

Times Series Assignment 3

12111603 Tan Zhiheng

2023-11-10

- Question 1
 - (a)
 - (b)
- Question 2
 - (a)
 - (b)
- Question 3
 - (a)
 - (b)
- Question 4
- Question 5
- Question 5
 - (a)
 - (b)
 - (c)
 - (d)

Question 1

(a)

- $\mathbb{E}X_t = -2t + \mathbb{E}(W_t + 0.5W_{t-1}) = -2t.$

$$\gamma(t+h, t) = \text{Cov}(X_{t+h}, X_t) = \text{Cov}(-2t - 2h + W_{t+h} + 0.5W_{t+h-1}, -2t + W_t + 0.5W_{t-1})$$

$$\bullet \quad = \begin{cases} 1.25\sigma^2, & h = 0 \\ 0.5\sigma^2, & h = \pm 1 \\ 0, & \text{else} \end{cases}$$

The times series is non-stationary because the mean function is related to time t .

(b)

We can obtain that

$$Y_t = \nabla X_t = X_t - x_{t-1} = -2 + W_t - 0.5W_{t-1} - 0.5W_{t-2}.$$

Thus,

- $\mathbb{E}Y_t = -2 + \mathbb{E}(W_t - 0.5W_{t-1} - 0.5W_{t-2}) = -2.$

$$\bullet \quad \gamma(t+h, t) = \text{Cov}(Y_{t+h}, Y_t) = \text{Cov}(-2 + W_{t+h} - 0.5W_{t+h-1} - 0.5W_{t+h-2}, -2 + W_t - 0.5W_{t-1} - 0.5W_{t-2})$$
$$= \begin{cases} 1.5\sigma^2, & h = 0 \\ -0.25\sigma^2, & h = \pm 1 \\ -0.5\sigma^2, & h = \pm 2 \\ 0, & \text{else} \end{cases}$$

The times series Y_t is stationary because it is a $MA(2)$ with shift, whose characteristic function has no roots inside the unit ball on the complex plane.

Question 2

Rewrite the times series as follows:

$$\phi(B)X_t = \theta(B)W_t$$

where $\phi(z) = 1 - z + 0.25z^2$ and $\theta(z) = 1 - 0.25z$.

(a)

Since the roots of $\phi(z)$ are 2's, whose modules are greater than 1, then we can say the times series are casual.

(b)

By the formula $\psi_j = \theta_j + \phi_1\psi_{j-1} + \dots + \phi_p\psi_{j-p}$, we can figure out that

- $\psi_1 = 0.75$
- $\psi_2 = 0.5$
- $\psi_k = \psi_{k-1} - 0.25\psi_{k-2}$ for all $k \geq 3$

Question 3

Rewrite the times series as follows:

$$\phi(B)X_t = \theta(B)W_t$$

where $\phi(z) = 1 - z + 0.25z^2$ and $\theta(z) = 1 - 0.25z$.

(a)

Since the root of $\theta(z)$ is 4, whose module is greater than 1, then we can say the times series are invertible.

(b)

Since $\phi(B)X_t = \theta(B)W_t$ and θ is invertible, then we have

$$W_t = \theta(B)^{-1}\phi(B)X_t = (\theta^{-1}\phi)(B)X_t.$$

Because

$$\begin{aligned}\theta^{-1}(z)\phi(z) &= \frac{1}{1 - 0.25z}(1 - 0.5z)^2 \\ &= (1 + \frac{1}{4}z + (\frac{1}{4}z)^2 + \dots)(1 - z + \frac{1}{4}z^2)\end{aligned}$$

Then we have

- $\pi_1 = -0.75$
- $\pi_2 = \frac{1}{4^2}$
- $\pi_3 = \frac{1}{4^3}$
- $\pi_k = \frac{1}{4^k}$ for all $k \geq 2$

Question 4

Since $h = 10 > p = 3$ in the $AR(p = 3)$ model, then we maintain that the best linear prediction of X_{11} becomes

$$X_{11}^{10} = Proj(X_{11}|X_1, \dots, X_{10}) = \sum_{j=1}^3 \phi_j X_{11-j} = \phi_1 X_{10} + \phi_2 X_9 + \phi_3 X_8.$$

Given $x_{10} = -0.74$, $x_9 = -3.5$ and $x_8 = 3.0$, then by the equation above we can make a prediction of X_{11} is $\hat{X}_{11} = 2.352$.

