

STAT 345/445 Lecture 7

Section 2.1: Distributions of Functions of a Random Variable

Functions of random variables

Sometimes we want to transform a random variable.

For example:

- If X is the temperature in Fahrenheit, what is the distribution of the temperature in Celsius?

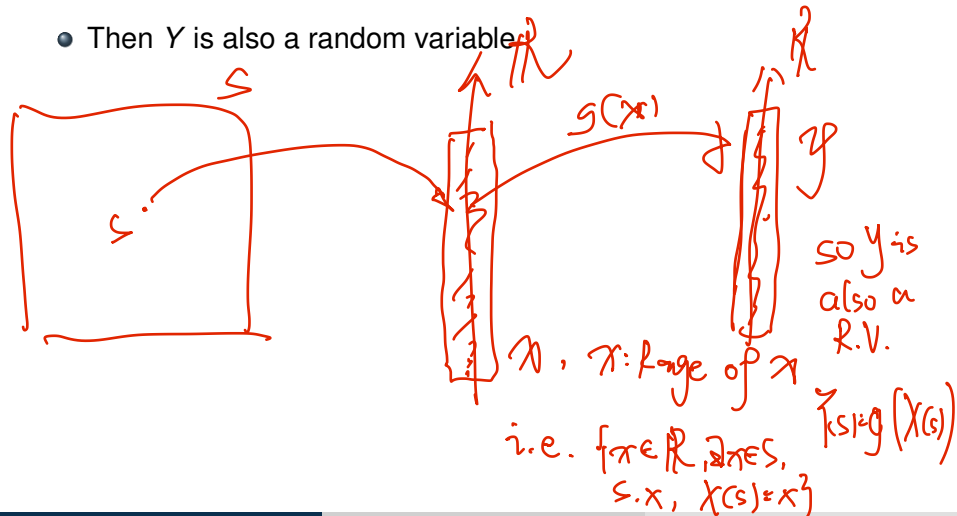
$$Y = (X - 32)\frac{5}{9}$$

- If X and Y denote height and weight, what is the distribution of the BMI?

$$B = \frac{X}{Y^2}$$

Functions of random variables

- Let X be a random variable and let $g(\cdot)$ be a function.
- Then Y is also a random variable



Functions of random variables

pdf or pmf/pdf of y

- What is the distribution of the random variable $Y = g(X)$?
- Have the cdf $F_X(x)$ or pmf/pdf $f_X(x)$ of X
 - want to find the cdf $F_Y(y)$ or pmf/pdf $f_Y(y)$ of Y .

notice that $g^{-1}(y)$ may not be just one value

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y))$$

Let's review *inverse mappings*...

Inverse mapping

- For a function $g(x) : \mathcal{X} \rightarrow \mathcal{Y}$ we define an **inverse mapping** as

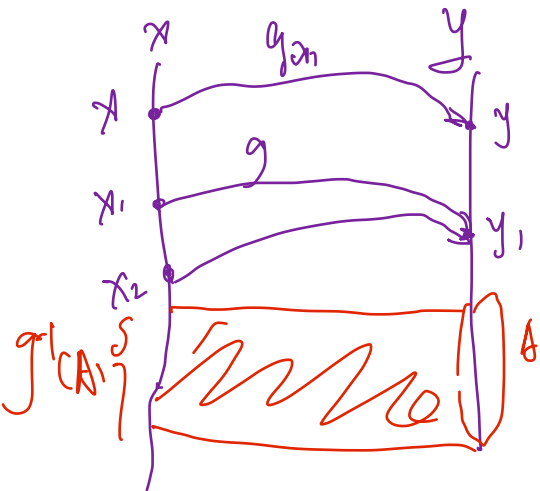
$$g^{-1}(A) = \{x \in \mathcal{X} : g(x) \in A\}$$

for any set $A \subset \mathcal{Y}$

- Note that $g^{-1}(A) \subset \mathcal{X}$
- In particular:

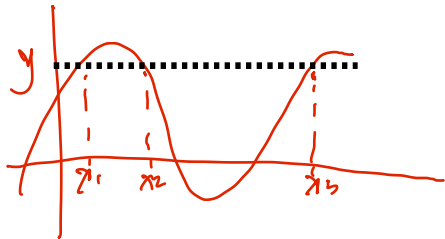
$$g^{-1}(\{y\}) = \{x \in \mathcal{X} : g(x) = y\}$$

- Can still be a set in \mathcal{X} rather than just one number
- Usually just write $g^{-1}(y)$

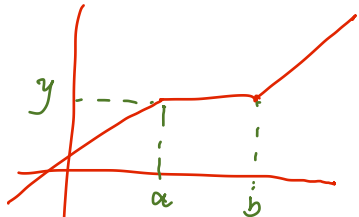


$$g^{-1}(\{y_1\}) = g^{-1}(y_1)$$

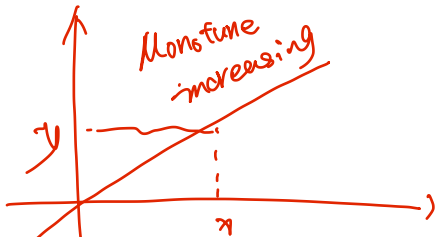
$$= \{x_1, x_2\}$$



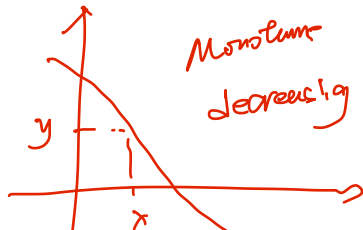
$$g^{-1}(y) = \{x_1, x_2, x_3\}$$



$$g^{-1}(y) = [a, b]$$



$$g^{-1}(y) = x$$



$$g^{-1}(y) = x$$

More on inverse mapping

- A function $g(x) : \mathcal{X} \rightarrow \mathcal{Y}$ is a **one-to-one** function if and only if $\forall y \in \mathcal{Y}$ we have

$$g^{-1}(\{y\}) = \{x\}$$

- Can write $g^{-1}(y) = x$
- ***Strictly monotone*** functions are one-to-one

Probability of a transformation

- Let X be a random variable in (S, \mathcal{B}, P) and let $Y = g(X)$.
- Probabilities for Y can be obtained from probabilities of X and the inverse mapping $g^{-1}(\cdot)$

- In general

$A \subset Y$

$$P(Y \in A) = P(X \in g^{-1}(A))$$

In particular:

$$F_Y(y) = P(Y \leq y) \\ = P(X \in g^{-1}(-\infty, y])$$

- Will look at discrete and continuous variables separately

Discrete random variables

- Let X be a discrete random variable and let $Y = g(X)$ for some function $g(\cdot)$.
- Then Y is a discrete random variable
- Then

$$\begin{aligned}
 f_Y(y) &= P(Y=y) = P(g(X)=y) = P(X \in g^{-1}(y)) \\
 &= P(X \in \{x: g(x)=y\}) \\
 &= \sum_{x \in g^{-1}(y)} P(X=x) = \sum_{x \in g^{-1}(y)} f_X(x)
 \end{aligned}$$

and

$$F_Y(y) = \sum_{v \leq y} f_Y(v)$$

Discrete example 1

- Let $X \sim \text{Binom}(n, p)$, i.e. X has pmf

$$f_X(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & \text{for } x \in \{0, 1, 2, \dots, n\} \\ 0 & \text{otherwise} \end{cases}$$

- X can be thought of as the number of successes in n independent Bernoulli trials
- What is the distribution of $Y = n - X$?

