

# Time Series Analysis using R

A.Tharwinvikram Nadar

20/11/2020

## Motivation:

To study the Britannia stock price data perform various time series model and it's analysis

```
library(ggplot2)
library(forecast)
library(tseries)
library(ISLR)
raww_data = read.csv("/home/rocky/Desktop/Vikram2019-20/dataset/Stock-Data/BRITANNIA.csv",
                     ,header = TRUE)
```

In data we see Open close and VWAP prices of the stock.

To build a forecasting model for the closing price of the Britannia stock data

```
head(raww_data,10)
```

##	Date	Symbol	Series	Prev.Close	Open	High	Low	Last	Close
## 1	2000-01-03	BRITANNIA	EQ	703.25	705.0	759.50	705.00	758.00	756.90
## 2	2000-01-04	BRITANNIA	EQ	756.90	710.0	770.00	710.00	740.00	754.55
## 3	2000-01-05	BRITANNIA	EQ	754.55	755.0	759.00	705.00	740.00	735.30
## 4	2000-01-06	BRITANNIA	EQ	735.30	740.0	794.15	740.00	770.00	785.65
## 5	2000-01-07	BRITANNIA	EQ	785.65	808.0	848.50	798.00	848.50	848.50
## 6	2000-01-10	BRITANNIA	EQ	848.50	900.0	916.40	865.00	916.40	912.20
## 7	2000-01-11	BRITANNIA	EQ	912.20	920.0	920.00	839.25	865.00	853.75
## 8	2000-01-12	BRITANNIA	EQ	853.75	900.0	900.00	860.55	890.95	882.70
## 9	2000-01-13	BRITANNIA	EQ	882.70	890.0	920.00	875.00	885.00	881.40
## 10	2000-01-14	BRITANNIA	EQ	881.40	872.5	880.00	864.00	870.00	869.65
##	VWAP	Volume	Turnover	Trades	Deliverable.Volume	X.Deliverble			
## 1	741.01	7512	5.566488e+11	NA		NA		NA	
## 2	742.52	8135	6.040391e+11	NA		NA		NA	
## 3	739.92	6095	4.509784e+11	NA		NA		NA	
## 4	788.83	19697	1.553756e+12	NA		NA		NA	
## 5	827.53	33107	2.739708e+12	NA		NA		NA	
## 6	905.42	29575	2.677784e+12	NA		NA		NA	
## 7	858.02	20635	1.770516e+12	NA		NA		NA	
## 8	885.18	9312	8.242756e+11	NA		NA		NA	
## 9	898.96	19526	1.755313e+12	NA		NA		NA	
## 10	873.91	15675	1.369847e+12	NA		NA		NA	

```
tail(raww_data,10)
```

##	Date	Symbol	Series	Prev.Close	Open	High	Low	Last
## 5153	2020-09-17	BRITANNIA	EQ	3844.50	3848.0	3890.95	3768.75	3803.00
## 5154	2020-09-18	BRITANNIA	EQ	3815.65	3837.7	3839.75	3774.85	3804.00
## 5155	2020-09-21	BRITANNIA	EQ	3797.50	3790.0	3795.00	3613.50	3643.50
## 5156	2020-09-22	BRITANNIA	EQ	3629.30	3660.0	3678.00	3540.05	3590.00
## 5157	2020-09-23	BRITANNIA	EQ	3588.00	3618.9	3652.90	3562.80	3630.00
## 5158	2020-09-24	BRITANNIA	EQ	3624.90	3590.0	3655.00	3560.20	3611.00
## 5159	2020-09-25	BRITANNIA	EQ	3612.75	3640.0	3716.95	3615.00	3703.75
## 5160	2020-09-28	BRITANNIA	EQ	3686.40	3710.9	3778.00	3689.00	3731.05
## 5161	2020-09-29	BRITANNIA	EQ	3737.35	3769.0	3796.00	3701.65	3715.05
## 5162	2020-09-30	BRITANNIA	EQ	3736.85	3734.0	3825.00	3714.05	3795.50

##	Close	VWAP	Volume	Turnover	Trades	Deliverable.Volume	X.Deliverble
## 5153	3815.65	3845.28	741010	2.849393e+14	49657	152843	0.2063
## 5154	3797.50	3801.11	947698	3.602309e+14	41566	631679	0.6665
## 5155	3629.30	3687.23	604282	2.228128e+14	43918	251550	0.4163
## 5156	3588.00	3590.40	670648	2.407896e+14	45992	179945	0.2683
## 5157	3624.90	3611.53	405856	1.465763e+14	30346	61928	0.1526
## 5158	3612.75	3612.21	517316	1.868652e+14	38631	152823	0.2954
## 5159	3686.40	3668.71	507368	1.861385e+14	33389	133845	0.2638
## 5160	3737.35	3737.45	390640	1.459997e+14	23905	96348	0.2466
## 5161	3736.85	3756.82	449330	1.688051e+14	24309	126432	0.2814
## 5162	3798.15	3789.81	535771	2.030472e+14	32948	170100	0.3175

## Some Terminology

**VWAP(volume weighted average price)**-In finance, volume-weighted average price is the ratio of the value traded to total volume traded over a particular time horizon. It is a measure of the average price at which a stock is traded over the trading horizon.

**Volume** - In the context of a single stock trading on a stock exchange, the volume is commonly reported as the number of shares that changed hands during a given day. The transactions are measured on stocks, bonds, options contracts, futures contracts and commodities.

Here `tsclean` function is used to identify and replace outliers and missing values

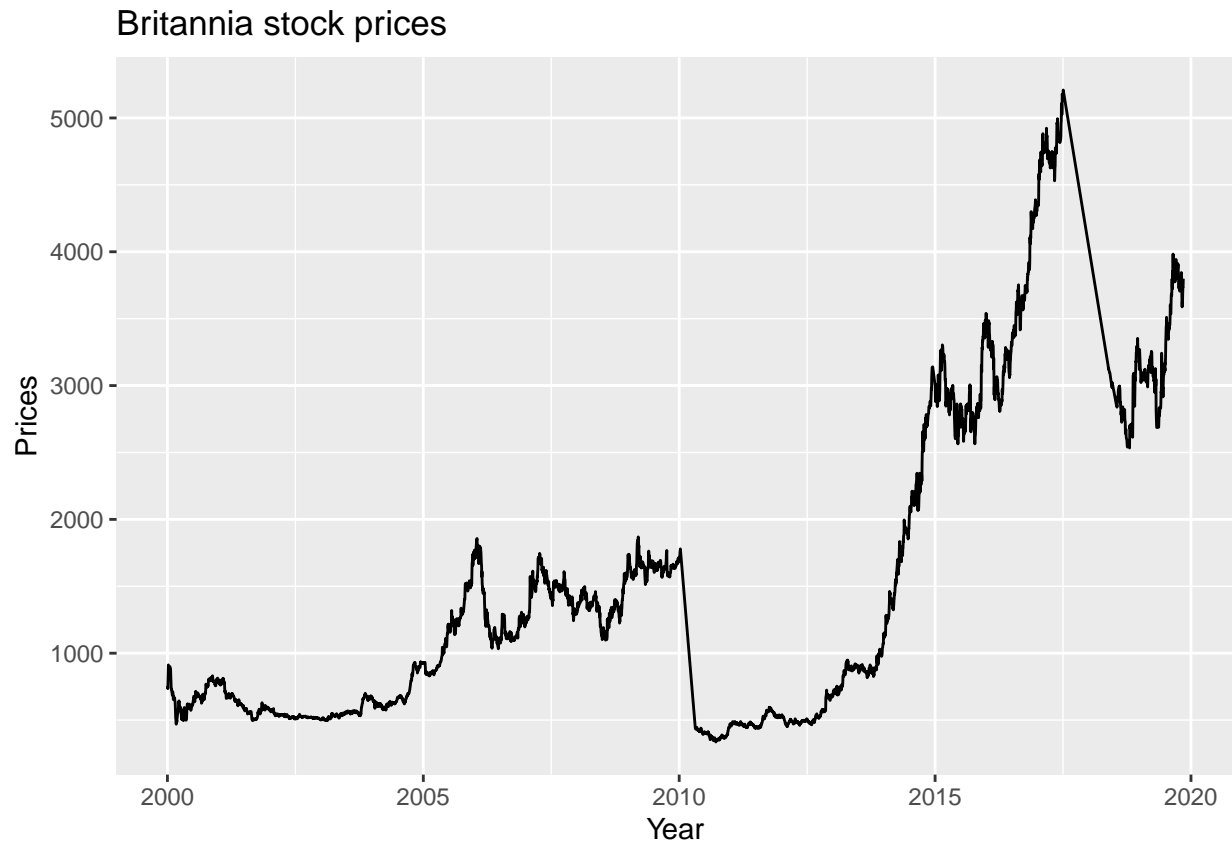
Here frequency is set to 260 because in a year stock market functions for 260 days

```
clean_dataaa = tsclean(raww_data$Close)
my_time = ts(clean_dataaa,start = 2000,frequency = 260)
```

## Plotting the time series

```
autoplot(my_time)+ggtitle("Britannia stock prices")+xlab("Year")+ylab("Prices")
```

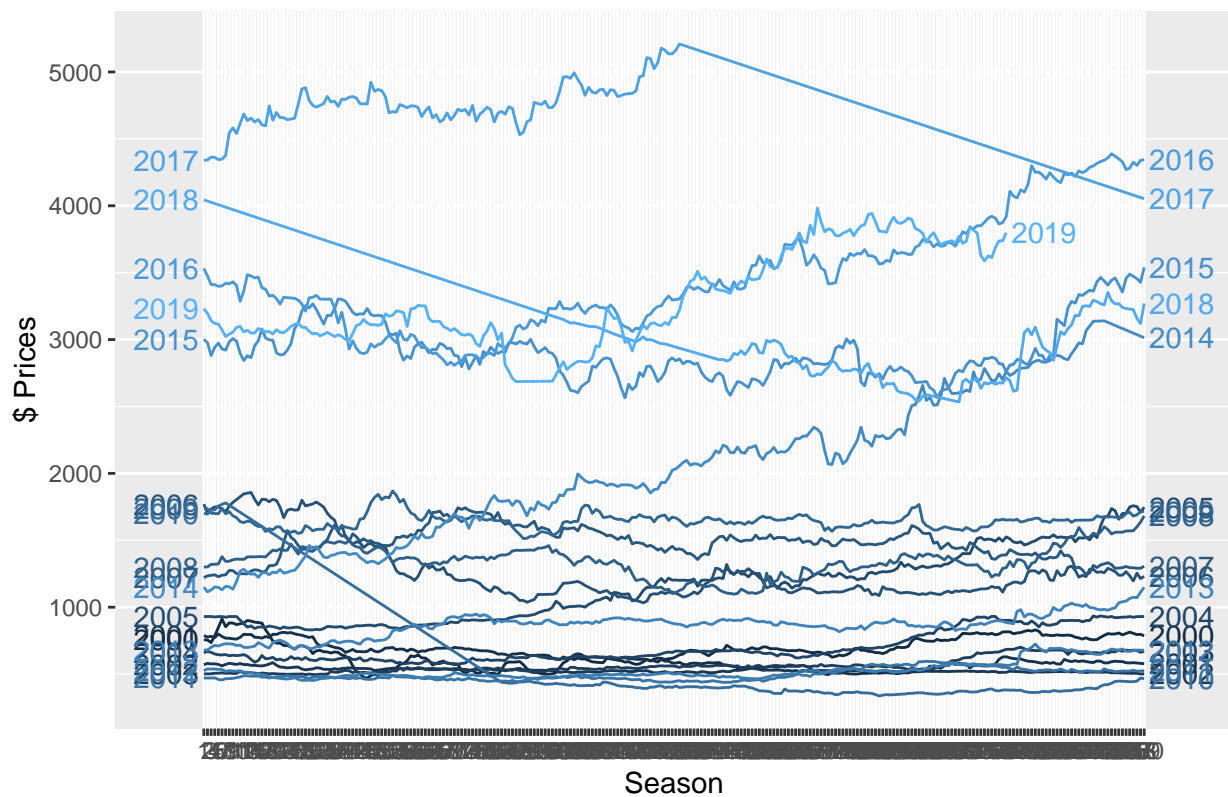
```
## Warning in is.na(main): is.na() applied to non-(list or vector) of type 'NULL'
```



