

ECE HW#4 Kefu Zhu

Question 7

(A)

Is the process X_n a MC?

The process X_n is a Markov chain because we can write
$$\begin{cases} X_{n+1} = 0, & X_n = 0 \\ X_{n+1} = \sum_{i=1}^{X_n} D_i, & X_n > 0 \end{cases}$$

which indicate that the number of women in the $n + 1$ generation only depends on the number of women in the n generation \Leftrightarrow memory less property of markov chain.

What are the transition probabilities $P(X_{n+1} = j|X_n = 0)$ and $P(X_{n+1} = j|X_n = 1)$?

Because once the total number of women becomes zero at some point in time, the number will stay as zero and won't increase in the future. We have

$$P_{0j} = P(X_{n+1} = j|X_n = 0) = \begin{cases} 0, & j \neq 0 \\ 1, & j = 0 \end{cases}$$

As stated before, because X_n is a Markov chain, we can write

$$P_{1j} = P(X_{n+1} = j|X_n = 1) = P(\sum_{i=1}^{X_n} D_i = j|X_n = 1) = P(D_i = j) = p_j$$

What are the transition probabilities into state $X_n = 0$?

The probability of the number of women becomes zero at some point given there are still i women in the last generation is equivalent to the probability that every woman has no daughter, which can be represented as

$$P(X_{n+1} = 0|X_n = i) = \prod_{k=1}^i P(D_k = 0) = p_0^i$$

Is the probability $P(X_{n+1} = i|X_n = i)$ of a state transitioning into itself strictly positive? Is this MC recurrent?

As stated before, $P(X_{n+1} = i|X_n = i) = 1$, for $i = 0$

For $i > 0$, we know that one way of having the same number of women for two generations in a row is that

every woman has exactly one daughter, with probability p_1^i . Therefore, we have

$$P(X_{n+1} = i | X_n = i) \geq p_1^i > 0, \text{ for } i > 0$$

Since $P(X_{n+1} = 0 | X_n = 0) = 1$, the state 0 is recurrent. Because $\forall i \neq 0, p_0^i > 0$, all states that is not state 0 are transient (Because there is a probability that state i goes to state 0 and then stays there forever \rightarrow never come back to state i).

The MC is not recurrent.

(B)

Is X_{rN} not a MC?

For the special case, state 0, which means that there are no women of type r in the current generation, can happens because of two scenarios

1. Type r has not been created so far $\rightarrow P(X_{rm} > 0 | X_{rn} = 0) = 0$
2. Type r existed in the past but is now extinct $\rightarrow P(X_{rm}) > 0$

Since the transition probabilities to other states starting from state 0 depend on the past scenarios, this process must not be a Markov chain

Given $X_{rm} > 0$, is the process $X_{r,n:\infty} = X_{rm}, X_{r,n+1}, \dots$ a MC?

Because we are conditioning on $X_{rm} > 0$, we eliminate the other scenario for the state 0. The process now is a Markov chain. Now we have

$$P_{0j} = \begin{cases} 1, & j = 0 \\ 0, & j \neq 0 \end{cases}$$

$P_{1j} = (1 - q)p_j$, which indicate the probability of having exactly j daughters without the mutation

What is the value for P_{i0} ?

Consider one scenario of P_{i0} ,

$P_{10} = p_0 + (1 - p_0)q$ because the number of women who has the mutation becomes extinct if one of the following two mutually exclusive scenarios happens:

1. woman has no daughters
2. woman has daughters but her daughters has type r

Because of independence, $P_{i0} = P_{10}^i = (p_0 + (1 - p_0)q)^i, i \geq 0$

Is $P_{ii} > 0$? Is this MC recurrent

Same logic as stated in part A, only state 0 is recurrent. All other states are transient.

