

Analyzing the impact of Abduction on Education, Distress, and wages in Uganda.

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MSBA Data Analytics III

Propensity Score Matching

For this problem, I will analyze the data from:

Christopher Blattman and J Annan. 2010. “The consequences of child soldiering” *Review of Economics and Statistics* 92 (4):882-898

The data are from a panel survey of male youth in war-afflicted regions of Uganda. The authors wanted to estimate the impact of forced military service on various outcomes. They focus on Uganda because there were a significant number of abductions of young men into Lord’s Resistance Army.

Blattman and Annan describe the abductions as follows:

Abduction was large-scale and seemingly indiscriminate; 60,000 to 80,000 youth are estimated to have been abducted and more than a quarter of males currently aged 14 to 30 in our study region were abducted for at least two weeks. Most were abducted after 1996 and from one of the Acholi districts of Gulu, Kitgum, and Pader.

Youth were typically taken by roving groups of 10 to 20 rebels during night raids on rural homes. Adolescent males appear to have been the most pliable, reliable and effective forced recruits, and so were disproportionately targeted by the LRA. Youth under age 11 and over 24 tended to be avoided and had a high probability of immediate release. Lengths of abduction ranged from a day to ten years, averaging 8.9 months in our sample. Youth who failed to escape were trained as fighters and, after a few months, received a gun. Two thirds of abductees were forced to perpetrate a crime or violence. A third eventually became fighters and a fifth were forced to murder soldiers, civilians, or even family members in order to bind them to the group, to reduce their fear of killing, and to discourage disobedience.

In this problem I will look at the effect of abduction on *educ* (years of education). The *abd* variable is the treatment in this case. Note that *educ*, *distress*, and *logwage* are all outcomes/post-treatment variables.

Variables	Description
abd	abducted by the LRA (the treatment)
c_ach - c_pal	Location indicators (each abbreviation corresponds to a subdistrict; i.e. ach = Acholibur, etc.)
age	age in years
fthr_ed	father’s education (years)
mthr_ed	mother’s education (years)
orphan96	indicator if parent’s died before 1996
hh_fthr_frm	indicator if father is a farmer
hh_size96	household size in 1996
educ	years of education
distress	index of emotional distress (0-15)

Variables	Description
logwage	log of average daily wage earned in last 4 weeks

1. Calculate the naive Average Treatment Effect (ATE) of abduction on education (educ), distress (distress), and wages (logwage). Do this by running three separate regressions.

Note, the predictor (explanatory) variable is 'abd', and we are regressing it (using it to predict) against education, distress and wages. (i.e. to see the effects that 'abd' has/had on each individual response variable - educ, distress, and wages).

```
blattman <- read.csv("~/Downloads/Fall 2023 - MSBA /MSBA 650/ICA 3/blattman.csv")
library(modelsummary)
```

```
## Version 2.0.0 of 'modelsummary', to be released soon, will introduce a
## breaking change: The default table-drawing package will be 'tinytable'
## instead of 'kableExtra'. All currently supported table-drawing packages
## will continue to be supported for the foreseeable future, including
## 'kableExtra', 'gt', 'huxtable', 'flextable, and 'DT'.
##
## You can always call the 'config_modelsummary()' function to change the
## default table-drawing package in persistent fashion. To try 'tinytable'
## now:
##
## config_modelsummary(factory_default = 'tinytable')
##
## To set the default back to 'kableExtra':
##
## config_modelsummary(factory_default = 'kableExtra')
```

```
reg1 <- lm(educ ~abd, data = blattman)
reg2 <- lm(distress ~ abd, data=blattman)
reg3 <- lm(log.wage ~ abd, data = blattman)
modelsummary(list("Education" = reg1, "Distress" = reg2, "Log wage" = reg3), stars=TRUE)
```

	Education	Distress	Log wage
(Intercept)	7.416*** (0.172)	3.759*** (0.144)	4.205*** (0.219)
abd	-0.595** (0.218)	0.593** (0.182)	0.272 (0.277)
Num.Obs.	741	741	741
R2	0.010	0.014	0.001
R2 Adj.	0.009	0.013	0.000
AIC	3672.4	3406.0	4028.4
BIC	3686.2	3419.8	4042.2
Log.Lik.	-1833.183	-1699.990	-2011.182
F	7.457	10.611	0.961
RMSE	2.87	2.40	3.65