Question 5

Suppose $X_t = \phi X_{t-1} + W_t$, which is an $AR(1)$ model and assume that $|\phi| \neq 1$.

Let Y be the linear prediction of X_2 given X_1 and X_3 , then Y can be rewritten as the form of $Y = \alpha X_1 + \beta X_3$ where α and β are two constants to be solved.

Since the best linear prediction is essentially a projection, then we can say that the difference between the variable to be predicted and its projection is perpendicular to the plane spanned by given variables, which means

$$(X_2 - Y) \perp Span\{X_1, X_3\}.$$

Thus, we can establish our equation systems as follows:

$$\begin{cases} Cov(X_1, X_2 - Y) = \mathbb{E}[X_1(X_2 - Y)] = \gamma(1) - \alpha\gamma(0) - \beta\gamma(2) = 0 \\ Cov(X_3, X_2 - Y) = \mathbb{E}[X_3(X_2 - Y)] = \gamma(1) - \alpha\gamma(2) - \beta\gamma(0) = 0 \end{cases}$$

Invoke that $\gamma(h) = \frac{\phi^h}{1-\phi^2}$ in $AR(1)$ model, we can solve the equation systems above to obtain that $\alpha = \beta = \frac{\phi}{1+\phi^2}$.

Consequently, the best linear prediction of X_2 given X_1 and X_3 is

$$\frac{\phi}{1+\phi^2}(X_1 + X_3).$$

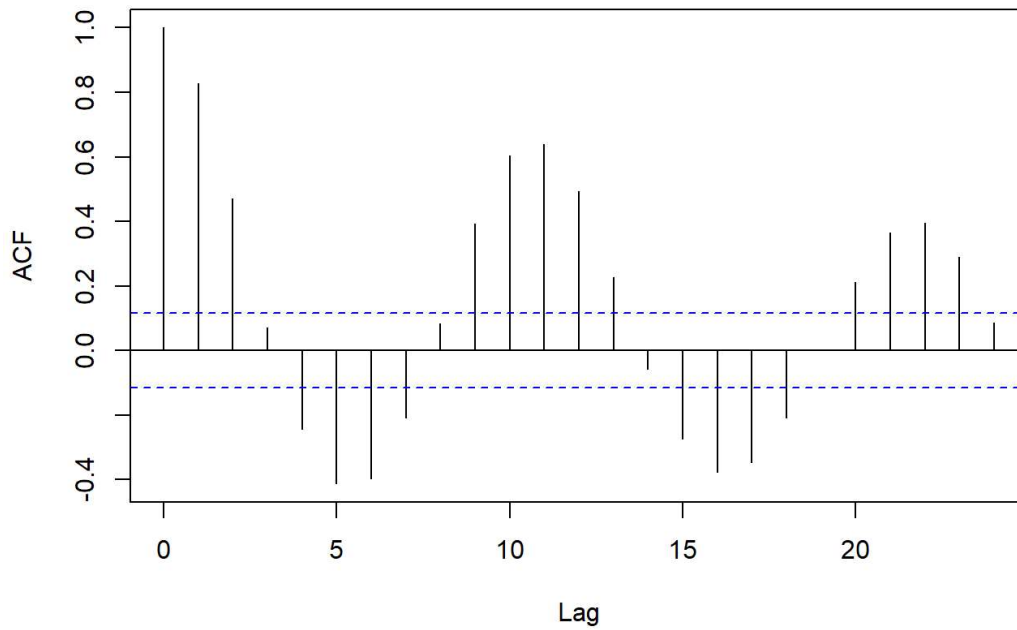
Question 5

(a)

```
data <- read.csv("C:/Users/Lenovo/Desktop/sunspot(2).dat")

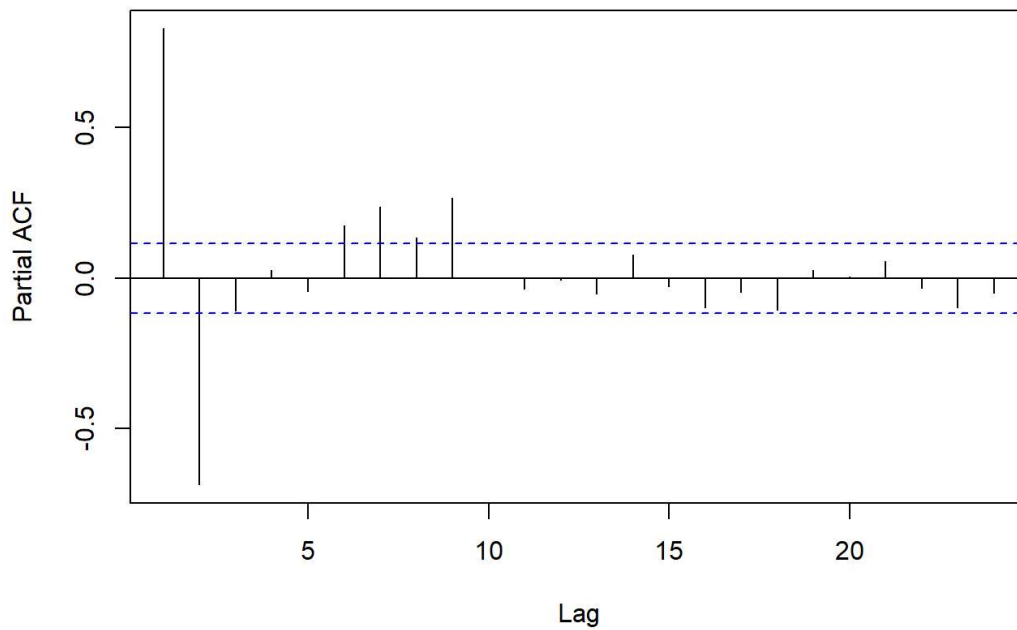
data$X5.00 <- sqrt(data$X5.00)
ts_data <- ts(data)
acf(ts_data, main = "ACF Plot")
```

ACF Plot



```
pacf(ts_data, main = "PACF Plot")
```

PACF Plot



(b)

There does not exist any truncation in the ACF plot, which means it cannot be an $MA(q)$ model. However, we see a cut off at $h = 2$ in the PACF plot, which indicates that we can use an $AR(2)$ model to fit it.

