Probability I - Lecture 2

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Contents

| 1 | Expectation | 3 |
|---|-----------------------------|---|
| 2 | Higher Moments 2.1 Variance | 4 |
| 3 | Independence | 7 |

1 Expectation

Definition 1. If X is a discrete random variable, then the expectation (**expected value**) is:

$$E[X] = \sum_{k \in range(X)} k \cdot P[X = k].$$

aka a space average. If X is a continuous random variable with pdf $f_{x}(t)$,

$$E[X] = \int_{-\infty}^{\infty} t \cdot f_x(t) dt.$$

Remark. Examples

- $E[Bernoulli(\theta)] = 0 \cdot (1-p) + 1 \cdot p = p$
- Let Y be the outcome of a fair 3 sided die. $E[Y] = \frac{1+2+3+4}{4} = 2.5$.
- M = maximum of two fair 4-sided dice. Let M be random variable representing the maximum of two fair dice rolls (independent). The $\Omega = \{(1,1),\ldots\}$. Also, P is uniform for dice rolls. $E[M] = 1 \cdot P[M=1] + 2 \cdot P[M=2] + 3 \cdot P[M=3] + 4 \cdot P[M=4] = 1 \cdot \frac{3}{16} + 2\frac{3}{16} + 3\frac{5}{16} + 4\frac{7}{16}$
- $R \sim Geo(p)$ geometric distribution. Define $P[R=k] = p(1-p)^k$ and $k \in \{1, 2, 3, \ldots\}$.
- $x \sim \exp(\lambda)$ i.e. $f_x(t) = \lambda e^{-\lambda t}$ when $t \ge 0$ and 0 otherwise.

Theorem 1. If X and Y are two random variables defined on the same space and $a, b \in \mathbb{R}$. Then,

$$E[aX + bY] = aE[X] + bE[Y].$$

Proof. Discrete case

$$\begin{split} E[Z] &= \sum_{k \in range(z)} k \cdot P[Z = k] \\ &= \sum_{r \in range(X), t \in range(y)} (ar + bt) P[Z = ar + bt] \\ &= \sum_{r,t} (ar + bt) \cdot P[X = r, Y = t] \quad \text{Note r,t is shorthand} \\ &= \sum_{r} \sum_{t} ar P[X = r, Y = t] + \sum_{r} \sum_{t} bt \cdot P[X = r, Y = t] \\ &= \sum_{r} ar \cdot (\sum_{t} P[X = r, Y = t) + \sum_{t} bt \sum_{r} P[X = r, Y = t] \\ &= a \sum_{r} r \cdot P[X = r] + b \sum_{t} t \cdot P[Y = t] \\ &= a E[X] + b E[Y]. \end{split}$$

This is useful! For example, back to card dealing, assuming we agin deal 3 cards from 52 cards. Observe A=X+Y+Z, we can use $E[A]=E[X]+E[Y]+E[Z]=\frac{3}{13}$

2 Higher Moments

 $E[x^k] = \text{kth moment of X}.$

2.1 Variance

Definition 2. The second moment known as the **variance**: $E[x^2]$:

$$var(x) = E[x^2] - E[x]^2.$$

Lemma 1.

$$var(x) = E[(X - E[X])^2].$$

Remark. The two methods of computation are equivalent but different. For example:

$$E[Ber(p)^2] = p, E[Ber(p)] = p^2 \rightarrow var(Ber(p)) = p - p^2.$$

Then,

$$Y = Ber(p) - E[Ber(p)] = Ber(p) - p.$$

where Y is the distribution:

$$P[Y = p - 1] = pP[Y = -p] = 1 - p.$$

and the distribution of Y^2 :

$$P[Y^2 = (1-p)^2] = p^2 P[Y^2 = p^2] = 1-p.$$

so when we compute we get the same answer:

$$var(Ber(p)) = E[(Ber(p) - E[Ber(p)])] = E[Y^2]$$

= $(1 - p)^2 \cdot p + p^2 \cdot (1 - p)$
= $p(1 - p)$
= $p - p^2$.

Proof. Let $\mu = E[X]$

$$\begin{split} E[(X-\mu)^2] &= E[X^2 - 2\mu X + \mu^2] \\ &= E[X^2] - 2\mu E[X] + \mu^2 \\ &= E[X^2] - 2\mu^2 + \mu^2 \\ &= E[X^2] - \mu^2 \\ &= E[X^2] - E[X]^2. \end{split}$$

Remark. Variance can be thought of as a measure of how random a R.V. is.

