

K436765

01. Consider the linear model

Let us consider the estimation of β .
Show or explain in detail, one way
or another, how the maximum
likelihood estimator

$$\hat{\beta} = (X'X)^{-1}X'Y \quad \text{is obtained.}$$

Let us consider the estimation of
unknown parameters of the linear
model, $y = X\beta + \epsilon$, $\epsilon \sim N(0, \sigma^2 I)$

Since $y \sim N(X\beta, \sigma^2 I)$ the
likelihood function $L(\beta, \sigma^2/y)$ &
log likelihood function $l(\beta, \sigma^2/y)$
have the form

$$L(\beta, \sigma^2/y) = \frac{1}{\sqrt{(2\pi)^n \sigma^2 I}} \times \exp \left\{ -\frac{(y - X\beta)'(y - X\beta)}{2\sigma^2} \right\}$$

$$l(\beta, \sigma^2/y) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{(y - XB)'(y - XB)}{2\sigma^2}$$

The maximum likelihood estimators

$\hat{\beta} = (\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p)$ of the parameters vector β & $\hat{\sigma}^2$ of the scalar $\sigma^2 > 0$ can be obtained by finding the maximum of the log-likelihood function $l(\beta, \sigma^2/y)$ with respect to β, σ^2 by the process of differentiation.

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$$\hat{\beta}, \hat{\sigma}^2 = \arg \max_{\beta, \sigma^2} l(\beta, \sigma^2/y) = \arg \max_{\beta, \sigma^2} \left(-\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} (y - XB)'(y - XB) \right)$$

The local maximum (β, σ^2) with respect to β occurs at point, where the gradient vector (partial derivatives)

$$\frac{\partial l}{\partial \beta} = \begin{pmatrix} \frac{\partial l}{\partial \beta_0} \\ \frac{\partial l}{\partial \beta_1} \\ \frac{\partial l}{\partial \beta_p} \end{pmatrix}$$

$$\frac{\partial l}{\partial \beta} = \frac{\partial \left(-\frac{1}{2\sigma^2} (y - XB)'(y - XB) \right)}{\partial \beta}$$

$$= \left(-\frac{1}{2\sigma^2} \right) \frac{\partial (y'XB)}{\partial \beta} + \left(-\frac{1}{2\sigma^2} \right) \frac{\partial (X'XB)}{\partial \beta}$$

$$= \left(-\frac{1}{2\sigma^2} \right) (-2X'y) + \left(-\frac{1}{2\sigma^2} \right) 2X'XB$$

$$= -\frac{1}{2\sigma^2} (X'y - X'XB) = 0$$

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Since $\frac{\partial l}{\partial \beta} = 0$ for any $\beta^2 > 0$ if
 & only if the normal equations

$$X'X\beta = X'y \quad \text{holds for } \beta$$

The maximum likelihood estimator $\hat{\beta}$ is
 the solution to the normal equations

$$X'X\hat{\beta} = X'y$$

$$\text{i.e. } \hat{\beta} = (X'X)^{-1}X'y.$$

2 In generalized linear model, the likelihood equations can written ~~for~~ form

$$\frac{\partial l(\beta, \phi)}{\partial \beta_j} = \sum_{i=1}^n \frac{(y_i - \mu_i)}{\text{var}(\phi_i)} \eta_j' \left(\frac{\partial \mu_i}{\partial \eta_i} \right) = 0$$

$$j = 0, 1, 2, \dots, n$$

Consider now the single log-linear ~~inverse~~ Inverse Gaussian model with

$$Y_i \sim \text{IG}(\mu, \phi)$$

$$\log(\mu_i) = \eta_i' \beta$$

What kind of more simplified form the likelihood equations have in this case? That is, what form $\frac{\partial l(\beta)}{\partial \beta}$ has in the

single Inverse Gaussian model? By using the likelihood equations, find the maximum likelihood estimator $\hat{\beta}$

Hence, $\log(\mu_i) = \eta_i = \beta_0$

$$Y_i \sim \text{IG}(\mu_i, \phi)$$

$$\frac{\partial l(\beta_0)}{\partial \beta_0} = \sum_{i=1}^n \frac{(y_i - \mu_i)}{\phi \mu_i^2} \mu_i^{1-1}$$

$$\mu_i = e^{\beta_0} = \frac{1}{\phi} \sum_{i=1}^n \frac{(y_i - \mu_i)}{\mu_i}$$

$$\frac{\partial l(\beta_0)}{\partial \beta_0} = \frac{1}{\phi} \sum_{i=1}^n \frac{(y_i - e^{\beta_0})}{e^{\beta_0}}$$

For the MLE

$$\frac{1}{\phi} \sum_{i=1}^n \frac{(y_i - e^{\beta_0})}{e^{\beta_0}} = 0$$

$$\sum_{i=1}^n (y_i - e^{\beta_0}) = 0$$

$$\sum_{i=1}^n y_i - n e^{\beta_0} = 0$$

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$$\beta_0 = \log \left(\frac{1}{n} \sum_{i=1}^n y_i \right)$$

Thus

$$\hat{\beta}_0 = \log(5)$$

This is the MLE for β_0 , for simple non-linear Inverse Gaussian model.

03

let us assume $\Gamma(\mu, \phi)$

Consider the model

$$\log(\mu_i) = \beta_0 + \beta_1 x_i$$

let the estimate of the parameter

β_0, β_1, ϕ be a) $\hat{\beta}_0 = 1$

$$\hat{\beta}_1 = 0.5$$

$$\hat{\phi} = 0.1$$

when

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$$X = \begin{pmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \\ 1 & 5 \\ 1 & 6 \\ 1 & 7 \\ 1 & 8 \\ 1 & 9 \\ 1 & 10 \end{pmatrix}$$

calculate the estimated
covariance matrix

$$\text{cov}(\hat{\beta}) = (X'WX)^{-1}$$