

## Solution for CDA2015

- **Solution to 3.6.**

Living arrangement	Stage of breast cancer		Total
	Local	Advanced	
Alone	59 (72) (-2.638731)	85 (72) (2.6387308)	144 32.58%
With spouse	109 (104.5) (0.8574376)	100 (104.5) (-0.857438)	209 47.29%
With others	53 (44.5) (2.0164045)	36 (44.5) (-2.016405)	89 20.14%
Total	221 50%	221 50%	442

From SAS output,  $X^2 \approx 8.3292$ , which is approximated by  $\chi^2_{df}$  distribution with  $df = 2$ , and  $p$ -value = 0.0155; Also  $G^2 \approx 8.3752$ , approximated by  $\chi^2_{df}$  distribution with  $df = 2$ , and  $p$ -value = 0.0152.

- **Solution to 3.18.**

Drug	Nervousness		Total
	Yes	No	
Loratadine	4 (2.4258) (1.2186962)	184 (185.57) (-1.218715)	188 30.57%
Placebo	2 (3.3806) (-0.994638)	260 (258.62) (0.9946141)	262 42.26%
Chlorpheniramine	2 (2.1935) (-0.154409)	168 (167.81) (0.1543916)	170 27.42%
Total	8 1.29%	612 98.71%	620

(a). No.  $X^2 = 1.6234 \sim \chi^2(2)$  and  $p = 0.4441$ ;  $G^2 = 1.5528 \sim \chi^2(2)$  and  $p = 0.4600$ .

Alternative method. Since the number of observations in some cells are smaller than 5, we can also use Fisher exact test. By SAS output, the  $p$ -value for Fisher exact test is about 0.395, so that we can not reject the independence assumption, i.e. no inferential evidence that nervousness depends on drug.

Note: both methods are correct!

(b). (i)  $\hat{\theta} = \frac{n_{11}n_{22}}{n_{12}n_{21}} \approx 2.826$ ,  $\hat{\sigma}(\log \hat{\theta}) = \left( \frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}} \right)^{1/2} \approx 0.871$ . For  $\alpha = 0.05$ ,  $\text{CI}_\alpha$  for  $\log \theta$  is:

$$[\log(\hat{\theta}) - z_{\alpha/2}\hat{\sigma}(\log(\hat{\theta})), \log(\hat{\theta}) + z_{\alpha/2}\hat{\sigma}(\log(\hat{\theta}))] = [-0.668, 2.745]$$

and hence  $\text{CI}_\alpha$  for  $\theta$  is:

$$[e^{-0.668}, e^{2.745}] = [0.513, 15.565].$$

(ii)  $\hat{\pi}_1 = 4/188$ ,  $\hat{\pi}_2 = 2/262$ ,  $\hat{\pi}_1 - \hat{\pi}_2 \approx 0.01364$ ,  $\hat{\sigma}(\hat{\pi}_1 - \hat{\pi}_2) = \left( \frac{\hat{\pi}_1(1-\hat{\pi}_1)}{n_1} + \frac{\hat{\pi}_2(1-\hat{\pi}_2)}{n_2} \right)^{1/2} \approx 0.0118$ . So  $\text{CI}_\alpha$  for  $\pi_1 - \pi_2$  is:  $(-0.0094, 0.0368)$ .

- **Solution to 3.26.** Recall that, as  $n \rightarrow \infty$ ,  $\sqrt{n}(\hat{\pi}_i - \pi_i) \xrightarrow{d} N(0, \pi_i(1 - \pi_i))$  and  $\text{Cov}(\hat{\pi}_i, \hat{\pi}_j) = -\pi_i\pi_j/n$  for  $i \neq j$ . Then by delta method,

$$\begin{aligned} \sqrt{n}(g(\hat{\pi}) - g(\pi)) &= \sum_i \frac{\partial g(\pi)}{\partial \pi_i} \sqrt{n}(\hat{\pi}_i - \pi_i) + o_P(1) \\ &= \sum_i -\frac{\eta_i}{\delta^2} \sqrt{n}(\hat{\pi}_i - \pi_i) + o_P(1) \\ &\xrightarrow{d} N(0, \sigma^2), \end{aligned}$$

where

$$\sigma^2 = \sum_i \frac{\eta_i^2}{\delta^4} \pi_i(1 - \pi_i) - 2 \sum_{i < j} \sum \frac{\eta_i \eta_j}{\delta^4} \pi_i \pi_j = \left[ \sum_i \pi_i \eta_i^2 - \left( \sum_i \pi_i \eta_i \right)^2 \right] / \delta^4.$$

□

- **Solution to 3.30.** Recall the notation

$$\hat{\pi}_1 = y_1/n_1, \quad \hat{\pi}_2 = y_2/n_2, \quad \hat{\pi} = (y_1 + y_2)/(n_1 + n_2), \quad n = n_1 + n_2.$$

Then

$$z^2 = \frac{(y_1/n_1 - y_2/n_2)^2}{\left( \frac{y_1+y_2}{n_1+n_2} \right) \left( 1 - \frac{y_1+y_2}{n_1+n_2} \right) \left( \frac{1}{n_1} + \frac{1}{n_2} \right)} = \frac{(n_2 y_1 - n_1 y_2)^2 (n_1 + n_2)}{n_1 n_2 (y_1 + y_2) (n_1 + n_2 - y_1 - y_2)}.$$

For the table

$y_1$	$n_1 - y_1$
$y_2$	$n_2 - y_2$

$$X^2 = \frac{(y_1 - n_1 \hat{\pi})^2}{n_1 \hat{\pi}} + \frac{(n_1 - y_1 - n_1(1 - \hat{\pi}))^2}{n_1(1 - \hat{\pi})} + \frac{(y_2 - n_2 \hat{\pi})^2}{n_2 \hat{\pi}} + \frac{(n_2 - y_2 - n_2(1 - \hat{\pi}))^2}{n_2(1 - \hat{\pi})}.$$

Simple calculation implies that  $X^2 = z^2$ . □