```
library(forecast)
```

```
## Registered S3 method overwritten by 'quantmod':  
##   method      from  
## as.zoo.data.frame zoo
```

```
model <- arima(ts_data, order = c(2, 0, 0))
summary(model)
```

```
##
## Call:
## arima(x = ts_data, order = c(2, 0, 0))
##
## Coefficients:
##          ar1          ar2    intercept
##          1.4029    -0.6916         6.3471
## s.e.    0.0425     0.0425     0.2390
##
## sigma^2 estimated as 1.353:  log likelihood = -447.12,   aic = 902.25
##
## Training set error measures:
##
##              ME      RMSE      MAE    MPE  MAPE      MASE      ACF1
## Training set 0.00195595 1.163115 0.9109023 -Inf   Inf  0.6511489 -0.07022899
```

Thus, we maintain that the $AR(2)$ model is

$$X_t = 1.40X_{t-1} - 0.69X_{t-2} + 6.35 + W_t,$$

where $W_t \sim WN(0, 1.353)$.

(c)

```
forecast_result <- forecast(model, h = 4)

print(forecast_result)
```

```
##      Point Forecast    Lo 80    Hi 80    Lo 95    Hi 95
## 285      5.692866 4.202274 7.183459 3.4132023 7.972531
## 286      5.133391 2.565413 7.701369 1.2060073 9.060774
## 287      5.096921 1.900942 8.292901 0.2090931 9.984750
## 288      5.432706 2.010790 8.854621 0.1993380 10.666073
```

- $X_{285}^{284} \approx 5.69$, whose 95% CI is approximately [3.41, 7.97]
- $X_{286}^{284} \approx 5.13$, whose 95% CI is approximately [1.21, 9.06]
- $X_{287}^{284} \approx 5.10$, whose 95% CI is approximately [0.21, 9.98]
- $X_{288}^{284} \approx 5.43$, whose 95% CI is approximately [0.20, 10.67]

