported having fit a random effects model using the Comprehensive Meta-Analysis software^[In the case of the first meta-analysis, we were informed by the first author of the meta-analysis that they used the Comprehensive Meta-Analysis software in response to an email asking for further information. Our email’s text as well as all the reproducibility-relevant information contained in the meta-analysts’ response can be found in the document “Email\_to\_authors” on our OSF project, which per default estimates between-study heterogeneity via the Der-Simonian-Laird method (DerSimonian & Laird, 1986), we used these settings for all our analyses, too.

### 2.3.1 Meta-analysis 1.

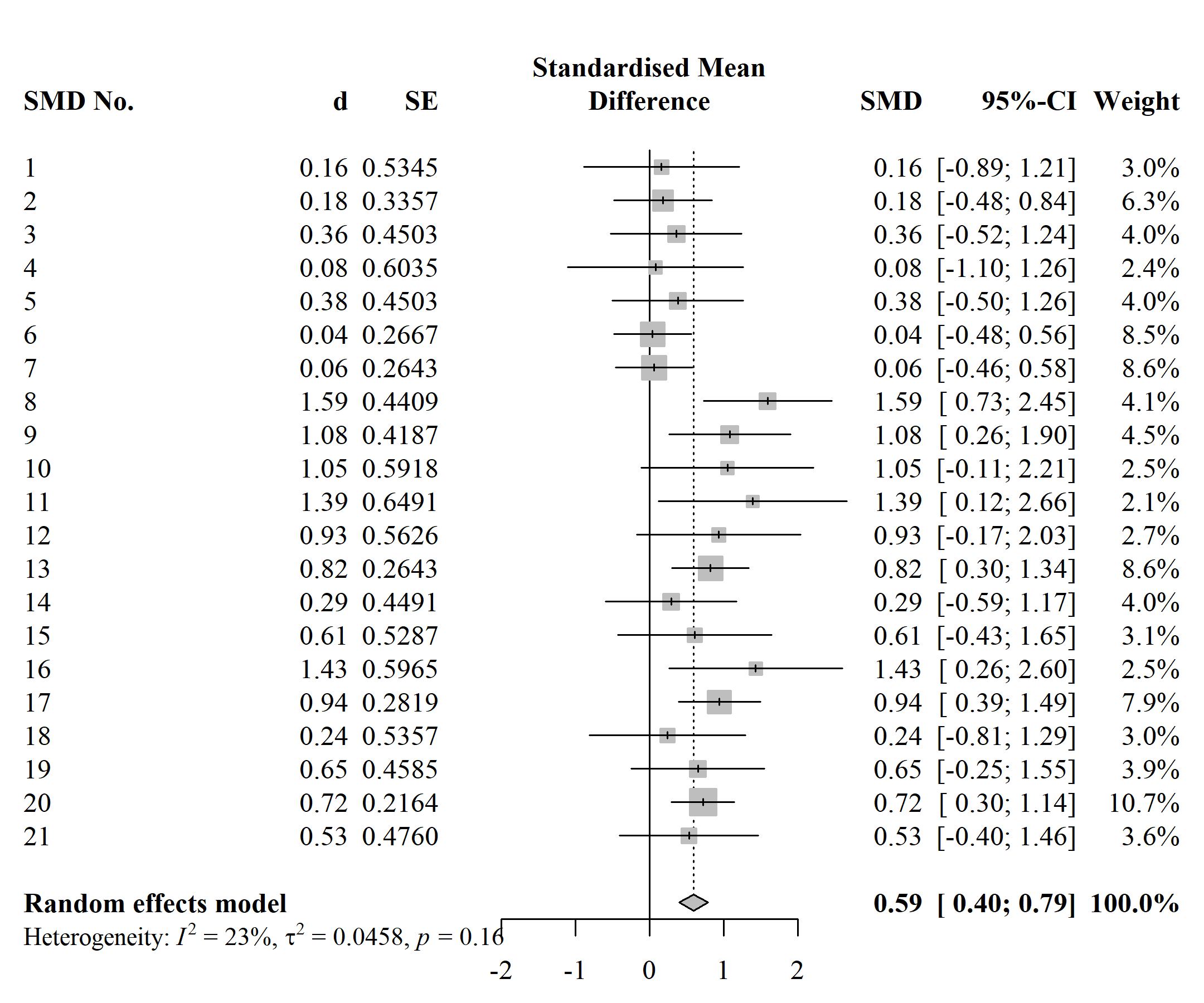
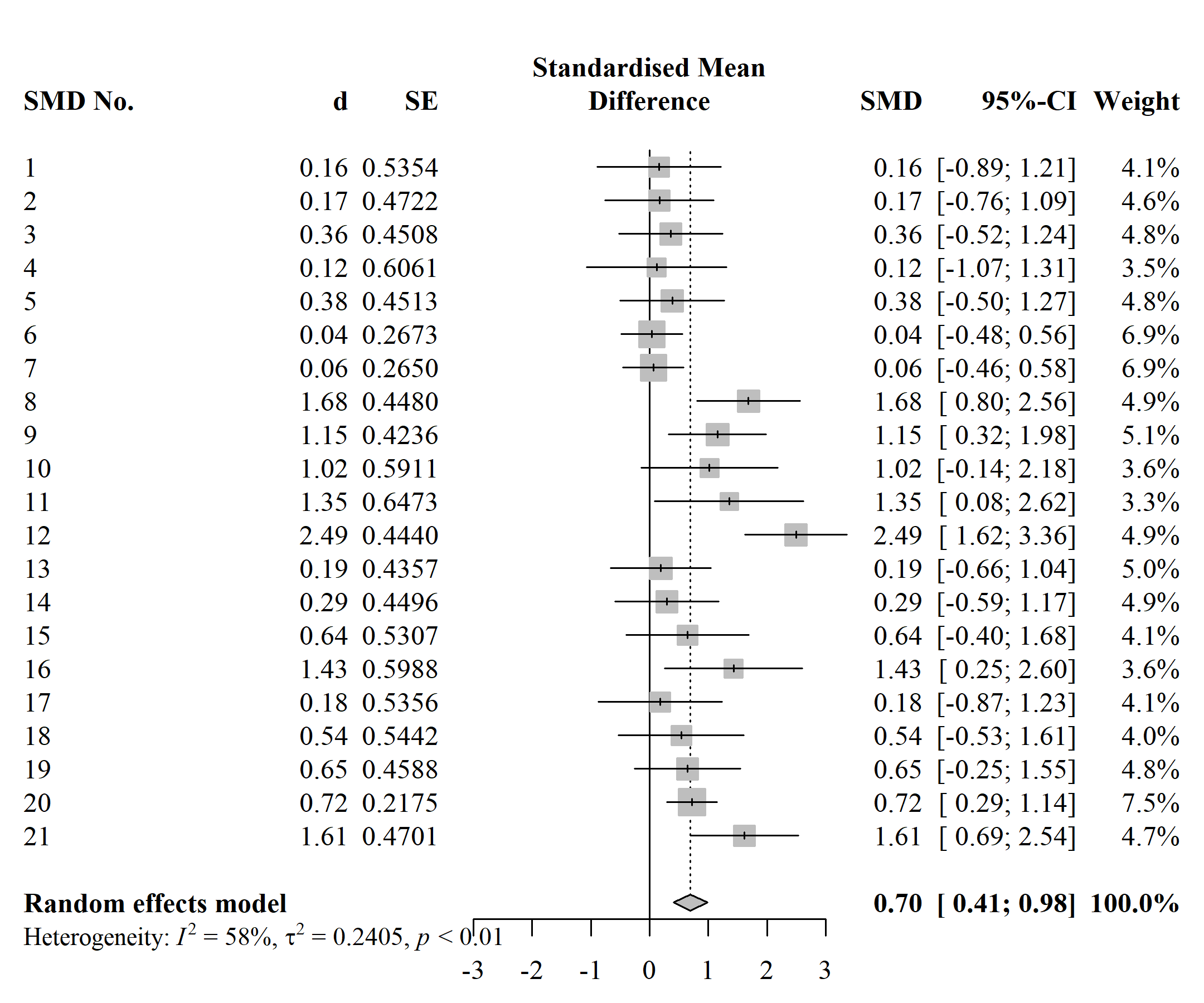
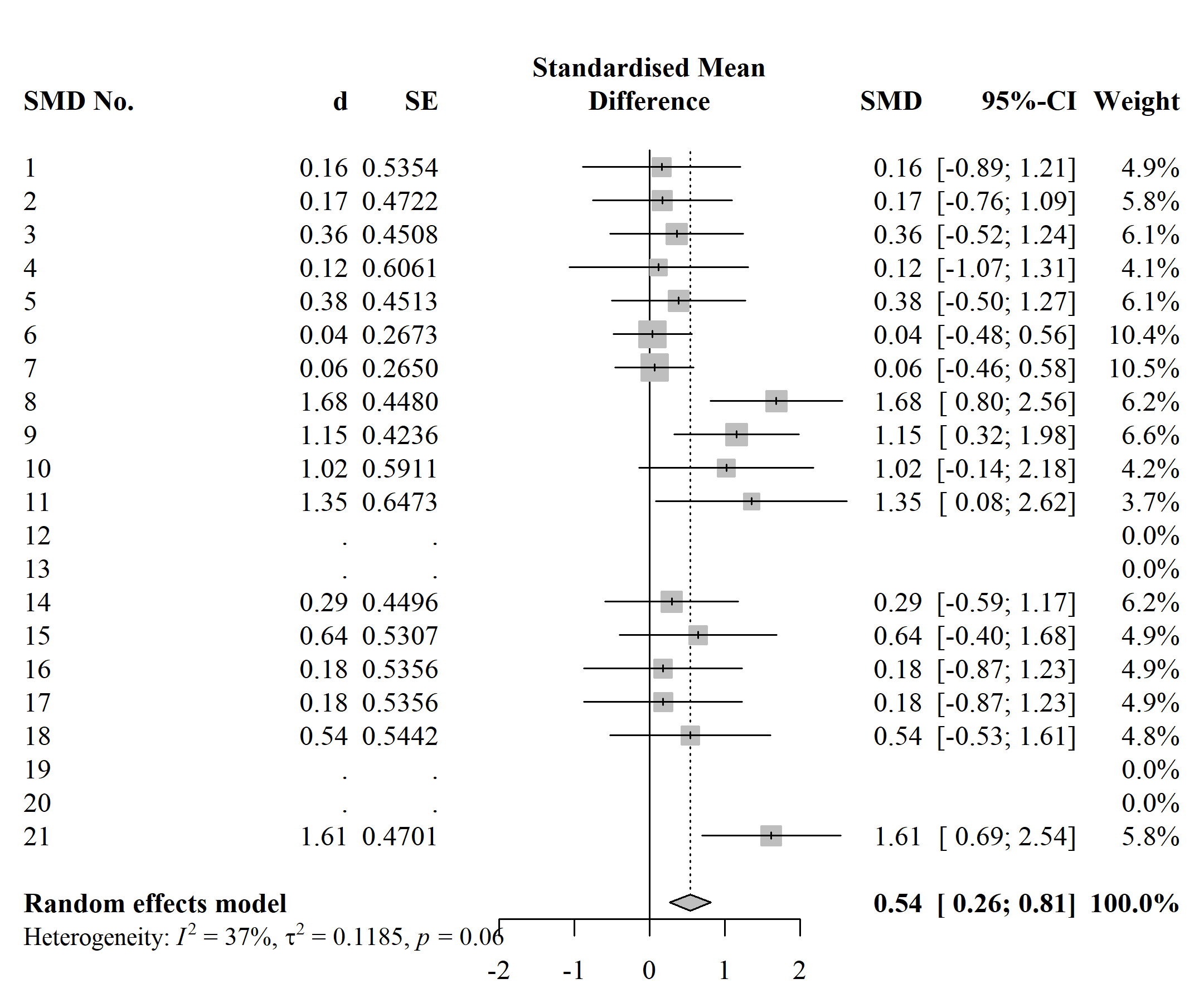
The authors of this meta-analysis reported the following reproducibility-relevant information/data:

