

# MATH 3235: Probability Theory

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# Chapter 1

## Events and Probabilities

### 1.1 Probability Spaces

**Definition 1.1.** A *probability space* is a triple  $(\Omega, \mathcal{F}, \mathbb{P})$  where

- $\Omega$  is called the *sample space* (the set of all possible outcomes of a random experiment);
- $\mathcal{F} \subseteq \mathcal{P}(\Omega)$ , called the *event space*,<sup>1</sup> is nonempty and must satisfy:
  - (i) if  $A \in \mathcal{F}$ , then  $A^c \in \mathcal{F}$ ,
  - (ii) if  $A_1, A_2, \dots \in \mathcal{F}$ , then  $\bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$ ;
- $\mathbb{P}$  is a probability measure on  $(\Omega, \mathcal{F})$  (to be defined later).

**Remark.** In general, when  $\Omega$  is finite or countably infinite, one takes  $\mathcal{F} = \mathcal{P}(\Omega)$ .

**Proposition 1.1.** *We always have  $\emptyset, \Omega \in \mathcal{F}$ .*

*Proof.* Since  $\mathcal{F} \neq \emptyset$ , there exists some event  $A \in \mathcal{F}$ . Then we get  $A^c \in \mathcal{F}$  and  $\Omega = A \cup A^c \in \mathcal{F}$  by the complement and union properties of  $\mathcal{F}$ . Finally  $\emptyset = \Omega^c \in \mathcal{F}$  by the complement property.  $\square$

**Definition 1.2.** A *probability measure* on  $(\Omega, \mathcal{F})$  is a function  $\mathbb{P} : \mathcal{F} \rightarrow [0, \infty)$  such that

- (i)  $\mathbb{P}(\Omega) = 1$ ,
- (ii) and  $\mathbb{P}(\bigcup_{k=1}^{\infty} A_k) = \sum_{k=1}^{\infty} \mathbb{P}(A_k)$  whenever  $A_1, A_2, \dots \in \mathcal{F}$  are pairwise disjoint.<sup>2</sup>

**Proposition 1.2.** *The following properties hold for any probability measure  $\mathbb{P}$  on  $(\Omega, \mathcal{F})$ :*

- (1) *For any  $A \in \mathcal{F}$ , we have  $\mathbb{P}(A^c) = 1 - \mathbb{P}(A)$ .*
- (2) *Let  $A, B \in \mathcal{F}$  with  $A \subseteq B$ . Then  $\mathbb{P}(A) \leq \mathbb{P}(B)$ .*
- (3) *Let  $A, B, C \in \mathcal{F}$ . Then*

$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).$$

*This is the principle of inclusion-exclusion.*

*Proof.* (1) Observe that  $A \cup A^c = \Omega$  and  $A \cap A^c = \emptyset$ , so  $1 = \mathbb{P}(\Omega) = \mathbb{P}(A \cup A^c) = \mathbb{P}(A) + \mathbb{P}(A^c)$ .

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<sup>1</sup>The elements of  $\mathcal{F}$  are called *events*. Events with cardinality 1 are called *elementary*.

<sup>2</sup>i.e.  $A_i \cap A_j = \emptyset$  whenever  $i \neq j$ .

(2) Write  $B = A \cup (B \setminus A)$ .<sup>3</sup> Since  $A \cap (B \setminus A) = \emptyset$ , we have  $\mathbb{P}(B) = \mathbb{P}(A) + \mathbb{P}(B \setminus A) \geq \mathbb{P}(A)$ .<sup>4</sup>

(3) Left as an exercise. Follow similar ideas as in (2).  $\square$

**Remark.** Observe that property (2) implies  $\mathbb{P}(A) \leq \mathbb{P}(\Omega) = 1$  since any  $A \subseteq \Omega$ .

**Example 1.2.1.** Pick a point uniformly at random from the unit square  $\Omega = [0, 1] \times [0, 1]$  and record its coordinates. Then the probability of the point being inside a fixed shape  $S \subseteq \Omega$  is  $|S|$ , the area of  $S$ .

**Remark.** Note that  $\mathbb{P}$  only satisfies *countable* additivity. For instance let  $\Omega = [0, 1]$  and  $\mathbb{P}$  be the uniform measure on  $\Omega$ . Then  $\Omega = \bigcup_{x \in [0, 1]} \{x\}$  and  $\mathbb{P}(\{x\}) = 0$  for every  $x \in [0, 1]$ , but  $\mathbb{P}(\Omega) = 1$ . This is because the union  $\bigcup_{x \in [0, 1]} \{x\}$  is uncountable.

