

Queuing Theory

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Queueing theory

M/M/1 queue

Multiserver queues

Networks of queues

- ▶ Queuing theory is concerned with the (boring) issue of waiting
 - ⇒ Waiting is boring, queuing theory not necessarily so
- ▶ “Customers” arrive to receive “service” by “servers”
 - ⇒ Between arrival and start of service wait in queue
- ▶ Quantities of interest (for example)
 - ⇒ Number of customers in queue ⇒ L (for length)
 - ⇒ Time spent in queue ⇒ W for (wait)
- ▶ Queues are a pervasive application of CTMCs

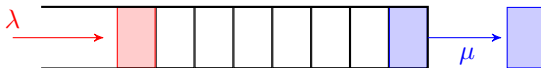


Where do queues appear?

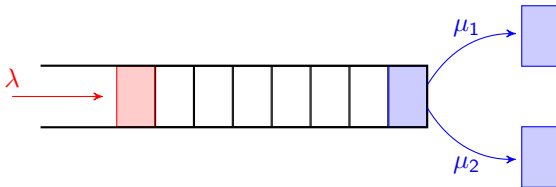
- ▶ Queues are fundamental to the analysis of (public) **transportation**
 - ▶ Wait to enter a highway \Rightarrow Customers = cars
 - ▶ **Q**: Subway travel times, subway or buses?
 - ▶ **Q**: Infrequent big buses or frequent small buses?
- ▶ Packet traffic in **communication networks**
 - ▶ Route determination, congestion management
 - ▶ Real-time requirements, delays, resource management
- ▶ **Logistics** and operations research
 - ▶ Customers = raw materials, components, final products
 - ▶ Customers in queue = products in storage = inactive capital
- ▶ **Customer service**
 - ▶ **Q**: How many representatives in a call center? Call center pooling

Examples of queues

- ▶ Simplest rendition \Rightarrow Single queue, single server, infinite spots
 - \Rightarrow Simpler if arrivals and services are Poisson \Rightarrow M/M/1 queue
 - \Rightarrow Limiting number of spots not difficult \Rightarrow Losses appear

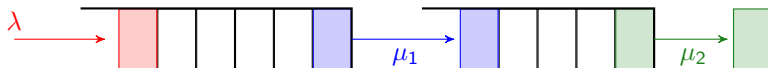


- ▶ Multi-server queues \Rightarrow Single queue, many servers
 - \Rightarrow M/M/c queue \Rightarrow c Poisson servers (i.e., exp. service times)

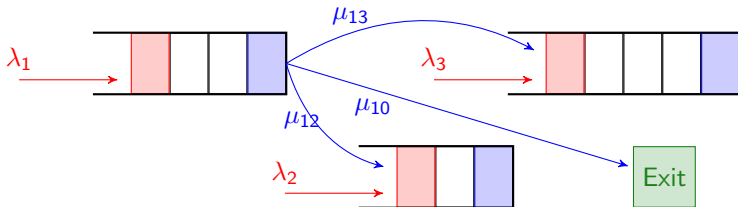


- **Groups of interacting queues** \Rightarrow Applications become interesting

Ex: A queue tandem



- Can have **arrivals at different points** and **random re-entries**



- Batch service and arrivals, loss systems (not considered)

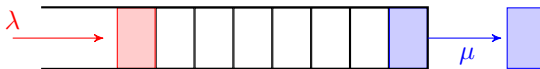
Queuing theory

M/M/1 queue

Multiserver queues

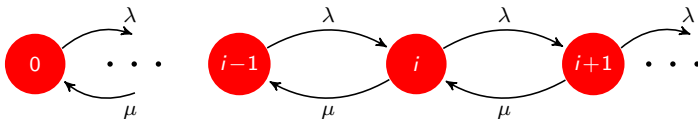
Networks of queues

- Arrival and service processes are Poisson \Rightarrow Birth & death process
 - a) Customers arrive at an average rate of λ per unit time
 - b) Customers are serviced at an average rate of μ per unit time
 - c) Interarrival and inter-service time are exponential and independent



- Hypothesis of Poisson arrivals is reasonable
- Hypothesis of exponential service times not so reasonable
 - \Rightarrow Simplifies the analysis. Otherwise, study a M/G/1 queue
- Steady-state behavior (systems operating for a long time)
 - \Rightarrow Q: Limit probabilities for the M/M/1 system?

