

Homework 1

蒋翌坤 20307100013

Problem 1

We have $f_{X_{(1)}, X_{(n)}}(x, y) = n(n-1)[F(y) - F(x)]^{n-2}f(x)f(y)$

Let $u = y - x, v = \frac{1}{2}(x + y)$, then $|J| = 1$

$$f_{R,V}(u, v) = n(n-1)[F(v + \frac{u}{2}) - F(v - \frac{u}{2})]^{n-2}f(v - \frac{u}{2})f(v + \frac{u}{2})$$

Here, $F(x) = \frac{x}{a}, f(x) = \frac{1}{a}$ for $0 < x < a$. Therefore, the joint pdf of R and V are

$$f_{R,V}(u, v) = \begin{cases} n(n-1)[\frac{v+\frac{u}{2}}{a} - \frac{v-\frac{u}{2}}{a}]^{n-2} \frac{1}{a} \frac{1}{a} = \frac{n(n-1)}{a^n} u^{n-2}, & 0 < u < a, 0 < v < a \\ 0, & \text{otherwise} \end{cases}$$

Problem 2

Draw two samples from $N(0, a)$ as X, Y , order them, we get $X_{(1)}, X_{(2)}$. Therefore, $Z = X_{(1)}$. So, the distribution of Z is the same as $X_{(1)}$:

$$f_Z(z) = 2[1 - F(z)]^{2-1}f(z) = 2f(z)[1 - F(z)]$$

where $F, f \sim N(0, a)$

Problem 3

$Y = d \tan X$, where $X \sim U(0, \pi/2)$, d is a constant. Therefore, the pdf of Y is

$$f_Y(y) = \frac{2}{\pi} \cdot \frac{1}{1 + \frac{y^2}{d^2}} \cdot \frac{1}{d} = \frac{2}{\pi d(1 + \frac{y^2}{d^2})} = \frac{2}{\pi d(1 + \frac{y^2}{d^2})}, 0 < y < \infty$$

$$\begin{aligned} E(Y) &= \int_0^\infty y \cdot \frac{2}{\pi d(1 + \frac{y^2}{d^2})} dy = \frac{2}{\pi} \int_0^\infty \frac{\frac{y}{d}}{1 + \frac{y^2}{d^2}} dy \\ &\stackrel{u=y/d}{=} \frac{2d}{\pi} \int_0^\infty \frac{u}{1 + u^2} du = \frac{2d}{\pi} \cdot \frac{1}{2} \log(1 + u^2) \Big|_0^\infty = \infty \end{aligned}$$

Problem 4

a

Let $W_1 = X_1^2/\sigma^2 \sim \chi_1^2, W_2 = X_2^2/\sigma^2 \sim \chi_1^2$, we have $Y_1 = \sigma^2(W_1 + W_2), Y_2^2 = \frac{W_1}{W_1 + W_2}$
 $W_1 = Y_1 Y_2^2/\sigma^2, W_2 = Y_1(1 - Y_2^2)/\sigma^2$,

$$|J| = \left| \begin{array}{cc} \frac{Y_2^2}{\sigma^2} & \frac{2Y_1 Y_2}{\sigma^2} \\ \frac{1 - Y_2^2}{\sigma^2} & -\frac{2Y_1 Y_2}{\sigma^2} \end{array} \right| = \frac{2Y_1 |Y_2|}{\sigma^4}$$

The joint distribution of W_1, W_2 is

$$f_{W_1, W_2}(w_1, w_2) = \frac{1}{2} \cdot \frac{1}{\sqrt{w_1 w_2}} \cdot \frac{1}{\Gamma(1/2)^2} \exp\left(-\frac{w_1 + w_2}{2}\right)$$

