

Homework 7

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1. $\hat{\beta}, e$ are both linear functions of y . Since $y \sim N(X\beta, \sigma^2 I)$, so $\hat{\beta}, e$ are jointly normal. $\text{Cov}(\hat{\beta}, e) = (X^T X)^{-1} X^T \text{Cov}(y)(I - H)^T = \sigma^2 (X^T X)^{-1} X^T (I - X(X^T X)^{-1} X^T) = 0$. So $\hat{\beta} \perp e$

□

2. We already know $\hat{\beta}, e$ are normal.

$$\mathbb{E}(\hat{\beta}) = \beta, \text{Cov}(\hat{\beta}) = (X^T X)^{-1} X^T \sigma^2 I (X^T X)^{-1} X^T = \sigma^2 (X^T X)^{-1}, \text{ so } \hat{\beta} \sim N(\beta, \sigma^2 (X^T X)^{-1})$$

$e = (I - H)y = (I - H)(X\beta + \epsilon) = (I - H)\epsilon$, $e^T e = \epsilon^T (I - H)\epsilon$. $\text{rank}(I) = n$ and $\text{rank}(H) = p$, so $\text{rank}(I - H) = n - p$. Here $(I - H)^2 = (I - H)$, $\epsilon \sim N(0, \sigma^2 I)$, then $\frac{\epsilon}{\sigma} \sim N(0, I)$, so $\frac{e^T e}{\sigma^2} = \frac{\epsilon^T (I - H)\epsilon}{\sigma^2} \sim \chi^2(n - p)$

□

3. $D = A(X^T X)^{-1} A^T = AX^{-1}(AX^{-1})^T$. For any nonzero vector $x \in \mathbb{R}^m$,

$$x^T D x = x^T A X^{-1} (A X^{-1})^T x = ((A X^{-1})^T x)^T (A X^{-1})^T x \geq 0.$$

For $x^T D x$ to be zero, $(A X^{-1})^T x = 0 \Rightarrow A = A X^{-1} X = 0$. However, $\text{rank}(A) = m \neq 0$. Therefore, $x^T D x > 0$, and D is positive definite.

□

4. We have $\sum_{i=1}^n \hat{y}_i = \sum_{i=1}^n \sum_{j=1}^p x_{ij} \hat{\beta}_j = \sum_{j=1}^p \hat{\beta}_j \sum_{i=1}^n x_{ij} = \sum_{j=1}^p \hat{\beta}_j n \bar{x}_j$. Note that we get $\hat{\beta}_j$ under the constraint $0 = \sum_{i=1}^n x_{ic}(y_i - \sum_{j=1}^p \hat{\beta}_j x_{ij})$ for $c = 1, 2, \dots, p$, and $x_{i1} = 1$. We have $n\bar{y} = n \sum_{j=1}^p \hat{\beta}_j \bar{x}_j$. So, $\sum_{i=1}^n \hat{y}_i = n\bar{y}$, $\sum_{i=1}^n (y_i - \hat{y}_i) = 0$ and $\sum_{j=1}^p \hat{\beta}_j \sum_{i=1}^n x_{ij}(y_i - \hat{y}_i) = \sum_{i=1}^n \hat{y}_i (y_i - \hat{y}_i) = 0$. Therefore,

$$\begin{aligned} \sum_{i=1}^n (y_i - \bar{y})^2 &= \sum_{i=1}^n (y_i - \hat{y}_i + \hat{y}_i - \bar{y})^2 \\ &= \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2 + 2 \sum_{i=1}^n (y_i - \hat{y}_i)(\hat{y}_i - \bar{y}) \\ &= \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2 + 2 \left(\sum_{i=1}^n \hat{y}_i (y_i - \hat{y}_i) - \bar{y} \left(\sum_{i=1}^n (y_i - \hat{y}_i) \right) \right) \\ &= \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2 \end{aligned}$$

□

5. $\mathbb{E}(A\hat{\beta}) = u$ and $\text{Cov}(A\hat{\beta}) = A\sigma^2(X^T X)^{-1} A^T = \sigma^2 D$, so $A\hat{\beta} \sim N(u, \sigma^2 D)$. So, $(A\hat{\beta} - u)^T \frac{D^{-1}}{\sigma^2} (A\hat{\beta} - u) \sim \chi^2(m)$. $\frac{\text{SSE}}{\sigma^2} = \frac{e^T e}{\sigma^2} \sim \chi^2(n - p)$. Also, SSE only depends on e and $(A\hat{\beta} - u)^T D^{-1} (A\hat{\beta} - u)$ only depends on $\hat{\beta}$. They are independent. Therefore,

$$\frac{(A\hat{\beta} - u)^T D^{-1} (A\hat{\beta} - u)/m}{\text{SSE}/(n - p)} = \frac{(A\hat{\beta} - u)^T \frac{D^{-1}}{\sigma^2} (A\hat{\beta} - u)/m}{\frac{\text{SSE}}{\sigma^2}/(n - p)} \sim \frac{\chi^2(m)/m}{\chi^2(n - p)/(n - p)} \sim F(m, n - p)$$

□