

$D_c$  = the 2-out-of-3 system is functioning

$$= (A_1 \cap A_2) \cup (A_1 \cap A_3) \cup (A_2 \cap A_3)$$

## Axioms and Implications

[0,1]

**Definition.** A probability  $\mathbb{P}$  is a rule (or function) which assigns a number to each event, and which satisfies the following **axioms** (or properties):

- Axiom 1:  $\mathbb{P}(A) \geq 0$ .
- Axiom 2:  $\mathbb{P}(\Omega) = 1$ . one of the possible outcomes occur
- Axiom 3: **Sum Rule:** For any disjoint  $A_1, A_2, \dots$

$$\mathbb{P}(\cup_j A_j) = \sum_j \mathbb{P}(A_j).$$

Some consequences of the axioms:

- Consequence 1: If  $A \subseteq B$  then  $\mathbb{P}(A) \leq \mathbb{P}(B)$



- Consequence 2:  $\mathbb{P}(\emptyset) = 0$   
like 3: Take  $A = \Omega$ ,  $\Omega^c = \emptyset$

- Consequence 3:  $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ .

see below

- Consequence 4:  $0 \leq \mathbb{P}(A) \leq 1$ .

like 1:  $\emptyset \subseteq A \subseteq \Omega$

- Consequence 5:  $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$ . \*

If  $A$  and  $B$  are disjoint, this reduces to Axiom 3.

Proving these involves cleverly using the axioms, above.

**Example:** Prove Consequence 3.

$$\Omega = A \cup A^c$$

$$1 = \mathbb{P}(\Omega) \quad (\text{Axiom 2})$$

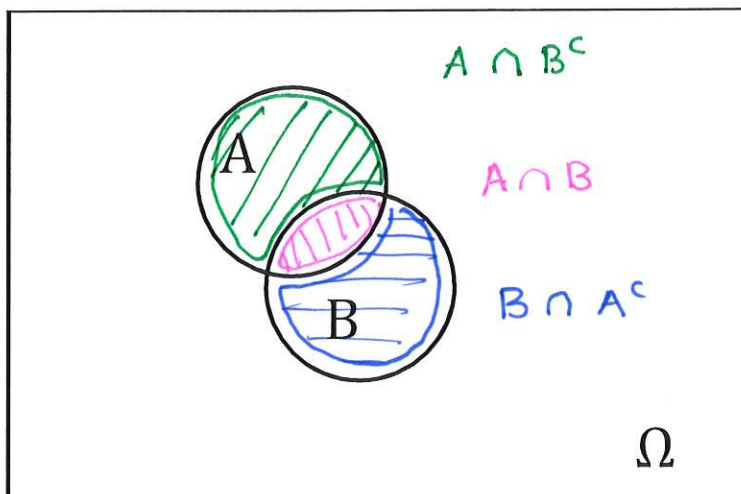
$$= \mathbb{P}(A \cup A^c)$$

$$= \mathbb{P}(A) + \mathbb{P}(A^c) \quad (\text{Axiom 3}).$$

Now rearrange this equation.  $\rightarrow \mathbb{P}(A^c) = 1 - \mathbb{P}(A)$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

**Question:** Visually, what does Consequence 5 mean?



$$P(A) = P(A \cap B^c) + P(A \cap B)$$

$$P(B) = P(B \cap A^c) + P(A \cap B)$$

Note that these simple rules of probability are highly similar to those one would use to measure length, area, volume, weight, etc.

**Example:** Consider a fair, two-sided coin, with sample space  $\Omega = \{H, T\}$  (here, the word *fair* means  $P(\{H\}) = P(\{T\}) = 1/2$ ). Write out all possible events and verify that the third probability axiom is satisfied in the situation  $A_1 = \{H\}$ ,  $A_2 = \{T\}$ .

