

1. Consider the data in file [Alba.txt](#).

```
> data<-read.table("Alba.txt", sep="\t", dec=".", header=TRUE)
> head(data)
  Dose Herbicide DryMatter
1    1 Glyphosate    4.7
2    1 Glyphosate    4.6
3    1 Glyphosate    4.1
4    1 Glyphosate    4.4
```

Data are from an experiment, comparing the potency of the two herbicides glyphosate and bentazone in white mustard *Sinapis alba*.  
Dose - a numeric vector containing the dose in g/ha.  
Herbicide - a factor with levels Bentazone Glyphosate (the two herbicides applied).  
DryMatter - a numeric vector containing the response (dry matter in g/pot).  
Christensen, M. G. and Teicher, H. B., and Streibig, J. C. (2003)  
Linking fluorescence induction curve and biomass in herbicide screening,  
Pest Management Science, 59, 1303-1310.

Denote the variables as  $Y = \text{DryMatter}$ ,  $X_1 = \text{Dose}$ , and  $X_2 = \text{Herbicide}$ .

- (a) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the models

$$\begin{aligned}\mathcal{M}_{\text{identity}} : \quad & \mu_i = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\log} : \quad & \log(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\text{inverse}} : \quad & \frac{1}{\mu_i} = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}, \\ \mathcal{M}_{\frac{1}{\mu^2}} : \quad & \frac{1}{\mu^2} = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1},\end{aligned}$$

where index  $j$  is related to the categories of  $X_2$ . Which model fits the best to the data? You may use the mean square error value  $\text{MSE} = \frac{\sum_{i=1}^n (y_i - \hat{\mu}_i)^2}{n}$  as one method of choosing the model. Return the maximum likelihood estimate  $\hat{\beta}_1$  of your chosen model as your solution for this question.

(2 points)

- (b) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathcal{M} : \quad g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}.$$

Choose the link function  $g$  based on your solution to (a). Calculate the maximum likelihood estimate for the expected value  $\mu_{i_*}$  when  $X_1 = 50$  and  $X_2 = \text{Glyphosate}$ . (1 point)

- (c) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathcal{M}: g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}.$$

Choose the link function  $g$  based on your solution to (a). Calculate the 95% confidence interval estimate for the expected value  $\mu_{i*}$  when  $X_1 = 50$  and  $X_2 = \text{Glyphosate}$ . Particularly, what is your obtained lower bound of the confidence interval?

(1 point)

- (d) Let us assume  $Y_i \sim N(\mu_i, \sigma^2)$ . Consider the model

$$\mathcal{M}: g(\mu_i) = \beta_0 + \beta_1 x_{i1} + \alpha_j + \gamma_j x_{i1}.$$

Choose the link function  $g$  based on your solution to (a). Create 80 % prediction interval for new observation  $Y_f$ , when  $X_1 = 50$  and  $X_2 = \text{Glyphosate}$ . Particularly, what is your obtained lower bound of the prediction interval?

(2 points)

2. Consider the following data set `ratstime.txt`:

```

time poison treat
1 0.31      I      A
2 0.82      I      B
3 0.43      I      C
4 0.45      I      D
5 0.45      I      A
6 1.10      I      B
.
.
```

Effect of toxic agents on rats

Description

An experiment was conducted as part of an investigation to combat the effects of certain toxic agents.

A data frame with 48 observations on the following 3 variables.

time  
survival time in tens of hours

poison  
the poison type - a factor with levels I II III

treat  
the treatment - a factor with levels A B C D

The response variable is  $Y = \text{time}$  and the explanatory variables are  $X_1 = \text{poison}$  and  $X_2 = \text{treat}$ .

- (a) Let us assume  $Y_i \sim N(\mu_{jh}, \sigma^2)$  and let us model the data with the main effect model

$$g(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h.$$

Based on the mean square error value  $MSE = \frac{\sum_{i=1}^n (y_i - \hat{\mu}_{jh})^2}{n}$ , which link function fits best to the data:

- i.  $\mu_{jh} = \beta_0 + \beta_j + \alpha_h$ ,
- ii.  $\log(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h$ ,
- iii.  $\frac{1}{\mu_{jh}} = \beta_0 + \beta_j + \alpha_h$ ?