- ▶ Define CTMC by identifying states $Q(t)$ with queue lengths
⇒ Transition rates $q_{i,i+1} = \lambda$ for all i , and $q_{i,i-1} = \mu$ for $i \neq 0$
- ▶ Recall that first of two exponential times is exponentially distributed
⇒ Mean transition times are $\nu_i = \lambda + \mu$ for $i \neq 0$ and $\nu_0 = \lambda$



- ▶ Limit distribution equations (Rate out of j = Rate into j)

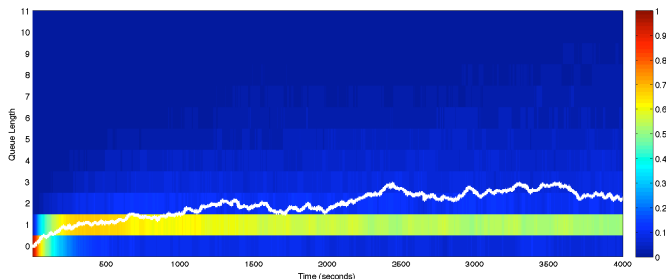
$$\lambda P_0 = \mu P_1, \quad (\lambda + \mu) P_i = \lambda P_{i-1} + \mu P_{i+1}$$

Queue length as a function of time

- ▶ Simulation for $\lambda = 30$ customers/min, $\mu = 40$ services/min
- ▶ Probability distribution estimated by sample averaging with $M = 10^5$

$$P(Q(t) = k) \approx \frac{1}{M} \sum_{i=1}^M \mathbb{I}\{Q_i(t) = k\}$$

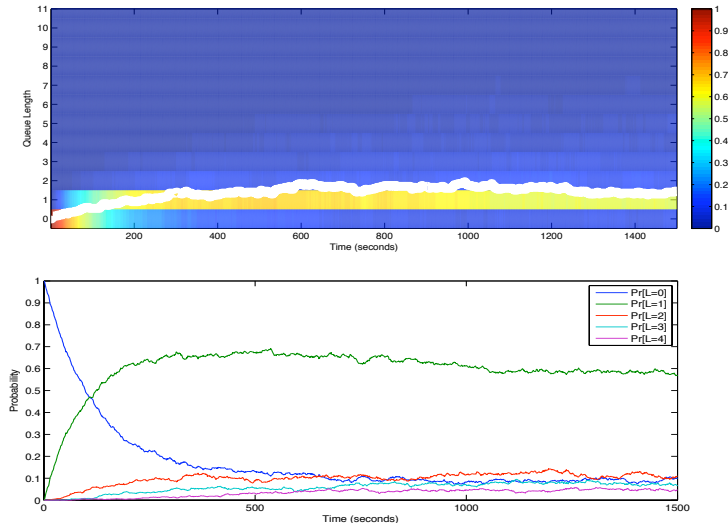
- ▶ Steady state (in a probabilistic sense) reached in around 10^3 mins.



- ▶ Queue length vs. time. Probabilities are color coded
⇒ Mean queue length shown in white