**Definition 1.3.** Let  $\Omega$  be finite and  $\mathcal{F} = \mathcal{P}(\Omega)$ . The uniform probability on  $(\Omega, \mathcal{F})$  is the one such that

$$\mathbb{P}(\{\omega\}) = \frac{1}{\text{card } \Omega} \quad \text{for all } \omega \in \Omega.$$

**Proposition 1.3.** Let  $\mathbb{P}$  be the uniform probability on a finite set  $\Omega$  and let  $A \in \mathcal{F}$ . Then

$$\mathbb{P}(A) = \frac{\text{card } A}{\text{card } \Omega}.$$

*Proof.* Note that  $A$  is finite since  $\Omega$  is and so we may enumerate its elements as  $A = \{\omega_1, \omega_2, \dots, \omega_n\}$ , where  $n = \text{card } A$ . Then the sets  $\{\omega_i\}_{i=1}^n$  are pairwise disjoint and thus we have

$$\mathbb{P}(A) = \mathbb{P}\left(\bigcup_{i=1}^n \{\omega_i\}\right) = \sum_{i=1}^n \mathbb{P}(\{\omega_i\}) = \sum_{i=1}^n \frac{1}{\text{card } \Omega} = \frac{n}{\text{card } \Omega} = \frac{\text{card } A}{\text{card } \Omega},$$

which is the desired result.  $\square$

## 1.2 Conditional Probability

**Definition 1.4.** Let  $B \in \mathcal{F}$  such that  $\mathbb{P}(B) > 0$ . Then the *conditional probability* of  $A$  given  $B$ , written  $\mathbb{P}(A|B)$ , is given by

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}.$$

**Remark.** The intuition is that the extra information gained by knowing the occurrence of  $B$  should update our computation of the probability of  $A$ .

**Remark.** Another way to think about conditional probability is a restriction of the sample space to  $B$ .

**Example 1.4.1.** Suppose a family has two children, one of which is a girl. What is the probability the other is a girl? Define the sample space to be

$$\Omega = \{(B, G), (B, B), (G, G), (G, B)\}.$$

<sup>3</sup>Note that  $B \setminus A \in \mathcal{F}$  since  $B \setminus A = B \cap A^c = (B^c \cup A)^c$ .

<sup>4</sup>Since  $\mathbb{P} : \mathcal{F} \rightarrow [0, \infty)$ , we have  $\mathbb{P}(B \setminus A) \geq 0$ .

Note that each elementary event is equally likely, i.e.

$$\mathbb{P}(\{(B, G)\}) = \mathbb{P}(\{(G, B)\}) = \mathbb{P}(\{(B, B)\}) = \mathbb{P}(\{(G, G)\}) = \frac{1}{4}.$$

Let  $A = \{\text{both of them are } G\} = \{(G, G)\}$  and  $B = \{\text{one is a girl}\} = \{(G, B), (B, G), (G, G)\}$ . Then

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(\{(G, G)\})}{\mathbb{P}(\{(G, B), (B, G), (G, G)\})} = \frac{1/4}{3/4} = \frac{1}{3}.$$

**Remark.** If we instead condition on the event that one of them is a girl born on a Monday, then the probability changes! Carry out the calculation, and it should be  $3/7$  for a 7-day week.

**Example 1.4.2.** A drug test is 98% accurate, i.e. a drug user tests positive 98% of the time and a non-drug user tests negative 98% of the time. Among a given population, it is known 2% of people use drugs. Suppose I pick a person at random in the population and this person tests positive. What is the probability that the person is a drug user? Define the events

$A = \text{the person is a drug user}$  and  $B = \text{the person tests positive}$ .

Then the goal is to compute  $\mathbb{P}(A|B)$ . The 98% accuracy assumption implies that  $\mathbb{P}(B|A) = 0.98$ . Now

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|A)\mathbb{P}(A)}{\mathbb{P}(B)}.$$

Note that  $B = (B \cap A) \cup (B \cap A^c)$  and this is a disjoint union, so

$$\begin{aligned} \mathbb{P}(B) &= \mathbb{P}(B \cap A) + \mathbb{P}(B \cap A^c) = \mathbb{P}(B|A)\mathbb{P}(A) \\ &\quad + \mathbb{P}(B|A^c)\mathbb{P}(A^c) = 0.98(0.02) + 0.02(0.98) = 2(0.98)(0.02). \end{aligned}$$

Here we noted that the 98% accuracy of the test also implies that  $\mathbb{P}(B|A^c) = 0.02$ . Thus we get

$$\mathbb{P}(A|B) = \frac{0.98(0.02)}{2(0.98)(0.02)} = \frac{1}{2}.$$

Compute as an exercise that  $\mathbb{P}(A^c|B^c) = 0.996$ .

**Remark.** This test is designed clear non-drug users, not to identify drug users.

## 1.3 Bayes' Theorem

**Definition 1.5.** A *partition* of  $\Omega$  is a collection or sequence of events  $\{B_k\}_{k=1}^{\infty}$  such that

$$B_i \cap B_j = \emptyset \text{ for } i \neq j \quad \text{and} \quad \Omega = \bigcup_{k=1}^{\infty} B_k.$$

**Remark.** For any event  $A$ , observe that

$$A = A \cap \Omega = A \cap \left( \bigcup_{k=1}^{\infty} B_k \right) = \bigcup_{k=1}^{\infty} (A \cap B_k).$$

