Detection of regulated genes in microarray data

Vincent Detours

IRIBHM – Université Libre de Bruxelles

vdetours@ulb.ac.be

A typical experiment: gene selection

Up-regulated genes are selected as follows

- 1. Compute expression ratio of tumor vs. healthy tissue for each patient
- 2. Select genes with fold-change >2 in at least 8/12=2/3 of the patients

A typical experiment: the data

The samples

- 12 PTC samples
- 12 patient-matched adjacent healthy tissues

The platform & processing

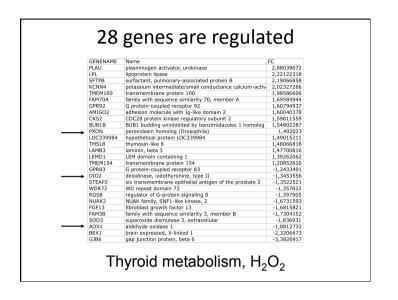
- Affymetrix[®] chips, 20,000 genes
- MAS 5.0 normalization

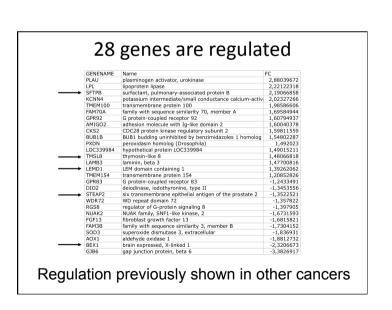
2

28 genes are regulated

GENENAME	Name	FC
PLAU	plasminogen activator, urokinase	2,88039672
LPL	lipoprotein lipase	2,22122318
SFTPB	surfactant, pulmonary-associated protein B	2,19066858
KCNN4	potassium intermediate/small conductance calcium-activ	2,02327266
TMEM100	transmembrane protein 100	1,98586606
FAM70A	family with sequence similarity 70, member A	1,69584944
GPR92	G protein-coupled receptor 92	1,60794937
AMIGO2	adhesion molecule with Ig-like domain 2	1,60040378
CKS2	CDC28 protein kinase regulatory subunit 2	1,59811559
BUB1B	BUB1 budding uninhibited by benzimidazoles 1 homolog	1,54802287
PXDN	peroxidasin homolog (Drosophila)	1,492023
LOC339984	hypothetical protein LOC339984	1,49015211
TMSL8	thymosin-like 8	1,48066818
LAMB3	laminin, beta 3	1,47700816
LEMD1	LEM domain containing 1	1,39262062
TMEM154	transmembrane protein 154	1,20852826
GPR83	G protein-coupled receptor 83	-1,2433491
DIO2	deiodinase, iodothyronine, type II	-1,3453556
STEAP2	six transmembrane epithelial antigen of the prostate 2	-1,3522521
WDR72	WD repeat domain 72	-1,357822
RGS8	regulator of G-protein signaling 8	-1,397905
NUAK2	NUAK family, SNF1-like kinase, 2	-1,6731593
FGF13	fibroblast growth factor 13	-1,6815821
FAM3B	family with sequence similarity 3, member B	-1,7304152
SOD3	superoxide dismutase 3, extracellular	-1,836931
AOX1	aldehyde oxidase 1	-1,8812732
BEX1	brain expressed, X-linked 1	-2,3206673
GJB6	gap junction protein, beta 6	-3,3826917

GENENAME	Name	FC
PLAU	plasminogen activator, urokinase	2.88039672
LPL	lipoprotein lipase	2,22122318
SFTPB	surfactant, pulmonary-associated protein B	2.19066858
KCNN4	potassium intermediate/small conductance calcium-active	2.02327266
TMEM100	transmembrane protein 100	1,98586606
FAM70A	family with sequence similarity 70, member A	1,69584944
GPR92	G protein-coupled receptor 92	1,60794937
AMIGO2	adhesion molecule with Ig-like domain 2	1,60040378
CKS2	CDC28 protein kinase regulatory subunit 2	1,59811559
BUB1B	BUB1 budding uninhibited by benzimidazoles 1 homolog	1,54802287
PXDN	peroxidasin homolog (Drosophila)	1,492023
LOC339984	hypothetical protein LOC339984	1,49015211
TMSL8	thymosin-like 8	1,48066818
LAMB3	laminin, beta 3	1,47700816
LEMD1	LEM domain containing 1	1,39262062
TMEM154	transmembrane protein 154	1,20852826
GPR83	G protein-coupled receptor 83	-1,2433491
DIO2	deiodinase, iodothyronine, type II	-1,3453556
STEAP2	six transmembrane epithelial antigen of the prostate 2	-1,3522521
WDR72	WD repeat domain 72	-1,357822
RGS8	regulator of G-protein signaling 8	-1,397905
NUAK2	NUAK family, SNF1-like kinase, 2	-1,6731593
FGF13	fibroblast growth factor 13	-1,6815821
FAM3B	family with sequence similarity 3, member B	-1,7304152
SOD3	superoxide dismutase 3, extracellular	-1.836931
AOX1	aldehyde oxidase 1	-1,8812732
BEX1	brain expressed, X-linked 1	-2,3206673
GJB6	gap junction protein, beta 6	-3,3826917