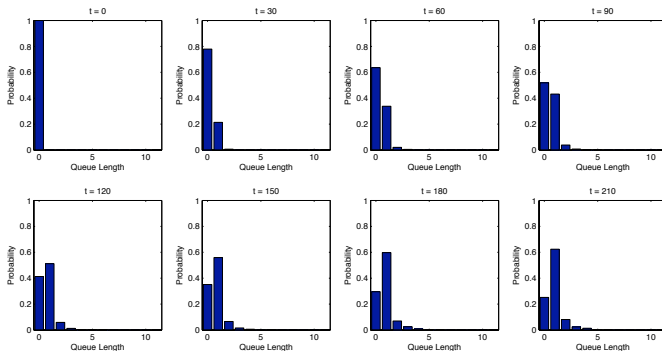
Close up on initial times

- Probabilities settle at their equilibrium values



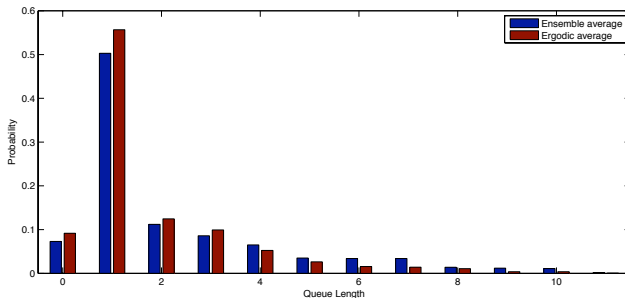
Another view of queue length evolution

- Cross-sections of queue length probabilities at different times



- Compare ensemble averages for large t with ergodic averages

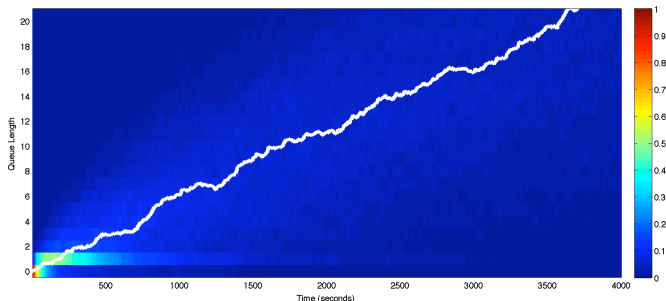
$$T_i(t) = \frac{1}{t} \int_0^t \mathbb{I}\{Q(\tau) = i\} d\tau$$



- They are approximately equal, as they should (equal as $t \rightarrow \infty$)

A non stable queue

- ▶ All former observations valid for stable queues ($\lambda < \mu$)
- ▶ Simulation for $\lambda = 60$ customers/min and $\mu = 40$, customers/min
 - ⇒ Queue length grows unbounded
 - ⇒ Probability of small number of customers in queue vanishes
 - ⇒ Actually CTMC transient, $P_i \rightarrow 0$ for all i



- ▶ Queue length vs. time. Probabilities are color coded
 - ⇒ Mean queue length shown in white

Solution of limit distribution equations

- ▶ Start expressing all prob. in terms of P_0 . Define traffic intensity $\rho := \lambda/\mu$
- ▶ Repeat process done for birth and death process
- ▶ Equation for $P_0 \Rightarrow \lambda P_0 = \mu P_1$
- ▶ Sum eqs. for P_1 and $P_0 \Rightarrow \begin{aligned} \lambda P_0 &= \mu P_1 \\ (\lambda + \mu)P_1 &= \lambda P_0 + \mu P_2 \end{aligned} \Rightarrow \lambda P_1 = \mu P_2$
- ▶ Sum result and eq. for $P_2 \Rightarrow \begin{aligned} \lambda P_1 &= \mu P_2 \\ (\lambda + \mu)P_2 &= \lambda P_1 + \mu P_3 \end{aligned} \Rightarrow \lambda P_2 = \mu P_3$
- ▶ Sum result and eq. for $P_i \Rightarrow \begin{aligned} \lambda P_{i-1} &= \mu P_i \\ (\lambda + \mu)P_i &= \lambda P_{i-1} + \mu P_{i+1} \end{aligned} \Rightarrow \lambda P_i = \mu P_{i+1}$
- ▶ From where it follows $\Rightarrow P_{i+1} = (\lambda/\mu)P_i = \rho P_i$ and recursively $P_i = \rho^i P_0$

- ▶ The sum of all probabilities is 1 (use geometric series formula)

$$1 = \sum_{i=0}^{\infty} P_i = \sum_{i=0}^{\infty} \rho^i P_0 = \frac{P_0}{1 - \rho}$$

- ▶ Solve for P_0 to obtain

$$P_0 = 1 - \rho, \quad \Rightarrow P_i = (1 - \rho)\rho^i$$

⇒ Valid for $\lambda/\mu < 1$, if not CTMC is transient (queue unstable)