$$P_{ii} \geq ((1 - q)p_i)^i > 0, i \geq 0$$

The MC is not recurrent

(C)

```
% Set simulation parameter
X_o = 100;
max_t = 50;
max_types = 1000;

% Set stochastic process parameters
mu = 1.05;
q = 10^-2;
% Initialize empty matrix with all zeros for storing population size
X=zeros(max_types, max_t);
% Initialize empty matrix with all zeros for storing number of types
number_of_types=zeros(1, max_t);
% Initialize population (1 people)
X(1:X_o,1) = 1;
% Initialize first generation
number_of_types(1)=X_o;
number_of_extinct_types=zeros(1,max_t);

% Start simulation
for n=2:max_t
    disp('n = '+string(n))
    number_of_types(n)=number_of_types(n-1);
    for type = 1:number_of_types(n-1);
        for i = 1:X(type,n-1)
            % Daughter/Not Daughter
            daughters = poissrnd(mu,1,1);
            % Mutation/No mutation
            mutation = binornd(1,q,1,1);

            % If there is a mutation
            if mutation
                number_of_types(n) = number_of_types(n)+1;
                X(number_of_types(n),n) = daughters;
            % Otherwise
```

```

        else
            X(type,n) = X(type,n) + daughters;
        end
    end

    % Add the extinct type
    if X(type,n)== 0
        number_of_extinct_types(n)=number_of_extinct_types(n)+1;
    end
end
end
end

```

```

# The total number of mitochondrial DNA types by generation n
number_of_types

# The total number of extinct types by generation n
number_of_extinct

```

(D)

