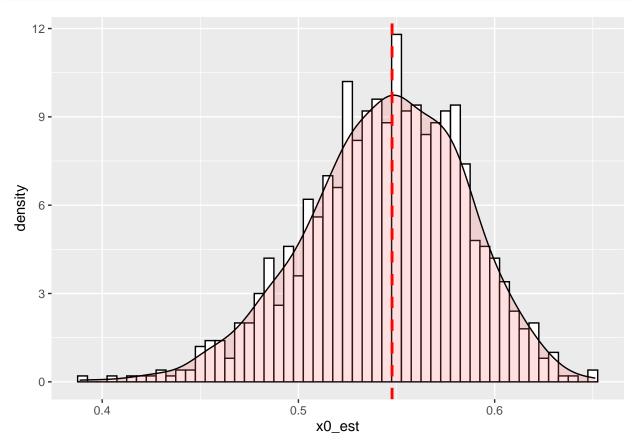
```
set.seed(1234)
n = 50
x = runif(n, 0, 1)
noise = rnorm(n, 0, 0.2)
y = rep(1,n) + x^2 + noise
data = as.data.frame(cbind(x, y))
data$x2 = data$x^2
model = lm(y \sim x2 + x, data = data)
summary(model)
##
## Call:
## lm(formula = y \sim x2 + x, data = data)
##
## Residuals:
##
       Min
                                             Max
                  1Q
                      Median
                                     3Q
## -0.42676 -0.15765 -0.04785 0.11788 0.57410
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                 0.8487
                            0.1054
                                    8.051 2.14e-10 ***
## x2
                 0.5004
                            0.4840
                                      1.034
                                               0.306
## x
                 0.5330
                            0.4907
                                     1.086
                                               0.283
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.2198 on 47 degrees of freedom
## Multiple R-squared: 0.6156, Adjusted R-squared: 0.5993
## F-statistic: 37.64 on 2 and 47 DF, p-value: 1.747e-10
Write a function for taking a sample and delivering an estimate for \hat{x}0.
bootstrap = function(data){
  choices = sample(1:n, size=n, replace=T)
  bsdata = as.data.frame(cbind(data$x[choices], data$x2[choices], data$y[choices]))
  colnames(bsdata) = c("x", "x2", "y")
 bs_model = lm(y \sim x2 + x, data=bsdata)
  w0 = bs model$coefficients[1]
 w2 = bs_model$coefficients[2]
  w1 = bs_model$coefficients[3]
  x0_hat = (-w1 + sqrt(w1^2 - 4*w2*(w0-1.3))) / (2*w2)
  return(x0_hat)
}
```

Generate a sample from this model, and compute \hat{x} 0 for this sample.

Simulate a 1000 realizations from this model, to approximate the sampling distribution of $\hat{}$ x0.

```
x0_est = vector()
for (i in 1:1000){
  n = 50
  x = runif(n, 0, 1)
  noise = rnorm(n, 0, 0.2)
  y = rep(1,n) + x^2 +noise
  data = as.data.frame(cbind(x, y))
```

```
data$x2 = data$x^2
  model = lm(y \sim x2 + x, data = data)
  w0 = model$coefficients[1]
  w2 = model$coefficients[2]
  w1 = model$coefficients[3]
  x0_hat = (-w1 + sqrt(w1^2 - 4*w2*(w0-1.3))) / (2*w2)
  x0_est[i] = x0_hat
library(ggplot2)
x0_est = as.data.frame(x0_est)
ggplot(x0_est, aes(x=x0_est)) +
    geom_histogram(aes(y=..density..),
                   binwidth=0.005,
                   colour="black", fill="white") +
    geom_density(alpha=.2, fill="#FF6666")+
  geom_vline(aes(xintercept=sqrt(0.3)), # Ignore NA values for meany
               color="red", linetype="dashed", size=1)# Overlay with transparent density plot
```

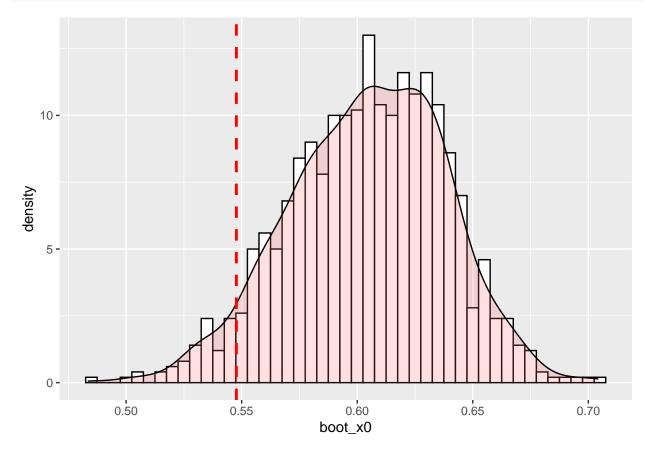


Using just your original sample, compute the bootstrap distribution of \hat{x} 0.

```
boot_x0 = vector()
print(sqrt(0.3))

## [1] 0.5477226

for (i in 1:1000){
  boot_x0[i] = bootstrap(data)
```



Compare these two distributions, and in particular their standard deviations.

```
## x0_est
## x0_est 0.04009306
print(sqrt(var(as.vector(boot_x0))))
```

boot_x0 ## boot_x0 0.03355649

print(sqrt(var(as.vector(x0_est))))

Here, we can see that the bootstrap result has less variance than the simulation result does, though with larger bias.