(number of failures instead number of success)

Example 1:

$$g(x) = n - x$$

$$x = \{0, 1, 2, \dots, n\}$$

$$y = \{0, 1, \dots, n\}$$

$$g^{-1}(y) = \{x: \overset{g(x)}{n-x=y}\} = \{n-y\}$$

i.e. a set of one value.

$$P_Y(y) = \sum_{x \in \{y\}} P_X(x) = P_X(n-y) = \binom{n}{n-y} p^{ny} (1-p)^{n-(n-y)}$$

could stop here

$$\binom{n}{n-y} = \frac{n!}{(n-y)! y!}$$

$$= \binom{n}{y} p^{n-y} (1-p)^y$$

$$\frac{n!}{y! (n-y)!}$$

put for Binomial
(n, 1-p)

Discrete example 2

- Let $X \sim \text{Binom}(10, p)$. What is the distribution of $Y = |X - 5|$?

$$X = \{0, 1, 2, \dots, 10\}$$

$$g(x) = |x - 5|$$

$$Y = \{0, 1, 2, 3, 4, 5\}$$

$$y=0, \text{ only one } x: x=5$$

$$\text{If } y = |x - 5| \text{ then}$$

$$\text{either } y = x - 5 \text{ or } y = -x + 5 \Rightarrow x = y + 5 \text{ or } x = 5 - y$$

$$\text{For } y=0: g^{-1}(0) = \{5\}, \text{ for } y=1, 2, 3, 4, 5; g^{-1}(y) = \{y+5, 5-y\}$$

↓

$$f_y(0) = f_x(5) = \binom{10}{5} p^5 (1-p)^5$$

for $y = 1, \dots, 5$,

$$\begin{aligned} f_x(y) = f_x(y+5) + f_x(5-y) &= \binom{10}{y+5} p^{y+5} (1-p)^{10-y-5} \\ &\quad + \binom{10}{5-y} p^{5-y} (1-p)^{10-5+y} \end{aligned}$$

~~Continuous random variables~~

Monotone transf

- It's easiest to deal with *monotone* functions g :

Increasing: $u > v \Rightarrow g(u) > g(v)$

Decreasing: $u > v \Rightarrow g(u) < g(v)$

- The **support** of a distribution (or random variable) is defined as

$$\mathcal{X} = \{x : f_X(x) > 0\} \quad (1)$$

$$\text{and let } \mathcal{Y} = \{y : \exists x \in \mathcal{X} \text{ such that } g(x) = y\} \quad (2)$$

Support of y

Support of x

- If g is monotone on \mathcal{X} then it is *one-to-one* and *onto* from \mathcal{X} to \mathcal{Y} .
 - Uniquely pairs an x to one y
 - Get an inverse function: $g^{-1}(y) = x$

cdf – method

Theorem ("cdf-method")

Let X be a random variable with cdf $F_X(x)$ and let $Y = g(X)$. Then

(a) If g is an increasing function on \mathcal{X} then

$$F_Y(y) = F_X(g^{-1}(y))$$

(b) If g is a decreasing function on \mathcal{X} *and* X is continuous, then

$$F_Y(y) = 1 - F_X(g^{-1}(y))$$

In general: *i.e. for both continuous and discrete*

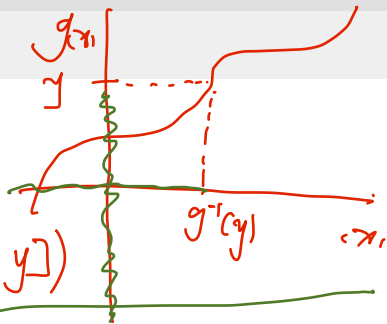
$$F_Y(y) = 1 - F_X(g^{-1}(y)) + P(X = g^{-1}(y))$$

cdf – method, proof

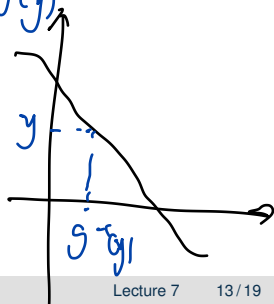
$$(a) F_Y(y) = F_X(g^{-1}(y))$$

$$F_Y(y) = P(Y \leq y) = P(Y \in (-\infty, y])$$

$$\begin{aligned} Y \leq y &\Leftrightarrow X \leq g^{-1}(y) \\ &= P(X \in (-\infty, g^{-1}(y)]) \\ &= P(X \leq g^{-1}(y)) \\ &= F_X(g^{-1}(y)) \end{aligned}$$