The events are  $\emptyset, \{H\}, \{T\}, \{H, T\}$

$$P(A_1 \cup A_2) = P(\Omega) = 1$$

and since  $A_1$  and  $A_2$  are disjoint

$$P(A_1) + P(A_2) = \frac{1}{2} + \frac{1}{2} = 1$$

**Question:** In the above example,  $P(\Omega) = 1$  and  $P(A_1) = 1/2$ . Is it true that

$$P(\Omega \cup A_1) = P(\Omega) + P(A_1) = 1 + \frac{1}{2}?$$

Why/why not?

This is **FALSE** because  $A_1$  and  $\Omega$  are not disjoint

In fact,

$$\Omega \cup A_1 = \Omega \quad P(\Omega \cup A_1) = P(\Omega) = 1.$$

## Discrete Sample Spaces

If  $\Omega$  is *countable* it is called a **discrete** sample space; otherwise it is called a **continuous** sample space.

Let  $\Omega$  be a discrete sample space, e.g.  $\Omega = \{a_1, a_2, \dots, a_n\}$ .

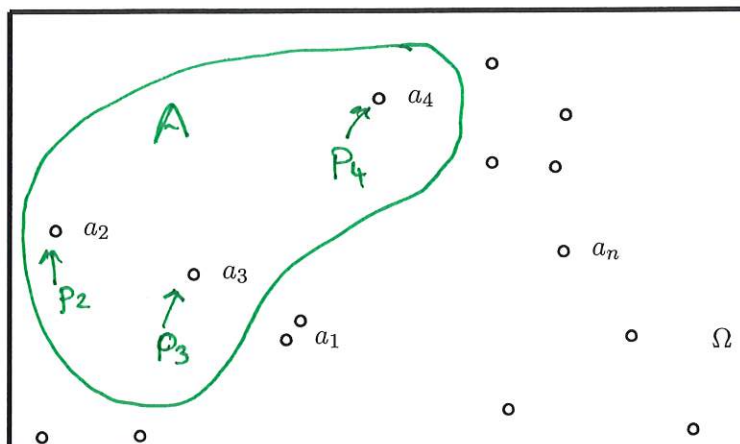


Figure 2.1: A finite sample space  $\Omega = \{a_1, a_2, \dots, a_n\}$ .

Let  $\mathbb{P}(\{a_i\}) = p_i$ , for  $i = 1, \dots, n$ , and define

$$\mathbb{P}(A) = \sum_{i: a_i \in A} p_i, \text{ for all } A \subset \Omega.$$

Then  $\mathbb{P}$  is a probability measure.

Thus we can *specify*  $\mathbb{P}$  by specifying only the probabilities of the elementary events  $\{a_i\}$ .

**Example.** Experiment: throw a fair die.

Sample space:

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

Define  $\mathbb{P}$  by:

$$\text{for each } i=1, \dots, 6 \quad p_i = 1/6 \quad \mathbb{P}(A) = \frac{\# \text{ elements in } A}{6} = \frac{|A|}{6} \quad \leftarrow \text{cardinality of } A.$$

This completely specifies/models the experiment. For example, the probability of getting an even number is

$$\mathbb{P}(\{2, 4, 6\}) = \frac{3}{6} = \frac{1}{2}$$

**Example.** We draw two cards from a full deck of 52 cards. What is the probability of drawing at least one Ace?

Give all the cards a number from 1 to 52. Draw the cards one-by-one. Write a possible outcome as  $(x, y)$ ,  $x \neq y$ .