From visual inspection we don't see trend and seasonality. We observe there is a sudden price drop around the year 2010. We will try plotting some seasonality plots to check seasonality

```
ggseasonplot(my_time, year.labels=TRUE, year.labels.left=TRUE, continuous = TRUE) +  
  ylab("$ Prices") +  
  ggtitle("Seasonal plot:Britannia stock prices ")  
  
## Warning in is.na(ylab): is.na() applied to non-(list or vector) of type 'NULL'
```

Seasonal plot:Britannia stock prices

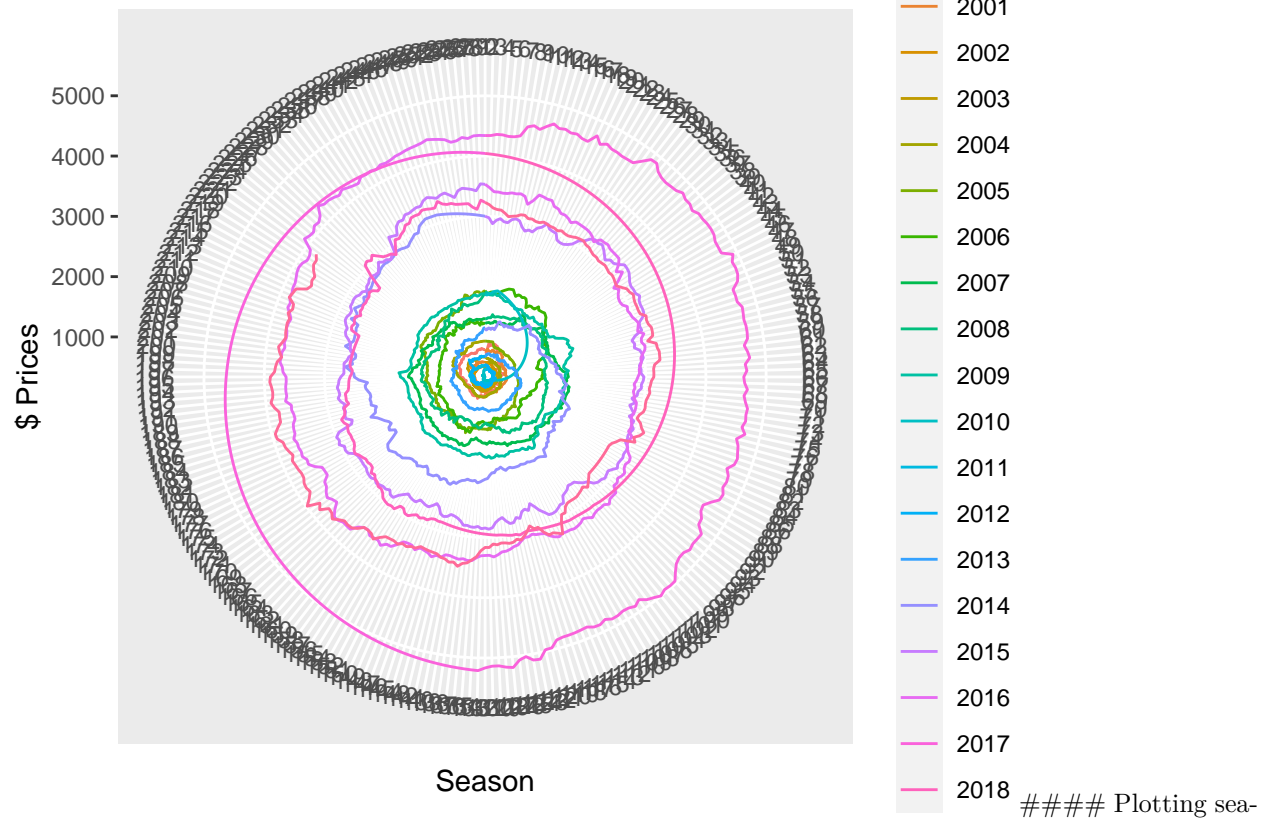


#### Plotting Seasonal Polar plot

```
ggseasonplot(my_time, polar=TRUE) + ylab("$ Prices") +  
  ggtitle("Polar seasonal plot: Britannia stock prices")
```

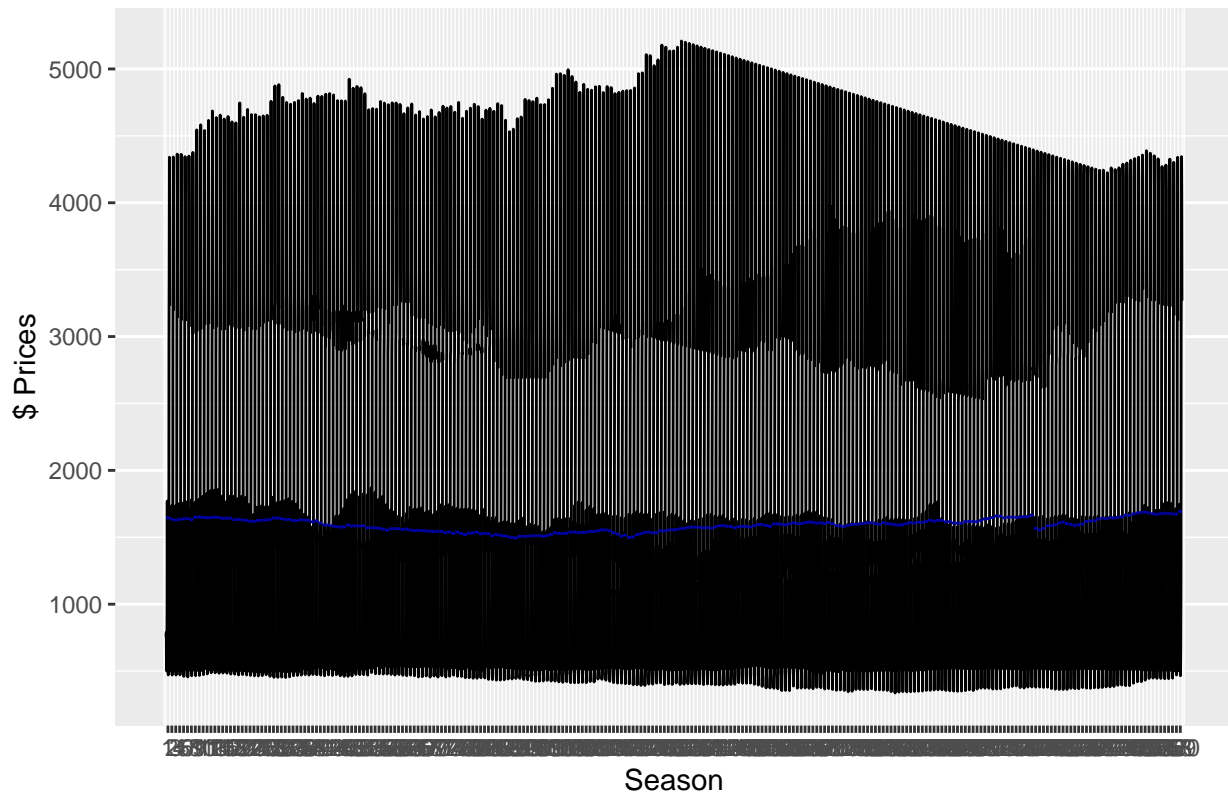
## Warning in is.na(ylab): is.na() applied to non-(list or vector) of type 'NULL'

Polar seasonal plot: Britannia stock prices



```
ggsubseriesplot(my_time)+ylab("$ Prices") +
  ggtitle("Seasonal plot:Britannia stock prices ")
```

Seasonal plot:Britannia stock prices



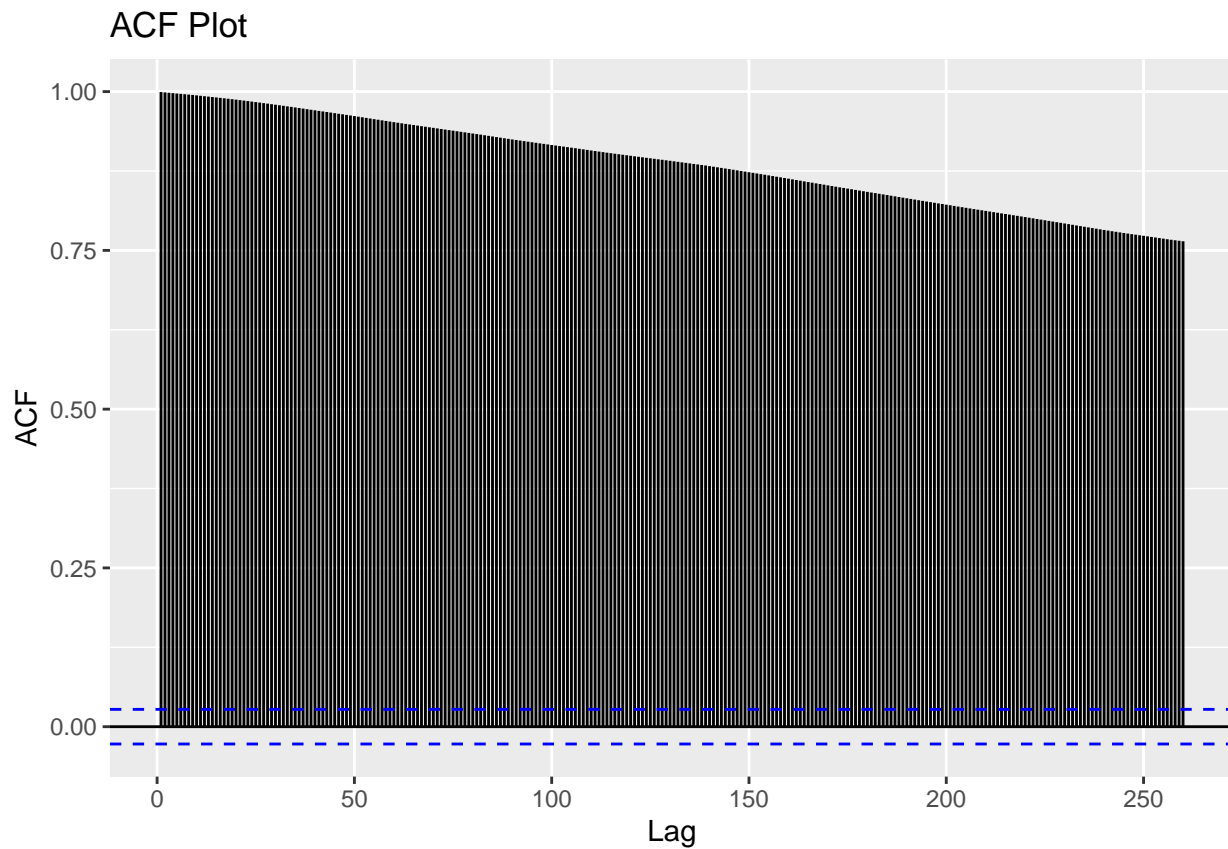
#### Hence from the seasonal plots presence of seasonality in the series is not ensured.

### Autocorrelation function:

Let  $x_{\{t\}}$  be a series  $s$  and  $t$  be a point in the series then the auto correlation function is defined as  $\rho(s, t) = \frac{\gamma(s, t)}{\sqrt{\gamma(s, s)\gamma(t, t)}}$ . The ACF measures the linear predictability of the series at time  $t$ , say  $x_{\{t\}}$ , using only the value of  $x_{\{s\}}$ .