* Standardised ES measure is SMD
* Group sample sizes for each primary SMD calculated. In case of crossover studies, they reported the total sample size for both the control and treatment groups.
* All primary SMDs calculated
* CI lower and upper limits for each SMD
* The outcome measure for each SMD (e.g., “Total latency score in JHFT”)
* What the two compared groups represented (e.g., control group: “sham before inpatient daily rehabilitation at retention,” treatment group: “ctDCS on cH before inpatient daily rehabilitation at retention”)
* Random effects model
* The pooled SMD

The following information was missing:

* Type of SMD used (Cohen’s vs. Hedge’s )
* Whether a different method of standardisation was used for the crossover studies (since there are no two groups whose SDs can be pooled to standardise the mean difference)
* Sampling variances of the SMDs (they had to be extracted from the funnel plot)
* Whether sampling variances were calculated differently for the SMDs corresponding to crossover trials
* Which type of values were used to compute each SMD (e.g., means and SDs vs. -value and s)
* Which exact values were used and where they were found (e.g., -value reported on page line or means and SDs reported in Figure )
* Rationale for choosing outcomes
* Enough details about the outcome used so as to leave no room for ambivalence (e.g., “Upper Limb FMA” instead of just “FMA” when the primary study reported both “Upper Limb FMA” and “Total )
* Which between-study heterogeneity estimator was used
* Software used to run the meta-analysis

Table 6 summarises the reproducibility status and classification of each primary SMD. Six SMDs were successfully reproduced following the procedure as described in the meta-analysis or the seemingly standard procedure. Five further SMDs were approximated. Following a deviating procedure, three more SMDs could be reproduced.

Using the SMDs and the sample sizes reported in the meta-analysis along with the sampling variances extracted from the funnel plot (MAM 3), all meta-analytic estimates were successfully reproduced. This MAM was thus treated as the “reported” variety and compared to the two other MAMs. All three MAMs are depicted as forest plots in Figure 3. The pooled SMDs from MAMs 1 and 2 deviate from the reported pooled SMD by 0.05 and 0.11 SDs, respectively. More noteworthy differences are observable in the heterogeneity estimates and . 