Then, the joint distribution of Y_1, Y_2 is

$$\begin{aligned} f_{Y_1, Y_2}(y_1, y_2) &= \frac{1}{2} \cdot \frac{\sigma^2}{|y_1 y_2| \sqrt{1 - y_2^2}} \cdot \frac{1}{\Gamma(1/2)^2} \exp\left(-\frac{y_1}{2\sigma^2}\right) \cdot \frac{2y_1 |y_2|}{\sigma^4} \\ &= \frac{1}{\sigma^2} \exp\left(-\frac{y_1}{2\sigma^2}\right) \cdot \frac{1}{\Gamma(1/2)^2} \cdot \frac{1}{\sqrt{1 - y_2^2}} \end{aligned}$$

b

$$\frac{Y_1}{\sigma^2} = X_1^2/\sigma^2 + X_2^2/\sigma^2 \sim \chi_2^2, \text{ so}$$

$$f_{Y_1}(y_1) = \frac{1}{\sigma^2} \exp\left(-\frac{y_1}{2\sigma^2}\right), y_1 > 0$$

$$\begin{aligned} f_{Y_2}(y_2) &= \int_0^\infty f_{Y_1, Y_2}(y_1, y_2) dy_1 = \frac{1}{\Gamma(1/2)^2} \cdot \frac{1}{\sqrt{1 - y_2^2}} \int_0^\infty \frac{1}{\sigma^2} \exp\left(-\frac{y_1}{2\sigma^2}\right) dy_1 \\ &= \frac{1}{\Gamma(1/2)^2} \cdot \frac{1}{\sqrt{1 - y_2^2}} \end{aligned}$$

$$f_{Y_1, Y_2}(y_1, y_2) = f_{Y_1}(y_1) f_{Y_2}(y_2)$$

Therefore, Y_1, Y_2 are independent.

A geometric interpretation is that consider a right triangle whose two sides are X_1 and X_2 ($X_1, X_2 > 0$), then Y_1 is the square of the hypotenuse, and Y_2 is the sine of the angle opposite to X_1 . Since Y_1 and Y_2 are independent, we know that the length of the hypotenuse and the angle are independent.

Problem 5

Let $\delta = \sqrt{\sum_{i=1}^n a_i^2}$, $\mathbf{C} = \begin{pmatrix} \frac{a_1}{\delta} & \cdots & \frac{a_n}{\delta} \\ c_{21} & \cdots & c_{2n} \\ \vdots & \ddots & \vdots \\ c_{n1} & \cdots & c_{nn} \end{pmatrix}$, and \mathbf{C} orthogonal. Let $\mathbf{W} = \mathbf{C}\mathbf{X}$. We have

$$\mathbf{W}^T \mathbf{W} = \mathbf{X}^T \mathbf{C}^T \mathbf{C} \mathbf{X} = \mathbf{X}^T \mathbf{X}$$

$$W_1 = \frac{a_1}{\delta} X_1 + \cdots + \frac{a_n}{\delta} X_n = \frac{1}{\delta} \sum_{i=1}^n a_i X_i$$

$$E(W_1) = \frac{1}{\delta} \sum_{i=1}^n a_i E(X_i) = \frac{1}{\sqrt{\sum_{i=1}^n a_i^2}} \sum_{i=1}^n a_i^2 = \sqrt{\sum_{i=1}^n a_i^2}$$

$$\text{Var}(W_1) = \frac{1}{\delta^2} \sum_{i=1}^n a_i^2 \text{Var}(X_i) = \frac{1}{\sum_{i=1}^n a_i^2} \sum_{i=1}^n a_i^2 = 1$$

Therefore, $W_1 \sim N(\sqrt{\sum_{i=1}^n a_i^2}, 1)$

When $i \geq 2$,

$$E(W_i) = E\left(\sum_{k=1}^n c_{ik} X_k\right) = \sum_{k=1}^n c_{ik} a_k = \delta \sum_{k=1}^n c_{ik} \frac{a_k}{\delta} = 0$$

$$\text{Var}(W_i) = \sum_{k=1}^n c_{ik}^2 = 1$$

Therefore, $W_i \sim N(0, 1)$, $W_2^2 + \dots + W_n^2 \sim \chi_{n-1}^2$.

Let $Y = W_1 \sim N(\sqrt{\sum_{i=1}^n a_i^2}, 1)$, $Z = W_2^2 + \dots + W_n^2 \sim \chi_{n-1}^2$, we have

$$X_1^2 + \dots + X_n^2 = W_1^2 + W_2^2 + \dots + W_n^2 = Y^2 + Z$$

Y depends on W_1 and Z depends on W_2, \dots, W_n , and W_1, W_2, \dots, W_n are independent. Therefore, Y and Z are independent.