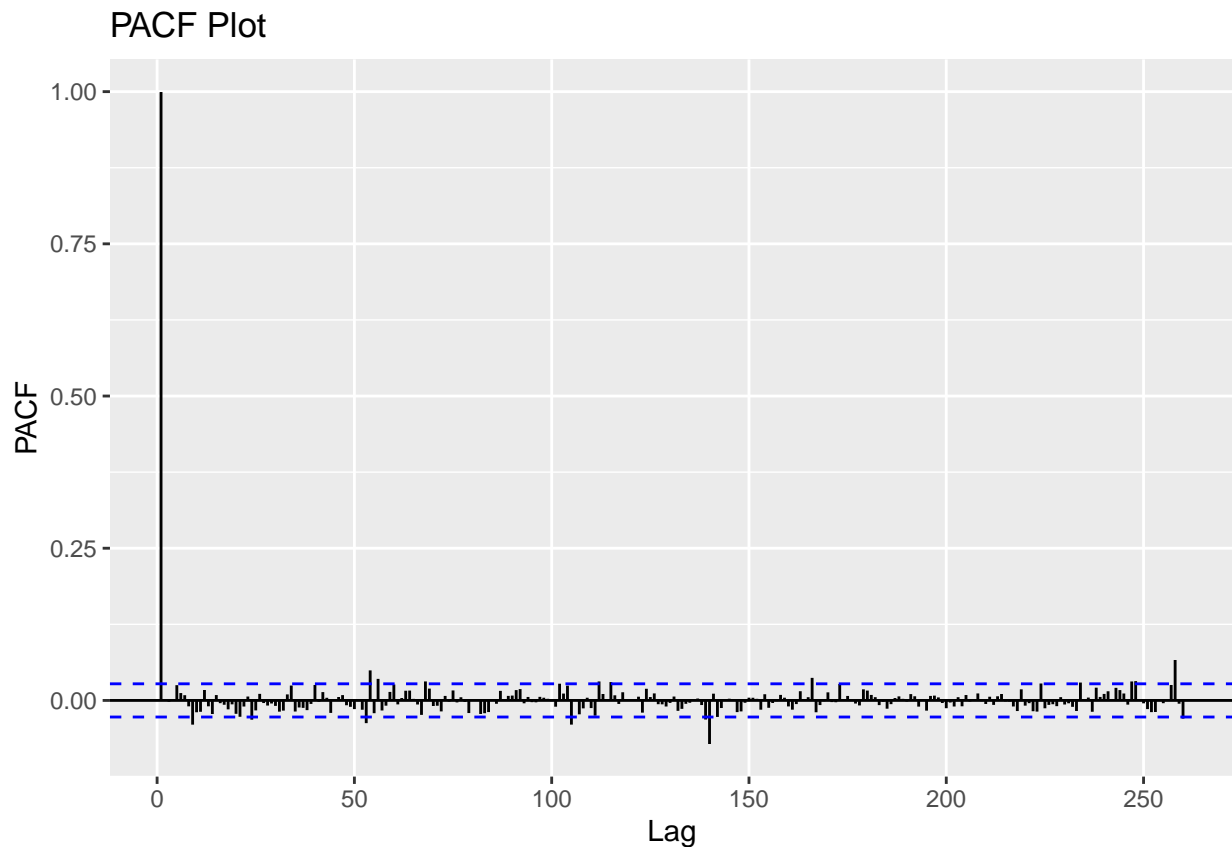
From ACF plot we can say that the data is highly auto correlated and there is decreasing trend present in the data

```
ggAcf(my_time, lag.max = 260) + ggtitle("ACF Plot")
```



PACF plot talks about the coorelation of consecutive points of the series. From the plot we see that lag 1, 53 and other lag outside the dotted line are statically significant and the rest are statically signifcant to 0

```
ggPacf(my_time, lag.max = 260) + ggtitle("PACF Plot")
```



#### Now we will split the data. #### Splitting the data into train and test data

```
train_data = head(my_time,round(length(my_time)*0.8))
h = length(my_time)-length(train_data)
test_data = tail(my_time,h)
```

Now we will try to build basic models on our series

- 1) Average method
- 2) Naive method
- 3) Seasonal Naive
- 4) Random walk drift

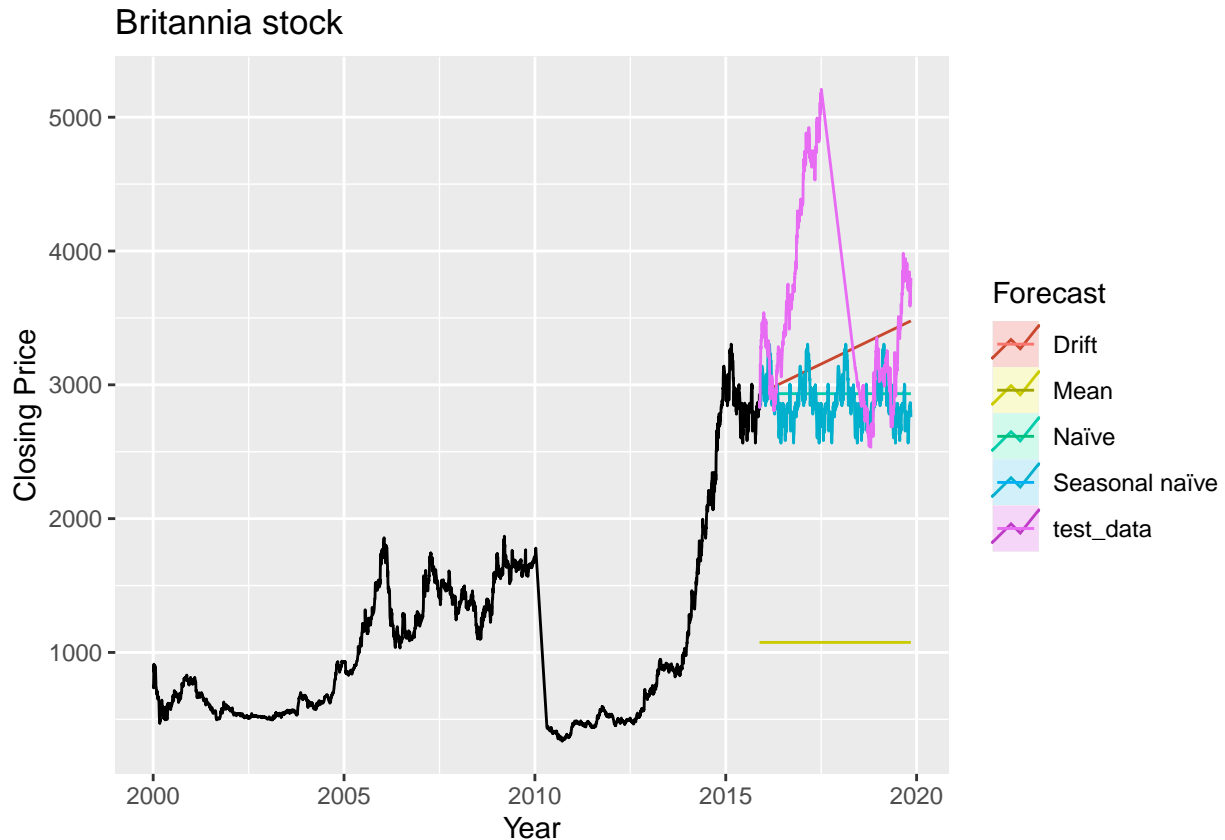
Building all the models in the series

```
model_1 = meanf(train_data, h=h)
model_2 = rwf(train_data, h=h)
model_3 = rwf(train_data, drift=TRUE, h=h)
model_4 = snaive(train_data, h=h)
```



```
autoplot(train_data) +
  autolayer(model_1,series="Mean", PI=FALSE) +
  autolayer(model_2,series="Naïve", PI=FALSE) +
  autolayer(model_3,series="Drift", PI=FALSE) +
  autolayer(model_4,series="Seasonal naïve", PI=FALSE)+
  ggtitle("Britannia stock ") +
  xlab("Year") + ylab("Closing Price") +
  guides(colour=guide_legend(title="Forecast"))+autolayer(test_data)
```

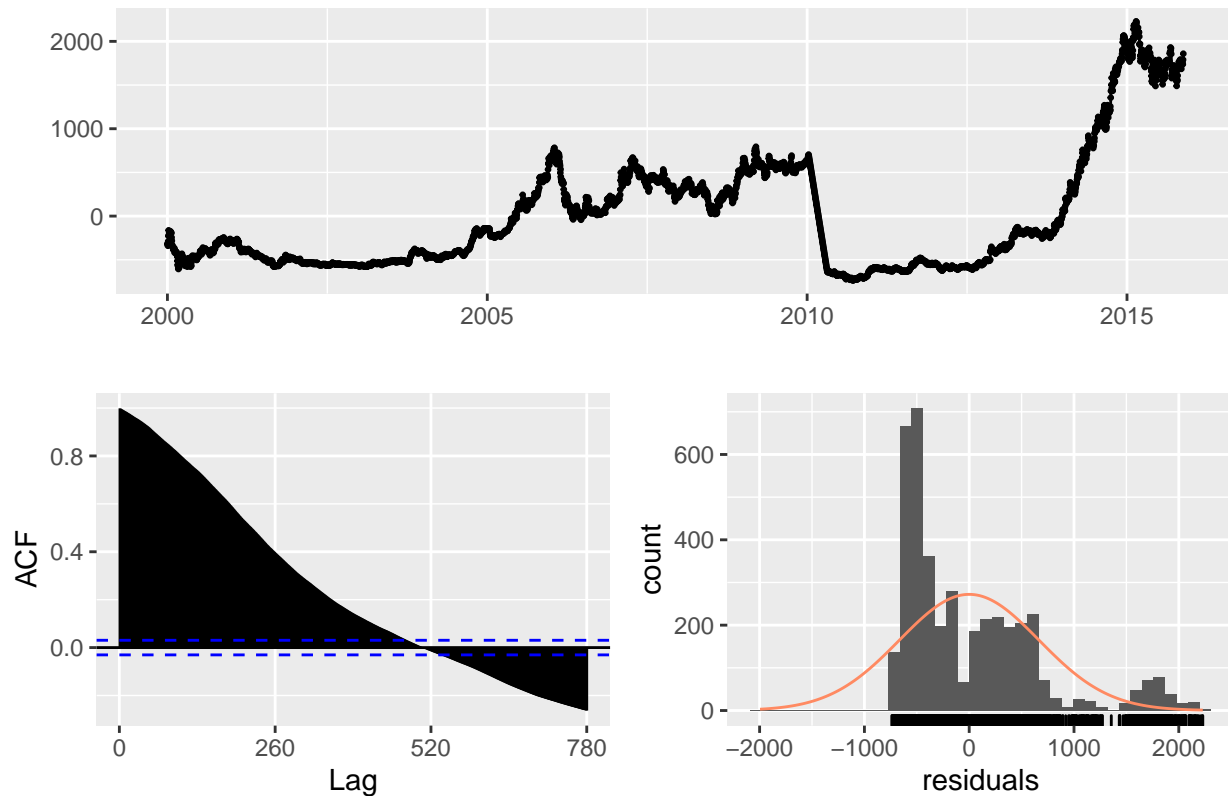
```
## Warning in is.na(main): is.na() applied to non-(list or vector) of type 'NULL'
```



```
#### From visual inspection we see that drift seem to perform well as it captures some the test data
#### Now we will perform residue analysis for each model to see which performs better ## Residue
analysis of model 1
```

```
checkresiduals(model_1)
```

## Residuals from Mean



```
##
##  Ljung-Box test
##
## data:  Residuals from Mean
## Q* = 654480, df = 519, p-value < 2.2e-16
##
## Model df: 1.   Total lags used: 520
```

