

9.3. RECTANGULAR (OR UNIFORM) DISTRIBUTION

Definition. A random variable X is said to have a continuous rectangular (uniform) distribution over an interval (a, b) , i.e., $(-\infty < a < b < \infty)$, if its p.d.f. is given by :

$$f(x; a, b) = \begin{cases} \frac{1}{b-a}, & \text{if } a < x < b \\ 0, & \text{otherwise} \end{cases} \quad \dots (9.19)$$

Remarks 1. a and b , ($a < b$) are the two parameters of the distribution. The distribution is called uniform distribution on (a, b) since it assumes a constant (uniform) value for all x in (a, b) .

2. The distribution is also known as rectangular distribution, since the curve $y = f(x)$ describes a rectangle over the x -axis and between the ordinates at $x = a$ and $x = b$.

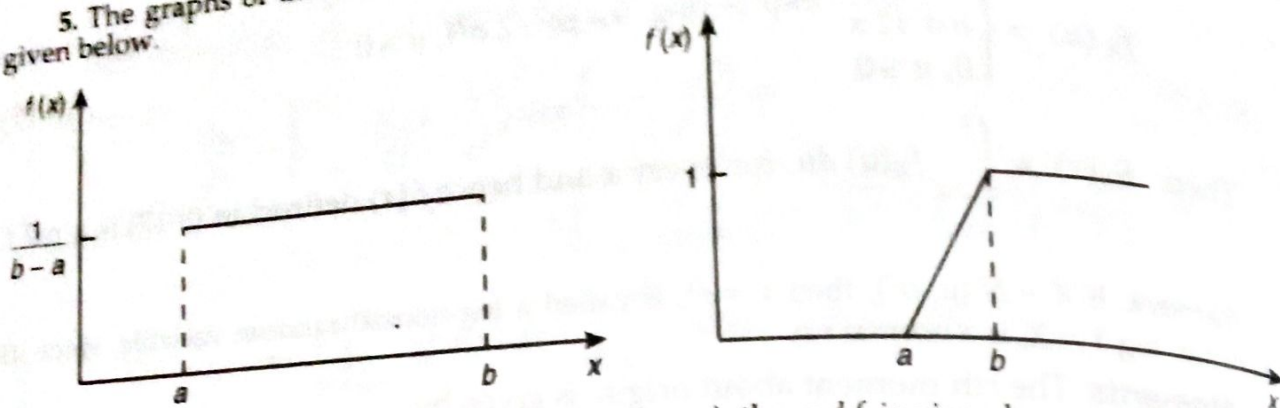
3. A uniform or rectangular variate X on the interval (a, b) is written as : $X \sim U[a, b]$ or $X \sim R[a, b]$.

4. The cumulative distribution function $F(x)$ is given by :

$$F(x) = \begin{cases} 0 & , \quad x \leq a \\ \frac{x-a}{b-a} & , \quad a < x < b \\ 1 & , \quad x \geq b \end{cases} \quad \dots (9.19a)$$

Since $F(x)$ is not continuous at $x = a$ and $x = b$, it is not differentiable at these points. Thus $\frac{d}{dx} F(x) = f(x) = \frac{1}{b-a} \neq 0$, exists everywhere except at the points $x = a$ and $x = b$ and consequently p.d.f. $f(x)$ is given by (9.19).

5. The graphs of uniform p.d.f. $f(x)$ and the corresponding distribution function $F(x)$ are given below.



6. For a rectangular or uniform variate X in $(-a, a)$, the p.d.f. is given by :

$$f(x) = \begin{cases} \frac{1}{2a}, & -a < x < a \\ 0, & \text{otherwise} \end{cases} \quad \dots (9.19b)$$

9.3.1. Moments of Rectangular Distribution. Let $X \sim U[a, b]$.

$$\mu'_r = \int_a^b x^r f(x) dx = \frac{1}{b-a} \int_a^b x^r dx = \frac{1}{b-a} \left(\frac{b^{r+1} - a^{r+1}}{r+1} \right) \quad \dots (9.20)$$

In particular

$$\text{Mean} = \mu'_1 = \frac{1}{b-a} \left(\frac{b^2 - a^2}{2} \right) = \frac{b+a}{2} \quad \dots (9.20a)$$

$$\text{and} \quad \mu'_2 = \frac{1}{b-a} \left(\frac{b^3 - a^3}{3} \right) = \frac{1}{3} (b^2 + ab + a^2)$$

$$\therefore \text{Variance} = \mu'_2 - \mu_1'^2 = \frac{1}{3} (b^2 + ab + a^2) - \left\{ \frac{1}{2} (b+a) \right\}^2 = \frac{1}{12} (b-a)^2 \quad \dots (9.20b)$$