- ▶ Expression coincides with non-concurrent queue in discrete time
 - ⇒ Not surprising. Continuous time \approx discrete time with small Δt
 - ⇒ For small Δt non-concurrent hypothesis is accurate
- ▶ Present derivation “much cleaner,” though

- ▶ To compute **expected queue length** $\mathbb{E}[L]$ use limit probabilities

$$\mathbb{E}[L] = \sum_{i=0}^{\infty} iP_i = \sum_{i=0}^{\infty} i(1-\rho)\rho^i$$

- ▶ Latter is derivative of geometric sum ($\sum_{i=0}^{\infty} ix^i = x/(1-x)^2$). Then

$$\mathbb{E}[L] = (1-\rho) \times \frac{\rho}{(1-\rho)^2} = \frac{\rho}{1-\rho}$$

- ▶ Recall $\lambda < \mu$ or equivalently $\rho < 1$ for queue stability
 \Rightarrow If $\lambda \approx \mu$ queue is stable but $\mathbb{E}[L]$ becomes very large

- ▶ Customer arrives, L in queue already. **Q**: Time spent in queue?
 - ⇒ Time required to service these L customers
 - ⇒ Plus time until arriving customer is served

- ▶ Let T_1, T_2, \dots, T_{L+1} be these times. Queue wait $\Rightarrow W = \sum_{i=1}^{L+1} T_i$
- ▶ Expected value (condition on $L = \ell$, then expectation w.r.t. L)

$$\mathbb{E}[W] = \mathbb{E}\left[\sum_{i=1}^{L+1} T_i\right] = \mathbb{E}\left[\mathbb{E}\left[\sum_{i=1}^{\ell+1} T_i \mid L = \ell\right]\right]$$

- ▶ $L = \ell$ “not random” in inner expectation \Rightarrow interchange with sum

$$\mathbb{E}[W] = \mathbb{E}\left[\sum_{i=1}^{L+1} \mathbb{E}[T_i]\right] = \mathbb{E}[(L+1)\mathbb{E}[T_i]] = \mathbb{E}[L+1]\mathbb{E}[T_i]$$

- Use expression for $\mathbb{E}[L]$ to evaluate $\mathbb{E}[L + 1]$ as

$$\mathbb{E}[L + 1] = \mathbb{E}[L] + 1 = \frac{\rho}{1 - \rho} + 1 = \frac{1}{1 - \rho}$$

- Substitute expressions for $\mathbb{E}[L + 1]$ and $\mathbb{E}[T_i] = 1/\mu$

$$\mathbb{E}[W] = \frac{1}{\mu} \times \frac{1}{1 - \rho} = \frac{1}{\mu - \lambda}$$

- Recall $\lambda =$ arrival rate. Former may be written as

$$\mathbb{E}[W] = \frac{1}{\lambda} \times \frac{\rho}{1 - \rho} = (1/\lambda)\mathbb{E}[L]$$

- ▶ For M/M/1 queue have just seen $\Rightarrow \mathbb{E}[L] = \lambda \mathbb{E}[W]$
 \Rightarrow Expression referred to as **Little's law**
- ▶ **True** even if arrivals and departures are **not Poisson** (not proved)
- ▶ **Expected nr.customers in queue** = **arrival rate** \times **expected wait**