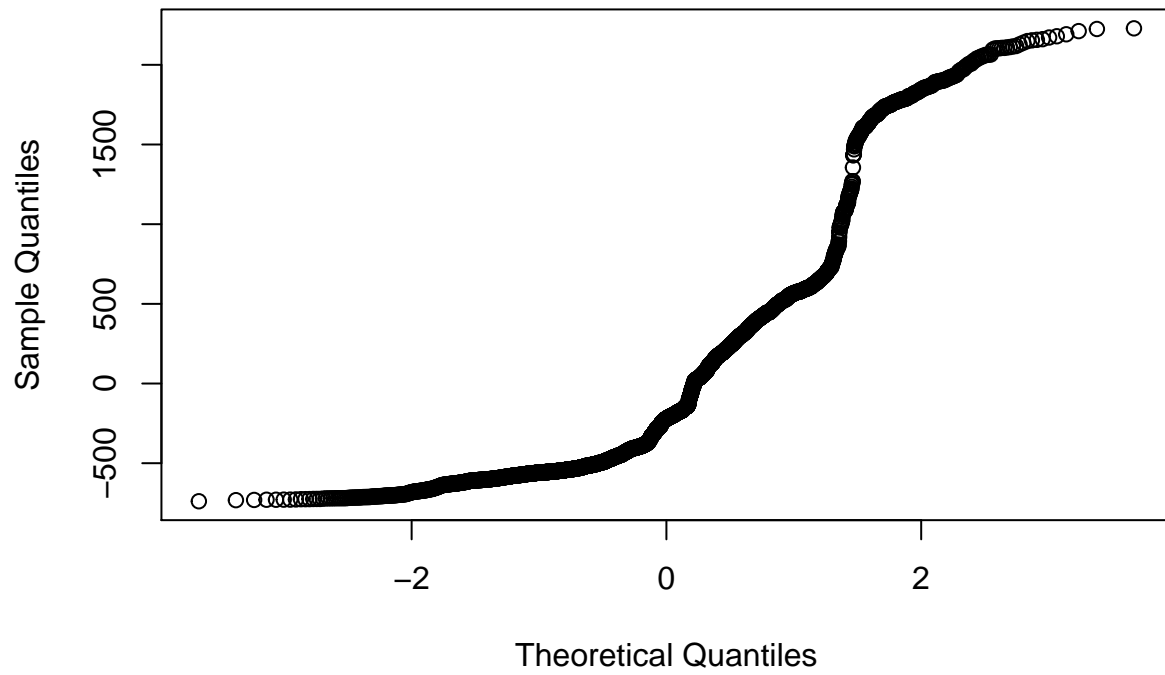
pvalue of Ljung-Box test  $\leq 0.05$  hence we reject the null hypothesis therefore there is no co-relation in the residuals therefore it is not stationary

```
shapiro.test(model_1$residuals)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  model_1$residuals
## W = 0.83878, p-value < 2.2e-16
```

```
qqnorm(model_1$residuals)
```

### Normal Q-Q Plot

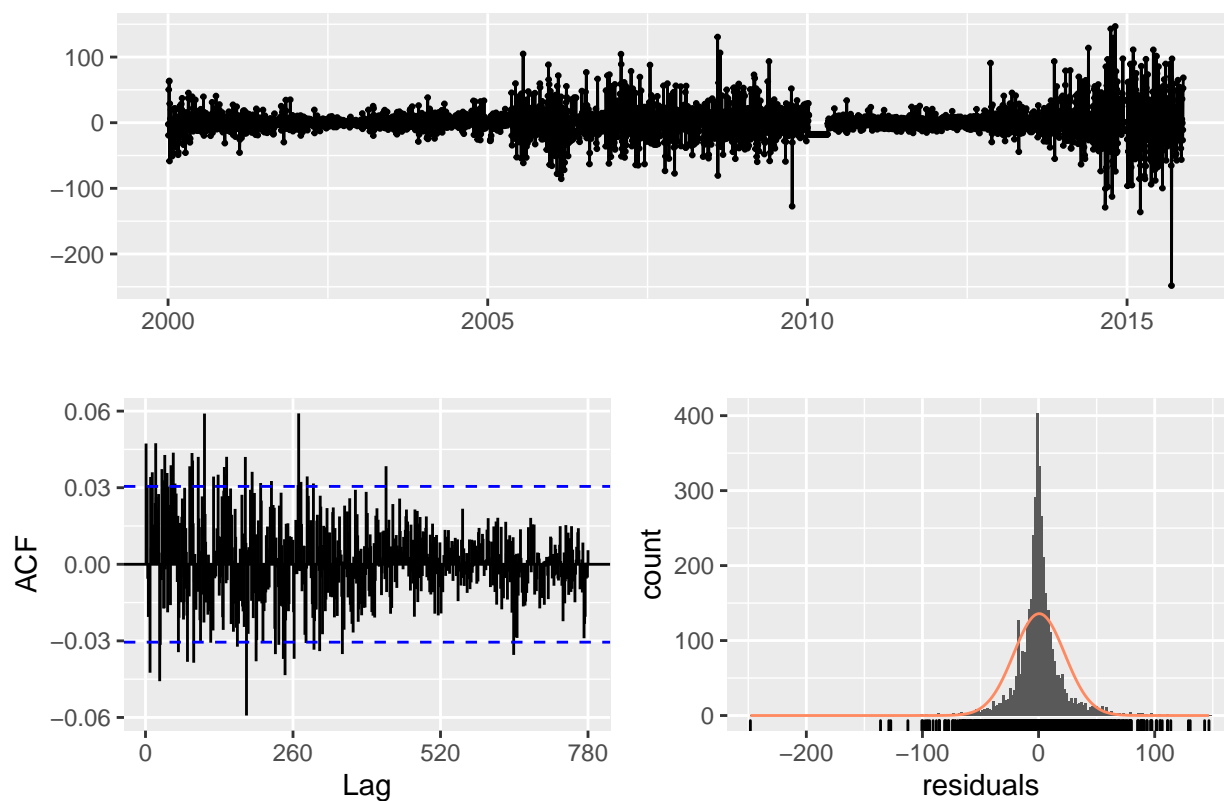


#### pvalue of Shapiro-Wilk test  $\leq 0.05$  hence we reject the null hypothesis therefore residuals not normally distributed

### Residue analysis of model 2

```
checkresiduals(model_2)
```

## Residuals from Random walk



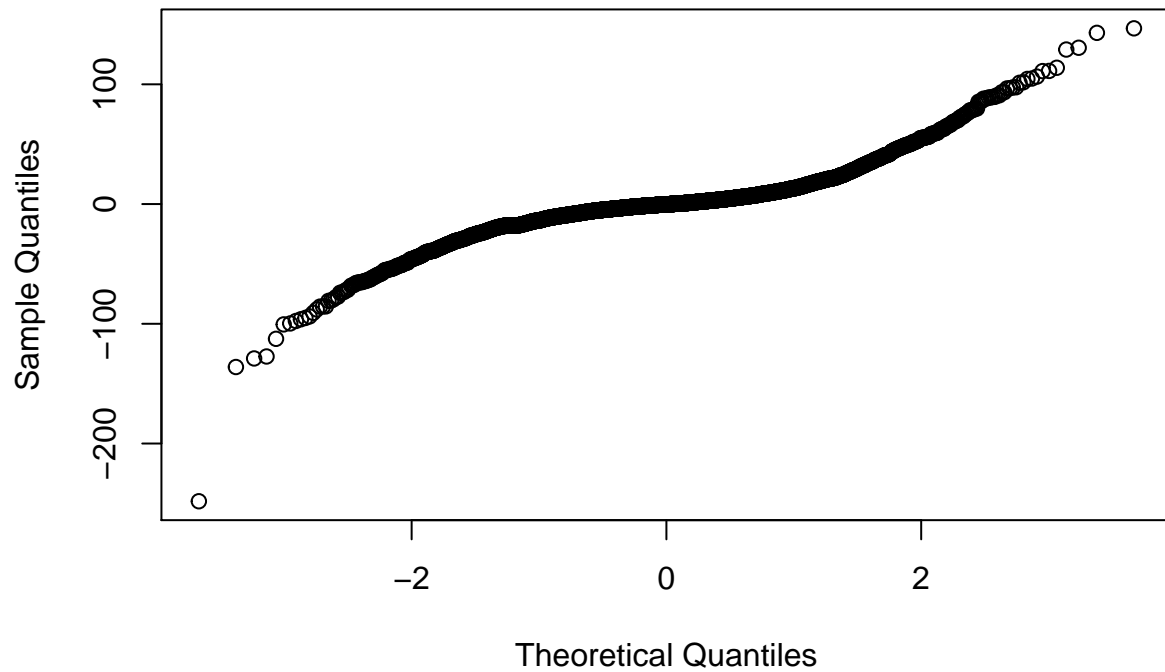
```
##
##  Ljung-Box test
##
## data:  Residuals from Random walk
## Q* = 764.1, df = 520, p-value = 1.459e-11
##
## Model df: 0.   Total lags used: 520
```

```
shapiro.test(model_2$residuals)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  model_2$residuals
## W = 0.86267, p-value < 2.2e-16
```

```
qqnorm(model_2$residuals)
```

### Normal Q-Q Plot

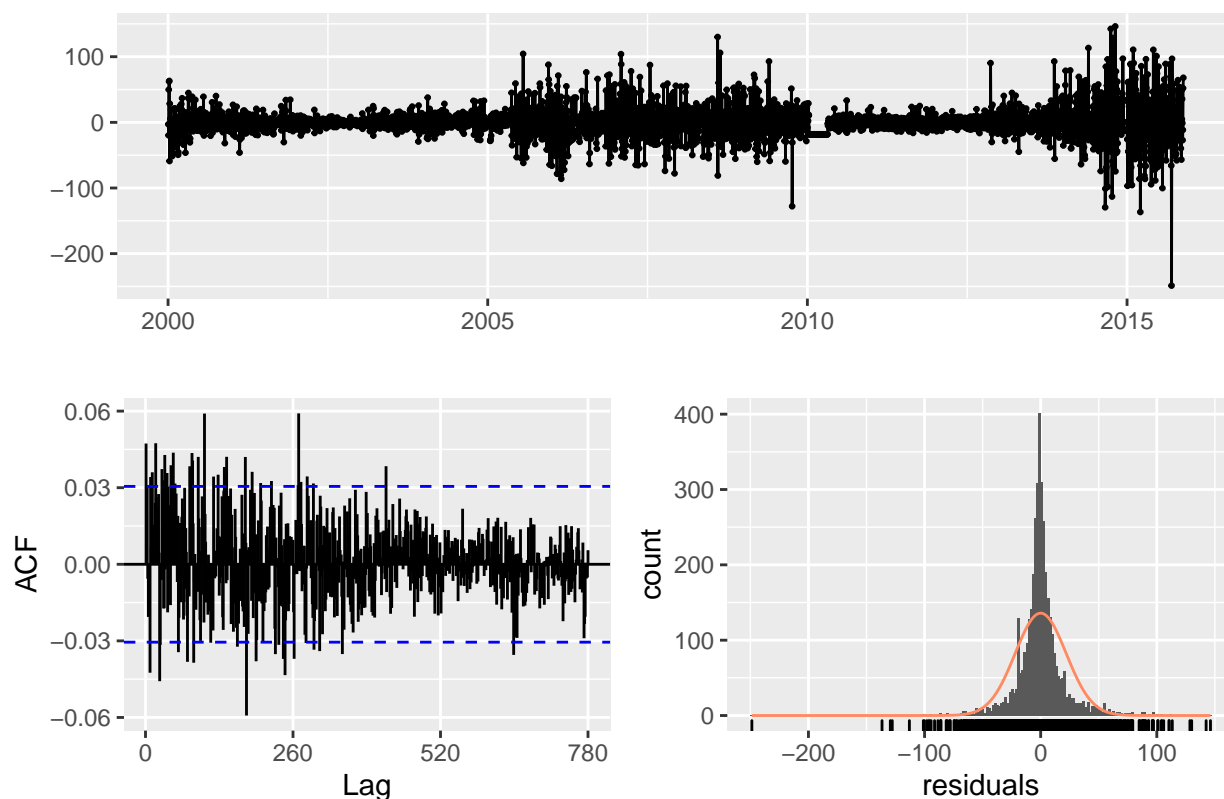


#### pvalue of Ljung-Box test  $\leq 0.05$  hence we reject the null hypothesis therefore there is no co-relation in the residuals therefore it is not stationary. pvalue of Shapiro-Wilk test  $\leq 0.05$  hence we reject the null hypothesis therefore residuals not normally distributed