(2 points)

- (b) Distributional assumption could be either  $Y_i \sim N(\mu_{jh}, \sigma^2)$ ,  $Y_i \sim \text{Gamma}(\mu_{jh}, \phi)$ , or  $Y_i \sim \text{IG}(\mu_{jh}, \phi)$ . Choose the link function  $g$  based on your solution to (a). Based on your analysis, which one is most suitable in this case?

- i.  $Y_i \sim N(\mu_{jh}, \sigma^2)$ ,
- ii.  $Y_i \sim \text{Gamma}(\mu_{jh}, \phi)$ ,
- iii.  $Y_i \sim \text{IG}(\mu_{jh}, \phi)$ .

(1 point)

- (c) Regardless what was your solution to the question (a) and (b), let us assume  $Y_i \sim \text{IG}(\mu_{jh}, \phi)$  and consider the model

$$\log(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h.$$

Create 95 % confidence interval for  $\mu_{jh}$ , when

$$x_{i*1} = II, \quad x_{i*2} = B.$$

Particularly, what is your obtained lower bound of the confidence interval?

(1 point)

- (d) Regardless what was your solution to the question (a) and (b), let us assume  $Y_i \sim \text{Gamma}(\mu_{jh}, \phi)$  and consider the model

$$\log(\mu_{jh}) = \beta_0 + \beta_j + \alpha_h.$$

Create 80 % prediction interval for new observation  $Y_f$ , when

$$x_{f1} = II, \quad x_{f2} = B.$$

Particularly, what is your obtained lower bound of the prediction interval?

(2 points)

3. (a) In case of generalized linear model  $g(\mu_i) = \beta_0 + \beta_1 x_i$ , find the inverse function  $g^{-1}$  (that is, solve what form the expected value  $\mu_i$  has), when the link function  $g$  is

- i.  $\sqrt{\mu_i} = \beta_0 + \beta_1 x_i$ ,
- ii.  $\frac{1}{\mu_i^2} = \beta_0 + \beta_1 x_i$ ,
- iii.  $\log\left(\frac{\mu_i}{1-\mu_i}\right) = \beta_0 + \beta_1 x_i$ .

(2 points)

- (b) Let us assume  $Y_i \sim IG(\mu_i, \phi)$ . Consider the model

$$\log(\mu_i) = \beta_0 + \beta_1 \log(x_i).$$

Let the estimates of the parameters  $\beta_0, \beta_1, \phi$  be as  $\hat{\beta}_0 = 1, \hat{\beta}_1 = 0.5, \tilde{\phi} = 0.05$ .

- i. Calculate the maximum likelihood estimate for the expected value  $\mu_i$  when  $x_i = 5$ .
- ii. Calculate the Pearson residual

$$o_i = \frac{y_i - \hat{\mu}_i}{\sqrt{\widehat{\text{Var}}(Y_i)}},$$

when  $x_i = 5$  and the observed value is  $y_i = 12$ .

(2 points)

- (c) In generalized linear models, the likelihood equations can be written in form

$$\frac{\partial l(\boldsymbol{\beta}, \phi)}{\partial \beta_j} = \sum_{i=1}^n \frac{(y_i - \mu_i)}{\text{Var}(Y_i)} \cdot x_{ij} \cdot \left( \frac{\partial \mu_i}{\partial \eta_i} \right) = 0, \quad j = 0, 1, 2, \dots, p.$$

Consider now the simple Gamma model with

$$Y_i \sim \text{Gamma}(\mu_i, \phi),$$

$$\mu_i = \eta_i = \beta_0.$$

What kind of more simplified form the likelihood equations have in this case? That is, what form  $\frac{\partial l(\beta_0)}{\partial \beta_0}$  has in the simple Gamma model? By using the likelihood equations, find the maximum likelihood estimator  $\hat{\beta}_0$ .

(2 points)