### 2.3.2 Meta-analysis 2.

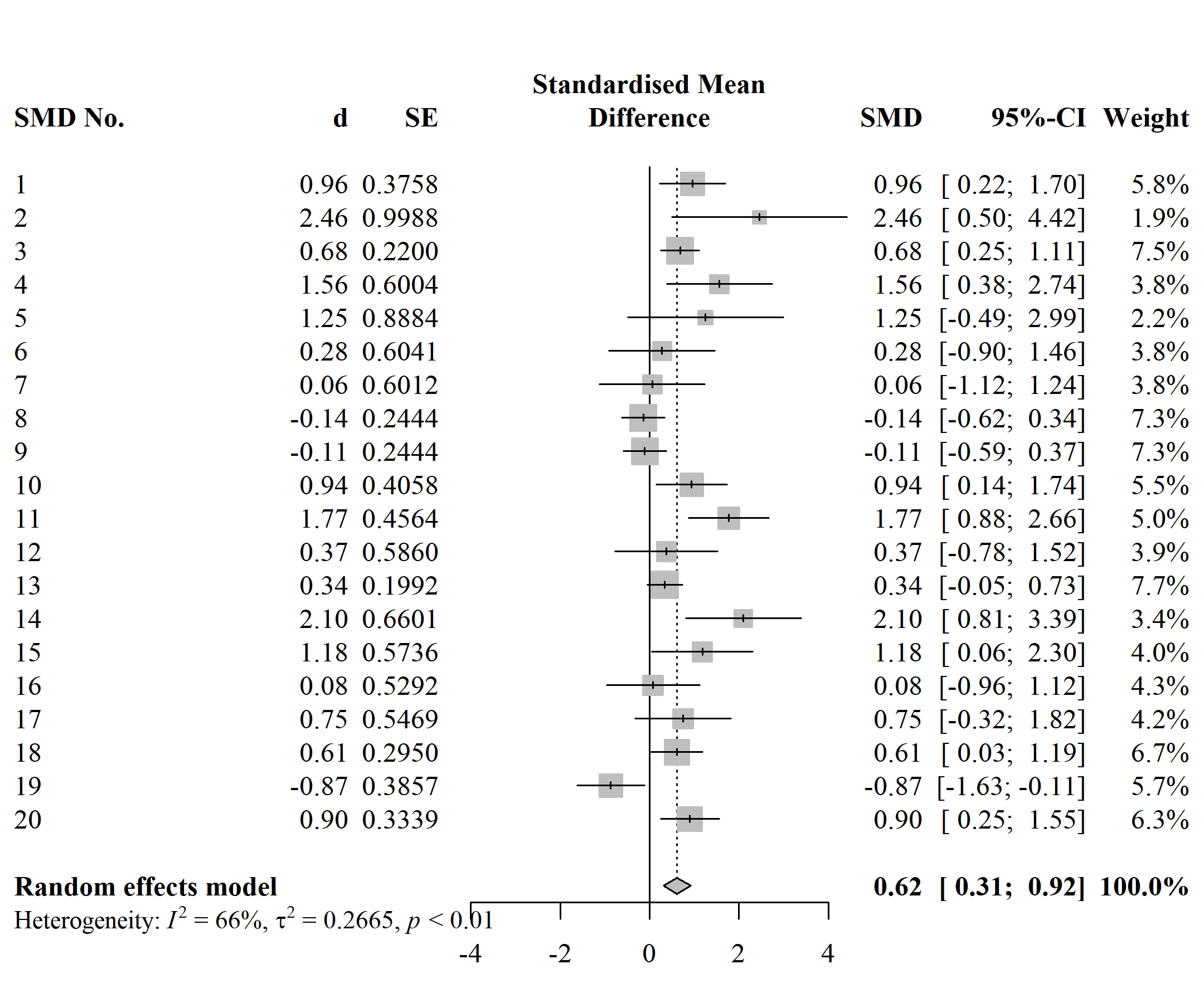
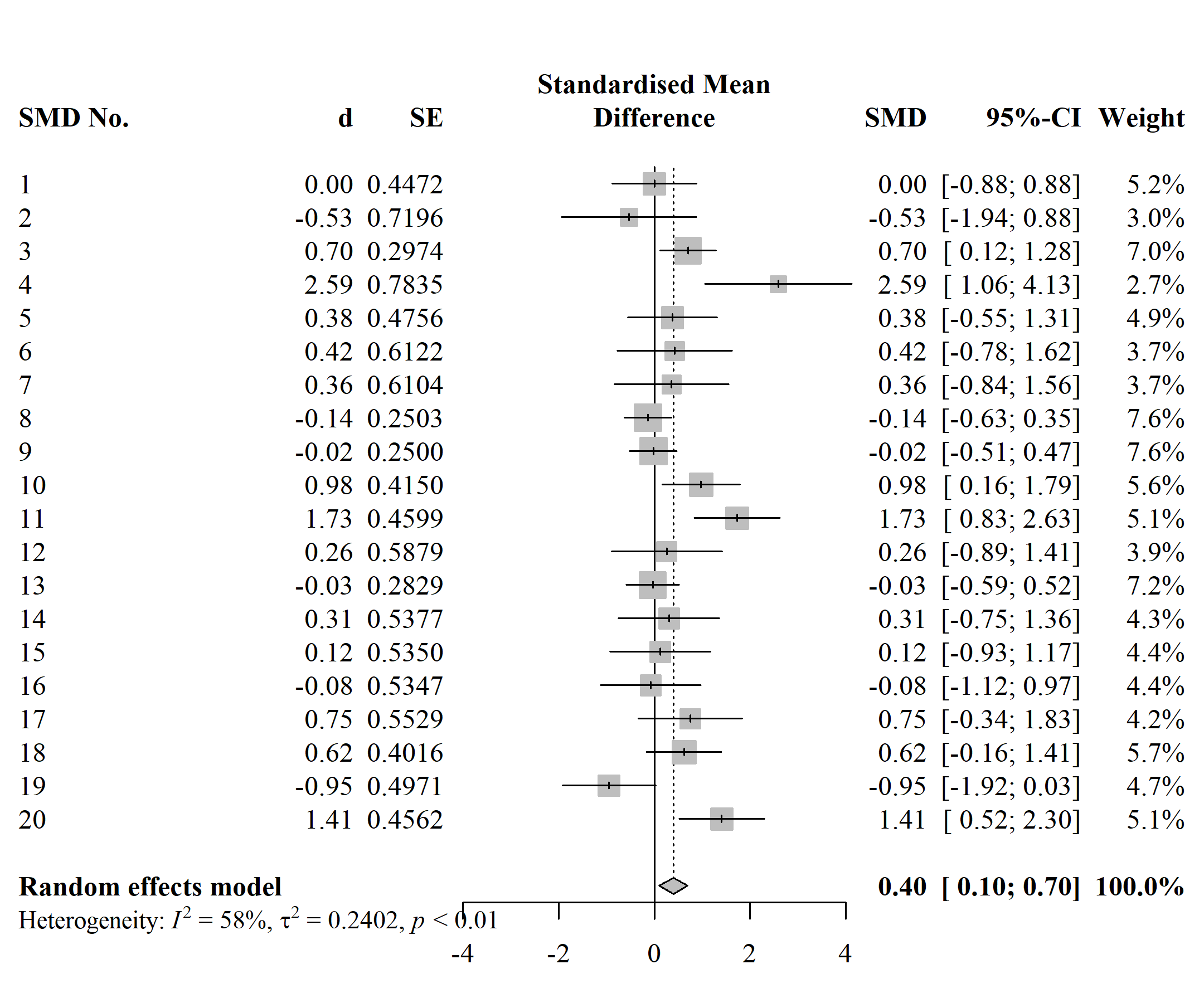
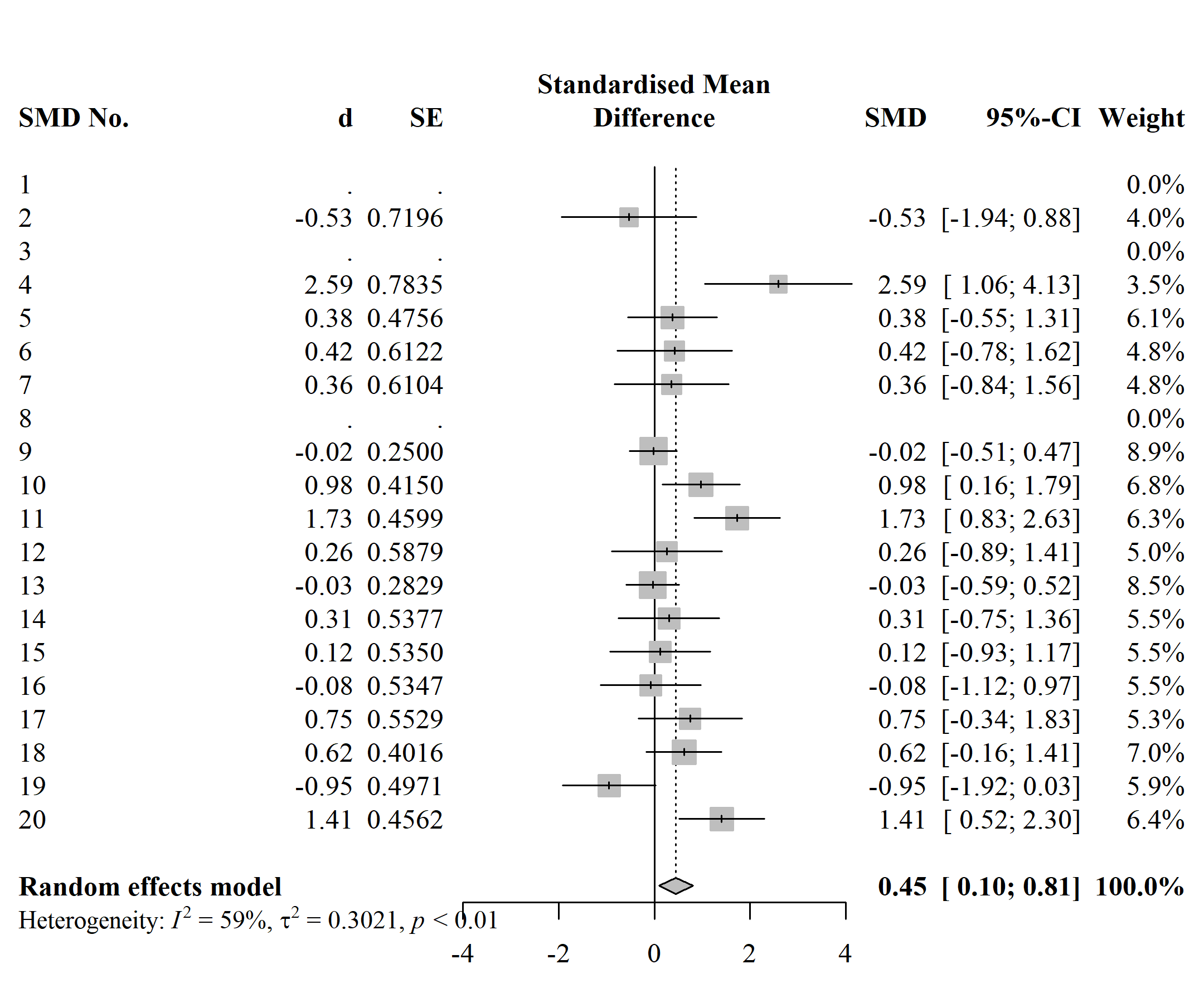
The following reproducibility-relevant information/data were reported:

* Standardised ES measure is SMD
* Group sample sizes (s) for each primary SMD calculated. In case of crossover studies, they reported the total sample size () for both the control and treatment groups.
* All primary SMDs calculated
* CI lower and upper limits for each SMD
* The outcome measure for each SMD
* SMDs were based on comparisons between cathodal tDCS group and sham group at post
* Random effects model
* The pooled SMD
* Software used was Comprehensive Meta-Analysis

The following information was missing:

* Type of SMD used
* Whether a different method of standardisation was used for the crossover studies
* Sampling variances of the SMDs (depicted in the funnel plot)
* Whether sampling variances were calculated differently for the SMDs corresponding to crossover trials
* Which type of values were used to compute each SMD
* Which exact values were used and where they were found
* Rationale for choosing outcomes
* Enough details about the outcome used so as to leave no room for ambivalence
* Which between-study heterogeneity estimator was used

Table 7 summarises the reproducibility status and classification of each primary SMD. One SMD was successfully reproduced following the procedure as described in the meta-analysis or the seemingly standard procedure. Three further SMDs were approximated. Following a deviating procedure, two more SMDs were reproducible and one approximated.

All meta-analytic estimates were successfully reproduced in MAM 3, whose output was again treated as “reported.” The three MAMs are depicted in Figure 4. The pooled SMDs from MAMs 1 and 2 deviate from the reported pooled SMD by 0.17 and 0.22 SDs, respectively. All three models displayed a very high degree of heterogeneity. 

### 2.3.3 Meta-analysis 3.

The following reproducibility-relevant information/data were reported:

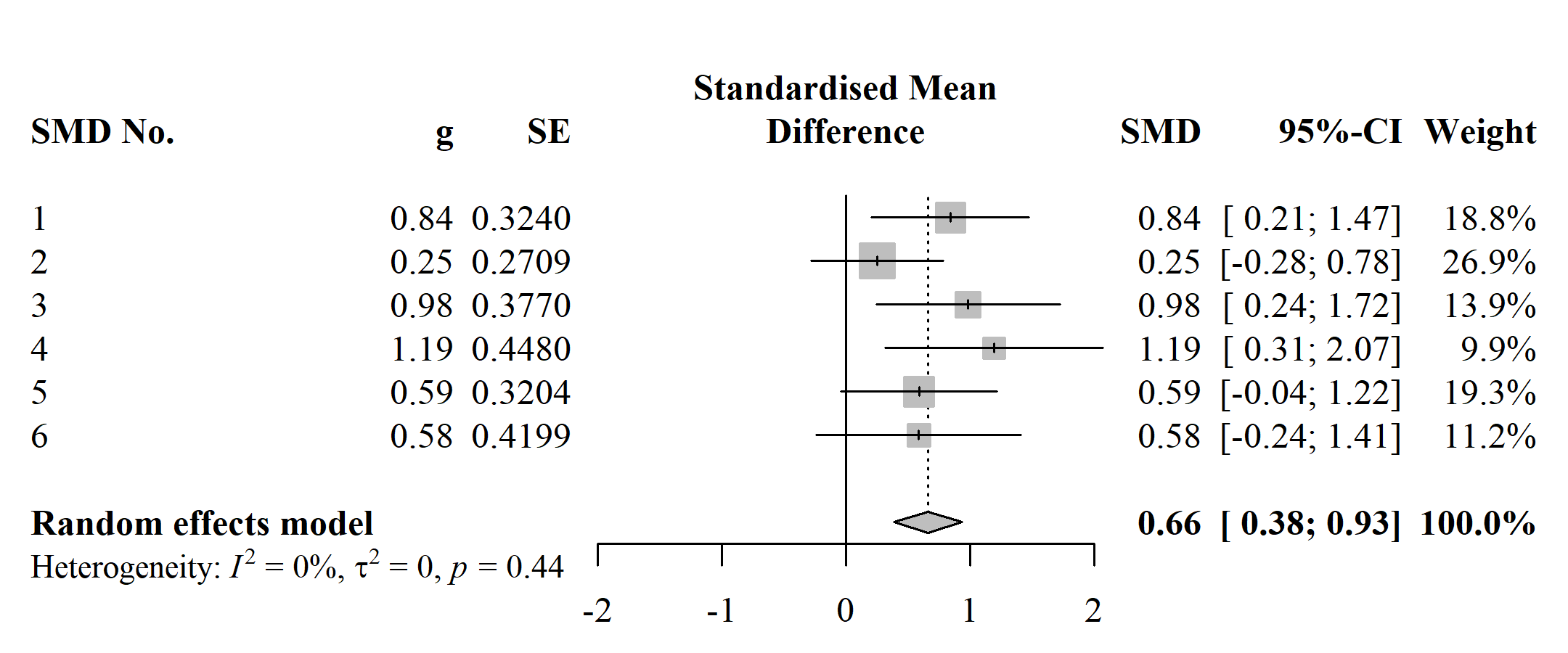
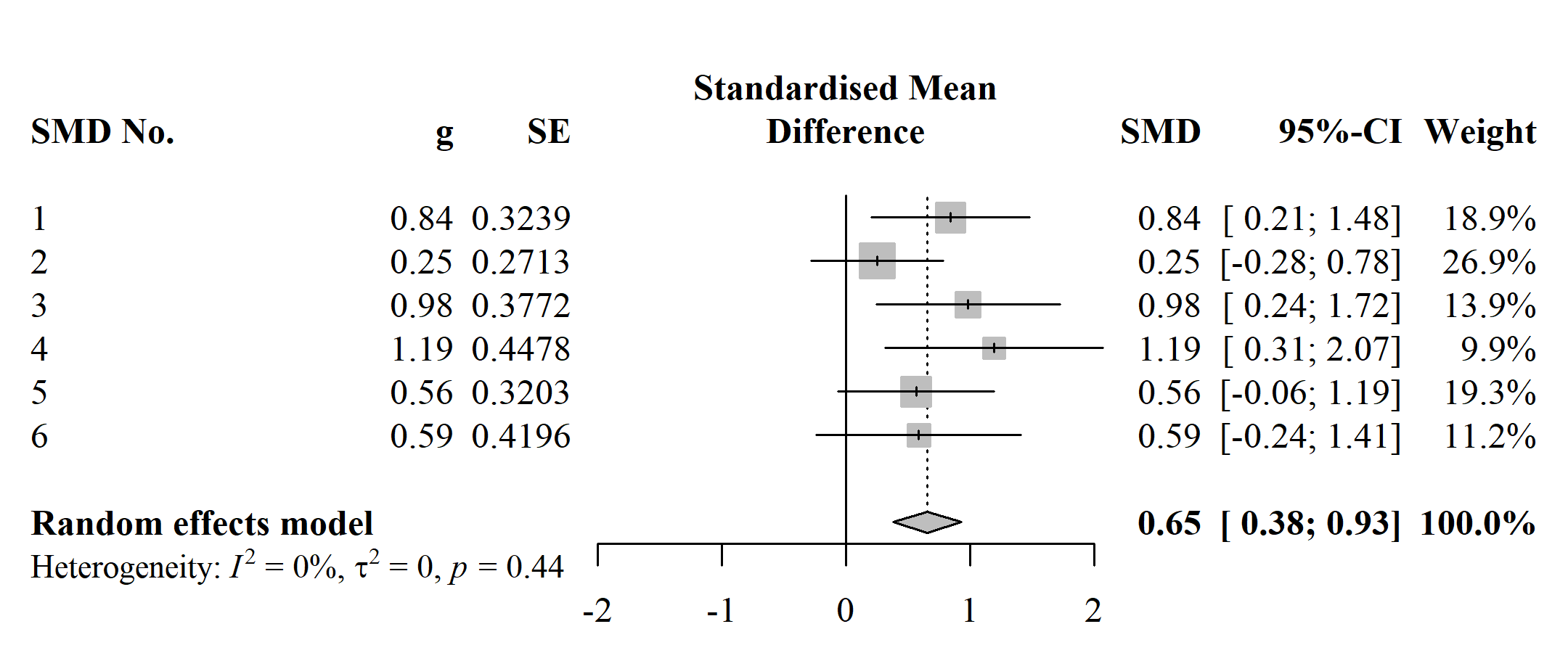
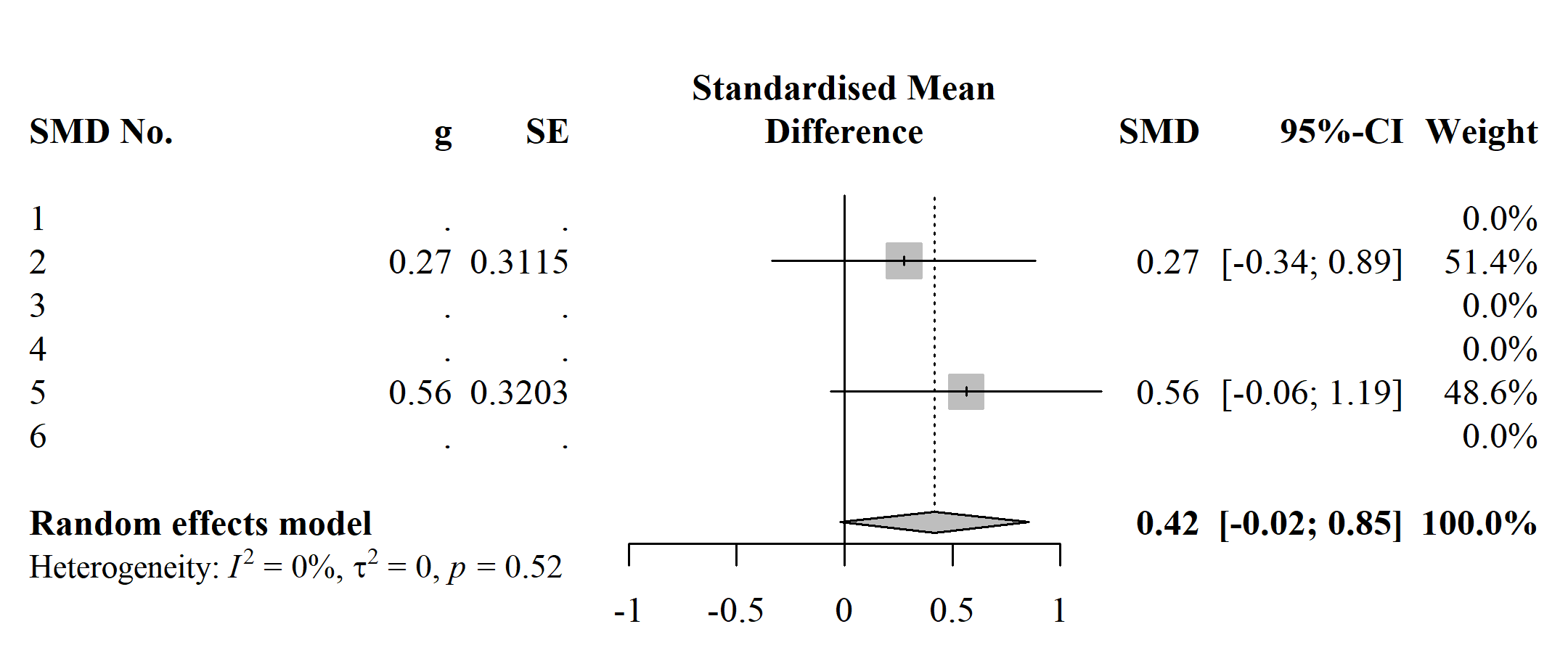
* Standardised ES measure is Hedges’
* Group sample sizes (s) for each primary SMD calculated. In case of crossover studies, they reported the total sample size () for both the control and treatment groups.
* All primary SMDs calculated
* CI lower and upper limits for each SMD
* The outcome measure used for each primary study was what the primary study defined as the primary outcome
* SMDs were based on comparisons between tDCS group and sham group at post
* The task learned by participants (which sometimes aided in finding the outcome used by the meta-analysts)
* The pooled SMD
* Software used was Comprehensive Meta-Analysis

The following information was missing:

* Whether a different method of standardisation was used for the crossover study
* Sampling variances of the SMDs
* Whether sampling variances were calculated differently for the SMD corresponding to the crossover trial
* Which type of values were used to compute each SMD
* Which exact values were used and where they were found
* Enough details about the outcome used so as to leave no room for ambivalence
* Which between-study heterogeneity estimator was used

Table 8 summarises the reproducibility status and classification of each primary SMD. One SMD was approximated following the procedure as described in the meta-analysis or the seemingly standard procedure. Following a deviating procedure, all 5 remaining SMDs could be reproduced or approximated.

Using the SMDs and the sample sizes reported in the meta-analysis along with the sampling variances calculated based on the reported CIs, all meta-analytic estimates were successfully reproduced. The output of this MAM (3) was thus again treated as the reported variety and used for comparison. The three MAMs are depicted in Figure 4. The pooled SMD from MAM 1 deviates from the reported pooled SMD by 0.24 SDs, whereas MAM 2 is almost identical to the MAM 3. Very little heterogeneity was observed in all three cases.



