

16.5 Exercises Lecture 5

16.5.1 Exercise 1: A generative model

Consider a generative classification model for K classes defined by prior class probabilities $p(C_k) = \pi_k$ and general class-conditional densities $p(\phi|C_k)$ where ϕ is the input feature vector. Suppose we are given a training data set $\{\phi_n, t_n\}$ where $n = 1, \dots, N$ and t_n is a binary target vector of length K that uses the 1-of- K coding scheme, so that it has components $t_{nj} = I_{jk}$ if pattern n is from class C_k . Assuming that the data points are drawn independently from this model, show that the maximum-likelihood solution for the prior probabilities is given by

$$\pi_k = \frac{N_k}{N} \quad (16.91)$$

where N_k is the number of data points assigned to class C_k .

16.5.2 Solution

We begin by writing down the likelihood function.

$$p(\{\phi_n, t_n\}|\pi_1, \pi_2, \dots, \pi_K) = \prod_{n=1}^N \prod_{k=1}^K [p(\phi_n|C_k)p(C_k)]^{t_{nk}} = \prod_{n=1}^N \prod_{k=1}^K [\pi_k p(\phi_n|C_k)]^{t_{nk}} \quad (16.92)$$

Hence we can obtain the expression for the logarithm likelihood:

$$\ln p = \sum_{n=1}^N \sum_{k=1}^K t_{nk} [\ln \pi_k + \ln p(\phi_n|C_k)] \propto \sum_{n=1}^N \sum_{k=1}^K t_{nk} \ln \pi_k \quad (16.93)$$

Since there is a constraint on π_k , so we need to add a Lagrange Multiplier to the expression, which becomes:

$$L = \sum_{n=1}^N \sum_{k=1}^K t_{nk} \ln \pi_k + \lambda \left(\sum_{k=1}^K \pi_k - 1 \right) \quad (16.94)$$

We calculate the derivative of the expression above with regard to π_k :

$$\frac{\partial L}{\partial \pi_k} = \sum_{n=1}^N \frac{t_{nk}}{\pi_k} + \lambda \quad (16.95)$$

And if we set the derivative equal to 0, we can obtain:

$$\pi_k = - \left(\sum_{n=1}^N t_{nk} / \lambda \right) = - \frac{N_k}{\lambda} \quad (16.96)$$

And if we perform summation on both sides with regard to k , we can see that:

$$1 = - \left(\sum_{k=1}^K N_k \right) / \lambda = - \frac{N}{\lambda} \quad (16.97)$$

Which gives $\lambda = -N$, and substitute it into 16.96, we can obtain the result.

16.5.3 Exercise 2: An example of Naive Bayes model

Consider a classification problem with K classes for which the feature vector ϕ has M components each of which can take L discrete states. Let the values of the components be represented by a 1-of- L binary coding scheme. Further suppose that, conditioned on the class C_k , the M components of ϕ are independent, so that the class-conditional density factorizes with respect to the feature vector components. Show that the quantities a_k given by :

$$a_k = \ln p(\mathbf{x}|C_k)p(C_k) \quad (16.98)$$

which appear in the argument to the softmax function describing the posterior class probabilities, are linear functions of the components of ϕ . Note that this represents an example of the naive Bayes model

16.5.4 Solution

Based on definition, we can write down

$$p(\phi, C_k) = \prod_{m=1}^M \prod_{l=1}^L \mu_{kml}^{\phi_{ml}} \quad (16.99)$$

Note that here only one of the value among $\phi_{m1}, \phi_{m2}, \dots, \phi_{mL}$ is 1, and the others are all 0 because we have used a 1-of- L binary coding scheme, and also we have taken advantage of the assumption that the M components of ϕ are independent conditioned on the class C_k . We substitute the expression above into

$$a_k = \ln p(\mathbf{x}|C_k)p(C_k) \quad (16.100)$$

which gives:

$$a_k = \sum_{m=1}^M \sum_{l=1}^L \phi_{ml} \cdot \ln \mu_{kml} + \ln p(C_k) \quad (16.101)$$

Hence it is obvious that a_k is a linear function of the components of ϕ

16.5.5 Exercise 3: A follow-up of exercise 1

Consider the classification model of exercise 1 and now suppose that the class-conditional densities are given by Gaussian distributions with a shared covariance matrix, so that

$$p(\phi, C_k) = N(\phi|\mu_k, \Sigma) \quad (16.102)$$

Show that the maximum likelihood solution for the mean of the Gaussian distribution for class C_k is given by

$$\mu_k = \frac{1}{N_k} \sum_{n=1}^N t_{nk} \phi_n \quad (16.103)$$

which represents the mean of those feature vectors assigned to class C_k . Similarly, show that the maximum likelihood solution for the shared covariance matrix is given by

$$\Sigma = \sum_{k=1}^K \frac{N_k}{N} S_k \quad (16.104)$$

where

$$S_k = \frac{1}{N_k} \sum_{n=1}^N t_{nk} (\phi_n - \mu_k)(\phi_n - \mu_k)^T \quad (16.105)$$

Thus Σ is given by a weighted average of the covariances of the data associated with each class, in which the weighting coefficients are given by the prior probabilities of the classes.

16.5.6 Exercise 4: Softmax and Sigmoid

Show that the softmax function is equivalent to a sigmoid in the 2-class case

16.5.7 Solution

$$\begin{aligned} \frac{\exp(w_1^T \mathbf{x})}{\exp(w_1^T \mathbf{x}) + \exp(w_0^T \mathbf{x})} &= \frac{1}{1 + \exp(w_0^T \mathbf{x}) / \exp(w_1^T \mathbf{x})} = \frac{1}{1 + \exp(w_0^T \mathbf{x} - w_1^T \mathbf{x})} = \\ &= \frac{1}{1 + \exp(-(w_1^T - w_0^T)^T \mathbf{x})} = \sigma(w^T \mathbf{x}) \end{aligned} \quad (16.106)$$

16.5.8 Exercise 5: A bit of algebra

Show that the derivative of the sigmoid function $\sigma(a) = (1 + e^{-a})^{-1}$ can be written as:

$$\frac{\partial \sigma(a)}{\partial a} = \sigma(a)(1 - \sigma(a)) \quad (16.107)$$

16.5.9 solution

$$\begin{aligned} \frac{\partial \sigma(a)}{\partial a} &= -\frac{1}{(1 + e^{-a})^2} \cdot e^{-a} \cdot (-1) = \frac{1}{1 + e^{-a}} \frac{e^{-a}}{1 + e^{-a}} = \sigma(a) \frac{1 + e^{-a} - 1}{1 + e^{-a}} = \sigma(a)(1 - \sigma(a)) \end{aligned} \quad (16.108)$$