1. Consider the data in file Alba.txt.

Data are from an experiment, comparing the potency of the two herbicides glyphosate and bentazone in white mustard Sinapis alba. Dose - a numeric vector containing the dose in g/ha. Herbicide - a factor with levels Bentazone Glyphosate (the two herbicides applied). DryMatter - a numeric vector containing the response (dry matter in g/pot).

Christensen, M. G. and Teicher, H. B., and Streibig, J. C. (2003) Linking fluorescence induction curve and biomass in herbicide screening, Pest Management Science, 59, 1303?1310.

Denote the variables as  $Y = \mathsf{DryMatter}$ ,  $X_1 = \mathsf{Dose}$ , and  $X_2 = \mathsf{Herbicide}$ .

(a) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the models

$$\begin{split} \mathcal{M}_{\text{identity}}: \quad & \mu_i = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\log}: \quad & \log(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\text{inverse}}: \quad & \frac{1}{\mu_i} = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\frac{1}{\mu^2}}: \quad & \frac{1}{\mu^2} = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \end{split}$$

where index j is related to the categories of  $X_2$ . Which model fits the best to the data? You may use the mean square error value  $\mathrm{MSE} = \frac{\sum_{i=1}^n (y_i - \hat{\mu}_i)^2}{n}$  as one method of choosing the model. Return the maximum likelihood estimate  $\hat{\beta}_1$  of your chosen model as your solution for this question.

(2 points)

(b) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathfrak{M}: \quad g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}.$$

Choose the link function g based on your solution to (a). Calculate the maximum likelihood estimate for the expected value  $\mu_{i_*}$  when  $X_1=50$  and  $X_2=\mathsf{Glyphosate}$ .

(c) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathfrak{M}: \quad g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_i + \gamma_i x_{i1}.$$

Choose the link function g based on your solution to (a). Calculate the 95% confidence interval estimate for the expected value  $\mu_{i_*}$  when  $X_1=50$  and  $X_2=$  Glyphosate. Particularly, what is your obtained lower bound of the confidence interval?

(1 point)

(d) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathcal{M}: \quad g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_i + \gamma_i x_{i1}.$$

Choose the link function g based on your solution to (a). Create 80 % prediction interval for new observation  $Y_f$ , when  $X_1 = 50$  and  $X_2 = \text{Glyphosate}$ . Particularly, what is your obtained lower bound of the prediction interval? (2 points)

2. Consider the following data set ratstime.txt:

	time	poison	treat
1	0.31	I	Α
2	0.82	I	В
3	0.43	I	C
4	0.45	I	D
5	0.45	I	Α
6	1.10	I	В

Effect of toxic agents on rats

Description

An experiment was conducted as part of an investigation to combat the effects of certain toxic agents.

A data frame with 48 observations on the following 3 variables.

time

survival time in tens of hours

poison

the poison type - a factor with levels I II III

treat

the treatment - a factor with levels A B C D

The response variable is Y = time and the explanatory variables are  $X_1 = \text{poison}$  and  $X_2 = \text{treat}$ .

(a) Let us assume  $Y_i \sim N(\mu_{jh}, \sigma^2)$  and let us model the data with the main effect model

$$g(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h.$$

Based on the mean square error value  $MSE = \frac{\sum_{i=1}^{n} (y_i - \hat{\mu}_{jh})^2}{n}$ , which link function fits best to the data:

- i.  $\mu_{jh} = \beta_0 + \beta_j + \alpha_h$ ,
- ii.  $\log(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h$ ,
- iii.  $\frac{1}{\mu_{jh}} = \beta_0 + \beta_j + \alpha_h$ ?

(2 points)

- (b) Distributional assumption could be either  $Y_i \sim N(\mu_{jh}, \sigma^2)$ ,  $Y_i \sim Gamma(\mu_{jh}, \phi)$ , or  $Y_i \sim IG(\mu_{jh}, \phi)$ . Choose the link function g based on your solution to (a). Based on your analysis, which one is most suitable in this case?
  - i.  $Y_i \sim N(\mu_{jh}, \sigma^2)$ ,
  - ii.  $Y_i \sim Gamma(\mu_{jh}, \phi)$ ,
  - iii.  $Y_i \sim IG(\mu_{ih}, \phi)$ .

(1 point)

(c) Regardless what was your solution to the question (a) and (b), let us assume  $Y_i \sim IG(\mu_{jh}, \phi)$  and consider the model

$$\log(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h.$$

Create 95 % confidence interval for  $\mu_{jh}$ , when

$$x_{i*1} = II, x_{i*2} = B.$$

Particularly, what is your obtained lower bound of the confidence interval? (1 point)

(d) Regardless what was your solution to the question (a) and (b), let us assume  $Y_i \sim Gamma(\mu_{jh}, \phi)$  and consider the model

$$\log(\mu_{ih}) = \beta_0 + \beta_i + \alpha_h.$$

Create 80 % prediction interval for new observation  $Y_f$ , when

$$x_{f1} = II, \qquad x_{f2} = B.$$

Particularly, what is your obtained lower bound of the prediction interval? (2 points)

3. (a) In case of generalized linear model  $g(\mu_i) = \beta_0 + \beta_1 x_i$ , find the inverse function  $g^{-1}$  (that is, solve what form the expected value  $\mu_i$  has), when the link function g is

i. 
$$\sqrt{\mu_i} = \beta_0 + \beta_1 x_i$$
,

ii. 
$$\frac{1}{\mu_i^2} = \beta_0 + \beta_1 x_i$$
,

iii. 
$$\log\left(\frac{\mu_i}{1-\mu_i}\right) = \beta_0 + \beta_1 x_i$$
.

(2 points)

(b) Let us assume  $Y_i \sim IG(\mu_i, \phi)$ . Consider the model

$$\log(\mu_i) = \beta_0 + \beta_1 \log(x_i).$$

Let the estimates of the parameters  $\beta_0, \beta_1, \phi$  be as  $\hat{\beta}_0 = 1, \hat{\beta}_1 = 0.5, \tilde{\phi} = 0.05$ .

- i. Calculate the maximum likelihood estimate for the expected value  $\mu_i$  when  $x_i = 5$ .
- ii. Calculate the Pearson residual

$$o_i = \frac{y_i - \hat{\mu}_i}{\sqrt{\widehat{\operatorname{Var}}(Y_i)}},$$

when  $x_i = 5$  and the observed value is  $y_i = 12$ .

(2 points)

(c) In generalized linear models, the likelihood equations can written in form

$$\frac{\partial l(\boldsymbol{\beta}, \phi)}{\partial \beta_j} = \sum_{i=1}^n \frac{(y_i - \mu_i)}{\operatorname{Var}(Y_i)} \cdot x_{ij} \cdot \left(\frac{\partial \mu_i}{\partial \eta_i}\right) = 0, \qquad j = 0, 1, 2 \dots p.$$

Consider now the simple Gamma model with

$$Y_i \sim Gamma(\mu_i, \phi),$$
  
 $\mu_i = \eta_i = \beta_0.$ 

What kind of more simplified form the likelihood equations have in this case? That is, what form  $\frac{\partial l(\beta_0)}{\partial \beta_0}$  has in the simple Gamma model? By using the likelihood equations, find the maximum likelihood estimator  $\hat{\beta}_0$ .

(2 points)