(d)

```
data2 <- read.csv("C:/Users/Lenovo/Desktop/sunspot2(2).dat")
ts_data2 <- ts(data2)
```

```
forecast_result$mean <- forecast_result$mean^2
forecast_result$lower <- forecast_result$lower^2
forecast_result$upper <- forecast_result$upper^2
```

```
data2[285:288, ]
```

```
## [1] 17.9 13.4 29.2 100.2
```

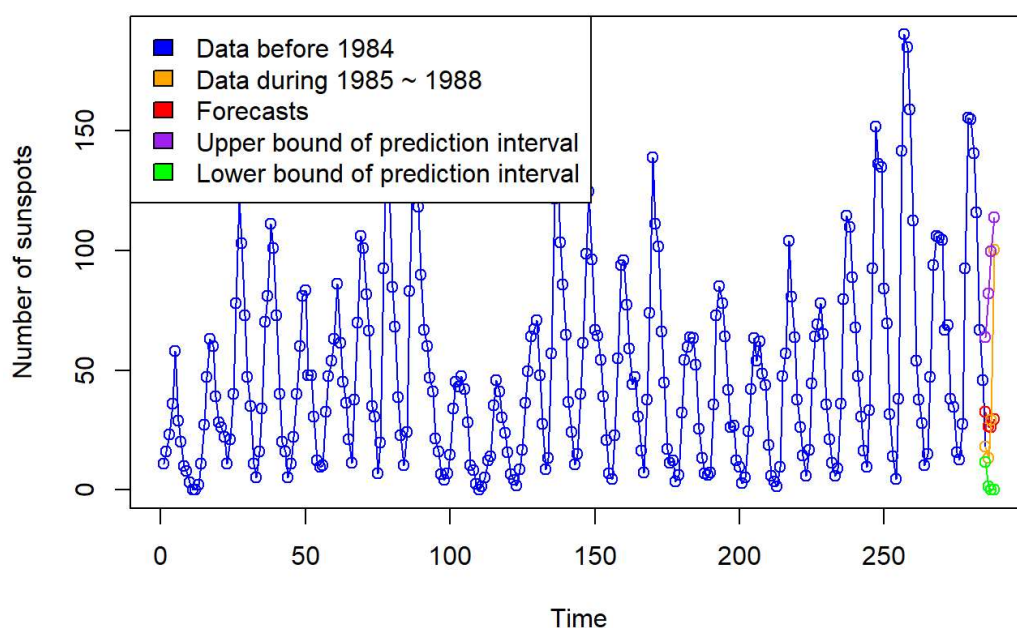
```

x1 <- 285:288
y1 <- c(17.9, 13.4, 29.2, 100.2)

plot(ts_data2, col = "blue", type = "o",
     main = "Plots of all data, forecasts and prediction intervals", xlab = "Time",
     ylab = "Number of sunspots")
legend("topleft", legend = c("Data before 1984", "Data during 1985 ~ 1988",
                             "Forecasts", "Upper bound of prediction interval",
                             "Lower bound of prediction interval"),
     fill = c("blue", "orange", "red", "purple", "green"), cex = 1
)
lines(forecast_result$mean, col = "red", type = "o")
lines(x1, y1, col = "orange", type = "o")
lines(x1, forecast_result$lower[, 2], col = "green", type = "o")
lines(x1, forecast_result$upper[, 2], col = "purple", type = "o")

```

Plots of all data, forecasts and prediction intervals



To have an explicit insight of the effect of our prediction, we filter out the figure during the time from 1950 to 1988.

```

plot(ts_data2, col = "blue", type = "o", xlim = c(250, 290),
     main = "Plots of all data, forecasts and prediction intervals", xlab = "Time",
     ylab = "Number of sunspots")
legend("topleft", legend = c("Data before 1984", "Data during 1985 ~ 1988",
                             "Forecasts", "Upper bound of prediction interval",
                             "Lower bound of prediction interval"),
     fill = c("blue", "orange", "red", "purple", "green"), cex = 1
)
lines(forecast_result$mean, col = "red", type = "o")
lines(x1, y1, col = "orange", type = "o")
lines(x1, forecast_result$lower[, 2], col = "green", type = "o")
lines(x1, forecast_result$upper[, 2], col = "purple", type = "o")

```

Plots of all data, forecasts and prediction intervals