+ p < 0.1, * p < 0.05, ** p < 0.01, *** p < 0.001

	Logistic Reg
(Intercept)	0.169 (0.478)
age	0.048** (0.016)
fthr_ed	-0.010 (0.023)
mthr_ed	-0.031 (0.027)
hh_fthr_frm	-0.241 (0.273)
hh_size96	-0.036+ (0.019)
orphan96	-0.016 (0.292)
Num.Obs.	741
AIC	978.7
BIC	1011.0
Log.Lik.	-482.370
F	2.708
RMSE	0.48

+ p < 0.1, * p < 0.05, ** p < 0.01, *** p < 0.001

2. Use a parametric model (Probit/Logit) to calculate the propensity scores for each person in the data to be abducted. Include whatever covariates or functions of covariates you think may be important.

```
logit.1 <- glm(abd ~ age+fthr_ed+mthr_ed+hh_fthr_frm+hh_size96+orphan96, C_ach+C_akw+C_ata+C_kma+C_oro+
modelsummary(list("Logistic Reg" = logit.1), stars = TRUE)
```

3. Use optimal match over the whole data set to estimate the ATE using propensity score matching. Do this for all three dependent variables.

```
library(MatchIt)
m.nn<-matchit(abd ~ age+fthr_ed+mthr_ed+hh_fthr_frm+hh_size96+orphan96, data = blattman, ratio=1, method="nn")
```

```
## Warning: Fewer control units than treated units; not all treated units will get
## a match.
```

```
# Get's the match dataset
nn.match<-match.data(m.nn)

reg4 <- lm(educ ~ abd, data = nn.match)
reg5 <- lm(distress ~ abd, data = nn.match)
reg6 <- lm(log.wage ~ abd, data = nn.match)

modelsummary(list("Education"=reg4,"Distress"=reg5,"Log Wages"=reg6),stars = TRUE, coef_rename = coef_r)
```

4. Use the cobalt package to make a “Love plot”. You can find information of the cobalt package here

Table 2: Propensity Score Matching Results

	Education	Distress	Log Wages
(Intercept)	7.416*** (0.174)	3.759*** (0.142)	4.205*** (0.222)
Abd	-0.484* (0.246)	0.790*** (0.201)	0.126 (0.314)
Num.Obs.	558	558	558
R2	0.007	0.027	0.000
R2 Adj.	0.005	0.025	-0.002
AIC	2777.1	2550.7	3049.0
BIC	2790.1	2563.7	3061.9
Log.Lik.	-1385.561	-1272.344	-1521.476
F	3.874	15.484	0.161
RMSE	2.90	2.37	3.70

+ p < 0.1, * p < 0.05, ** p < 0.01, *** p < 0.001

```
library(cobalt)
```

```
## cobalt (Version 4.5.4, Build Date: 2024-02-26)
```

```
##
```

```
## Attaching package: 'cobalt'
```

```
## The following object is masked from 'package:MatchIt':
```

```
##
```

```
## lalonde
```

```
b1<-bal.tab(abd ~ hh_size96 + hh_fthr_frm+orphan96+mthr_ed+fthr_ed+age,data=blattman,int = TRUE)
```

```
## Note: 's.d.denom' not specified; assuming "pooled".
```

```
v1<-var.names(b1, type = "vec", minimal = TRUE)
```

```
v1["hh_size96"]<-"Household Size in 1996"
```

```
v1["hh_fthr_frm"]<-"Father was Farmer"
```