### Residue analysis of model 3

```
checkresiduals(model_3)
```

## Residuals from Random walk with drift



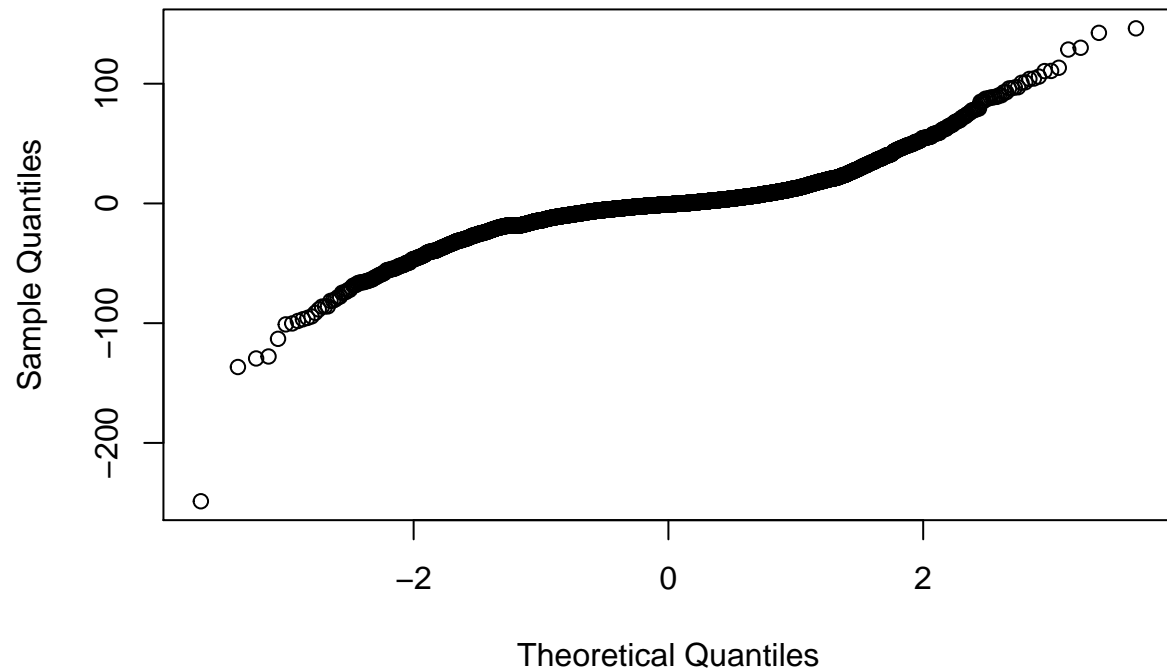
```
##
##  Ljung-Box test
##
## data:  Residuals from Random walk with drift
## Q* = 764.1, df = 519, p-value = 1.197e-11
##
## Model df: 1.   Total lags used: 520
```

```
shapiro.test(model_3$residuals)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  model_3$residuals
## W = 0.86267, p-value < 2.2e-16
```

```
qqnorm(model_3$residuals)
```

### Normal Q-Q Plot

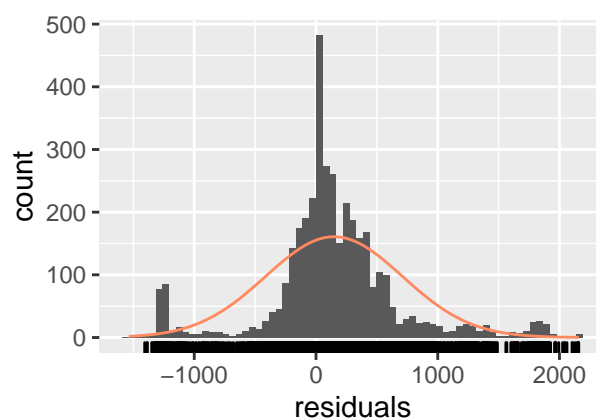
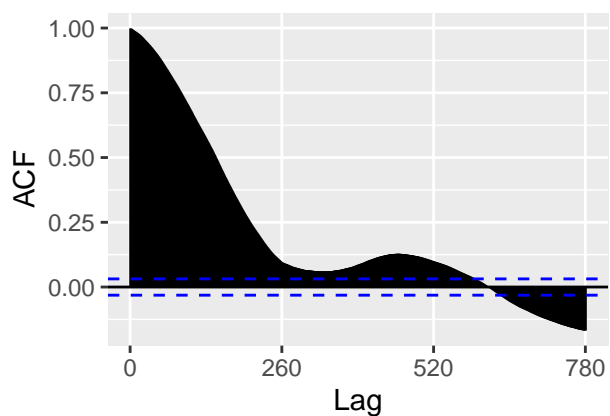
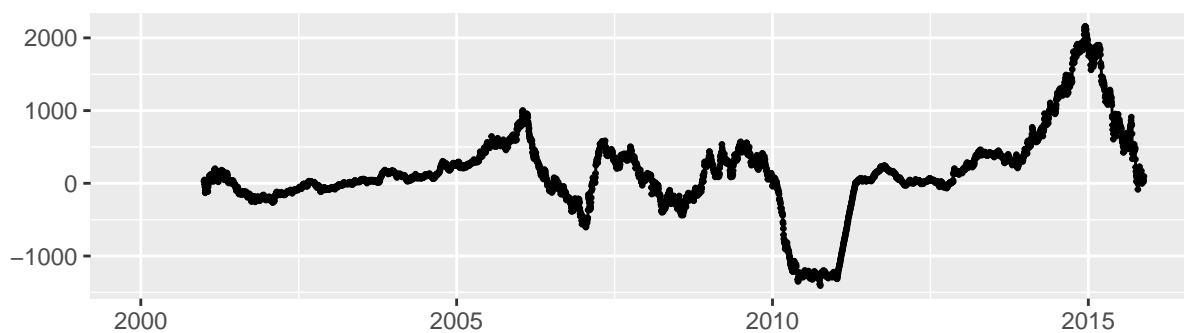


#### pvalue of Ljung-Box test  $\leq 0.05$  hence we reject the null hypothesis therefore there is no co-relation in the residuals therefore it is not stationary. pvalue of Shapiro-Wilk test  $\leq 0.05$  hence we reject the null hypothesis therefore residuals not normally distributed

### Residue analysis of model 4

```
checkresiduals(model_4)
```

## Residuals from Seasonal naive method



```
##
##  Ljung-Box test
##
## data:  Residuals from Seasonal naive method
## Q* = 418620, df = 520, p-value < 2.2e-16
##
## Model df: 0.   Total lags used: 520
```

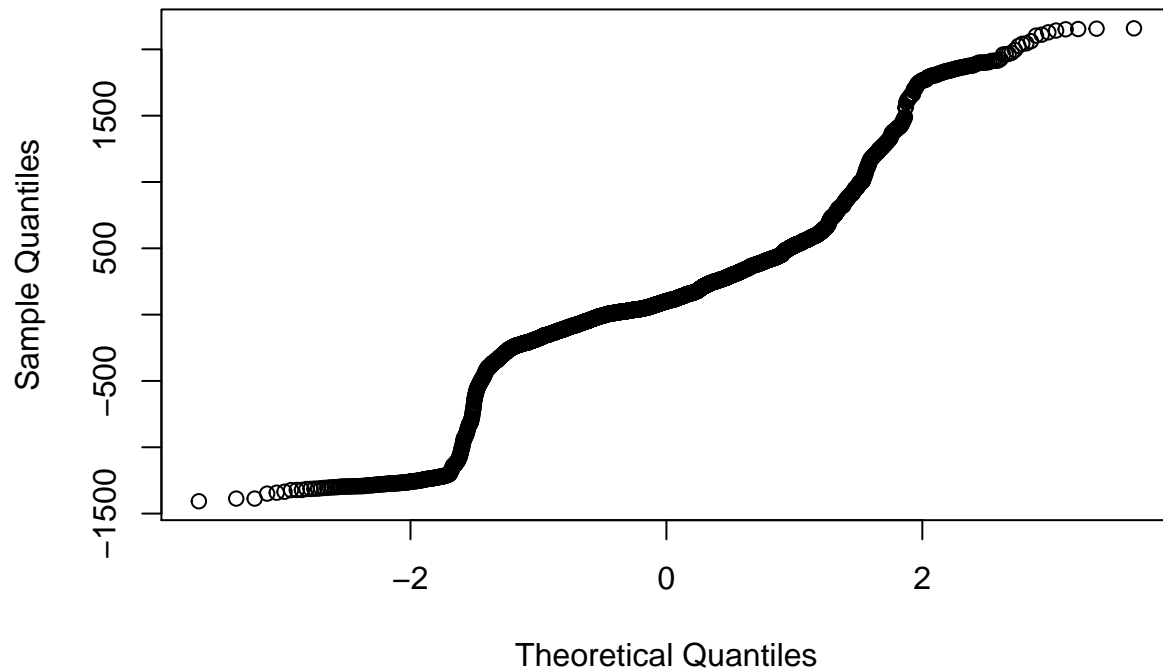
```
shapiro.test(model_4$residuals)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  model_4$residuals
## W = 0.90219, p-value < 2.2e-16
```

```
qqnorm(model_4$residuals)
```



## Normal Q-Q Plot



#### pvalue of Ljung-Box test  $\leq 0.05$  hence we reject the null hypothesis therefore there is no co-relation in the residuals therefore it is not stationary. pvalue of Shapiro-Wilk test  $\leq 0.05$  hence we reject the null hypothesis therefore residuals not normally distributed

## Stationarity of the series

We will use tranformation

```
log_data = log(my_time)
lambda = BoxCox.lambda(my_time)
Box_data = BoxCox(my_time, lambda = lambda)
```

```
autoplot(log_data)+ggtitle("Log transformation of the data")
```

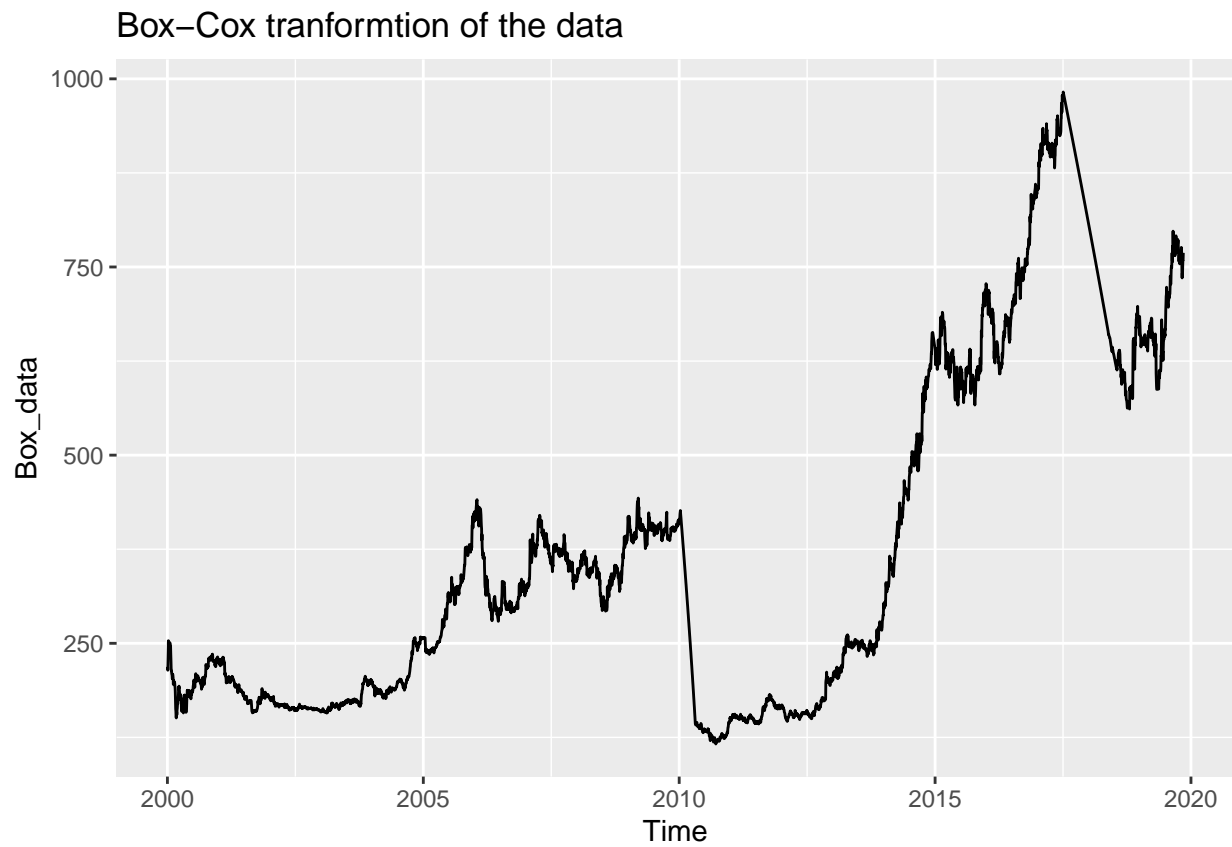
```
## Warning in is.na(main): is.na() applied to non-(list or vector) of type 'NULL'
```