## 2.4 Publication bias control

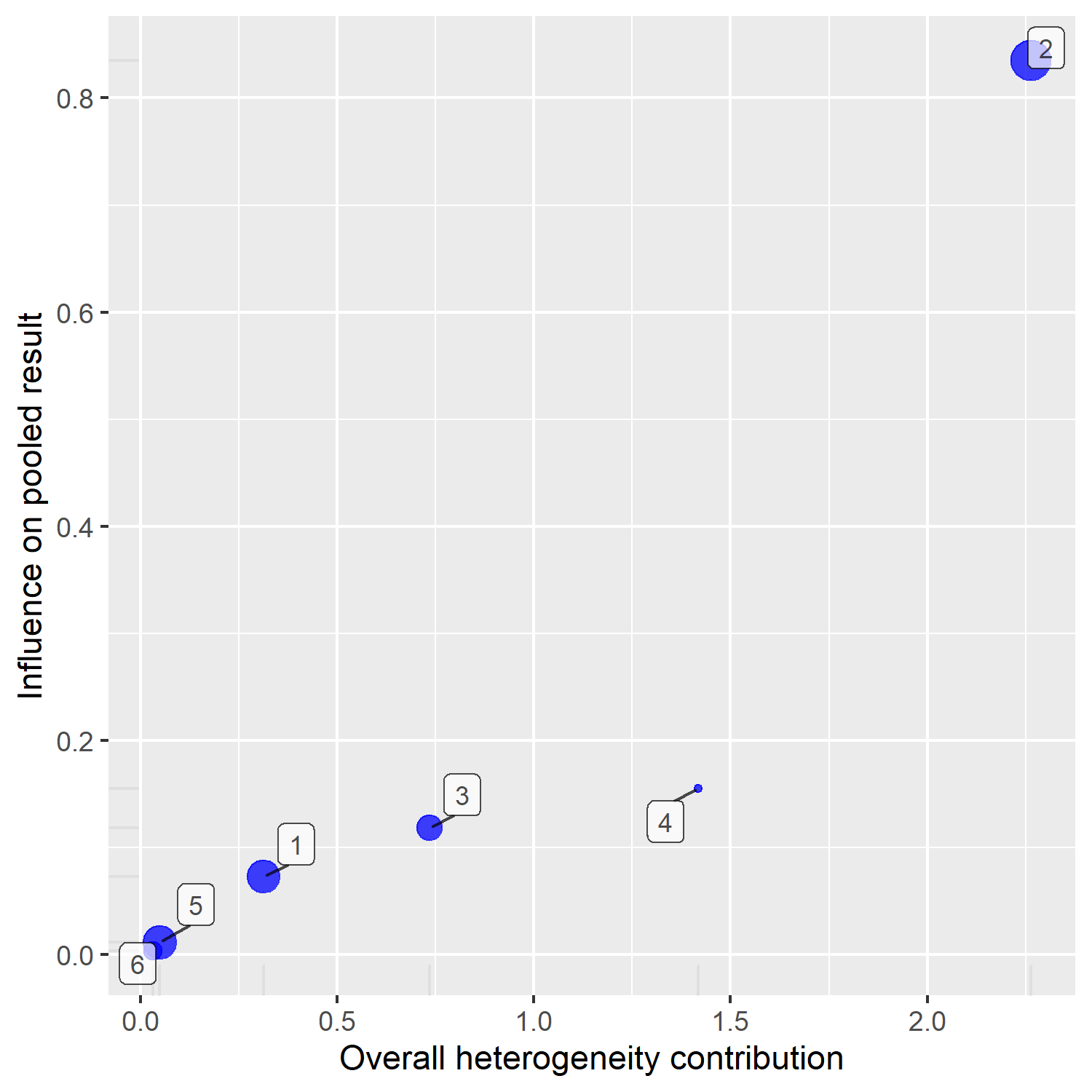
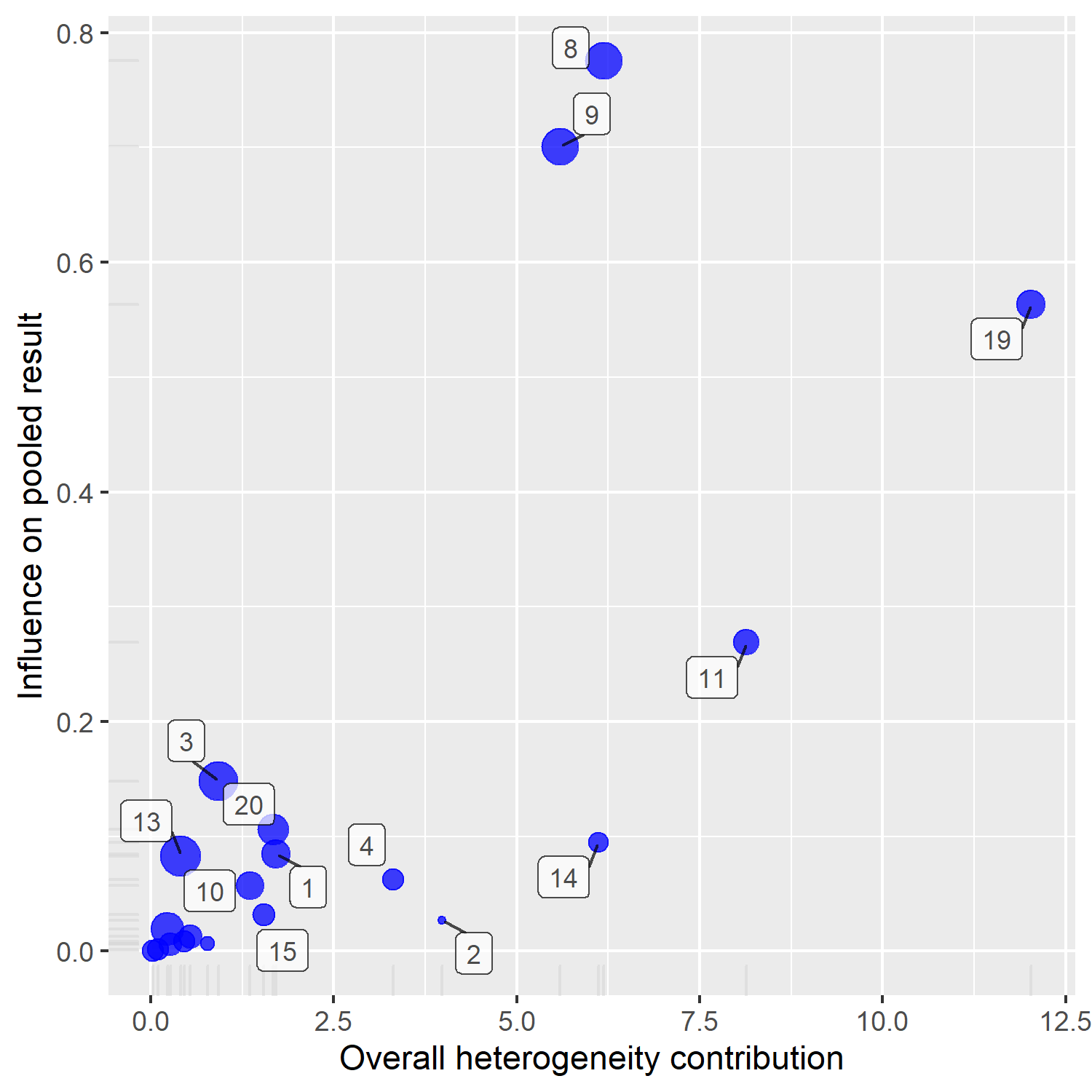
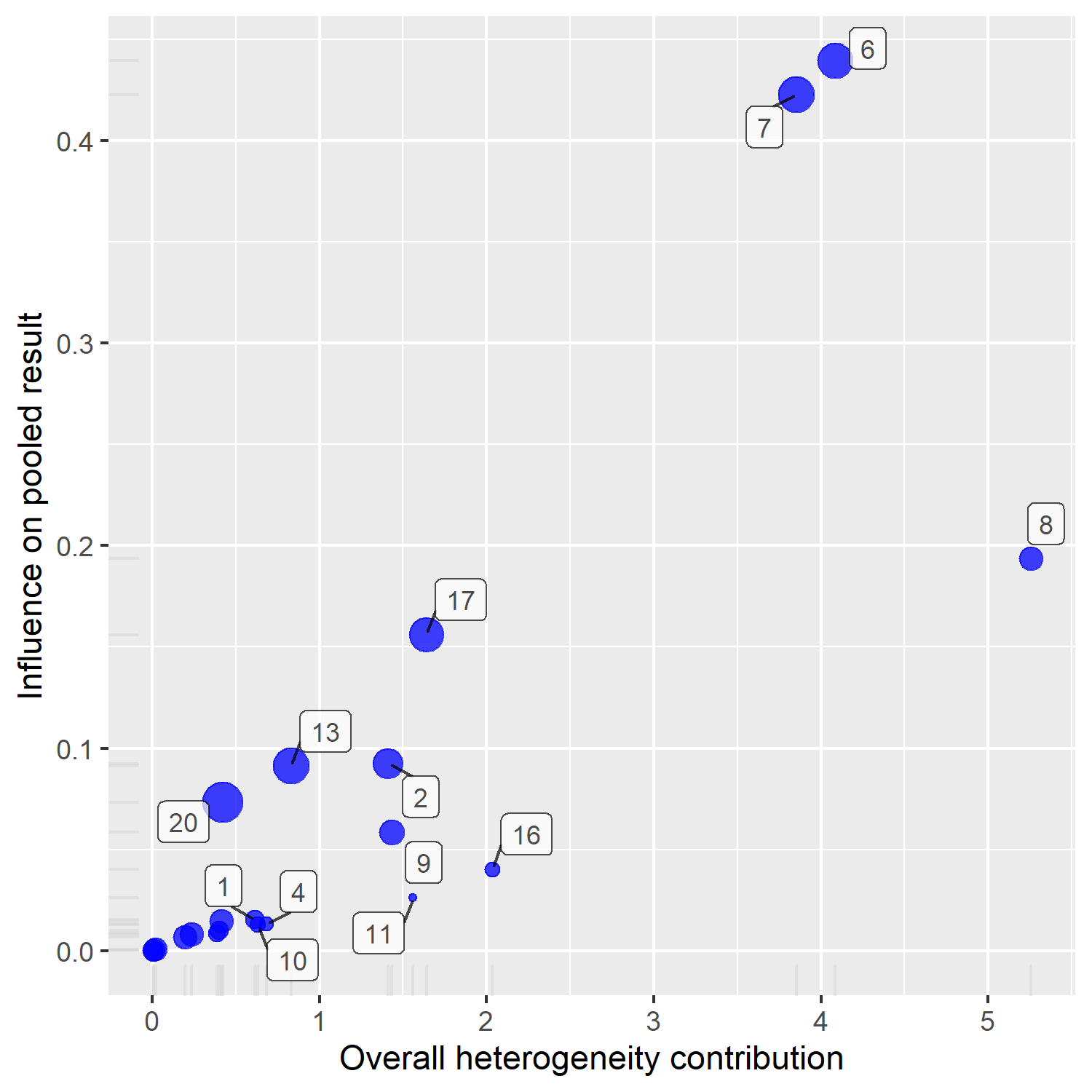
All three meta-analyses mentioned publication bias, although only meta-analysis 3 reported having taken measures to pre-emptively mitigate its effects: the authors report having searched ClinicalTrials.gov and ProQuest (which indexes theses and dissertations), contacted the authors of the primary studies to ask for more data, and not restricted their search to articles written in English.