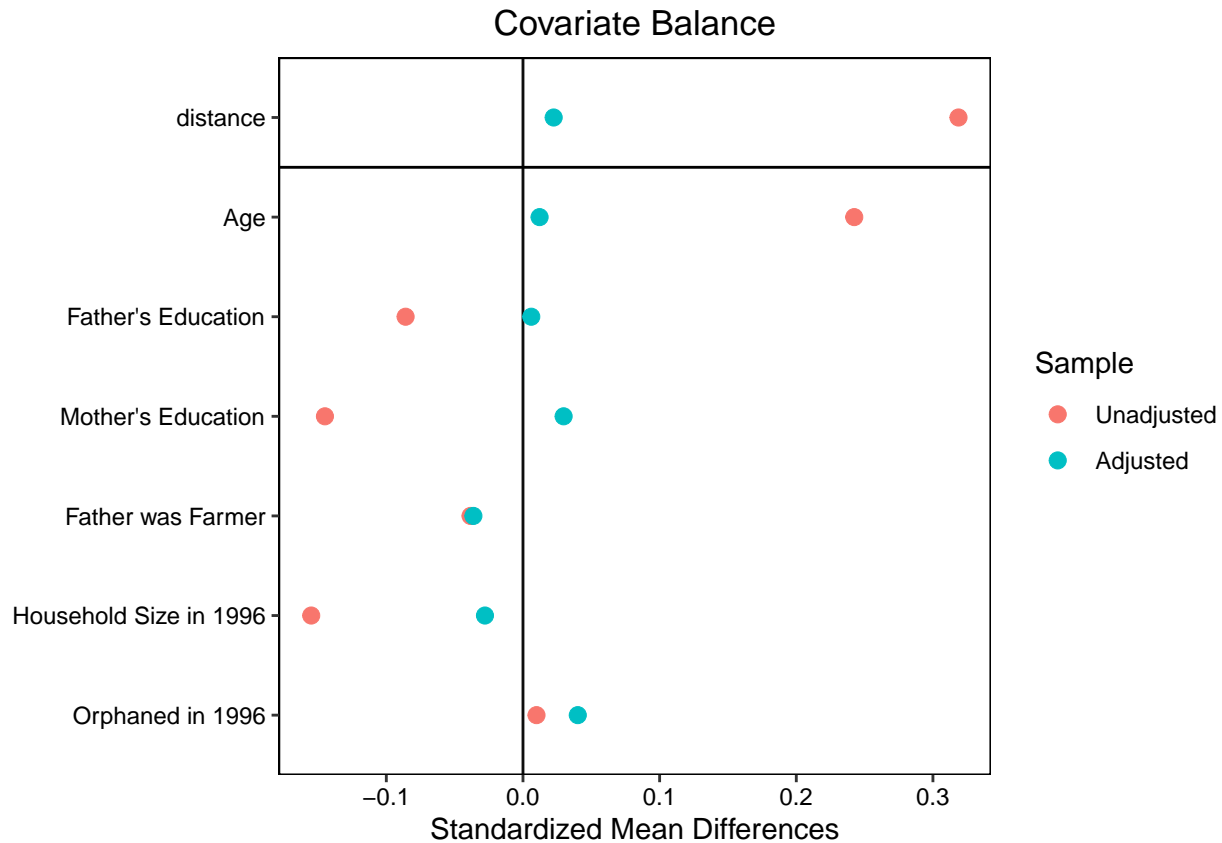
```
v1["orphan96"]<-"Orphaned in 1996"
```

```
v1["mthr_ed"]<-"Mother's Education"
```

```
v1["fthr_ed"]<-"Father's Education"
```

```
v1["age"]<-"Age"
```

```
love.plot(m.nn, binary = "std", var.names=v1) #use v1 for your variable names
```



Problem Set: BLP Methodology In this problem you will perform demand estimation using market level data. Run the following code in R

```
#install.packages("BLPestimatorR")
library(BLPestimatorR)
data(productData_cereal)
```

A table of market shares, prices, and characteristics of the top-selling brands of cereal in 1992 across several markets is now available in your environment. The data are aggregated from household-level scanner data (collected at supermarket checkout counters). We observe the following variables

price = price paid for the cereal const = just a column of 1's that you can ignore. sugar = how much sugar is in the cereal mushy = how mushy the cereal becomes with milk. share = market share of the cereal in that particular market. This number is between 0 and 1. cdid = tells you which market you are in. product_id = tells you which cereal is captured. IV1-IV20 = 20 constructed instrumental variables.

1. Find the market share of the outside good in every market. That is, sum all of the shares across all of the cereals for each market. You will notice that this number is less than 1. The market share of the outside option is equal to $1 - \text{total cereal market share in each market}$. (Hint: you can use the aggregate to sum up the cereal shares by market)

```
# We can use the function ave(variable, grouping variable, FUN = function(x) 1-sum(x))
productData_cereal$outside_share <- ave(productData_cereal$share, productData_cereal$cdid, FUN = function(x) 1-sum(x))
# this will create your new dependent variable (i.e. log(sj)-log(so))
productData_cereal$y <- log(productData_cereal$share)-log(productData_cereal$outside_share)
```

	OLS	FE	IV
(Intercept)	-2.993*** (0.112)		
sugar	0.046*** (0.004)	0.045*** (0.005)	0.045*** (0.005)
mushy	0.052 (0.052)	0.056 (0.043)	0.057 (0.043)
price	-10.120*** (0.880)	-9.319*** (0.760)	
fit_price			-9.236*** (0.792)
Num.Obs.	2256	2256	2256
R2	0.079	0.210	0.210
R2 Adj.	0.078	0.175	0.175
R2 Within		0.081	0.081
R2 Within Adj.		0.080	0.080
AIC	7068.7	6906.3	6906.3
BIC	7097.3	7461.3	7461.3
Log.Lik.	-3529.339		
F	64.593		
RMSE	1.16	1.07	1.07
Std.Errors		by: cdid	by: cdid
FE: cdid		X	X

