

Stat 443. Time Series and Forecasting.

Key topic: Some results on conditions for AR time series to be stationary; the conditions extend to ARMA.

Introduction to some tools for studying properties of ARMA

- Covariances of linear combinations (included in this set of slides for your review).
- Recursion equations from stochastic representation
- Equations for serial correlations with different lags
- Equations for serial conditional correlations (later)

Review for covariances

(a) If X, Y are random variables,

$$\text{Cov}(X, Y) = \text{Cov}(Y, X) = E[(X - \mu_X)(Y - \mu_Y)] = E[(X - \mu_X)Y].$$

(b) X is a random variable,

$L(\cdot) = \text{Cov}(X, \cdot) = E[(X - \mu_X)\cdot]$ is a linear operator over random variables; that is: if c_1, \dots, c_m are reals and Y_1, \dots, Y_m are random variables,

$$\text{Cov}(X, c_1 Y_1 + \dots + c_m Y_m) = \sum_{j=1}^m c_j \text{Cov}(X, Y_j)$$

(c) By symmetry of $\text{Cov}(\cdot, \cdot)$, then $\text{Cov}(\cdot, X)$ is a linear operator.

(d) $\text{Var}(X) = \text{Cov}(X, X)$: use the bilinear property of Cov to derive $\text{Var}(c_1 Y_1 + c_2 Y_2)$ when $X = c_1 Y_1 + c_2 Y_2$:

$$\begin{aligned}
 X &= c_1 Y_1 + c_2 Y_2 \\
 \text{Cov}(X, c_1 Y_1 + c_2 Y_2) &= c_1 \text{Cov}(X, Y_1) + c_2 \text{Cov}(X, Y_2) \\
 &= c_1 \text{Cov}(c_1 Y_1 + c_2 Y_2, Y_1) + c_2 \text{Cov}(c_1 Y_1 + c_2 Y_2, Y_2) \\
 &= c_1^2 \text{Cov}(Y_1, Y_1) + c_1 c_2 \text{Cov}(Y_2, Y_1) + c_2 c_1 \text{Cov}(Y_1, Y_2) + c_2^2 \text{Cov}(Y_2, Y_2) \\
 &= c_1^2 \text{Var}(Y_1) + c_2^2 \text{Var}(Y_2) + 2c_1 c_2 \text{Cov}(Y_1, Y_2) \\
 &= \text{Var}(c_1 Y_1 + c_2 Y_2)
 \end{aligned}$$

Extend to the result for covariance of linear combinations of random variables:

$$\text{Cov}\left(\sum_{i=1}^m c_i X_i, \sum_{j=1}^n d_j Y_j\right) = \sum_{i=1}^m \sum_{j=1}^n c_i d_j \text{Cov}(X_i, Y_j).$$

From this, get the variance of a single linear combination.

Next slides:

Derivation of conditions on the AR coefficients ϕ_1, ϕ_2, \dots in order that the AR(p) time series is stationary

Idea: Look at AR(1) and AR(2), from which general results for AR(p) can be conjectured/obtained.

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AR(1): let $\tilde{Y}_t = Y_t - \mu$. $\tilde{Y}_i = \phi\tilde{Y}_{i-1} + \epsilon_i$. Then

$$\begin{aligned}
 \tilde{Y}_t &= \phi\tilde{Y}_{t-1} + \epsilon_t = \phi(\phi\tilde{Y}_{t-2} + \epsilon_{t-1}) + \epsilon_t \\
 &= \phi^2\tilde{Y}_{t-2} + \phi\epsilon_{t-1} + \epsilon_t = \phi^2(\phi\tilde{Y}_{t-3} + \epsilon_{t-2}) + \phi\epsilon_{t-1} + \epsilon_t \\
 &= \phi^3\tilde{Y}_{t-3} + \phi^2\epsilon_{t-2} + \phi\epsilon_{t-1} + \epsilon_t \\
 &= \phi^k\tilde{Y}_{t-k} + \sum_{j=0}^{k-1} \phi^j \epsilon_{t-j} \\
 \tilde{Y}_{i+k} &= \phi^k\tilde{Y}_i + \sum_{j=0}^{k-1} \phi^j \epsilon_{i+k-j} \text{ shifting subscript index}
 \end{aligned}$$

If $|\phi| > 1$ and \tilde{Y}_i becomes much larger than 1 for some i , then the time series grows exponentially fast in absolute value.

If $|\phi| > 1$, the time series is not stationary.

If $\phi = \pm 1$, the variance of \tilde{Y}_{i+k} increases with k for fixed i , so the time series is not stationary.

Hence $|\phi| < 1$ is a necessary condition for (weak) stationarity.

AR(1) acf in the case of stationarity

AR(1): let $\tilde{Y}_t = Y_t - \mu$. $\tilde{Y}_i = \phi \tilde{Y}_{i-1} + \epsilon_i$.

$$\tilde{Y}_{i+k} = \phi^k \tilde{Y}_i + \sum_{j=0}^{k-1} \phi^j \epsilon_{i+k-j}$$

If time series is stationary, let $\sigma_Y^2 = \text{Var}(Y_t) = \text{Var}(\tilde{Y}_t)$. Apply covariance with \tilde{Y}_i to the above equation.

$$\begin{aligned}\gamma_k &:= \text{Cov}(\tilde{Y}_{i+k}, \tilde{Y}_i) = \phi^k \text{Cov}(\tilde{Y}_i, \tilde{Y}_i) + \sum_{j=0}^{k-1} \phi^j \text{Cov}(\epsilon_{i+k-j}, \tilde{Y}_i) \\ &= \phi^k \sigma_Y^2 + 0 \\ \rho_k &= \text{Cor}(\tilde{Y}_{i+k}, \tilde{Y}_i) = \gamma_k / \sigma_Y^2 = \phi^k\end{aligned}$$

Since a correlation is between -1 and 1 , this is consistent with the necessary condition for weak stationarity that $|\phi| < 1$.