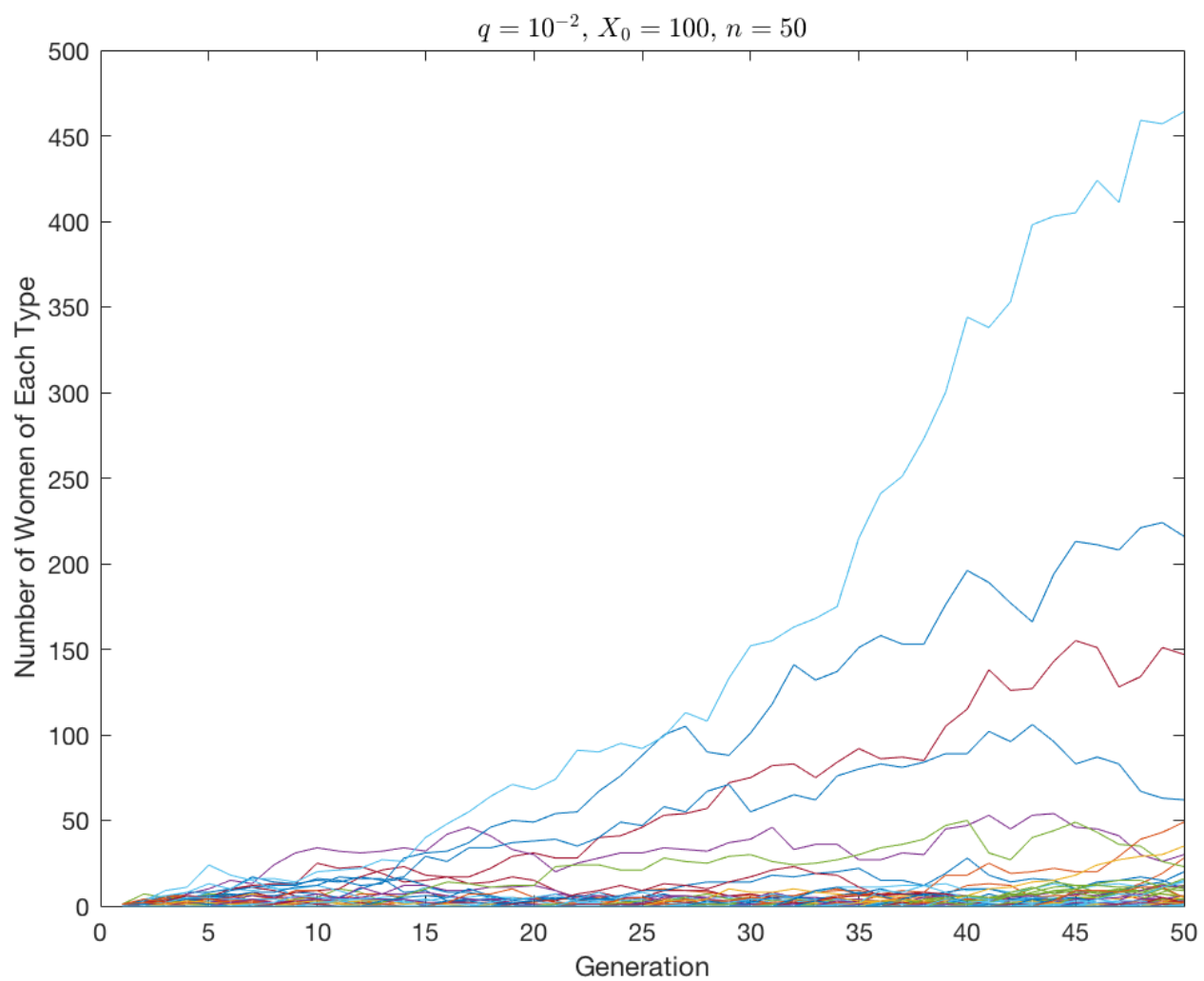
```

figure()
plot(1:max_t, X)
xlabel('Generation')
ylabel('Number of Women of Each Type')
title('$q=10^{-2}$, $X_0=100$, $n=50$', 'Interpreter', 'latex')

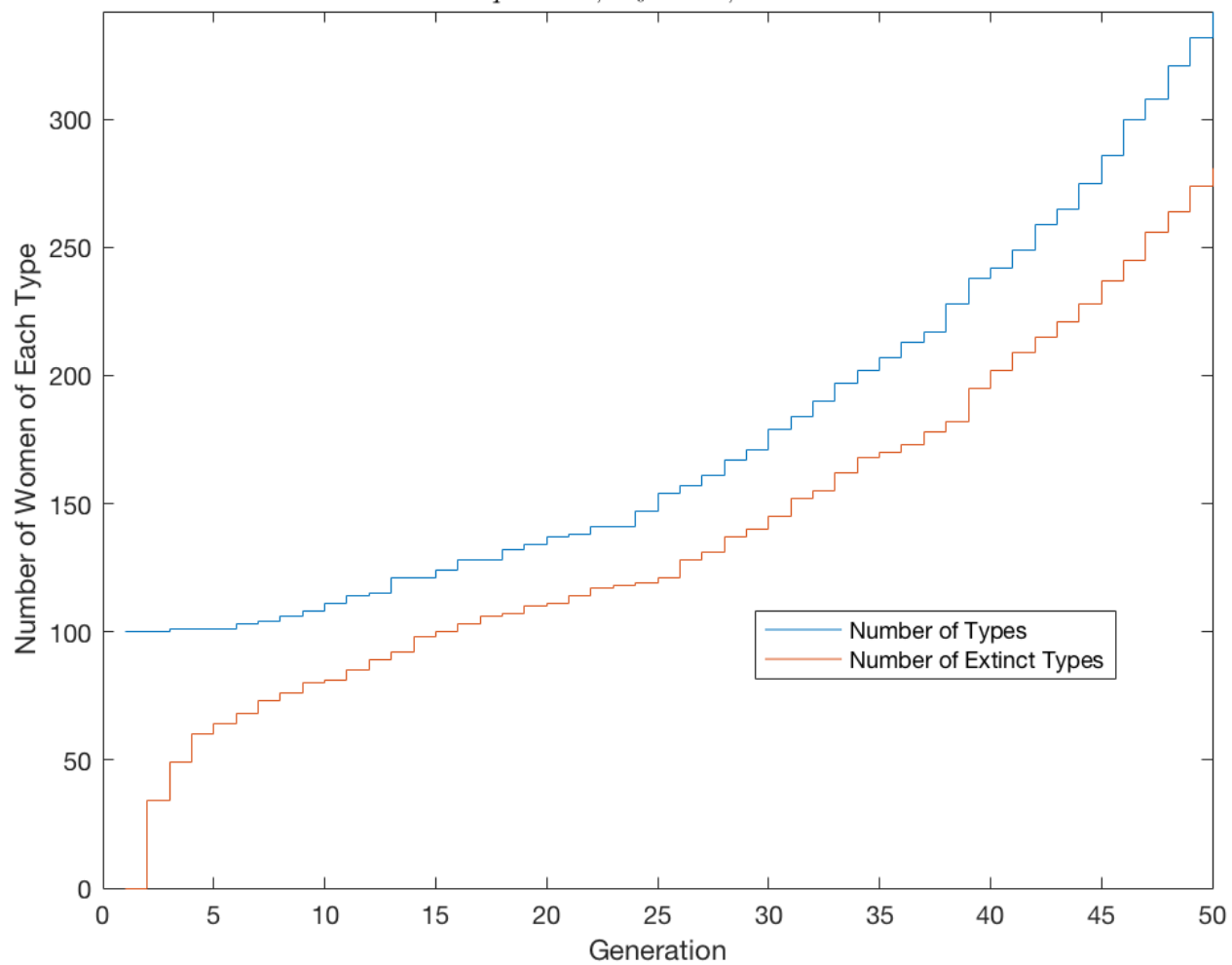
figure()
stairs(1:max_t, [number_of_types; number_of_extinct_types])
xlabel('Generation')
ylabel('Number of Women of Each Type')
title('$q=10^{-2}$, $X_0=100$, $n=50$', 'Interpreter', 'latex')