+ p < 0.1, * p < 0.05, ** p < 0.01, *** p < 0.001

2. Estimate the share regression using fat, sugar, calories, and price as explanatory variables using OLS

```
## In this section we are just going to run OLS on the linear demand curve
blp.reg.1<-lm(y~sugar+mushy+price, data=productData_cereal)
library(fixest)
blp.reg.2 <-feols(y~sugar+mushy+price |cdid , data=productData_cereal) # Include market fixed effects
```

3. 2SLS: Use the instrumental variables IV1 - IV10 to instrument for price

```
blp.reg.3 <- feols(y~sugar+mushy |cdid| price ~ IV1+IV2+IV3+IV4+IV5+IV6+IV7+IV8+IV9+IV10, data=productData_cereal)
modelsummary(list("OLS"=blp.reg.1, "FE"=blp.reg.2,"IV"=blp.reg.3), stars = TRUE)
```

4. 2SLS: perform the first stage F-stat test to judge the strength of your instruments and the second stage sargan test to see if these instruments are independent of the error term.

```
# Hint use the summary function with object bls.reg.3
summary(blp.reg.3)
```

```
## TSLS estimation, Dep. Var.: y, Endo.: price, Instr.: IV1, IV2, IV3, IV4, IV5, IV6, IV7, IV8, IV9, IV10
## Second stage: Dep. Var.: y
## Observations: 2,256
## Fixed-effects: cdid: 94
## Standard-errors: Clustered (cdid)
##           Estimate Std. Error   t value   Pr(>|t|)
```

```
## fit_price -9.235583    0.791762 -11.66459 < 2.2e-16 ***
## sugar      0.044858    0.005051  8.88123 4.7572e-14 ***
## mushy      0.056966    0.043007  1.32458 1.8856e-01
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## RMSE: 1.07114      Adj. R2: 0.175154
##                Within R2: 0.080861
## F-test (1st stage), price: stat = 7,401.1      , p < 2.2e-16 , on 10 and 2,243 DoF.
##                Wu-Hausman: stat =      0.31764, p = 0.573088, on 1 and 2,158 DoF.
##                Sargan: stat =      23.3      , p = 0.005464, on 9 DoF.
```

5. 2SLS: can you use a smaller set of instruments to get a better result? If so, then what instruments did you include? Report your results including the first stage F-stats and the overidentification test.

Yes you can. using a smaller set of instruments, specifically instruments IV1, IV2, IV4, IV5, IV6, IV9, IV10 helps you get a better result.

Hint: You will need to run the first stage regression and identify which instruments are significant. Try using only the significant instruments.

```
blp.reg.4 <- feols(y~sugar+mushy |cdid| price ~ IV1+IV2+IV4+IV5+IV6+IV9+IV10, data=productData_cereal)

labels <- c(
  sugar = "Sugar Amount",
  mushy = "Cereal Mushiness",
  price = "Cereal Price paid",
  fit_price = "Fit Price"
)

modelsummary(list("OLS"=blp.reg.1, "FE"=blp.reg.2,"IV"=blp.reg.3,"Restricted IV"=blp.reg.4), coef_renam
```

	OLS	FE	IV	Restricted IV
(Intercept)	−2.993*** (0.112)			
Sugar Amount	0.046*** (0.004)	0.045*** (0.005)	0.045*** (0.005)	0.045*** (0.005)
Cereal Mushiness	0.052 (0.052)	0.056 (0.043)	0.057 (0.043)	0.057 (0.043)
Cereal Price paid	−10.120*** (0.880)	−9.319*** (0.760)		
Fit Price			−9.236*** (0.792)	−9.245*** (0.791)
Num.Obs.	2256	2256	2256	2256
R2	0.079	0.210	0.210	0.210
R2 Adj.	0.078	0.175	0.175	0.175
R2 Within		0.081	0.081	0.081
R2 Within Adj.		0.080	0.080	0.080
AIC	7068.7	6906.3	6906.3	6906.3
BIC	7097.3	7461.3	7461.3	7461.3
Log.Lik.	−3529.339			
F	64.593			
RMSE	1.16	1.07	1.07	1.07
Std.Errors		by: cdid	by: cdid	by: cdid
FE: cdid		X	X	X

+ p < 0.1, * p < 0.05, ** p < 0.01, *** p < 0.001