Generalized additive models in R

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Today's plan

First session

- 1 Intro to Generalized Additive Models (GAMs) in R
- 4 Hands-on session

Coffee break

Second session

- Beyond mean modelling: GAMLSS and quantile GAMs
- Hands-on session

Intro to Generalized Additive Models (GAMs)

Structure:

- What is an additive model?
- Introducing smooth effects
- Big Data GAM methods
- Diagnostics and model selection tools
- GAM modelling using mgcv and mgcViz

Structure of the talk

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Regression setting:

- y is our response or dependent variable
- x is a vector of covariates or independent variables

In distributional regression we want a good model for Dist(y|x).

Model is $\mathsf{Dist}_m\{y|\theta_1(\mathbf{x}),\dots,\theta_q(\mathbf{x})\}$, where $\theta_1(\mathbf{x}),\dots,\theta_q(\mathbf{x})$ are param.

In a Gaussian model, the mean depends on the covariates

$$y|\mathbf{x} \sim N\{y|\mu = \theta(\mathbf{x}), \sigma^2\},$$

where $\mu = \mathbb{E}(y|\mathbf{x})$ and $\sigma^2 = \text{Var}(y)$.

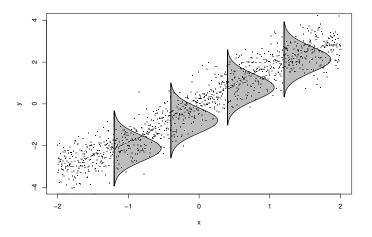


Figure: Gaussian model with variable mean.
In mgcv: gam(y~s(x), family=gaussian).

Generalized additive model (GAM):

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu(\mathbf{x}), \theta_2, \dots, \theta_p\},$$

where

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1} \Big\{ \sum_{i=1}^{m} f_i(\mathbf{x}) \Big\},\,$$

and g is the link function.

 f_i 's can be fixed, random or smooth effects with coefficients β .

Poisson GAM:

- $y|\mathbf{x} \sim \mathsf{Pois}\{y|\mu(\mathbf{x})\}$
- $\mathbb{E}(y|\mathbf{x}) = \mathsf{Var}(y|\mathbf{x}) = \exp\left\{\sum_{j=1}^{m} f_j(\mathbf{x})\right\}$
- $g = \log \text{ assures } \mu(\mathbf{x}) > 0$

Here $\mathbb{E}(y|\mathbf{x})$ and $Var(y|\mathbf{x})$ is implied by model...

... or we can have extra parameters for scale and shape.

Scaled Student's t GAM:

- $y|\mathbf{x} \sim \mathsf{ScaledStud}\{y|\mu(\mathbf{x}), \sigma, \nu\}$
- $\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = \sum_{i=1}^{m} f_i(\mathbf{x})$
- \bullet σ is scale parameter
- ν is shape parameter (degrees of freedom)
- $Var(y|\mathbf{x}) = \sigma^2 \frac{\nu}{\nu 2}$

Later we'll see models with multiple linear predictors, eg:

• $y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma(\mathbf{x})\}$

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Introducing smooth effects

Consider additive model

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1}\Big\{f_1(\mathbf{x}) + f_2(\mathbf{x}) + \cdots\Big\},\,$$

where

- $f_1(\mathbf{x}) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2$
- $f_2(\mathbf{x}) = f_2(x_2)$ is a non-linear smooth function

Smooth effects built using spine bases

$$f_2(x_2) = \sum_{k=1}^r \beta_k b_k(x_2)$$

where β_k are unknown coeff and $b_k(x_2)$ are known spline basis functions.

NB: we call $\sum_{i=1}^{m} f_i(\mathbf{x})$ linear predictor because it is linear in β .