9.3.2. M.G.F. of Rectangular Distribution is given by :

$$M_X(t) = \int_a^b e^{tx} f(x) dx = \int_a^b \frac{e^{tx}}{b-a} dx = \frac{e^{bt} - e^{at}}{t(b-a)}, t \neq 0 \quad \dots (9.20c)$$

9.3.3. Characteristic Function of Rectangular Distribution is given by :

$$\phi_X(t) = \int_a^b e^{itx} dx = \frac{e^{ibt} - e^{iat}}{it(b-a)}, t \neq 0. \quad \dots (9.20d)$$

9.3.4. Mean Deviation about Mean, η of Rectangular Distribution is given by :

$$\begin{aligned} \eta &= E |X - \text{Mean}| = \int_a^b |x - \text{Mean}| f(x) dx \\ &= \frac{1}{b-a} \int_a^b \left| x - \frac{a+b}{2} \right| dx = \frac{1}{b-a} \int_{-(b-a)/2}^{(b-a)/2} |t| dt, \text{ where } t = x - \frac{a+b}{2} \\ &= \frac{1}{b-a} \cdot 2 \int_0^{(b-a)/2} t dt = \frac{b-a}{4} \end{aligned} \quad \dots (9.20e)$$

Example 9.21. If X is uniformly distributed with mean 1 and variance $\frac{4}{3}$, find $P(X < 0)$.

Solution. Let $X \sim U[a, b]$, so that $p(x) = \frac{1}{b-a}$, $a < x < b$. We are given :

$$\text{Mean} = \frac{1}{2}(b+a) = 1 \Rightarrow b+a=2 \text{ and } \text{Var}(X) = \frac{1}{12}(b-a)^2 = \frac{4}{3} \Rightarrow b-a = \pm 4.$$

Solving, we get $a = -1$ and $b = 3$; ($a < b$). $\therefore p(x) = \frac{1}{4}$; $-1 < x < 3$

$$P(X < 0) = \int_{-1}^0 p(x) dx = \frac{1}{4} \Big|_x \Big|_{-1}^0 = \frac{1}{4}.$$

Example 9.22. Subway trains on a certain line run every half hour between mid-night six in the morning. What is the probability that a man entering the station at a random time during this period will have to wait at least twenty minutes?

Solution. Let the r.v. X denote the waiting time (in minutes) for the next train. Under the assumption that a man arrives at the station at random, X is distributed uniformly on $(0, 30)$, with p.d.f.,

$$f(x) = \begin{cases} \frac{1}{30}, & 0 < x < 30 \\ 0, & \text{otherwise} \end{cases}$$

The probability that he has to wait at least 20 minutes is given by :

$$P(X \geq 20) = \int_{20}^{30} f(x) dx = \frac{1}{30} \int_{20}^{30} 1. dx = \frac{1}{30}(30-20) = \frac{1}{3}.$$

Example 9.23. If X has a uniform distribution in $[0, 1]$, find the distribution (p.d.f.) of $Y = -2 \log X$. Identify the distribution also.

Solution. Let $Y = -2 \log X$. Then the distribution function G of Y is given by :

$$\begin{aligned} G_Y(y) &= P(Y \leq y) = P(-2 \log X \leq y) = P(\log X \geq -y/2) = P(X \geq e^{-y/2}) \\ &= 1 - P(X \leq e^{-y/2}) = 1 - \int_0^{e^{-y/2}} f(x) dx = 1 - \int_0^{e^{-y/2}} 1. dx = 1 - e^{-y/2} \\ \therefore g_Y(y) &= \frac{d}{dy} G(y) = \frac{1}{2} e^{-y/2}, 0 < y < \infty \end{aligned} \quad \dots (*)$$

[\because as X ranges in $(0, 1)$, $Y = -2 \log X$ ranges from 0 to ∞]

Remark. This example illustrates that if $X \sim U[0, 1]$, then $Y = -2 \log X$, has an exponential distribution with parameter $\theta = \frac{1}{2}$. [c.f. § 9.8] or $Y = -2 \log X$, has chi-square distribution with $n = 2$ degrees of freedom [c.f. Chapter 15].