5

Proposition 1. If $var(x) = 0 \Leftrightarrow P[x = a] = 1$ for some a i.e. X is not a random variable at all!

Proof. Assume var(x) = 0. Then this means $E[(X - \mu)^2] = 0$.

Remark. Example

 $X = \{1, 2, 3, 4\}$ with P uniform. $E[X] = \frac{5}{2}$.

$$\begin{split} E[X^2] &= 1 \cdot P[X^2 = 1] + 4 \cdot P[X^2 = 4] + 9 \cdot P[X^2 = 9] + 16 \cdot P[X^2 = 16]. \\ &= \frac{1 + 4 + 9 + 16}{4} = \frac{15}{2}. \end{split}$$

So,

$$var(X) = \frac{15}{2} - \frac{5}{2}^2 = \frac{5}{4}.$$

Remark. Example

 $U \sim uniform(0,2)$ continuous. $E[U] = \int_0^2 t \cdot \frac{1}{2}$ since uniform probability would be .5 across interval from 0 to 2.

$$var(u) = E[(U-1)^2] = \int_0^2 (t-1)^2 \cdot .5dt.$$
$$= \frac{1}{2} \int_0^2 t^2 - 2t + 1dt.$$

2 HIGHER MOMENTS

$$= \frac{1}{2} \left[\frac{t^3}{3} - t^2 + t \right]_0^2.$$
$$= \frac{1}{3}.$$

Proposition 2. 1. Variance is not preserved under scalar multiplication

$$var(aX) = E[(aX)^{2}] - E[aX]^{2}$$

$$= E[a^{2}X_{2}] - (aE[x])^{2}$$

$$= a^{2}(E[X^{2}] - E[X]^{2})$$

$$= a^{2} \cdot var(X).$$

2. variance is preserved under additivity

$$var(X + b) = var(X).$$

3. Variance is nonlinear

$$var(X + Y) \neq var(X) + (Y)$$
 in general.

Definition 3. We have a special term to measure the relationship of between the variance sof 2 random variables known as the **covariance**

$$cov(X,Y) = E[XY] - E[X]E[Y].$$

The matrix covariance formula for X,Y matrices is: $Cov(X,Y) = E[(X-E[X])(Y-E[Y])^T]$

Lemma 2.

$$Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y).$$

Proof.

$$\begin{split} E[(X+Y)^2] &= E[X^2] + 2E[XY] + E[Y^2] \\ E[X+Y]^2 &= E[X]^2 + 2E[X]E[Y] + E[Y]^2 \\ \mathrm{Var}(X+Y) &= (E[X^2] + 2E[XY] + E[Y^2]) - (E[X]^2 + 2E[X]E[Y] + E[Y]^2) \\ &= E[X^2] - E[X]^2 + 2(E[XY] - E[X]E[Y]) + E[Y^2] - E[Y]^2 \\ &= \mathrm{Var}(X) + 2\mathrm{Cov}(X,Y) + \mathrm{Var}(Y) \end{split}$$

Recall the example from the previous lecture

Note. For example, consider the discrete random variables x,y,z over $\Omega = \{1,2,3,4\}$ on a uniform probability measure:

$$X = \begin{cases} 1, w \in \{1, 2\} \\ 0, w \in \{3, 4\} \end{cases} .$$

$$Y = \begin{cases} 1, w \in \{3, 4\} \\ 0, w \in \{1, 2\} \end{cases} .$$

$$z = \begin{cases} 1, w \in \{1, 3\} \\ 0, w \in \{2, 4\} \end{cases} .$$

What is the variance of X + Y?

$$var(X+Y) = Var(x) + Var(y) + 2Cov(X,Y) = \frac{1}{4} + \frac{1}{4} + 2Cov(X,Y).$$

$$Cov(X,Y) = E[XY] - E[X]E[Y] = E[XY] \cdot \frac{1}{4} = -\frac{1}{4}.$$

$$var(X+Y) = \frac{1}{2} - 2 \cdot \frac{1}{4} = 0!$$

Definition 4. If Cov(X,Y) = 0 we say X,Y are uncorrelated

3 Independence

Definition 5. Let Ω be the outcome space and P the probability measure. $A, B \subseteq \Omega$ are events. A, B are **independent** if $P[A \cap B] = P[A] \cdot P[B]$. We say X, Y random variables on Ω are independent if $\{X = a\}$ and $\{Y = b\}$ are independent events for every a,b in range of X, Y respectively.