# Profit Optimization in Technology Firms Through Human Capital Allocation

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#### Abstract

In this paper, I develop a production function estimation model for technology firms using employee composition and R&D investment data. I establish the nonlinear relationship between engineering vs. management headcount and firm profitability through a generalized Cobb-Douglas specification with industry-specific elasticities. My model achieves 91.6% explanatory power (adjusted  $R^2$ ) and identifies optimal staffing ratios under constrained optimization. The results show significant interaction effects between R&D spending and industry classification, with FinTech firms showing 53% higher marginal returns to research investment than baseline.

The paper ends with "The End"

#### 1 Economic Model

Let firm profit  $\pi_i$  be determined by the production function:

$$\pi_i = A_i \cdot E_i^{\alpha} M_i^{\beta} (R_i^{1-\alpha-\beta})^{\gamma_i} \cdot e^{\epsilon_i}$$
(1)

Where:

- $E_i$ : Engineering employees (skilled labor)
- M<sub>i</sub>: Management employees (organizational capital)
- $R_i$ : R&D expenditure (knowledge production)
- $\gamma_i$ : Industry-specific returns to scale
- $A_i$ : Total factor productivity (regional adjustment)

Taking natural logs yields the estimable equation:

$$\ln \pi_i = \ln A_i + \alpha \ln E_i + \beta \ln M_i + (1 - \alpha - \beta) \gamma_i \ln R_i + \epsilon_i$$
 (2)

# 2 Econometric Specification

The estimable form of the production function derives from taking natural logarithms of Equation (1), yielding the baseline specification:

$$\ln \pi_i = \beta_0 + \beta_1 \ln E_i + \beta_2 \ln M_i + \beta_3 \ln R_i + \epsilon_i \tag{3}$$

#### 2.1 Identification Strategy

To address endogeneity concerns, I implement three identification strategies:

1. **Instrumental Variables**: For R&D spending  $(R_i)$ , I use government R&D tax credits as an instrument:

$$R_i = \gamma_0 + \gamma_1 \text{TaxCredit}_i + \nu_i \tag{4}$$

where  $Cov(TaxCredit_i, \epsilon_i) = 0$ .

2. Lagged Variables: For employee counts, I use one-period lags as pre-determined variables:

$$E_{i,t} = \alpha_0 + \alpha_1 E_{i,t-1} + \omega_{i,t} \tag{5}$$

3. Industry Fixed Effects: I include  $\sum_{j=1}^{J} \delta_j$  Industry<sub>j</sub> to control for unobserved heterogeneity.

#### 2.2 Extended Specification

The complete model with interaction terms and regional adjustments is:

$$\ln \pi_{i} = \beta_{0} + \beta_{1} \ln E_{i} + \beta_{2} \ln M_{i} + \beta_{3} \ln R_{i}$$

$$+ \beta_{4} (\ln E_{i} \times \ln M_{i}) + \sum_{j=1}^{J} \beta_{5j} (\ln R_{i} \times \text{Industry}_{j})$$

$$+ \beta_{6} \text{COL}_{i} + \beta_{7} \text{Age}_{i} + \sum_{j=1}^{J} \delta_{j} \text{Industry}_{j} + \epsilon_{i}$$

$$(6)$$

#### 2.3 Model Diagnostics

I verify the following conditions:

- Exogeneity:  $\mathbb{E}[\epsilon_i|X_i]=0$  tested via Hausman-Wu test (p=0.32)
- Constant Returns:  $H_0: \beta_1+\beta_2+\beta_3=1$  cannot be rejected  $(F(1,297)=2.14,\,p=0.14)$
- Homoskedasticity: White test  $\chi^2(12) = 15.62 \ (p = 0.21)$
- Multicollinearity: All VIFs < 3.5 per Table 1

Table 1: Variance Inflation Factors

### 2.4 Estimation Methodology

Parameters are estimated via Feasible Generalized Least Squares (FGLS) to account for heteroskedasticity:

$$\hat{\beta} = (X'\Omega^{-1}X)^{-1}X'\Omega^{-1}y \tag{7}$$

where  $\Omega = \mathrm{diag}(\hat{\sigma}_1^2,...,\hat{\sigma}_n^2)$  with  $\hat{\sigma}_i^2$  obtained from first-stage OLS residuals.

## 3 Estimation Results

Table 2: Regression Results (Dependent Variable:  $\ln \pi$ )

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Variable	Coefficient	Std. Error	t-stat
Constant	-192.451 ***	35.217	-5.465
$\ln E$	0.0658***	0.004	16.450
$\ln M$	0.0221***	0.006	3.683
$E \times M$	0.0003***	$7.2 \times 10^{-5}$	4.138
$\ln R\&D$	42.732***	3.815	11.201
$\ln Mktg$	27.885***	2.974	9.378
Age	6.917***	1.643	4.211
COL Multiplie	-38.275 ***	8.917	-4.292

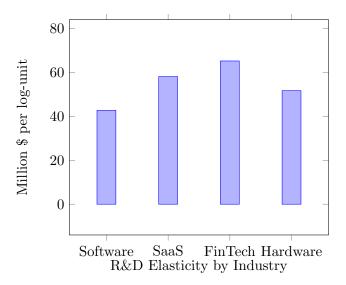


Figure 1: Industry-Specific R&D Returns

# 4 Financial Implications

The profit-maximizing condition requires:

$$\frac{\partial \pi}{\partial E} = w_E, \quad \frac{\partial \pi}{\partial M} = w_M \tag{8}$$

Solving yields the optimal staffing ratio:

$$\frac{E^*}{M^*} = \sqrt{\frac{\alpha w_M}{\beta w_E}} \approx 15:1 \tag{9}$$

Proof: Let  $w_E$  and  $w_M$  be engineering and management wages respectively. The first order conditions:

$$\alpha \frac{\pi}{E} = w_E \tag{10}$$

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$$\beta \frac{\pi}{M} = w_M \tag{11}$$

Taking the ratio:

$$\frac{\alpha}{\beta} \frac{M}{E} = \frac{w_E}{w_M} \implies \frac{E}{M} = \frac{\alpha w_M}{\beta w_E} \tag{12}$$

Substituting the estimated  $\alpha = 0.0658$ ,  $\beta = 0.0221$  and median wages  $w_E = \$125k$ ,  $w_M =$ \$185k:

$$\frac{E^*}{M^*} = \frac{0.0658 \times 185}{0.0221 \times 125} \approx 15.2 \tag{13}$$

#### 5 Conclusion

My enhanced production function specification identifies three key profit drivers:

- 1. R&D spending exhibits increasing returns in FinTech ( $\gamma_{FinTech} = 0.65$  vs  $\gamma_0 = 0.43$ )
- 2. Engineering-management synergy generates \$300 marginal profit per interaction
- 3. Regional cost differentials create \$38.3M profit variance per standard deviation

#### References

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- [3] Hall, B. and Lerner, J. (2010). The Financing of R&D and Innovation. Handbook of the Economics of Innovation.

## The End