All three meta-analyses report having inspected a funnel plot as a means to detect publication bias. Meta-analysis 1 addtionally used Fail-Safe . To correct for publication bias, meta-analysis 1 used trim-and-fill; meta-analysis 2 used trim-and-fill, Egger’s test, and Begg and Maxumdar’s rank correlation test (Begg & Mazumdar, 1994); meta-analysis 3 used Egger’s test. The authors of meta-analysis 1 concluded that the findings of the tests they used “support a minor publication bias conclusion” (Kang et al., 2016, p. 348). Similarly, the conclusion in meta-analysis 2 (Kang et al., 2018, p. 5) was “minimal publication bias in the studies used.” No clear conclusion was provided in meta-analysis 3.

Our publication bias testing routine (along with the associated decision rule pre-defined in the data analysis plan) indicated a concurring conclusion in the case of the first meta-analysis and the opposite conclusions for the two other meta-analyses. For meta-analysis 1, the estimate of the true effect produced by the three-parameter selection model (0.59) was virtually identical to the original. The estimate produced by the -curve was larger (0.70). Only the PET-PEESE intercepts (0.34 and 0.45, respectively) indicated that the random effects model based estimates might be overestimating the true effect. For meta-analysis 2, the selection model (0.20), -curve (0.18), and PEESE (0.21) estimates were much smaller than the original (0.62). The PET intercept (-0.12) was negative. Similar results were observed for the last meta-analysis: the estimates produced by the the -curve and PET-PEESE were 0.40, -0.71, and 0.01, respectively. Only the selection model yielded an estimate which is close to the one based on the random effects model (0.62).

### 2.4.1 Outlier/influential study analysis.

Only the third meta-analysis mentioned outliers or influential studies. They ran a leave-one-out analysis and reported that the main results did not change due to removing any one of the 6 studies they included. Our leave-one-out analysis indicated the presence of influential studies in all three meta-analyses (see Figure 6).



# 3 Discussion

The aim of this work was to evaluate the methodological quality of meta-analyses in tDCS-motor learning research with respect to reporting transparency, reproducibility, and publication bias control. We found the meta-analyses we studied to be lacking in all three points. Although the meta-analyses reported the content of most PRISMA (Moher et al., 2009) items, they described their methods so sparsely as to render reproducing their procedure impossible without lengthy detective work. None of the meta-analyses had pre-registration protocols, shared their data, or provided their data-analysis code. We failed to numerically reproduce the main pooled ES estimates reported in all meta-analyses when following the procedures they described. As to publication bias control, only meta-analysis 3 searched the grey literature and all three meta-analyses used “traditional” publication bias detection and correction tools without discussion of their assumptions or appropriateness.

The most notable finding is probably the high prevalence of discrepancies between how the meta-analysts reported having computed individual ESs and how they really did it. These discrepancies were most often in relation to the outcomes used. For example, there were multiple cases where the meta-analysis reported having used an outcome X whereas they actually used the outcome change in X from baseline. This was particularly perplexing when values for both outcomes were reported in the primary study. In general, all primary studies included in the meta-analyses reported values/tests for several outcomes and meta-analysis 3 was the only one to provide a rationale, albeit a vague one[[8]](#footnote-62), for why they chose the outcome they did for each primary study.

Attempting to reproduce the meta-analyses revealed further methodological issues which do not directly pertain to reporting quality: All three meta-analyses indiscriminately combined primary studies of different designs. For example, they computed SMDs using the same formula for both controlled and crossover designs, a procedure which neglects bias resulting from estimating sampling variances for crossover studies without accounting for carry-over effects or correlations between time points (Borenstein & Hedges, 2019; Madeyski & Kitchenham, 2018; Morris & DeShon, 2002)[[9]](#footnote-63). Another issue the meta-analysts neglected to account for is ES dependency (Gleser & Olkin, 2009), which is especially critical in the case of the first two meta-analyses as they derived multiple ESs from single studies.

Furthermore, although there are several ways to calculate an SMD besides the classic formula based on means and SDs (for an overview see Borenstein & Hedges, 2019), it can be argued that the meta-analyses we reviewed used methods that might have been a bit too creative. For example, -values based on median tests were used to calculate SMDs in two of the meta-analyses. Medians can replace means when they are derived from normally distributed data. However, primary studies usually report medians and interquartile ranges instead of means and SDs *because* their samples are not normally distributed (Higgins et al., 2019, p. 167). Besides, one would still need to impute the corresponding SDs. Similarly, Higgins et al. (2019) advise against combining SMDs based on post values and SMDs based on change from baseline as the SDs used to standardise them estimate different entities (p. 253).