## Log transformation of the data



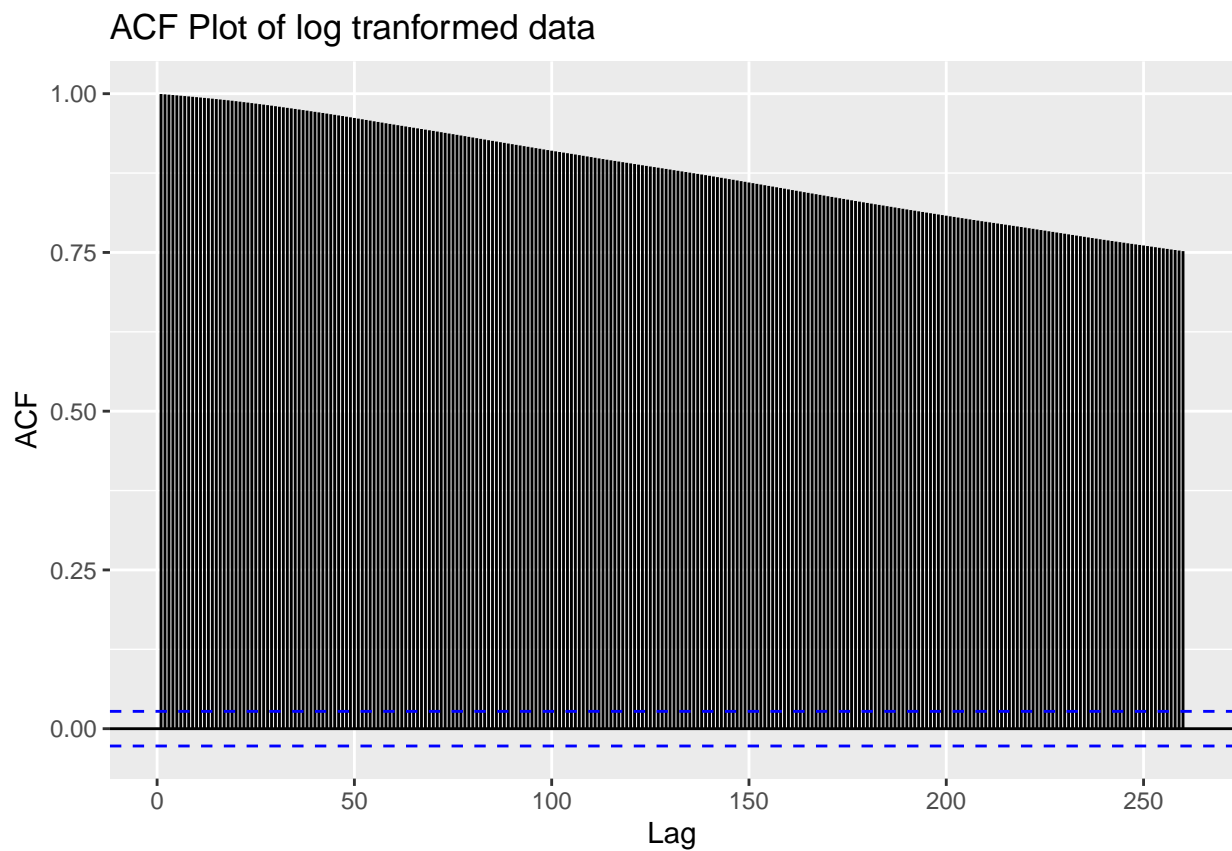
```
autoplot(Box_data)+ggtitle("Box-Cox tranformtion of the data")
```

```
## Warning in is.na(main): is.na() applied to non-(list or vector) of type 'NULL'
```

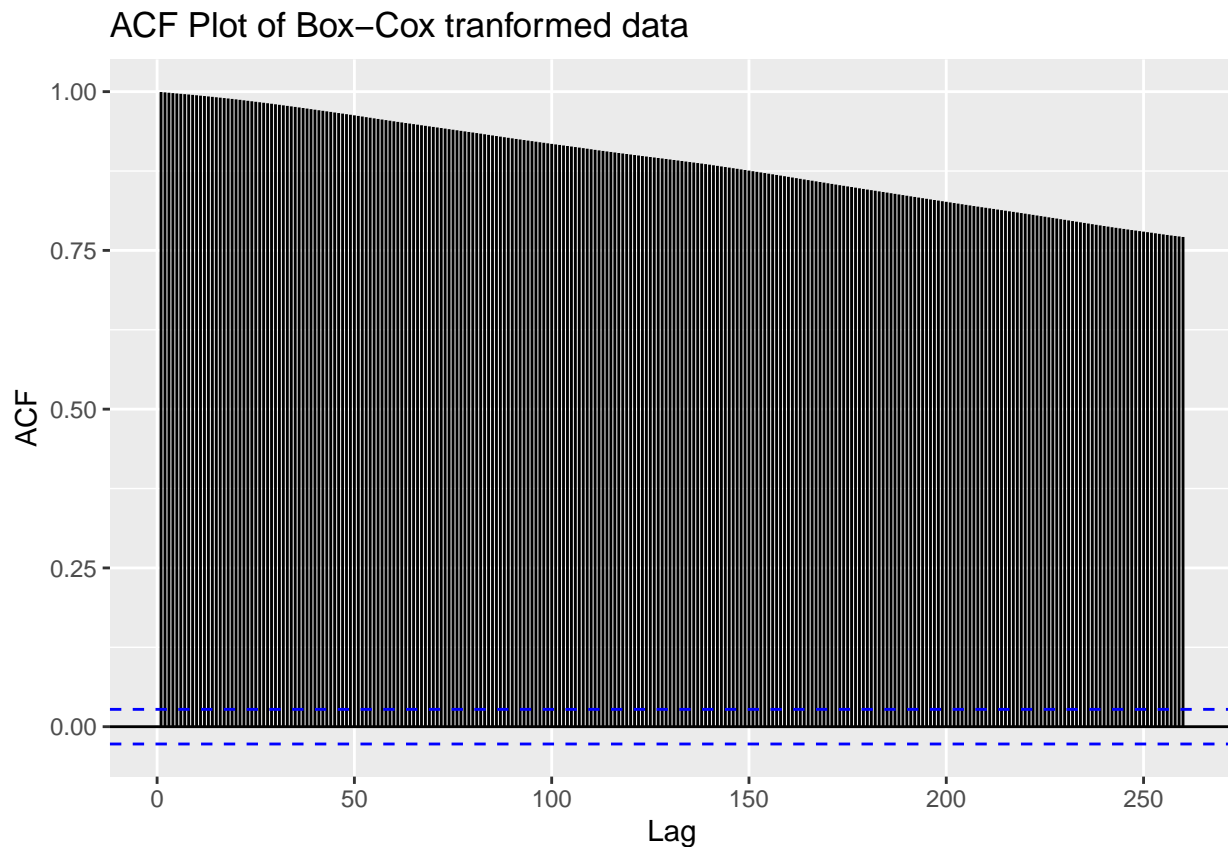


Now we will check the acf plot of tranformed data

```
ggAcf(log_data,lag.max = 260)+ggtitle("ACF Plot of log tranformed data")
```



```
ggAcf(Box_data,lag.max = 260)+ggtitle("ACF Plot of Box-Cox tranformed data")
```



#### Transformed data doesn't seem to work well

## Decomposing the data

The following two structures are considered for basic decomposition models:

Additive: = Trend + Seasonal + Random

Multiplicative: = Trend \* Seasonal \* Random

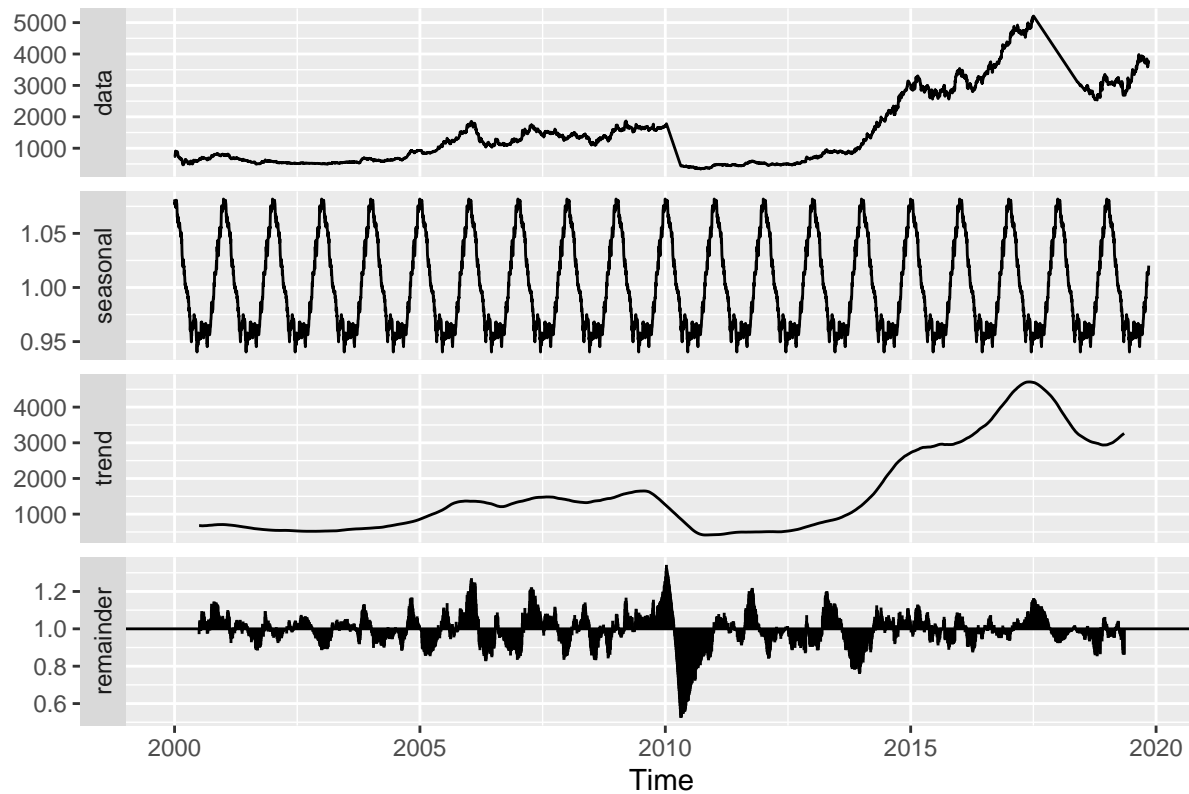
## How to Choose Between Additive and Multiplicative Decompositions

The additive model is useful when the seasonal variation is relatively constant over time.

