

# Homework 3

蒋翌坤 20307100013

## Problem 1

Sample's joint distribution is

$$f_{\mathbf{X}}(\mathbf{x}|\theta) = \prod_{i=1}^n e^{i\theta - x_i} \mathbb{I}(x_i > i\theta) = e^{-\sum_{i=1}^n x_i} e^{\frac{n(n+1)}{2}\theta} \prod_{i=1}^n \mathbb{I}\left(\frac{x_i}{i} > \theta\right) = \underbrace{e^{-\sum_{i=1}^n x_i}}_{h(\mathbf{x})} \cdot \underbrace{e^{\frac{n(n+1)}{2}\theta} \mathbb{I}(\min \frac{x_i}{i} > \theta)}_{g(T|\theta)}$$

So,  $f_{\mathbf{X}}$  can be factored into  $h(\mathbf{x})$  and  $g(T|\theta)$ . Thus,  $T = \min \frac{x_i}{i}$  is a sufficient statistic for  $\theta$ .

□

## Problem 2

To prove the independence of  $\bar{X}$  and  $X_{(n)} - X_{(1)}$ , we just need to prove the independence of  $\sum_{k=1}^n X_k$  and  $X_i - X_j$  for  $\forall i \neq j$ .

We know that  $X_i$ 's are jointly normal, so  $\sum_{k=1}^n X_k$  and  $X_i - X_j$  are bivariate normal.  $\text{Cov}(\sum_{k=1}^n X_k, X_i - X_j) = \text{Var}(X_i) - \text{Var}(X_j) = 0$ . Therefore,  $\sum_{k=1}^n X_k$  and  $X_i - X_j$  are independent for  $\forall i \neq j$ . Thus,  $\bar{X}$  and  $X_{(n)} - X_{(1)}$  are independent.

□

## Problem 3

Sample's joint distribution is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^n \frac{1}{2} e^{-|x_i - \theta|} = \frac{1}{2^n} e^{-\sum_{i=1}^n |x_i - \theta|}$$

For any samples  $\mathbf{X}$  and  $\mathbf{Y}$ , we have

$$\frac{f(\mathbf{x}|\theta)}{f(\mathbf{y}|\theta)} = \frac{e^{-\sum_{i=1}^n |x_i - \theta|}}{e^{-\sum_{i=1}^n |y_i - \theta|}} = e^{\sum_{i=1}^n |y_i - \theta| - \sum_{i=1}^n |x_i - \theta|}$$

We can see that  $\frac{f(\mathbf{x}|\theta)}{f(\mathbf{y}|\theta)}$  is independent of  $\theta$  iff  $X_{(i)} = Y_{(i)}$  for  $i = 1, \dots, n$ . So,  $T = (X_{(1)}, \dots, X_{(n)})$  are a minimal sufficient statistic for  $\theta$ , i.e., further reduction is not possible.

□

## Problem 4

Let  $Y = \log X$ ,  $f_Y(y|\alpha) = \alpha e^{y(\alpha-1)} e^{-e^{y\alpha}} e^y = \alpha e^{y\alpha} e^{-e^{y\alpha}}$ .

Let  $Z = \alpha Y$ .  $f_Z(z|\alpha) = \alpha e^z e^{-e^z} \frac{1}{\alpha} = e^z e^{-e^z}$ . We can see that  $Z$  is independent of  $\alpha$ .

So,  $\frac{\log X_1}{\log X_2} = \frac{Z_1}{Z_2}$  is independent of  $\alpha$ . Therefore,  $\frac{\log X_1}{\log X_2}$  is an ancillary statistic.

□