axis([0,50,0,number_of_types(end)])
legend('Number of Types', 'Number of Extinct Types', 'Location', 'Best')

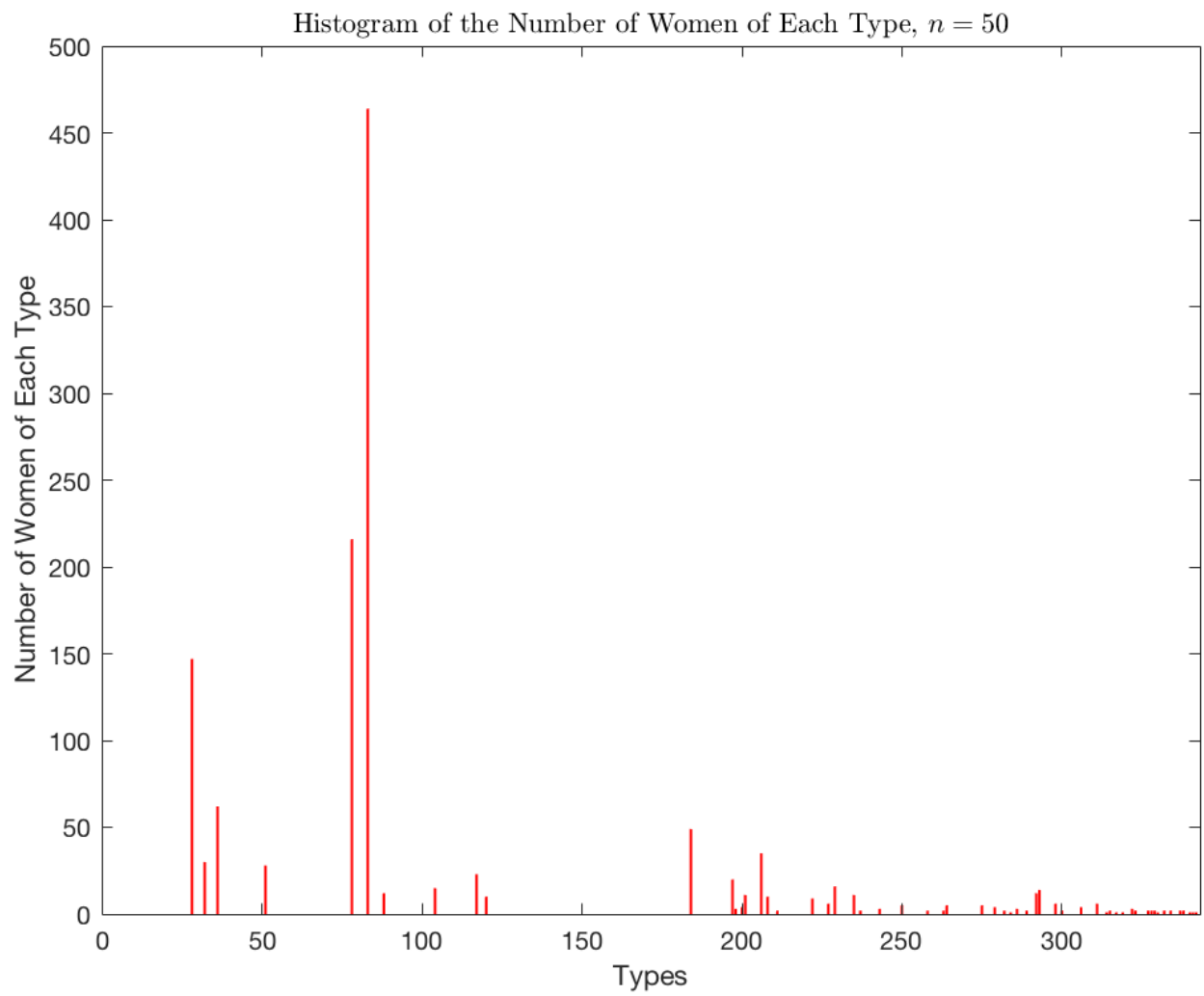
figure()
bar(1:number_of_types(end), X(1:number_of_types(end),max_t), 'r')
xlabel('Types')
ylabel('Number of Women of Each Type')
title('Histogram of the Number of Women of Each Type, $n=50$', 'Interpreter', 'latex')