Each elementary event  $\{(x, y)\}$  has the same probability

four suits, each suit has 13 cards  
2, ..., 10, J, Q, K, A

1st card is any one of 52	$\frac{1}{52 \times 51}$	2nd card is any one of remaining 51
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Let  $A$  be the event: "at least one Ace". Then,

$\mathbb{P}(A) = \sum_{(x,y) \in A} \frac{1}{52 \times 51} = \frac{ A }{52 \times 51}$
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We need to *count* how many elements are in  $A$ . Easier:  $|A^c| = 48 \times 47$ . Hence,

$A^c$  neither card  
is an ace

$\mathbb{P}(A) = 1 - \mathbb{P}(A^c) = 1 - \frac{ A^c }{52 \times 51} = 1 - \frac{48 \times 47}{52 \times 51} \approx 0.15$
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**Remark.** In many cases, as above, we can choose  $\Omega$  such that each elementary event is equally likely, i.e.  $\mathbb{P}(\{a_i\}) = 1/n$ .

This is sometimes known as the *Equally-Likely Principle*, or the *Equilikely Principle*.

**Question:** What if we choose the two cards *at the same time* (no order). Does that change the model? Does it change the probability?

sample space changes

**Example:** Consider tossing a fair coin twice (so that you sample from  $\{H, T\}$  with replacement). What is the probability of getting both one heads and one tails?

- Model I: Order recorded:

(H, H)

$$\Omega = \{HH, HT, TH, TT\}$$

$$A = \{HT, TH\}$$

$$\mathbb{P}(A) = \frac{|A|}{|\Omega|} = \frac{2}{2^2} = \frac{2}{4} = \frac{1}{2}$$

- Model II: Order ignored:

$\frac{1}{4}$        $\frac{1}{2}$        $\frac{1}{4}$

$$\tilde{\Omega} = \{\{H, H\}, \{H, T\}, \{T, T\}\}$$

$$\tilde{A} = \{\{H, T\}\}$$

$$\tilde{\mathbb{P}}(\tilde{A}) = \frac{1}{2}$$

is associated with  
HT and TH

**Question:** Did we apply the Equally-Likely Principle in Model I of the above example? **YES**  
Did we apply it in Model II? **NO**

**Example:** Suppose that we have three balls (labelled 1, 2, and 3) in an urn, and select (*without replacement*) two of the balls. Consider the event of 1 being chosen as one of the two balls.



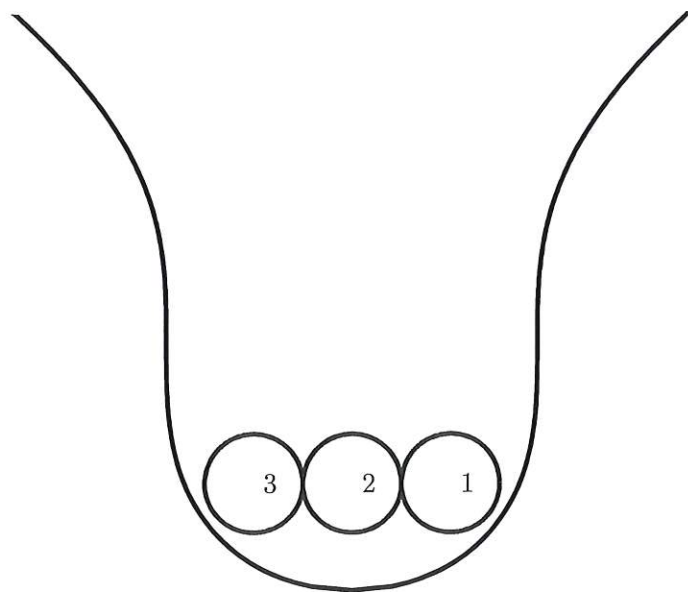


Figure 2.2: Balls 1, 2, and 3 in an urn.

- Model I: Order recorded:

$$\Omega = \{(1,2), (1,3), (2,1), (2,3), (3,1), (3,2)\}$$

Ball 1 is chosen  $A = \{(1,2), (1,3), (2,1), (3,1)\}$

$$\mathbb{P}(A) = \frac{|A|}{|\Omega|} = \frac{4}{6} = \frac{2}{3}$$

- Model II: Order ignored:

$$\tilde{\Omega} = \{\{1,2\}, \{1,3\}, \{2,3\}\}$$

Ball 1 chosen  $\tilde{A} = \{\{1,2\}, \{1,3\}\}$

$$\tilde{\mathbb{P}}(\tilde{A}) = \frac{|\tilde{A}|}{|\tilde{\Omega}|} = \frac{2}{3}$$

each outcome in  $\tilde{\Omega}$  is associated with two outcomes in  $\Omega$   
 Equally likely principle still holds.