All this was a statistical illusion!

Up-regulated genes were actually selected as follows

- 1. Randomly mix-up "tumor" and "healthy" sample labels.
- 2. Compute expression ratio of "tumor" vs. "patientmatched normal" tissue
- 3. Select genes with fold-change >2 in at least 8/12=2/3 of the patients

9

Improbable events are observed with high probability when thousands of observations are made

An example of multiple testing effect

- Assume that 1 person in 5,000 is above 2.1 meters high
- The presence of such person in this audience is unlikely
- But—by chance alone—we expect to find 4 in a stadium of 20,000

11

RT-PCR confirmations do *not* rule out statistical illusions, but statistical controls might

- Measurement quality is not the problem.
- The same illusion would occur if mislabeling had been unintentional or if labeling was correct but the tissue classes at hand were biologically similar
- The illusion results from a statistically inappropriate gene selection procedure
- Only statistical controls can prevent such illusions

10

Statistical significance may point to biological significance... or not

- Statistical significance, typically p-value, measures the probability that the observed event occurred by chance alone
- Some statistically significant events might be pointless from a biological standpoint
- Some biologically important events may not by statistically significant

Statistical hypothesis testing (explained in statistical jargon)

 A statistical test provides a mechanism for making quantitative decisions about a process or processes. The intent is to determine whether there is enough evidence to "reject" a conjecture or hypothesis about the process. The conjecture is called the null hypothesis.

13

Ranking regulated genes rests on three categories of technical choices

- 1. What test statistics to summarize information across samples?
- 2. What procedure to compute the distribution of that test statistics for non regulated genes?
- 3. What procedure to handle multiple testing?

15

Statistical hypothesis testing (explained to biologists)

- A statistical test aims at ruling out chance as the trivial explanation for the observations
- The 'statistics' is the output of any detection or measure assay
- The 'null distribution' is a negative control to make sure that the detection procedure being used does not detect something when there is nothing to be detected

14

Test statistics make differing trade-offs between biological and statistical significance

 The fold-change mean, assumes that large effects have stronger biological impacts, but it may be statistically confusing.

$$\mu_{\mathsf{M}}$$

 The t-statistics captures consistent, but possibly tiny fold-change variations

$$\frac{\mu_{\scriptscriptstyle M}}{\sqrt{\sigma_{\scriptscriptstyle M}^2}}$$

 The moderated t-statistics captures consistent variations, while discarding genes with tiny fold change

$$\frac{\mu_{M}}{\sqrt{s_{0} + \sigma_{M}^{2}}}$$

Other alternatives are possible...

The null distribution is the negative control of statisticians

- The null distribution is the distribution of the test statistics when no gene is regulated
- Under certain hypothesis the t- and other classical statistics have a null distribution with a known mathematical form
- The null distribution may be estimated by repeatedly computing the statistics over randomly mixed-up data

17

Technical options may be combined in various ways

Significance analysis of microarrays (SAM, Tusher *et al.*, 2001, *PNAS* 98, p5116) uses

- Moderated t-statistics
- Permutation-based estimate of the null distribution
- Control of the false discovery rate

19

p-values may be adjusted for multiple testing in various ways

- p-value (no adjustment): proportion of false positives among all genes
- Bonferroni correction: significance threshold = 0.05 / (number of tests)
- > Family-wise error rate (FWER): probability that one or more genes considered regulated are false positives
- > False discovery rate (FDR, q-value): proportion of false positives among genes considered regulated

18

Significance Analysis of Microarrays

SAM is among the most widely used gene selection procedure (R, and Excel add-on)