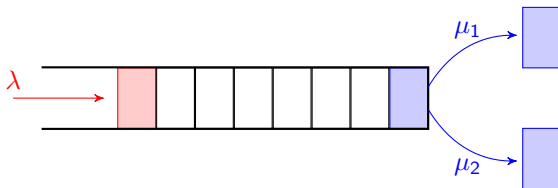
Queuing theory

M/M/1 queue

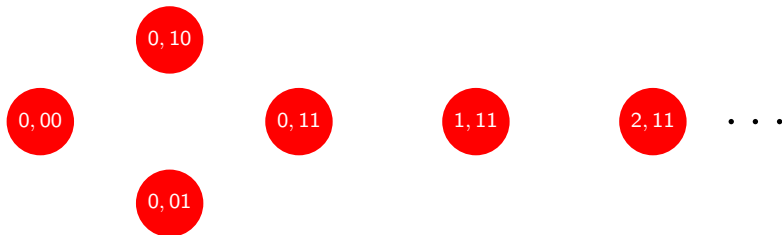
Multiserver queues

Networks of queues

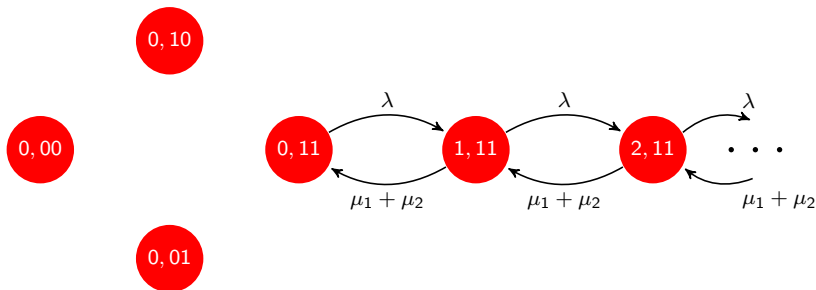
- ▶ Service offered by **two Poisson servers** with service rates μ_1 and μ_2
 - ⇒ Arrivals are Poisson with rate λ as in the M/M/1 queue
- ▶ When a server finishes serving a customer, serves next one in queue
 - ⇒ If queue is empty the server waits for the next customer
- ▶ If both servers are idle when a new customer arrives
 - ⇒ Service is performed by server 1 (simply by convention)



- ▶ When no customers are in line, need to distinguish servers' states
 - ▶ State $0, 00$ = no customers in queue, no customers being served
 - ▶ State $0, 10$ = no customers in queue, 1 customer served by server 1
 - ▶ State $0, 01$ = no customers in queue, 1 customer served by server 2
 - ▶ State $0, 11$ = no customers in queue, 2 customers in service
- ▶ When there are customers in line, both servers are busy
 - ▶ State $i, 11 = i > 0$ customers in queue and 2 customers in service
 - ▶ States $i, 01$, $i, 10$ and $i, 00$ are not possible for $i > 0$

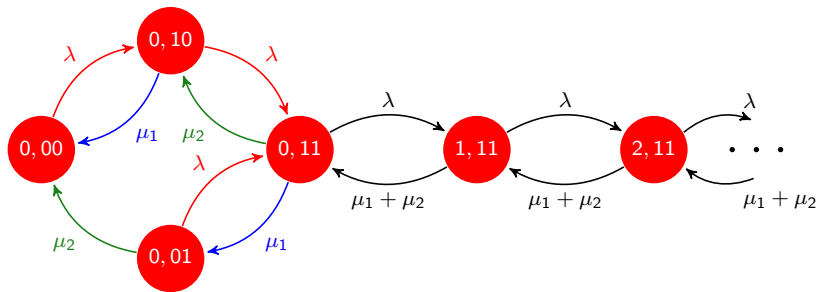


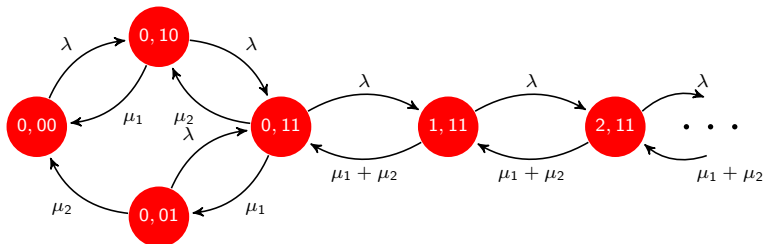
- ▶ Transition from $i, 11$ to $(i+1, 11)$ when arrival $\Rightarrow q_{i,11;(i+1),11} = \lambda$
- ▶ Transition from $i, 11$ to $(i-1, 11)$ when either server 1 or 2 finishes
 \Rightarrow First service completion by either server 1 or 2
- ▶ **Min. of two exponentials is exponential** $\Rightarrow q_{i,11;(i-1),11} = \mu_1 + \mu_2$



CTMC model: Transition rates (continued)