Example 9.24. Show that for rectangular distribution : $f(x) = \frac{1}{2a}$ - $a < x < a$,

m.g.f. about origin is $\frac{1}{at} (\sinh at)$. Also show that moments of even order are given by :

$$\mu_{2n} = \frac{a^{2n}}{(2n+1)}.$$

Solution. M.G.F. about origin is given by :

$$M_X(t) = E(e^{tX}) = \int_{-a}^a e^{tx} f(x) dx = \frac{1}{2a} \int_{-a}^a e^{tx} dx = \frac{1}{2a} \left[\frac{e^{tx}}{t} \right]_{-a}^a = \frac{1}{2at} (e^{at} - e^{-at}) = \frac{\sinh at}{at}$$

$$= \frac{1}{at} \left\{ at + \frac{(at)^3}{3!} + \frac{(at)^5}{5!} + \dots + \frac{(at)^{2n+1}}{(2n+1)!} + \dots \right\} = 1 + \frac{a^2 t^2}{3!} + \frac{a^4 t^4}{5!} + \dots + \frac{a^{2n} t^{2n}}{(2n+1)!} + \dots$$

Since there are no terms with odd powers of t in $M(t)$, all moments of odd order about origin vanish, i.e., μ'_{2n+1} (about origin) = 0.

In particular, μ'_1 (about origin) = 0, i.e., mean = 0

Thus μ'_r (about origin) = μ_r

(Since mean is origin.)

Hence $\mu_{2n+1} = 0$; $n = 0, 1, 2, \dots$, i.e., all moments of odd order about mean vanish.

The moments of even order are given by :

$$\mu_{2n} = \text{coefficient of } \frac{t^{2n}}{(2n)!} \text{ in } M(t) = \frac{a^{2n}}{(2n+1)!}$$

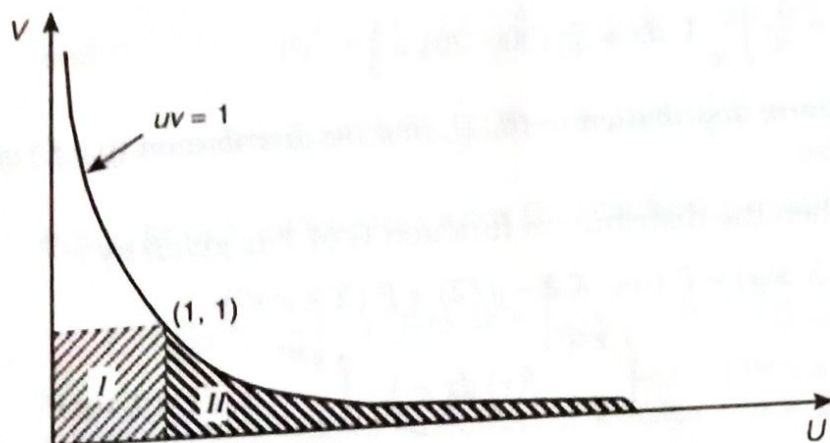
Example 9.25. If X_1 and X_2 are independent rectangular variates on $[0, 1]$, find the distributions of : (i) X_1/X_2 , (ii) $X_1 X_2$, (iii) $X_1 + X_2$, and (iv) $X_1 - X_2$.

Solution. We are given : $f_{X_1}(x_1) = f_{X_2}(x_2) = 1$; $0 < x_1 < 1$, $0 < x_2 < 1$.

Since X_1 and X_2 are independent, their joint p.d.f. is :

$$f(x_1, x_2) = f(x_1) f(x_2) = 1$$

(i) Let us transform to : $u = \frac{x_1}{x_2}$, $v = x_2$, i.e., $x_1 = uv$, $x_2 = v$



$$J = \frac{\partial (x_1, x_2)}{\partial (u, v)} = \begin{vmatrix} v & 0 \\ u & 1 \end{vmatrix} = v$$

$x_1 = 0$ maps to $u = 0$, $v = 0$

$x_1 = 1$ maps to $uv = 1$

(Rectangular hyperbola)

$x_2 = 0$ maps to $v = 0$, and

$x_2 = 1$ maps to $v = 1$.

The joint p.d.f. of U and V becomes :

$$g(uv) = f(x_1, x_2) |J| = v; 0 < u < \infty, 0 < v < \infty$$

To obtain the marginal distribution of U , we have to integrate out v .

$$\text{In region (I) : } g_1(u) = \int_0^1 v dv = \left[\frac{v^2}{2} \right]_0^1 = \frac{1}{2}, 0 \leq u \leq 1$$

$$\text{In region (II) : } g_1(u) = \int_0^{1/u} v dv = \left[\frac{v^2}{2} \right]_0^{1/u} = \frac{1}{2u^2}, 1 < u < \infty$$

$$\text{Hence the distribution } U = \frac{X_1}{X_2} \text{ is given by : } g(u) = \begin{cases} \frac{1}{2}, & 0 \leq u \leq 1 \\ \frac{1}{2u^2}, & 1 < u < \infty \end{cases}$$

(ii) Let $u = x_1 x_2, v = x_1 \Rightarrow x_1 = v, x_2 = \frac{u}{v}$ and $J = \begin{vmatrix} 0 & 1 \\ \frac{1}{v} & -\frac{u}{v^2} \end{vmatrix} = -\frac{1}{v}$

$x_1 = 0$ maps to $v = 0, x_1 = 1$ maps to $v = 1$

$x_2 = 0$ maps to $u = 0$, and $x_2 = 1$ maps to $u = v$

Moreover, $v = \frac{u}{x_2} \Rightarrow v \geq u$ (Since $0 < x_2 < 1$).