AR(2): let $\tilde{Y}_t = Y_t - \mu$. If stationary, then

$$\begin{aligned}
 \tilde{Y}_t &= \phi_1 \tilde{Y}_{t-1} + \phi_2 \tilde{Y}_{t-2} + \epsilon_t \\
 \sigma_Y^2 &= \phi_1^2 \sigma_Y^2 + \phi_2^2 \sigma_Y^2 + 2\phi_1 \phi_2 \rho_1 \sigma_Y^2 + \sigma_\epsilon^2 \\
 \sigma_\epsilon^2 &= (1 - \phi_1^2 - \phi_2^2 - 2\phi_1 \phi_2 \rho_1) \sigma_Y^2 > 0 \\
 0 &< (1 - \phi_1^2 - \phi_2^2 - 2\phi_1 \phi_2 \rho_1) \\
 \text{Cov}(\tilde{Y}_t, \tilde{Y}_{t-1}) &= \phi_1 \text{Cov}(\tilde{Y}_{t-1}, \tilde{Y}_{t-1}) + \phi_2 \text{Cov}(\tilde{Y}_{t-2}, \tilde{Y}_{t-1}) + \text{Cov}(\epsilon_t, \tilde{Y}_{t-1}) \\
 \rho_1 \sigma_Y^2 &= \phi_1 \sigma_Y^2 + \phi_2 \rho_1 \sigma_Y^2 + 0 \\
 \rho_1 &= \phi_1 / (1 - \phi_2) \in (-1, 1) \\
 \text{Cov}(\tilde{Y}_t, \tilde{Y}_{t-2}) &= \phi_1 \text{Cov}(\tilde{Y}_{t-1}, \tilde{Y}_{t-2}) + \phi_2 \text{Cov}(\tilde{Y}_{t-2}, \tilde{Y}_{t-2}) + \text{Cov}(\epsilon_t, \tilde{Y}_{t-2}) \\
 \rho_2 \sigma_Y^2 &= \phi_1 \rho_1 \sigma_Y^2 + \phi_2 \sigma_Y^2 + 0
 \end{aligned}$$

Examples of equations involving autocovariances/autocorrelations

If $E(\tilde{Y}_t) = 0$ for all t , then it looks like expected values can be 0 on both sides. The above equations come from

$$\begin{aligned}
 \text{Var}(\tilde{Y}_t) &= \text{Var}(\phi_1 \tilde{Y}_{t-1} + \phi_2 \tilde{Y}_{t-2} + \epsilon_t) \\
 \text{Cov}(\tilde{Y}_t, \tilde{Y}_{t-1}) &= \text{Cov}(\phi_1 \tilde{Y}_{t-1} + \phi_2 \tilde{Y}_{t-2} + \epsilon_t, \tilde{Y}_{t-1})
 \end{aligned}$$

One condition for weak stationarity of AR(2) is that $|\phi_2| < 1$, but $|\phi_1|$ can be > 1 . A proof of the ϕ_2 condition will be given later.

AR(p) extension: one condition for stationarity is that $|\phi_p| < 1$.

AR(2): if $|\phi_2| < 1$ and $\rho_1 = \phi_1/(1 - \phi_2) \in (-1, 1)$, and then

$$-1 < \phi_1/(1 - \phi_2) < 1$$

$$\phi_2 - 1 < \phi_1 < 1 - \phi_2$$

$$\phi_1 + \phi_2 < 1$$

$$\phi_2 - \phi_1 < 1$$

The three equations in blue define the triangular region of the (ϕ_1, ϕ_2) space for AR(2) to be stationary.

Exercise: draw the region (to be shown later).

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Pseudo-code for AR(p) forecast rmse; R function: `arp_fc = function(train,holdout,arvec,mu)`
Forecast is linear in the previous p observations (assuming stationary):

- input `train` with size n_{train} , `holdout` with size $n_{holdout}$, estimates $\hat{\phi} = (\hat{\phi}_1, \dots, \hat{\phi}_p)^T$ and $\hat{\mu}$.
- Join last p observations of `train` and the vector `holdout` to get a vector `z` of length $p + n_{holdout}$. Subtract $\hat{\mu}$ from each element: $\mathbf{z} \leftarrow \mathbf{z} - \hat{\mu}\mathbf{1}$.
- $\text{sse} \leftarrow 0$
- for i in $1, \dots, n_{holdout}$:
- $\mathbf{zprev} \leftarrow (z_{i+p-1}, \dots, z_{i+1}, z_i)^T$, $\mathbf{fc} \leftarrow \hat{\mu} + \hat{\phi}^T \mathbf{zprev}$ (same as $\hat{\mu} + \sum_{j=1}^p \hat{\phi}_j (y_{n_{train}+i-j} - \hat{\mu})$);
- $\text{fcvec}[i] \leftarrow \mathbf{fc}$; $\mathbf{yt} \leftarrow \mathbf{holdout}[i]$;
- $\text{fcerror} \leftarrow \mathbf{yt} - \mathbf{fc}$; $\text{sse} \leftarrow \text{sse} + \text{fcerror}^2$.
- end for; return $\text{rmse} = \sqrt{\text{sse}/n_{holdout}}$ and `fcvec`

Later to check if holdout set moving 1-step forecasts and rmse can be obtain via another call to `arima` with appropriate inputs

Summary

Start of tools for studying dependence properties of ARMA and ARIMA.

- Recursion equations from stochastic representation (slide 5)
- Equations for serial correlations with different lags (slide 7)
- Equations for serial conditional correlations (later)