```



$$q = 10^{-2}, X_0 = 100, n = 50$$





(E)

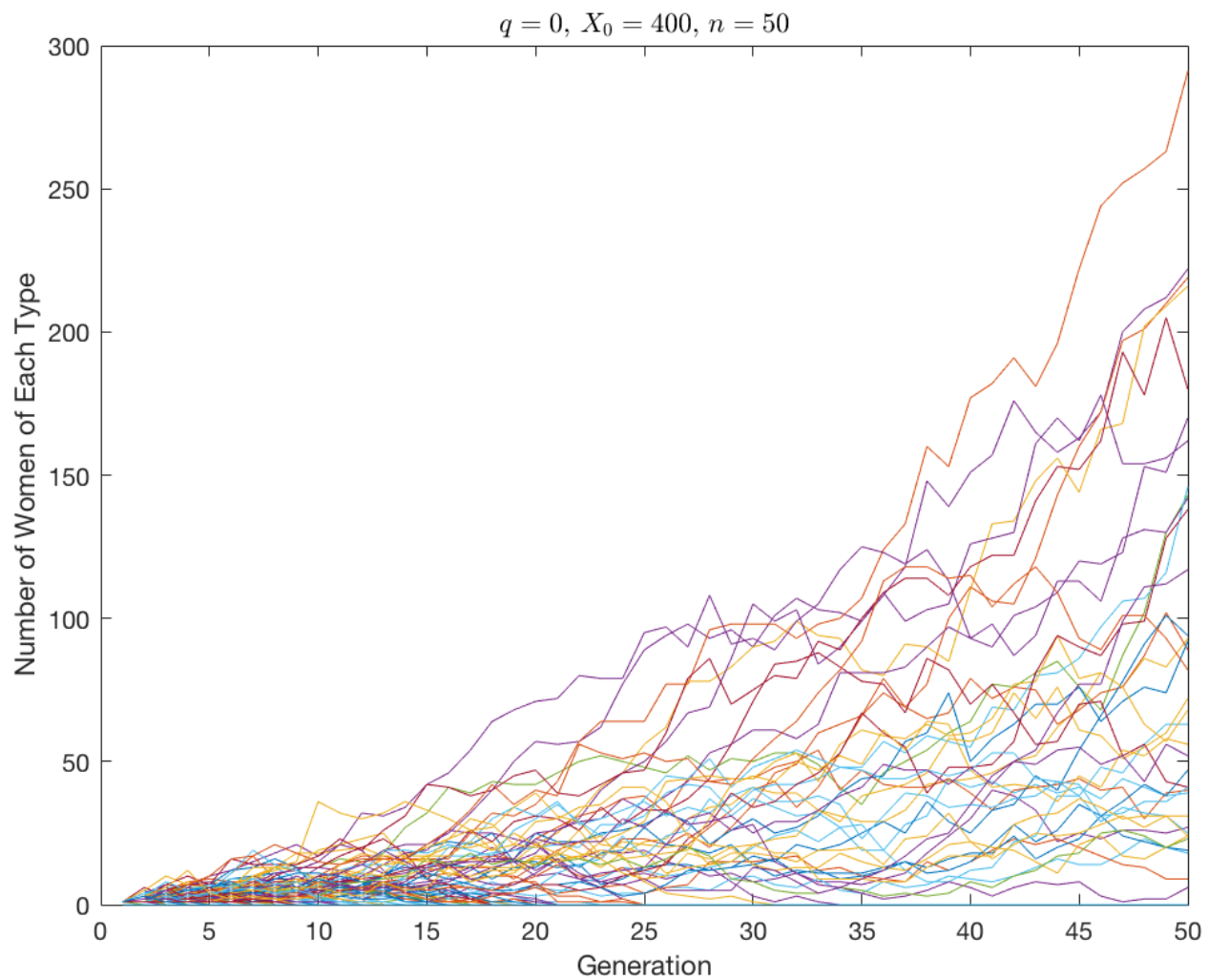
```

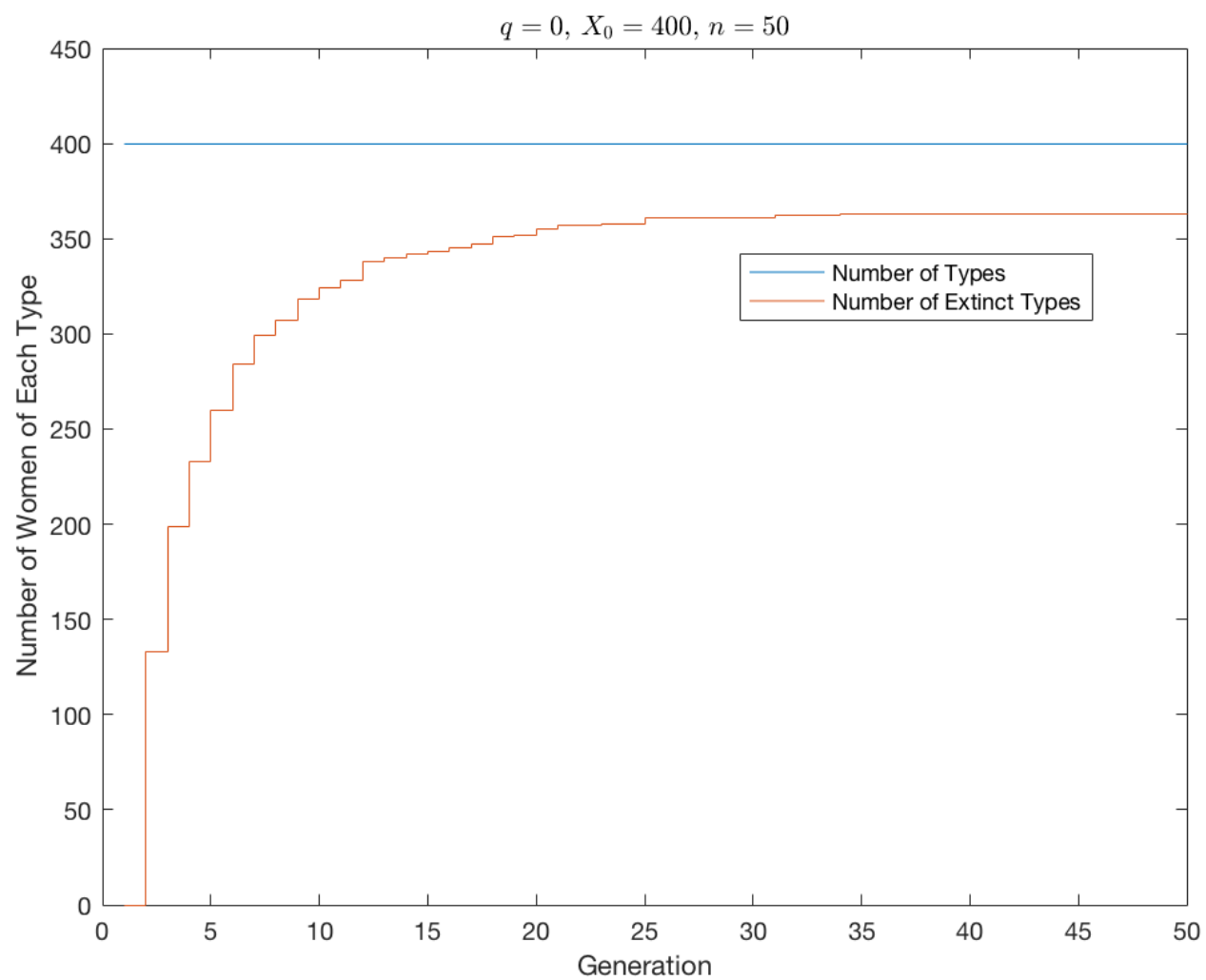
% Problem E
figure()
plot(1:max_t, X)
xlabel('Generation')
ylabel('Number of Women of Each Type')
title('$q=0, $ $X_{0}=400, $ $n=50$', 'Interpreter', 'latex')