It is flexible

- one class (deviation from no differential expression)
- paired/unpaired two-classes comparison
- F-test, i.e. *n* classes comparison
- Genes correlated with a continuous variable
- Genes associated with survival (Cox analysis)

The SAM score (two-classes analysis)

Relative difference (= modifyed t-test):

$$d(i) = \frac{\bar{x}_{\mathrm{I}}(i) - \bar{x}_{\mathrm{U}}(i)}{s(i) + s_0}$$

x is the expression level, so some percentile of the si across all genes

Gene-specific standard deviation $[a=(1/n_1+1/n_{11})/(n_1+n_{11}-2)]$

$$s(i) = \sqrt{a \left\{ \sum_{m} [x_{m}(i) - \bar{x}_{I}(i)]^{2} + \sum_{n} [x_{n}(i) - \bar{x}_{U}(i)]^{2} \right\}}$$

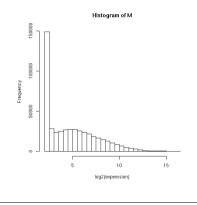
21

Tests on multiple genes are *not* independent

- · Cells and tissues functions are correlated
- Genes that are contiguous on chromosomes tend to be co-expressed
- Relative cell type abundance has a substantial influence on tissues gene expression

3

Normality assumptions are violated with microarray data



In fact distributions of expression values do not seems to fit simple mathemtical forms

Computing the null distribution for SAM's *d*-score

 The null hypothesis (i.e. no association beteween classes and gene expression) is modeled expermentally by scrambling this relationship

tumor

normal

B C D E

real data

permuted (i.e. scrambled) data (*B* permutations)



False positives are more of a problem than false negatives in microarray studies

- Most studies yield list of several hundred regulated genes
- there are ususally enough positive genes to get biological insights, and too much for experimental follow up
- So far, the issue of power has been paid little attention to
- There is a tool to compute the power of SAM analyses (http://www-stat.stanford.edu/~tibs/SAM/)

26

The number of samples needed to reliably detect differential expression is computable

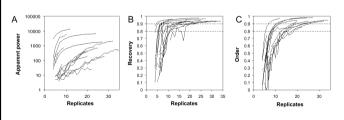


Fig. 3. Summary of results. Each line represents results for one data set shown in Table 1, at an FDR of 0.05. Not all of the 16 data sets are illustrated on these graphs, because some failed to meet criteria at this FDR (see our web site for more results). The plots are of the medical values for all trails. Error bars are omitted for clarity. The dashed lines in (B) and (C) indicate the 0.8 and o9 levels. (A) Plot of the number of genes selected (apparent power, the size of Szel). Note that the scale is logarithmic. (B) Recovery stability. (C) Order stability. Values below zero are not shown. Lareer versions of this and the other fluences are available as sunplementary data.

From Pavlidis et al., (2001), Bioinformatics 19, 1620-7

The *q*-value

- Storey & Tibshirani (2003, PNAS 100, p9440) propose a formal procedure to adjust p-values for multiple testing
- The q-value of a test is the fraction of false positive among all the tests with statistics as or more extreme than this test
- The procedure takes *p*-values as input (available as R package)

The q-value calculation, roughly

	Called significant	Called not significant	Total
Null true	F	$m_0 - F$	<i>m</i> ₀
Alternative true	T	$m_1 - T$	m_1
Total	S	m-S	m

FDR(t) = (# false positive)/(# called significant)

= E(F(t) / [F(t) + T(t)])

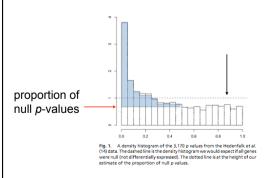
 \approx E(F(t)) / E(F(t) + T(t)) (because m is large)

 $\approx m_0 t / (\# p_i < t)$ (because *p*-values are uniformly distributed under true H_0)

t is the p-value threshold (also called α) m_0 can be estimated from the distribution of p-values

29

The trick: distribution of *p*-values contains information about the fraction of positive tests



1

The trick: distribution of *p*-values contains information about the fraction of positive tests

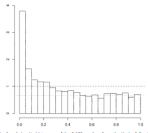


Fig. 1. A density histogram of the 3,170 p values from the Hedenfalk et al. (14) data. The dashed line is the density histogram we would expect if all genes were null (not differentially expressed). The dotted line is at the height of our