The joint p.d.f. of U and V is : $g(u, v) = f(x_1, x_2) |J| = \frac{1}{v}; 0 < u \leq v < 1$

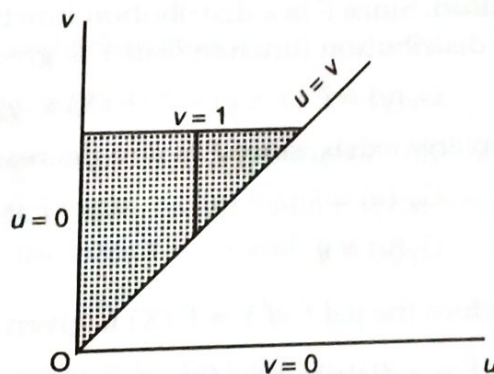
$\therefore g(u) = \int_u^1 g(u, v) dv = \int_u^1 \frac{1}{v} dv = \left| \log v \right|_u^1 = -\log u, 0 < u < 1$

(iii) and (iv). Let $u = x_1 + x_2, v = x_1 - x_2$

i.e., $x_1 = \frac{u+v}{2}$ $\left. \begin{array}{l} x_1 = 0 \Rightarrow u+v=0 \\ \text{i.e., } v=-u \\ x_2 = 0 \Rightarrow u-v=0 \\ \text{i.e., } v=u \end{array} \right\}$

$x_2 = \frac{u-v}{2}$ $\left. \begin{array}{l} x_1 = 1 \Rightarrow u+v=2 \\ x_2 = 1 \Rightarrow u-v=2 \end{array} \right\}$

and $J = \begin{vmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{vmatrix} = -\frac{1}{2}$



$\therefore g(u, v) = f(x_1, x_2) |J| = \frac{1}{2}, 0 < u < 2, -1 < v < 1$

In region (I)

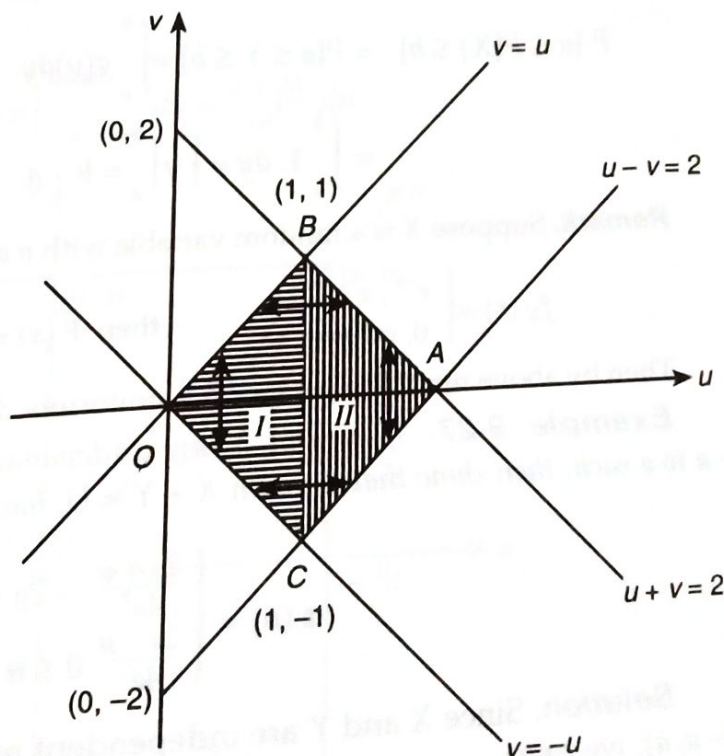
$$\begin{aligned} g_1(u) &= \int_{-u}^u \frac{1}{2} dv \\ &= \frac{1}{2} \left| v \right|_{-u}^u \\ &= u \end{aligned}$$

and in region (II),

$$\begin{aligned} g_2(u) &= \int_{u-2}^{2-u} \frac{1}{2} dv \\ &= \frac{1}{2} \left| v \right|_{u-2}^{2-u} \\ &= 2-u \end{aligned}$$

Distribution $U = X_1 + X_2$, is given

$$g(u) = \begin{cases} u, & 0 < u < 1 \\ 2-u, & 1 < u < 2 \end{cases}$$



For the distribution of V , we split the region as : OAB and OAC