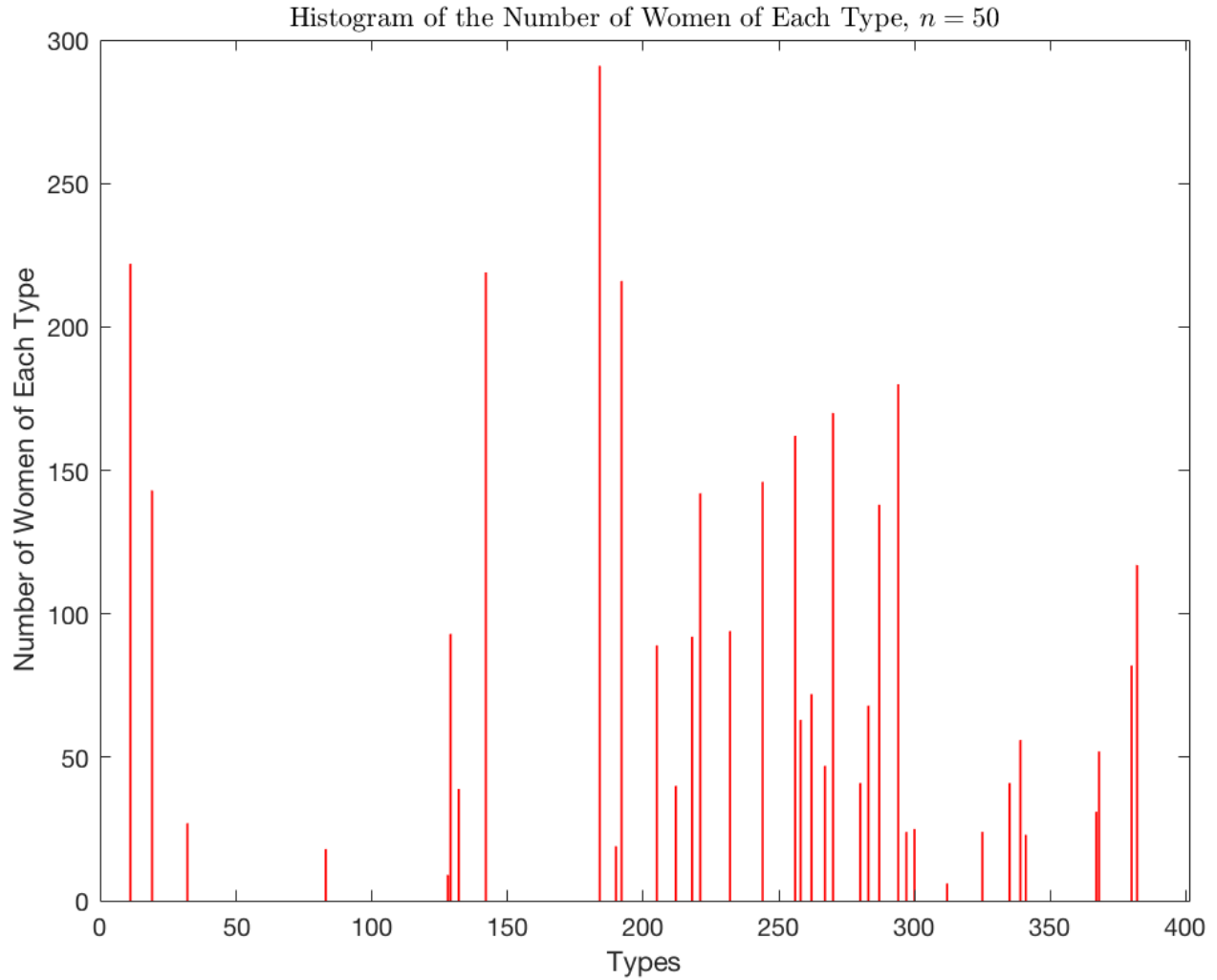
figure()
stairs(1:max_t, [number_of_types; number_of_extinct_types])
xlabel('Generation')
ylabel('Number of Women of Each Type')
title('$q=0, $ $X_{0}=400, $ $n=50$', 'Interpreter', 'latex')
axis([0, 50, 0, number_of_types(end)+50])
legend('Number of Types', 'Number of Extinct Types', 'Location', 'Best')

figure()
bar(1:number_of_types(end), X(1:number_of_types(end), max_t), 'r')
xlabel('Types')
ylabel('Number of Women of Each Type')
title('Histogram of the Number of Women of Each Type, $n=50$', 'Interpreter', 'latex')