Introducing smooth effects

B-splines:

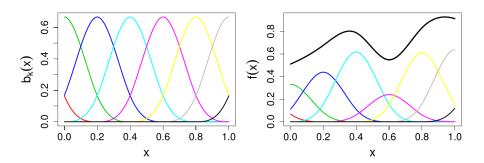
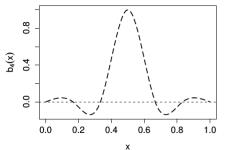
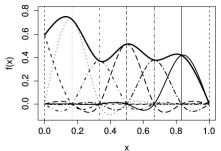


Figure: B-spline basis (left) and smooth (right).

$$s(x, bs = "cr", k = 20)$$

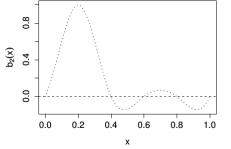


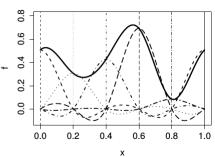


Cubic regression splines are related to the optimal solution to

$$\sum_{i=1}^{n} \{y_i - f(x_i)\}^2 + \gamma \int f''(x)^2 dx.$$

$$s(x, bs = "cc")$$

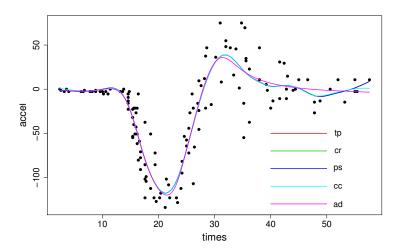




Cyclic cubic regression splines make so that

- $f(x_{min}) = f(x_{max})$
- $f'(x_{min}) = f'(x_{max})$

$$s(x, bs = "ad")$$



The wiggliness or smoothness of f(x) depends on x.

Based on thin plate regression splines basis.

Related to optimal solution to:

$$\sum_{i} \{y_i - f(x_i, z_i)\}^2 + \gamma \int f_{xx}^2 + 2f_{xz}^2 + f_{zz}^2 dx dz$$

A single smoothing parameter γ .

Isotropic: same smoothness along $x_1, x_2, ...$

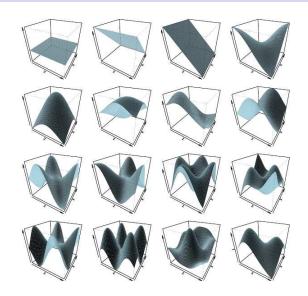
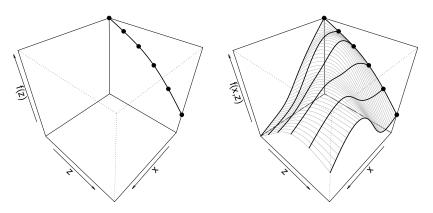


Figure: Rank 17 2D TPRS basis. Courtesy of Simon Wood.

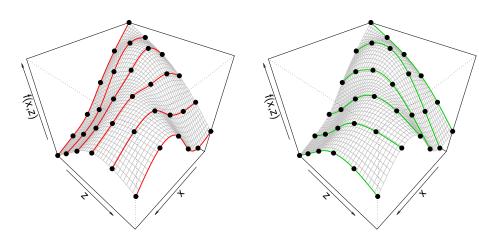
Isotropic effect of x_1 , x_2 are in same unit (e.g. Km).

If different units better use tensor product smooths te(x1, x2).

Construction: make a spline $f_z(z)$ a function of x by letting its coefficients vary smoothly with x



- x-penalty: average wiggliness of red curves
- z-penalty: average wiggliness of green curves



Can use (almost) any kind of marginal:

- te(x1, x2, x3) product of 3 cubic regression splines bases
- te(x1, x2, bs = c("cc", "cr"), k = c(10, 6))
- te(LO, LA, t, d=c(2,1), k=c(20,10), bs=c("tp","cc"))

Basis of te contains functions of the form $f(x_1)$ and $f(x_2)$.

To fit $f(x_1) + f(x_2) + f(x_1, x_2)$ separately use:

$$y \sim ti(x1) + ti(x2) + ti(x1, x2)$$

By-factor smooths

Approach (1) is s(x, by = subject), which means

- $\mu(x) = f_1(x) + ...$ if subject = 1
- $\mu(x) = f_2(x) + ...$ if subject = 2
- ...