The multiplicative model is useful when the seasonal variation increases over time.

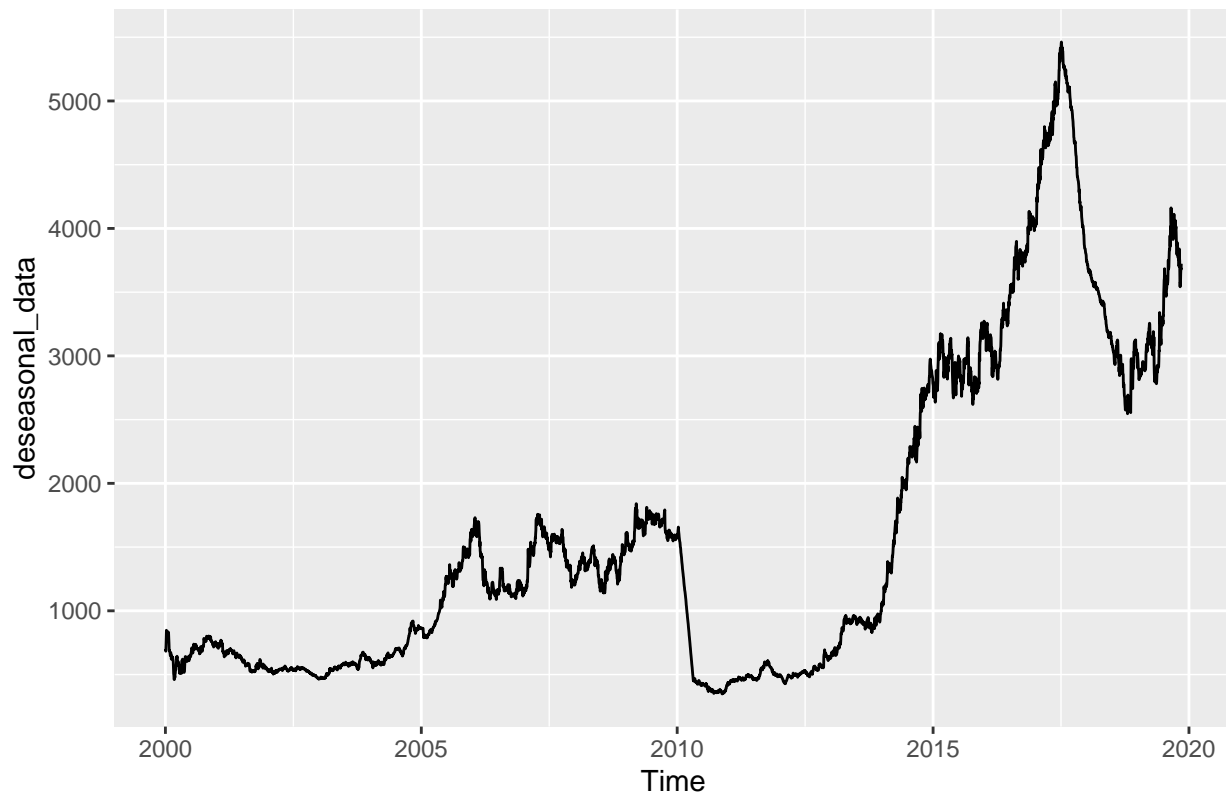
```
decomp_data = decompose(my_time,type = "multiplicative")
autoplot(decomp_data)
```

## Decomposition of multiplicative time series



```
deseasonal_data = my_time/decomp_data$seasonal  
autoplot(deseasonal_data)
```

```
## Warning in is.na(main): is.na() applied to non-(list or vector) of type 'NULL'
```



## Differencing ideas

lets try for differencing methods for the deseasonal data

```
ndiffs(deseasonal_data, test = "kpss")
```

```
## [1] 1
```

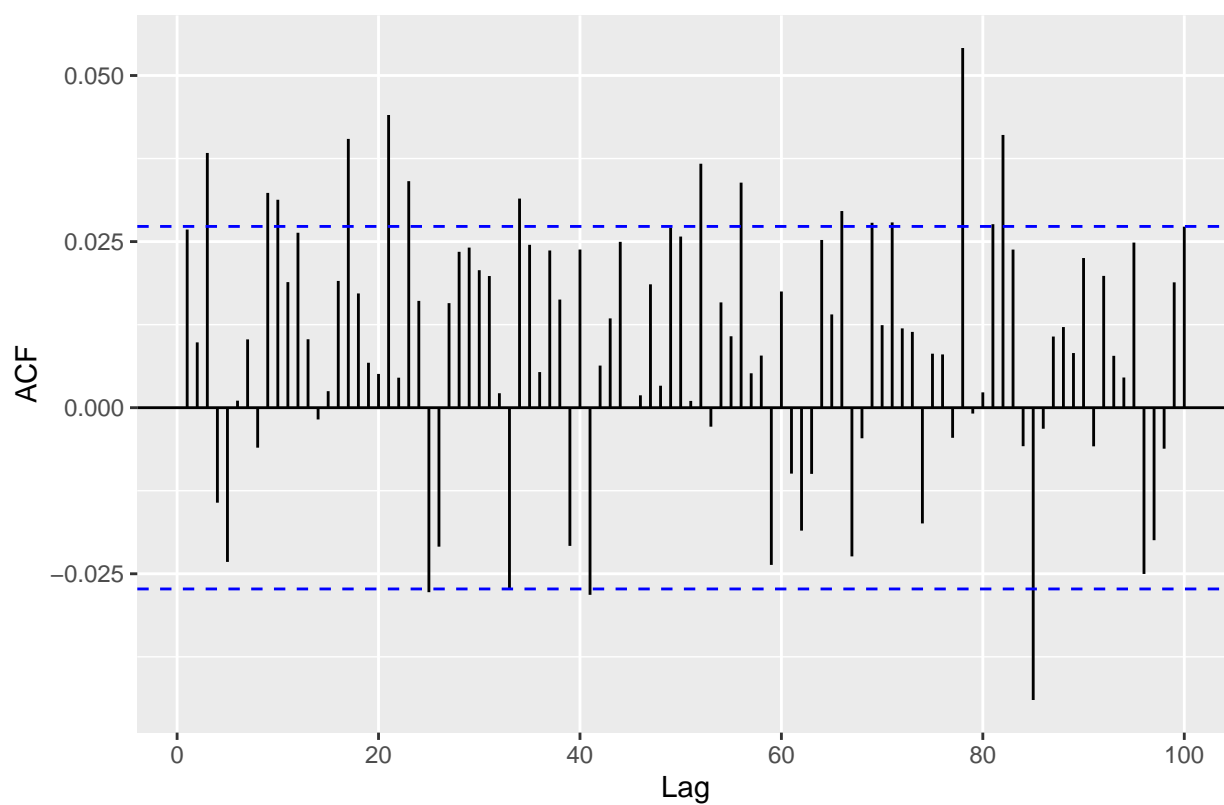
```
# using kpss test to find number of differencing required for the data
```

```
first_order = diff(deseasonal_data, 1)
```

Now we will look onto the acf plots of differenced data

```
ggAcf(first_order, lag.max = 100) + ggtitle("ACF Plot of first order difference of deseasonal data")
```

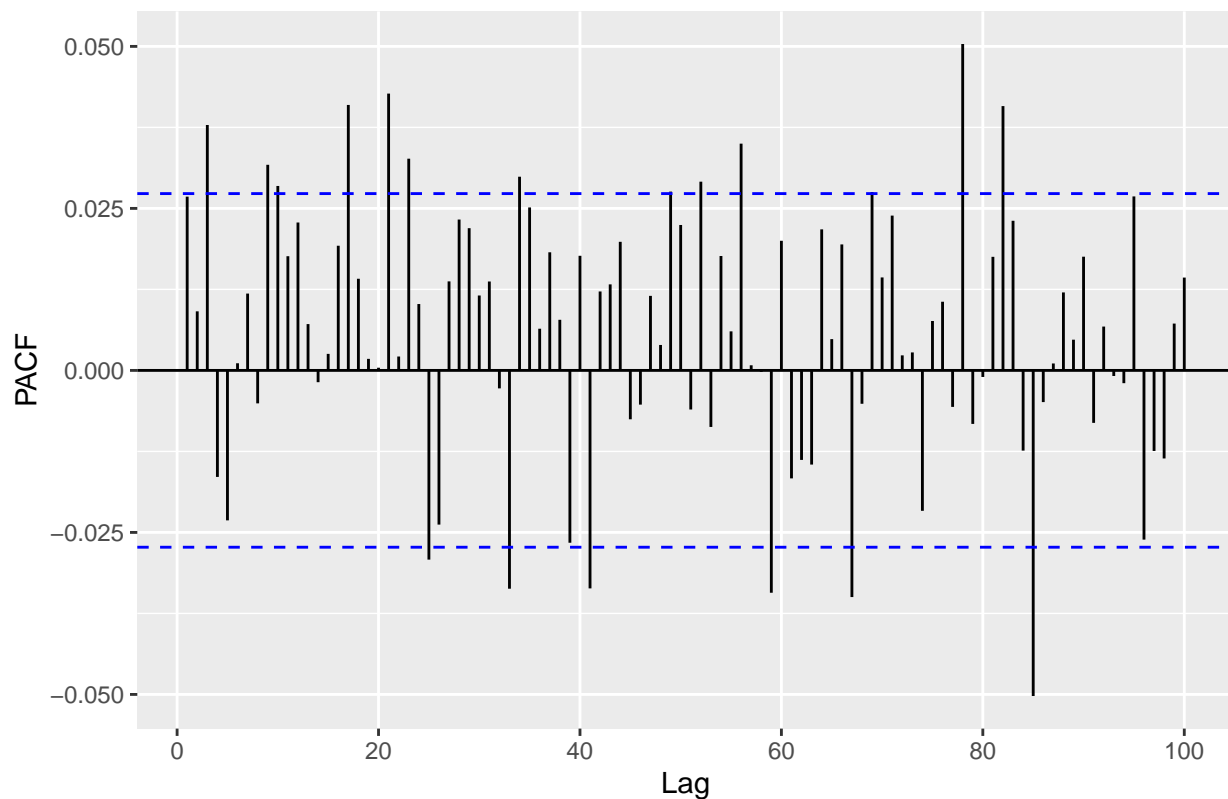
ACF Plot of first order difference of deseasonal data



```
ggPacf(first_order,lag.max = 100)+ggtitle("PACF Plot of first order difference of deseasonal data")
```