```





(F)

$$\because E(\sum_{i=1}^{X_{n-1}} D_i | X_{n-1} = k) = E(\sum_{i=1}^k D_i | X_{n-1} = k) = E(\sum_{i=1}^k D_i)$$

As defined in the question, $E(D_i) = v$

$$\therefore E(\sum_{i=1}^{X_{n-1}} D_i | X_{n-1} = k) = kv$$

By iterated expectations

$$E(X_n) = \sum_{k=0}^{\infty} E(X_n | X_{n-1} = k) P(X_{n-1} = k) = \sum_{k=0}^{\infty} kv \cdot P(X_{n-1} = k) = v \cdot E(X_{n-1})$$

$$\text{By iterations, } E(X_n) = v \cdot E(X_{n-1}) = v^2 \cdot E(X_{n-2}) = \dots = v^n \cdot E(X_0) = v^n \cdot X_0$$

For $v > 1$ ($v = \lambda$ in the part D and E), we expect to see an exponential increase in the expected value, which is similar to what we see in the simulation graph

Consider $X_0 = X_{r0}$, we then have $E(X_{rn}) = v_r^n \cdot E(X_{r0}) = v_r^n = (1 - q)^n v^n$

(G)

Based on Markov's inequality, $P(|X| \geq a) \leq \frac{E(X)}{a}$, we have $P(|X_{rn}| \geq a) \leq \frac{v_r^n}{a}$ for some value of a

If $v_r < 1$, then $\lim_{n \rightarrow \infty} P(X_{rn} \geq a) \leq \lim_{n \rightarrow \infty} \frac{v_r^n}{a} = 0$

$\Rightarrow \lim_{n \rightarrow \infty} P(X_{rn} < a) = 1 \Leftrightarrow \lim_{n \rightarrow \infty} P(X_{rn} = 0) = 1$

(H)

$\therefore P_e(j) = \begin{cases} 1, & v < 1 \text{ (Proved in part G)} \\ < 1, & v > 1 \text{ (Type } r \text{ could extinct)} \end{cases}$

\therefore By law of total probability, consider conditioning on $X_{r1} = j$

$$P(X_{r\infty} = 0 | X_{r1} = j, X_{r0} = 1) = P(X_{r\infty} = 0 | X_{r1} = j) = P(X_{r\infty} = 0 | X_{r0} = 1)^j = (P_e(1))^j$$

$$\therefore P(X_{r\infty} = 0 | X_{r1} = j) = (P_e(1))^j$$

\therefore Again, by law of total probability,

$$P_e(1) = \sum_{j=1}^{\infty} P(X_{r\infty} = 0 | X_{r1} = j, X_{r0} = 1) P(X_{r1} = j | X_{r0} = 1) = \sum_{j=1}^{\infty} p_j \cdot (P_e(1))^j$$

If in general, the starting generation have j individuals ($X_{r0} = j$), then by independence, we have

$$P_e(j) = (P_e(1))^j$$