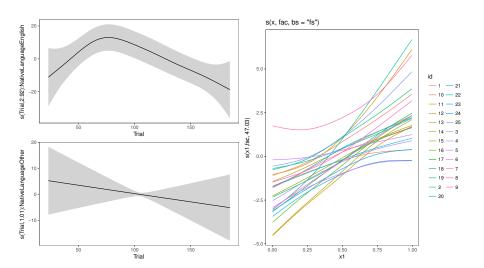
Approach (2) is s(x, subject, bs = "fs"), which means

- $\mu(x) = b_1 + f_1(x) + ...$ if subject = 1
- $\mu(x) = b_2 + f_2(x) + ...$ if subject = 2
- ...

where $b_1, b_2, \dots \sim N(0, \gamma_b \mathbf{I})$ are random effects.

In (1) each f_j has its own smoothing parameter.

In (2) all f_i 's have the same smoothing parameter.



Introducing smooth effects

In general

$$f(\mathbf{x}) = \sum_{k=1}^{r} \beta_k b_k(\mathbf{x}).$$

To determine complexity of $f(\mathbf{x})$:

- the basis rank r is large enough for sufficient flexibility
- ullet a complexity penalty on eta controls the wiggliness of the effects

GAM model fitting

 $\hat{oldsymbol{eta}}$ is the maximizer of **penalized** log-likelihood

$$\hat{\boldsymbol{\beta}} = \operatorname*{argmax}_{\boldsymbol{\beta}} \operatorname{PenLogLik}(\boldsymbol{\beta}|\boldsymbol{\gamma}) = \operatorname*{argmax}_{\boldsymbol{\beta}} \big\{ \underbrace{\widetilde{L_{\boldsymbol{y}}(\boldsymbol{\beta})}}_{\operatorname{goodness}} - \underbrace{\operatorname{Pen}(\boldsymbol{\beta}|\boldsymbol{\gamma})}_{\operatorname{penalize complexity}} \big\}$$

where:

- $L_{\nu}(\beta) = \sum_{i} \log p(y_{i}|\beta)$ is log-likelihood
- ullet Pen $(eta|\gamma)$ penalizes the complexity of the f_j 's
- $\gamma > 0$ smoothing parameters ($\uparrow \gamma \uparrow$ smoothness)

Concrete example $\mu(\mathbf{x}) = f(x_1) + g(x_2, x_3)$:

$$\mathsf{Pen}(oldsymbol{eta}|oldsymbol{\gamma}) pprox \gamma_1 \int f_{x_1x_1}^2 dx_1 + \gamma_2 \int g_{x_2x_2}^2 + g_{x_3x_3}^2 + 2g_{x_2x_3}^2 dx_2 dx_3$$

GAM model fitting

We use a hierarchical fitting framework:

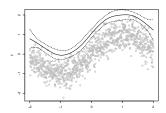
 $\ \, \textbf{ § Select } \gamma \ \, \text{determine smoothness} \\$

$$\hat{\gamma} = \mathop{\mathsf{argmax}}_{\gamma} \mathsf{LAML}(\gamma)$$

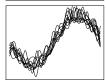
where LAML $(\gamma) \approx p(y|\gamma) = \int p(y,\beta|\gamma)d\beta$.

2 For fixed γ , estimate β to determine actual fit

$$\hat{eta} = \mathop{\mathsf{argmax}}_{eta} \mathop{\mathsf{PenLogLik}}(eta|\gamma).$$







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Structure of the talk

Structure:

- GAM model fitting
- Types of smooth effects
- Big Data methods

Big Data methods

Recall the GAM model structure

$$\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = g^{-1}\Big\{\sum_{j=1}^m f_j(\mathbf{x})\Big\}$$

Here $\mu(\mathbf{x}_i)$ can be written as $g^{-1}(\mathbf{X}_i\beta)$, where \mathbf{X}_i row of matrix \mathbf{X} having has n rows and

$$d = p + k_1 + \cdots + k_j + \cdots + k_m$$

columns.