- ▶ From 0,00 move to 0,10 on arrival $\Rightarrow q_{0,00;0,10} = \lambda$
- ▶ From 0,10 move to 0,11 on arrival $\Rightarrow q_{0,10;0,11} = \lambda$
- ▶ From 0,01 move to 0,11 on arrival $\Rightarrow q_{0,01;0,11} = \lambda$
- ▶ From 0,10 to 0,00 when server 1 finishes $\Rightarrow q_{0,10;0,00} = \mu_1$
- ▶ From 0,11 to 0,01 when server 1 finishes $\Rightarrow q_{0,11;0,01} = \mu_1$
- ▶ From 0,01 to 0,00 when server 2 finishes $\Rightarrow q_{0,01;0,00} = \mu_2$
- ▶ From 0,11 to 0,10 when server 2 finishes $\Rightarrow q_{0,11;0,10} = \mu_2$





- For states $i, 11$ with $i > 0$, eqs. are analogous to M/M/1 queue

$$(\lambda + \mu_1 + \mu_2)P_{i,11} = \lambda P_{(i-1),11} + (\mu_1 + \mu_2)P_{(i+1),11}$$

- For states $0, 11, 0, 10, 0, 01$ and $0, 00$ we have

$$(\lambda + \mu_1 + \mu_2)P_{0,11} = \lambda P_{0,10} + \lambda P_{0,01} + (\mu_1 + \mu_2)P_{1,11}$$

$$(\lambda + \mu_1)P_{0,10} = \lambda P_{0,00} + \mu_2 P_{0,11}$$

$$(\lambda + \mu_2)P_{0,01} = \mu_1 P_{0,11}$$

$$\lambda P_{0,00} = \mu_1 P_{0,10} + \mu_2 P_{0,01}$$

- System of linear equations \Rightarrow Solve numerically to find probabilities

- ▶ For large i behaves like M/M/1 queue with service rate $(\mu_1 + \mu_2)$
 - ⇒ Still, states with no queued packets are important
- ▶ M/M/c queue ⇒ c servers with rates μ_1, \dots, μ_c
 - ⇒ More cumbersome to analyze but no fundamental differences

Queuing theory

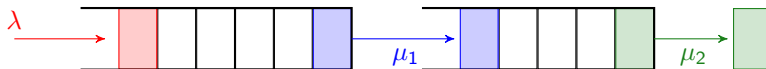
M/M/1 queue

Multiserver queues

Networks of queues

A queue tandem

- ▶ Customers arrive at system to receive two services
- ▶ They arrive at a rate λ and wait in queue 1 for service 1
⇒ Service 1 is performed at a rate μ_1
- ▶ After completions of service 1 customers move to queue 2
⇒ Service 2 is performed at a rate μ_2



- ▶ States (i, j) represent i customers in queue 1 and j in queue 2
- ▶ If both queues are empty ($i = j = 0$), only possible event is an **arrival**