PACF Plot of first order difference of deseasonal data



#### We observe significant difference but still we are not sure whether the series is stationry or not.  
 #### We will use ADF test to determine whether our diffrenced series is statinary or not

```
adf.test(first_order,alternative = "stationary")
```

```
## Warning in adf.test(first_order, alternative = "stationary"): p-value smaller  
## than printed p-value
```

```
##  
## Augmented Dickey-Fuller Test  
##  
## data: first_order  
## Dickey-Fuller = -14.669, Lag order = 17, p-value = 0.01  
## alternative hypothesis: stationary
```

p value is less than 0.05 hence hence series is stationary

## ARIMA model

From PACF plot of first order we see suggestive AR(3) and AR(9) so our intial models will ARIMA(3,1,0) ARIMA(9,1,0) as differenced lag is 1

```
fit_1 = Arima(deseasonal_data,order = c(3,1,0))  
fit_1
```

```
## Series: deseasonal_data  
## ARIMA(3,1,0)  
##
```

```

## Coefficients:
##          ar1      ar2      ar3
##      0.0267  0.0085  0.0383
## s.e.  0.0139  0.0139  0.0139
##
## sigma^2 estimated as 723.2:  log likelihood=-24310.87
## AIC=48629.75  AICc=48629.75  BIC=48655.94
fit_2 = Arima(deseasonal_data,order = c(9,1,0))
fit_2

## Series: deseasonal_data
## ARIMA(9,1,0)
##
## Coefficients:
##          ar1      ar2      ar3      ar4      ar5      ar6      ar7      ar8      ar9
##      0.0271  0.0094  0.0389 -0.0153 -0.0224  0.000  0.0121 -0.0055  0.0326
## s.e.  0.0139  0.0139  0.0139  0.0139  0.0140  0.014  0.0140  0.0140  0.0140
##
## sigma^2 estimated as 722.6:  log likelihood=-24305.71
## AIC=48631.42  AICc=48631.46  BIC=48696.91
fit_3 = Arima(deseasonal_data,order = c(3,1,1))
fit_3

## Series: deseasonal_data
## ARIMA(3,1,1)
##
## Coefficients:
##          ar1      ar2      ar3      ma1
##      -0.1585  0.0135  0.0409  0.1855
## s.e.   0.2387  0.0155  0.0139  0.2387
##
## sigma^2 estimated as 723.3:  log likelihood=-24310.57
## AIC=48631.14  AICc=48631.16  BIC=48663.89
fit_4 = Arima(deseasonal_data,order = c(9,1,1))
fit_4

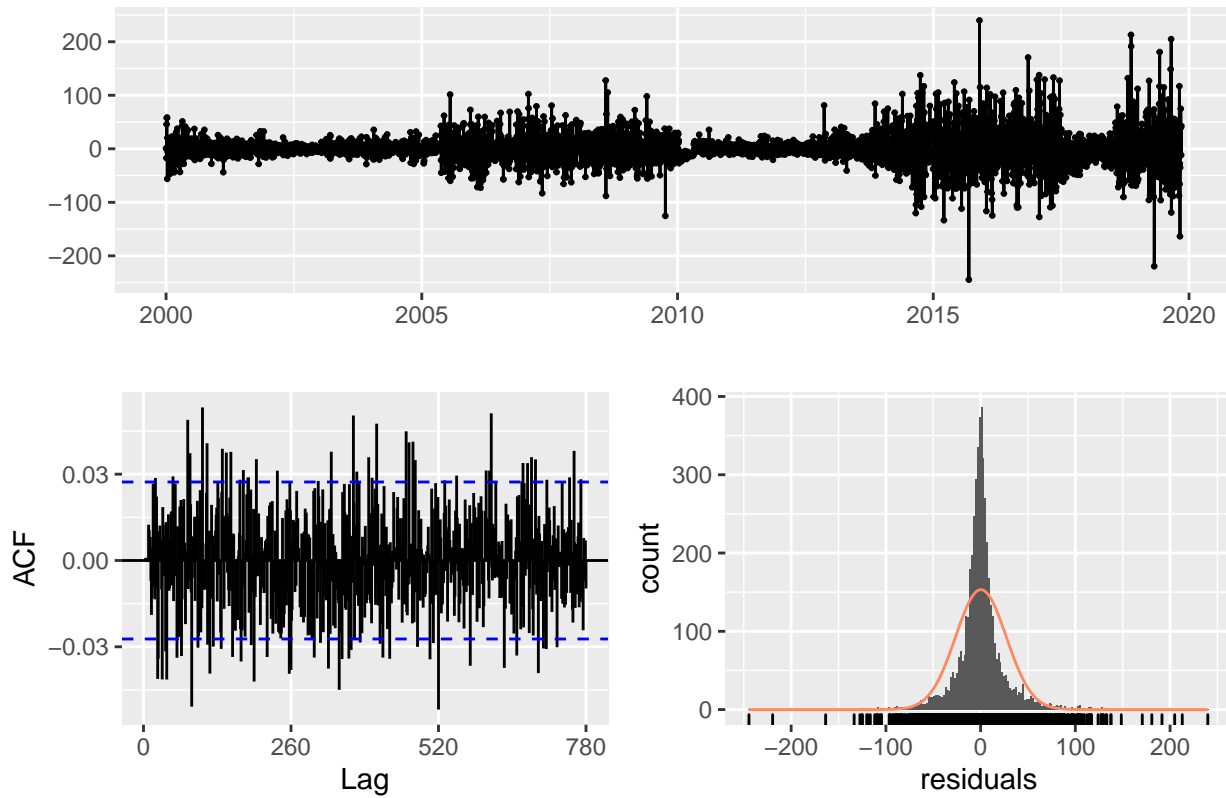
## Series: deseasonal_data
## ARIMA(9,1,1)
##
## Coefficients:
##          ar1      ar2      ar3      ar4      ar5      ar6      ar7      ar8
##      0.9911 -0.0164  0.0294 -0.0532 -0.0072  0.0225  0.0110 -0.0172
## s.e.  0.0188  0.0196  0.0196  0.0196  0.0196  0.0196  0.0196  0.0197
##          ar9      ma1
##      0.0240 -0.9696
## s.e.  0.0145  0.0127
##
## sigma^2 estimated as 719.4:  log likelihood=-24293.71
## AIC=48609.42  AICc=48609.47  BIC=48681.46

```

From the 4 models we AIC of model 4 seem to be less compared to other

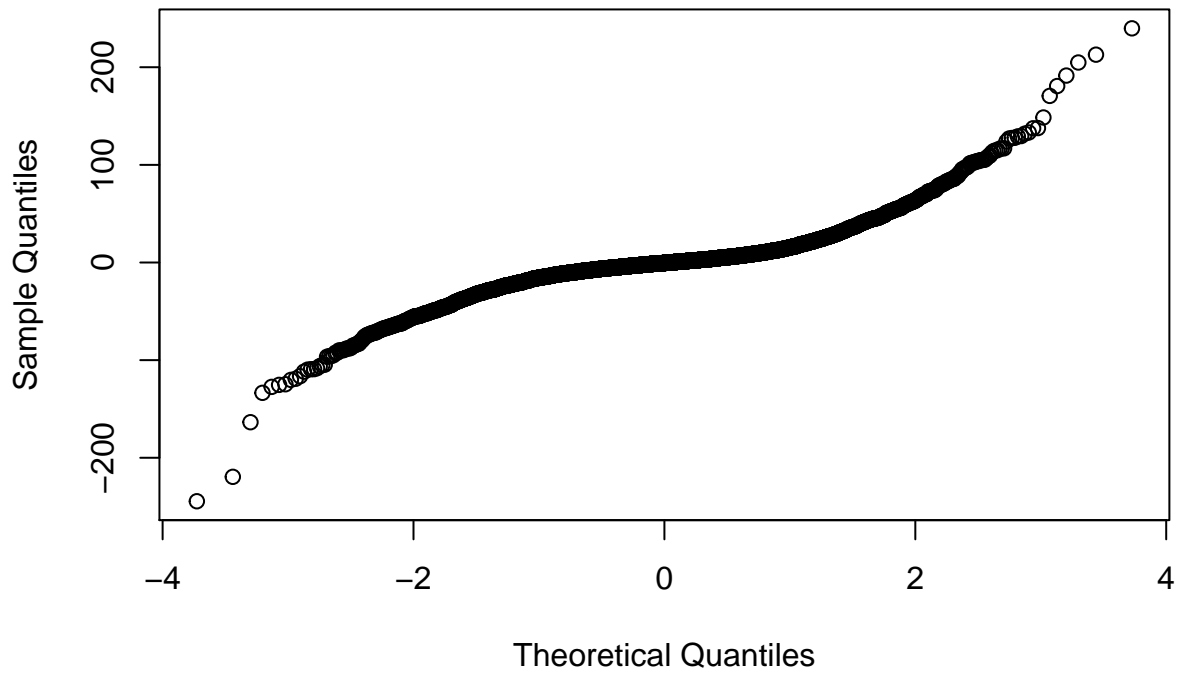
```
checkresiduals(fit_4)
```

### Residuals from ARIMA(9,1,1)



```
##  
##  Ljung-Box test  
##  
## data:  Residuals from ARIMA(9,1,1)  
## Q* = 902.93, df = 510, p-value < 2.2e-16  
##  
## Model df: 10.    Total lags used: 520  
qqnorm(fit_4$residuals)
```

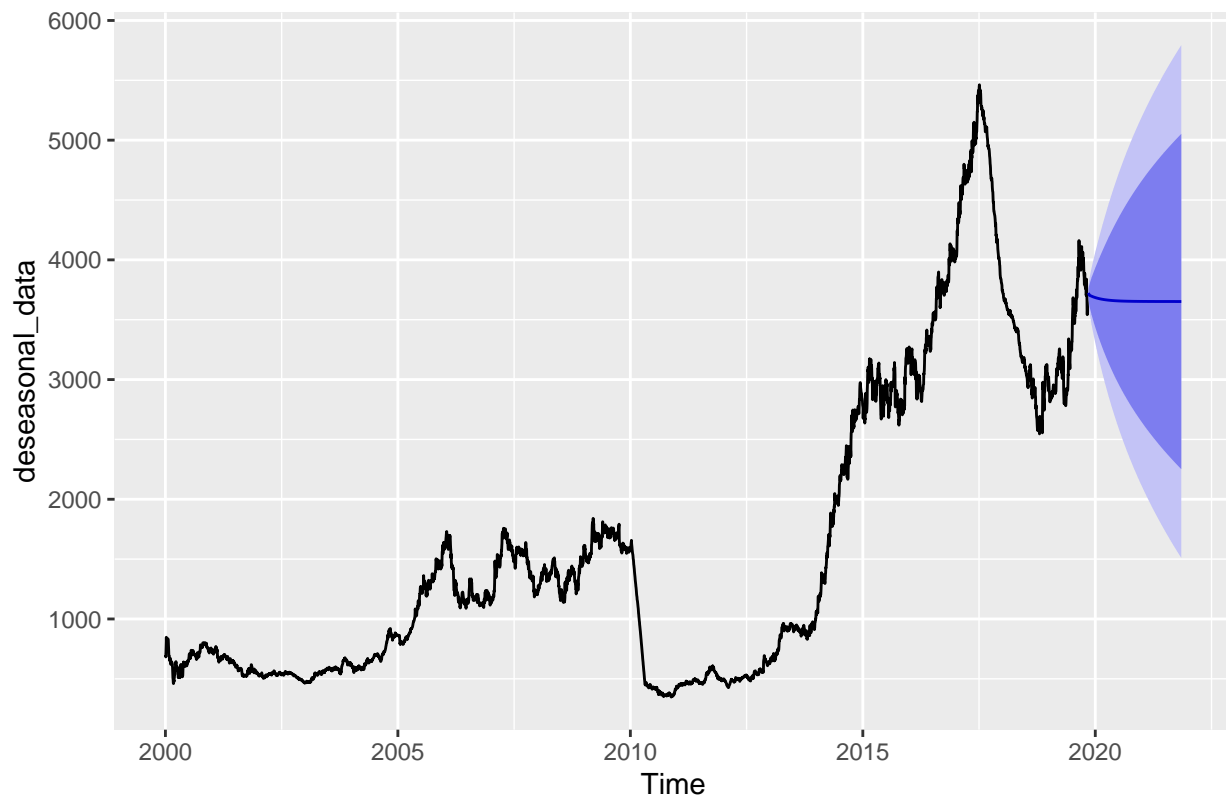
Normal Q-Q Plot



```
# normality of the ARIMA(9,1,1) is not acheived
```

```
autoplot(forecast(fit_4)) # forecast using ARIMA(9,1,1)
```

Forecasts from ARIMA(9,1,1)



```
fi = auto.arima(deseasonal_data,seasonal = FALSE)
fi

## Series: deseasonal_data
## ARIMA(0,1,1) with drift
##
## Coefficients:
##          ma1    drift
##          0.0264 0.5857
## s.e.    0.0138 0.3843
##
## sigma^2 estimated as 723.9:  log likelihood=-24313.77
## AIC=48633.54   AICc=48633.54   BIC=48653.19
```