Big Data methods

Bottom line: X can get very big, which causes problems:

- storing X takes too much memory
- \bullet computing things involving **X** (e.g. $\mathbf{X}^{\mathsf{T}}\mathbf{X}$) takes time

Solution implemented in mgcv::bam function:

• do not create **X** but only sub-blocks:

$$\mathbf{X} = \left[\begin{array}{cc} \mathbf{X}_{11} & \mathbf{X}_{12} \\ \mathbf{X}_{21} & \mathbf{X}_{22} \end{array} \right]$$

do not store them either, but create them when needed;

- any computation involving X is based on the blocks;
- use parallelization when possible;

Big Data methods

Further acceleration and memory savings by discretization.

Instead of having n unique rows of X discretize to b << n rows.

In mgcv:

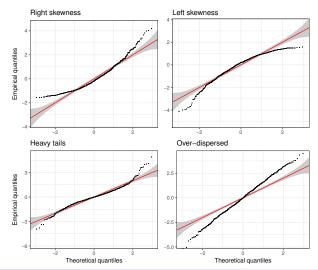
Structure of the talk

Structure:

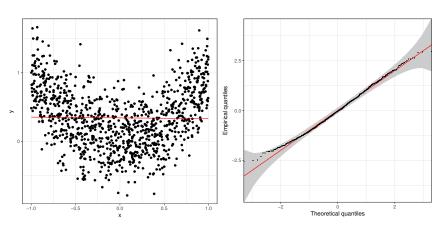
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In first hands-on session we'll use few basic diagnostics.

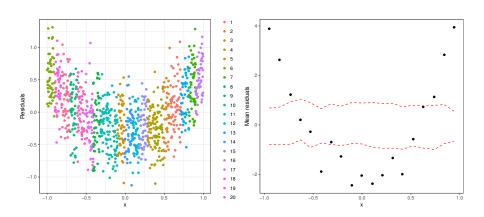
QQ-plots



Useful for choosing model $\operatorname{Dist}_m(y|\mathbf{x})$ (e.g. Poisson vs Neg. Binom.) Less useful for finding omitted variables and non-linearities.



Conditional residuals checks are more helpful here.



Recall structure of smooth effects:

$$f(\mathbf{x}) = \sum_{j=1}^k \beta_j b_j(\mathbf{x}).$$

where β shrunk toward zero by smoothness penalty.

Effective number of parameters we are using is < k.

Approximation is **Effective Degrees of Freedom** (EDF) < k.

By default k = 10 but this is arbitrary.

Exact choice of k not important, but it must not be too low.

Checking whether *k* is too low:

- 1 look at conditional residuals checks
- ② look at output of check(fit):

```
## k' edf k-index p-value

## s(wM) 9.00 8.60 0.91 <2e-16 ***

## s(wM_s95) 9.00 8.13 1.02 0.76

## s(Posan) 8.00 2.66 1.04 0.97
```

increase *k* and see if a **model selection criterion** improves

Model selection

General criterion is approximate Akaike Information Criterion (AIC):

$$AIC = \underbrace{-2 \log p(\mathbf{y}|\hat{\beta})}_{\text{goodness of fit}} + \underbrace{2\tau}_{\text{model complexity}}$$

where τ is EDF.

If $AIC_{m1} < AIC_{m2}$ choose model 1.

To select which effects to include we can also look at p-values:

```
summary(fit)
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 267.2004 75.4197 3.543 0.000405 ***
## Fl 6.2854 1.0457 6.010 2.20e-09 ***
## loc2 79.8459 80.4130 0.993 0.320858
## loc3 -71.2728 86.1725 -0.827 0.408284
```

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GAMs in mgcv and mgcViz

 ${\tt mgcv}$ is the recommended R package for fitting GAMs.

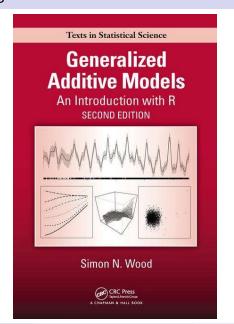
Today we'll work with mgcViz's interface.

mgcViz extends mgcv's tools for:

- plotting the estimated effects
- doing visual model checking

But most of the computation is done by mgcv.

Further reading



References I

Fasiolo, M., Y. Goude, R. Nedellec, and S. N. Wood (2017). Fast calibrated additive quantile regression. *arXiv preprint arXiv:1707.03307*.