## Counting

In the previous examples we drew two cards from a full deck, and balls from an urn, without replacement. In general, if choosing  $r$  objects from a collection of  $n$  objects, without replacement, then there are

$$n \times (n-1) \times \cdots \times (n-r+1)$$

ways of doing this. We write  ${}^n P_r$  for this quantity. (Check your calculator!) Another way to write this ( $n$  permutation  $r$ ) is

$${}^n P_r = \frac{n!}{(n-r)!},$$

where  $n!$  ("n factorial") is given by

$$n! = n \times (n-1) \times \cdots \times 2 \times 1.$$

Notice that the *order* of the letters matters.

**Example.** Two letters are chosen with replacement from the word *PING*. What is the sample space? How many outcomes are there?

$$\Omega = \{PP, PI, PN, PG, IP, II, IN, IG, NP, NI, NN, NG, GP, GI, GN, GG\}$$

$$|\Omega| = 4 \times 4 = 16$$

Notice that the *order* of the letters matters.

**Example.** How many three-letter words can be obtained from the letters in the word *SURFING*? This is sampling without replacement. Does the order matter?

$$|\Omega| = 7 \times 6 \times 5 = \frac{7!}{(7-3)!} = 210 \quad {}^7P_3$$

**Example.** Three people from the set {Joffrey, Balon, Robb, Stannis, Renly} must be chosen to go into a fighting pit. How many different combinations of combatants are there?

Suppose the order is important

$$|\Omega| = {}^5P_3 = 5 \times 4 \times 3 = 60$$

Some outcomes are equivalent though, so this is too many! Our event is a union of several outcomes.

Considering the opponents {Robb, Balon, Stannis}. This combination has been counted  ${}^3P_3 = 6$  times:

$$RBS, RSB, BRS, BSR, SRB, SBR$$

If  $N$  is the number of ways of choosing opponents, then  $N \times {}^3P_3 = {}^5P_3$ , so

$$N = {}^5P_3 / {}^3P_3 = \frac{5!}{(5-3)!3!} = 10 = {}^5C_3$$

In general, if choosing  $r$  objects from a collection of  $n$  objects, without replacement, then the number of combinations is:

$${}^nC_r = \frac{n!}{(n-r)!r!} = \binom{n}{r} \leftarrow \text{notation we will use.}$$

### Summary.

	Order matters	Order does not matter
With replacement	$n^r$	$\binom{n+r-1}{r}$
Without replacement	${}^nP_r = \frac{n!}{(n-r)!}$	${}^nC_r = \binom{n}{r} = \frac{n!}{(n-r)!r!}$

### Matlab code.

```

1 n = 6;
2 r = 2;
3 factorial(n)
4 ans =
5     720
6 nchoosek(n,r)
7 ans =
8     15
9 nPr = nchoosek(n,r)*factorial(r)
10 nPr =
11     30

```

**Example.** How many ways are there to order the letters in the word *INDOOROOPILLY*?

Notice that this situation does not fall into any of the above categories. In general, the number of permutations of  $n$  objects with  $n_1$  of type 1,  $n_2$  of type 2, et cetera, is given by

$$\frac{n!}{n_1!n_2!\dots n_k!},$$

where  $k$  is the number of types.

Hence, the number of ways to order the letters in *INDOOROOPILLY* is

$$13P13/(4P4*2P2*2P2) = 21621600$$

**Example.** Suppose that you have two red balls and three blue balls. How many *distinct* orderings of all balls are there?

$$4C1+4C2=10$$