Our results indicate that the methodological limitations found in meta-analyses in neighbouring fields also apply to tDCS-motor learning research. Like the reviews cited in Table 1, we found the meta-analyses to be highly under-reported. Although none of our reproductions led to a radical change in the pooled ES estimate, as was the case in Gøtzsche et al.’s (2007) and Ford et al.’s (2010) reviews, we had tremendous difficulties reproducing the analyses and found numerous indications of errors, as did Lakens et al. (2017) and Maassen et al. (2020). Like Banks et al. (2012), we found the meta-analyses we reviewed to mostly use outdated publication bias control methods of dubious validity.

The fact that our findings are unsurprising does not make them any less concerning. Together the three meta-analyses have been cited over 200 times and knowing the high status meta-analyses enjoy on the “hierarchy of evidence” (Evans, 2003), they are likely to be influential beyond the restricted realm of scientific publishing. Therefore, although we contribute no new information with regards to the effectiveness of tDCS for improving motor learning, we hope to have provided an incentive for consumers of research in this highly active field to interpret meta-analytic results with a grain of salt.

Inevitably, our work suffers from several limitations itself:

* We have defined our exclusion criteria based on mostly practical considerations. Their high restrictiveness has probably led to a sample that is not representative of the field at large. Our non-systematic literature search and study selection strategy can only have exasperated this issue. It also goes without saying that three meta-analyses is too small a sample to draw far-reaching conclusions.
* Although we had a mechanism in place to minimise the probability of data extraction errors on our part when evaluating reproducibility, it cannot be excluded that potential errors when extracting data for other variables influenced our results as data extraction and coding was not checked by others (Buscemi, Hartling, Vandermeer, Tjosvold, & Klassen, 2006; Jones, Remmington, Williamson, Ashby, & Smyth, 2005).
* Our treatment of the PRISMA guidelines was rather superficial in the sense that we made a binary decision whether the content of each item was reported in the meta-analysis or not. A much more nuanced evaluation is possible.
* We focused on reproducibility of data extraction and calculation of effect sizes because previous reviews demonstrated a high prevalence of mistakes in these steps. However, it cannot be excluded that other aspects we neglected such as study search and selection have a larger impact.
* Throughout the process of reproduction, we had to make subjective decisions that cannot be guaranteed to have been faultless. For example, it was not always trivial to judge whether our procedure for reproducing a certain primary SMD strictly followed the procedure purported to have been used by the meta-analysts.
* As mentioned before, publication bias testing remains an exceedingly active and contentious field of research and no consensus exists as of yet regarding which tests to use under which conditions. Any single test or constellation of tests used is thus inevitably arbitrary. One major problem all current methods suffer from is their mediocre performance when applied to a heterogeneous set of ESs. Two of the meta-analyses we reviewed synthesised such sets of ESs. Furthermore, the primary studies included in the meta-analyses are likely to have been selected based on different values than the ones the meta-analysts (and we) used for the publication bias tests (i.e., results of omnibus tests on several outcomes and not the pair-wise comparisons in one outcome at a single time point). This is relevant for the -curve and the selection model we used of which an important step is dividing the set of ESs into significant and non-significant ones.
* Due to our rather limited expertise with regards to the technical and practical aspects of tDCS research, it cannot be excluded that we have overlooked or misunderstood relevant tDCS-related methodological facets.

Future similar works may thus aim for a more fine grained and nuanced evaluation of reporting transparency which goes beyond the minimal requirements set by reporting guidelines, a more comprehensive reproducibility testing which is not restricted to data extraction and ES calculation as well as a thorough investigation of robustness towards changes in analytical decisions (especially data selection and outcomes used), and a more principled approach towards evaluating publication bias assessment.

Despite these limitations, we believe that our work provides further evidence that meta-analysis as a quantitative method of research synthesis is not as objective as one would like to believe. Countless researcher degrees of freedom (Wanous, Sullivan, & Malinak, 1989), combined with confirmation bias (Goodyear-Smith, van Driel, Arroll, & Del Mar, 2012) and perverse incentives fueled by a relentless race to acquire funding and promotion (Edwards & Roy, 2017) give rise to a cluster of “redundant, misleading, and conflicted” research products (Ioannidis, 2016, title).

Nevertheless, research synthesis is a vital part of the scientific endeavour and efforts to counteract the above problems should be promoted. Besides the guidelines which are constantly being updated and advancements in statistical procedures for controlling publication bias, many methodological procedures have been developed in the recent years to systematically account for the impact of different analytical decisions (Taylor & Munafò, 2016; e.g., Voracek et al., 2019). More emphasis is also being placed on flagship open science practices such as pre-registration and data and code sharing (Maassen et al., 2020; Page, Moher, et al., 2021).

However, to draw on the old “garbage in garbage out” principle (Borenstein et al., 2009b), it must be acknowledged that no amount of refinement in research synthesis methods will come very far if primary researchers do not step up their game in the quest for a healthy, reliable cumulative science. For tDCS research, efforts could include employing both mathematical (Lipka et al., 2021) and practical (e.g., a reporting check-list Buch et al., 2017) procedures with the aim of producing more reproducible and less heterogeneous findings.

More generally, in an ideal world, meta-analysis as a method for aggregating summary results from individual primary studies would be redundant. Meta-analysis of individual participant data (Riley, Lambert, & Abo-Zaid, 2010), more recently also referred to by the catchier term “mega-analysis” (Eisenhauer, 2021), would be the predominant quantitative research synthesis method as all primary studies would be sharing their data. Unfortunately, although on the rise (Christensen et al., 2020), open science practices remain relatively rare (Hardwicke et al., 2021; Hardwicke et al., 2020). Therefore, until we reach such a “scientific utopia” (Nosek, Spies, & Motyl, 2012), we should strive for the highest standards when conducting and evaluating opi of research synthesis.

# 4 References

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1. One review of meta-analyses in the organisational sciences found that 21 different analytical decisions have a negligible impact on the pooled effect size estimate (Aguinis, Dalton, Bosco, Pierce, & Dalton, 2011) [↑](#footnote-ref-20)
2. Three further relevant reviews in the biomedical sciences were identified after the study was done and could not have influenced our approach: two evaluating the methodological reproducibility of meta-analyses (Page et al., 2018; i.e., without attempting to actually reproduce them, Wayant, Page, & Vassar, 2019) and one testing the reproducibility of entire meta-analyses (i.e., including study search and screening, Ford, Guyatt, Talley, & Moayyedi, 2010). These reviews, too, reached the conclusion that meta-analysis reproducibility is mostly limited. [↑](#footnote-ref-21)
3. “trying multiple analyses to obtain statistical significance” (Simonsohn et al., 2014a, p. 534) [↑](#footnote-ref-23)
4. Vevea being a developer of two selection model varieties (Vevea & Hedges, 1995; Vevea & Woods, 2005) [↑](#footnote-ref-24)
5. The developers of the trim-and-fill method, which is widely regarded as being of little use for correcting bias, although potentially useful as a sensitivity analysis (Carter et al., 2019; Hilgard, Engelhardt, & Rouder, 2017; Moreno et al., 2009b; Simonsohn et al., 2014b; van Assen et al., 2015). [↑](#footnote-ref-25)
6. An ES estimate was considered successfully reproduced (or reproducible) if the reproduced ES equalled the one reported in the meta-analysis at the second decimal (e.g., 0.3324543 = 0.33) and approximated if the reproduced ES equalled the reported one at the first decimal (e.g., 0.1767492 0.22). [↑](#footnote-ref-33)
7. Sampling variances were not reported in any of the three meta-analyses. They were calculated based on SEs extracted from funnel plots in the case of the first two meta-analyses and based on CIs for the third meta-analysis [↑](#footnote-ref-34)
8. Namely that they used the outcome the respective primary study defined as their primary outcome. [↑](#footnote-ref-62)
9. Borenstein and Hedges (2019) go so far as to declare combining SMDs standardised by different types of SDs (e.g., between vs. within groups) downright verboten. [↑](#footnote-ref-63)