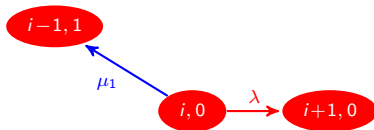
$$q_{00,10} = \lambda$$



- ▶ If queue 2 is empty might have **arrival** or completion of **service 1**

$$q_{i0,(i+1)0} = \lambda$$

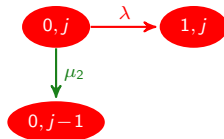
$$q_{i0,(i-1)1} = \mu_1$$



- If queue 1 is empty might have **arrival** or completion of **service 2**

$$q_{0j,1j} = \lambda$$

$$q_{0j,0(j-1)} = \mu_2$$

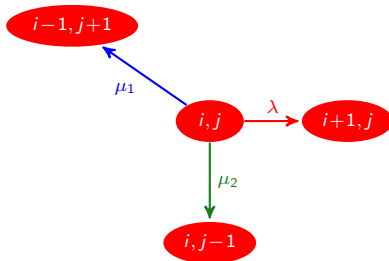


- If no queue is empty **arrival**, **service 1** and **service 2** possible

$$q_{ij,(i+1)j} = \lambda$$

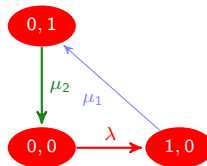
$$q_{ij,(i-1)(j+1)} = \mu_1$$

$$q_{ij,i(j-1)} = \mu_2$$



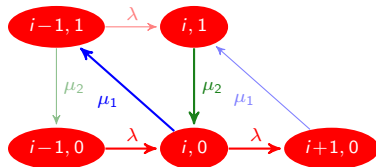
- ▶ Rate at which CTMC enters state $(i, j) =$ rate at which CTMC leaves (i, j)
- ▶ **State $(0, 0)$** - Both queues empty
- ▶ From $(0, 0)$ can go to $(1, 0)$
- ▶ Can enter $(0, 0)$ from $(0, 1)$

$$\lambda P_{00} = \mu_2 P_{01}$$



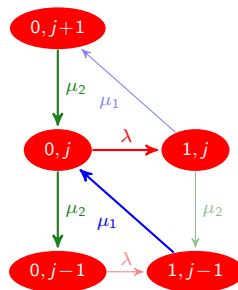
- ▶ **State $(i, 0)$** - Queue 2 empty
- ▶ From $(i, 0)$ go to $(i + 1, 0)$ or $(i - 1, 1)$
- ▶ Into $(i, 0)$ from $(i - 1, 0)$ or $(i, 1)$

$$(\lambda + \mu_1) P_{i0} = \lambda P_{(i-1)0} + \mu_2 P_{i1}$$



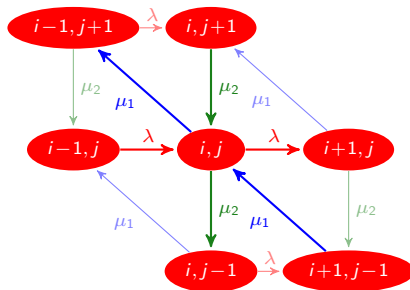
- **State $(0, j)$** - Queue 1 empty
- From $(0, j)$ go to $(1, j)$ or $(0, j - 1)$
- Into $(0, j)$ from $(1, j - 1)$ or $(0, j + 1)$

$$(\lambda + \mu_2)P_{0j} = \mu_1 P_{1(j-1)} + \mu_2 P_{0(j+1)}$$



- **State** (i, j) - Neither queue empty
- From (i, j) can go to $(i + 1, j)$, $(i - 1, j + 1)$ or $(i, j - 1)$
- Can enter (i, j) from $(i - 1, j)$, $(i + 1, j - 1)$ or $(i, j + 1)$

$$(\lambda + \mu_1 + \mu_2)P_{ij} = \lambda P_{(i-1)j} + \mu_1 P_{(i+1)(j-1)} + \mu_2 P_{i(j+1)}$$



- ▶ Direct substitution shows that balance equations are solved by

$$P_{ij} = \left(1 - \frac{\lambda}{\mu_1}\right) \left(\frac{\lambda}{\mu_1}\right)^i \left(1 - \frac{\lambda}{\mu_2}\right) \left(\frac{\lambda}{\mu_2}\right)^j$$

- ▶ Compare with expression for M/M/1 queue
 - ⇒ It behaves as **two independent M/M/1 queues**
 - ⇒ First queue has rates λ and μ_1
 - ⇒ Second queue has rates λ and μ_2
- ▶ Result can be generalized to networks of queues
 - ⇒ Important in transportation networks
 - ⇒ Also useful to analyze Internet traffic

- ▶ Queuing theory
- ▶ Customers and servers
- ▶ Queue length
- ▶ Time spent in queue
- ▶ M/M/1 queue
- ▶ Finite-capacity queue
- ▶ Multi-server queue
- ▶ Network of queues
- ▶ Queue tandem
- ▶ Poisson arrivals
- ▶ Exponential service times
- ▶ Balance equations
- ▶ Stable queue
- ▶ Traffic intensity
- ▶ Expected queue length
- ▶ Expected waiting time
- ▶ Little's law
- ▶ M/M/c queue
- ▶ Aggregate service rate
- ▶ Independent M/M/1 queues