Then we get

$$\mathbb{P}(A) = \mathbb{P}\left(\bigcup_{k=1}^{\infty} (A \cap B_k)\right) = \sum_{k=1}^{\infty} \mathbb{P}(A \cap B_k) = \sum_{k=1}^{\infty} \mathbb{P}(A|B_k)\mathbb{P}(B_k).$$

This is the *partition theorem* in the book (Grimmett and Welsh). Now observe that

$$\mathbb{P}(B_i|A) = \frac{\mathbb{P}(B_i \cap A)}{\mathbb{P}(A)} = \frac{\mathbb{P}(B_i \cap A)}{\sum_{k=1}^{\infty} \mathbb{P}(A|B_k)\mathbb{P}(B_k)} = \frac{\mathbb{P}(A|B_i)\mathbb{P}(B_i)}{\sum_{k=1}^{\infty} \mathbb{P}(A|B_k)\mathbb{P}(B_k)}$$

for each  $i = 1, 2, \dots$ . This is *Bayes' theorem*, which relates posterior probabilities to prior probabilities.

## 1.4 Conditional Independence

**Proposition 1.4.** *Let  $B$  be such that  $\mathbb{P}(B) > 0$ . Then  $Q : \mathcal{F} \rightarrow [0, 1]$  given by  $A \mapsto Q(A) = \mathbb{P}(A|B)$  is a probability measure.*

*Proof.* (i) Observe that

$$Q(\Omega) = \mathbb{P}(\Omega|B) = \frac{\mathbb{P}(\Omega \cap B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B)}{\mathbb{P}(B)} = 1.$$

(ii) Let  $\{A_k\}_{k=1}^{\infty} \subseteq \mathcal{F}$  be pairwise disjoint. Then observe that we have

$$\begin{aligned} Q\left(\bigcup_{k=1}^{\infty} A_k\right) &= \mathbb{P}\left(\bigcup_{k=1}^{\infty} A_k|B\right) = \frac{\mathbb{P}(\bigcup_{k=1}^{\infty} (A_k \cap B))}{\mathbb{P}(B)} \\ &= \frac{\sum_{k=1}^{\infty} \mathbb{P}(A_k \cap B)}{\mathbb{P}(B)} = \sum_{k=1}^{\infty} \mathbb{P}(A_k|B) = \sum_{k=1}^{\infty} Q(A_k). \end{aligned}$$

Thus  $Q$  is indeed a probability measure. □

**Definition 1.6.** Two events  $A$  and  $B$  are called *independent*, written  $A \perp\!\!\!\perp B$ , if  $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$ .

**Example 1.6.1.** Let  $A, B$  be events with  $\mathbb{P}(A) = 0.6$  and  $\mathbb{P}(B) = 0.8$ . Then  $0.4 \leq \mathbb{P}(A \cap B) \leq 0.6$ . This is because  $A \cap B \subseteq A$  implies  $\mathbb{P}(A \cap B) \leq \mathbb{P}(A) = 0.6$ . Also noting that  $A \cap B \subseteq \Omega$  and so  $\mathbb{P}(A \cap B) \leq \mathbb{P}(\Omega) = 1$  implies that

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cup B) \geq 0.6 + 0.8 - 1 = 0.4.$$

Note that if  $A$  and  $B$  were independent, then we can immediately conclude  $\mathbb{P}(A \cap B) = 0.6(0.8) = 0.48$ .

**Proposition 1.5.** *Assume  $A$  and  $B$  are events with  $0 < \mathbb{P}(A), \mathbb{P}(B) < 1$ . The following are equivalent:*

1.  $A$  and  $B$  are independent,
2.  $\mathbb{P}(A|B) = \mathbb{P}(A)$ ,
3.  $\mathbb{P}(B|A) = \mathbb{P}(B)$ ,
4.  $\mathbb{P}(A^c|B) = \mathbb{P}(A^c)$ ,
5. and  $\mathbb{P}(B^c|A) = \mathbb{P}(B^c)$ .

*Proof.* Note that  $\mathbb{P}(A|B)\mathbb{P}(B) = \mathbb{P}(A \cap B)$  and  $A \perp\!\!\!\perp B$  if and only if  $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$ . Then use cancellation since  $\mathbb{P}(B) \neq 0$ . Work out the rest as an exercise.  $\square$

**Definition 1.7.** We say that three events  $A, B, C$  are *independent* if  $A, B, C$  are pairwise independent<sup>5</sup> and  $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$ .

**Remark.** Pairwise independence does not imply independence. Consider flipping a fair coin twice. Let

$$A = \{\text{first flip is } T\}, \quad B = \{\text{second flip is } H\}, \quad C = \{\text{both flips are the same}\}.$$

Then  $A, B, C$  are pairwise independent but  $\mathbb{P}(A \cap B \cap C) = 0 \neq \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = 1/8$ .

## 1.5 Homework Problems

Problems #1, 2, 9, 10, 14, 16, 17, 19 from Grimmett and Welsh.

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<sup>5</sup>i.e.  $\mathbb{P}(X \cap Y) = \mathbb{P}(X)\mathbb{P}(Y)$  for any  $X, Y \in \{A, B, C\}$ .