$$(b) Y \leq y \Rightarrow X \geq g^{-1}(y)$$



$$F_Y(y) = P(Y \leq y) = P(X \geq g^{-1}(y))$$

$$= 1 - P(X < g^{-1}(y)) = 1 - P(X \leq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$$

if X is continuous

$$= 1 - P(X < g^{-1}(y)) - P(X = g^{-1}(y)) + P(X = g^{-1}(y))$$

$$= 1 - P(X \leq g^{-1}(y)) + P(X = g^{-1}(y)) = 1 - F_X(g^{-1}(y)) + P(X = g^{-1}(y))$$

Example: Exponential and Weibull

Let $X \sim \text{Expo}(1)$, i.e.

$$F_X(x) = \begin{cases} 0 & , x < 0 \\ 1 - e^{-x} & , x \geq 0 \end{cases}$$

Let $Y = X^\alpha$ for $\alpha > 0$. What is the distribution of Y ?

$$g(x) = x^\alpha, \alpha > 0 \text{ (constant)}$$

$$\text{if } x_1 < x_2, \text{ then } x_1^\alpha < x_2^\alpha$$

so g is monotone increasing. (only need increasing on \mathcal{X})

$$\text{or if } \frac{d}{dx} g(x) = \frac{d}{dx} x^\alpha > 0 \text{ for all } x > 0 \quad x \in [0, \infty)$$

Suppose of \tilde{F} ? $F_X(0) = 1 - e^0 = 0$, $\lim_{x \rightarrow \infty} F_X(x) = 1 - 0 = 1$

$$F_Y(y) = F_X(g^{-1}(y))$$

$$y = g(x) = x^2 \quad x \in [0, \infty) \Rightarrow y \in [0, \infty) = Y$$

$$y^{1/2} = x$$

$$\Rightarrow F_Y(y) = F_X(y^{1/2}) = 1 - e^{-y^{1/2}} \text{ for } y \in [0, \infty).$$

$$F_Y(y) = 0 \text{ for } y < 0$$

pdf – method

Theorem ("pdf method")

Let X be a continuous random variable with pdf $f_X(s)$ and let $Y = g(X)$ where g is a *monotone* function.

Suppose that $f_X(x)$ is continuous on \mathcal{X} and that $g^{-1}(y)$ has a continuous derivative on \mathcal{Y} .

Then the pdf of Y is given by

$$f_Y(y) = \frac{d}{dy} \int_{\mathcal{X}} \mathbb{I}_{\mathcal{X}}(g^{-1}(y))$$

$$= f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

$$= f_X(g^{-1}(y)) \frac{d}{dy} g^{-1}(y)$$

always positive, cuz increasing

$$\frac{d}{dx} h(g(x)) = h'(g(x)) g'(x)$$

Pdf method:

if $g(x)$ is increasing,

$$f_Y(y) = \frac{d}{dy} F_X(y) = \frac{d}{dy} F_X(g^{-1}(y))$$

Chain Rule.

If g is decreasing, $f_y(y) = \frac{d}{dy}(1 - F_x(g(y))) = -f_x(g(y)) \frac{d}{dy}g^{-1}(y)$

Note: $\frac{d}{dy}g^{-1}(y)$ is negative.

Therefore :

$$f_y(y) = f_x(g^{-1}(y)) \left| \frac{d}{dy}g^{-1}(y) \right| \text{ for } y \in Y$$

0 otherwise.

Example: Exponential and Uniform

Let $X \sim \text{Expo}(1)$, i.e.

$$f_X(x) = \begin{cases} e^{-x} & , x \geq 0 \\ 0 & , x < 0 \end{cases}$$

Let $Y = 1 - e^{-X}$. What is the distribution of Y ?

$$\text{So } g^{-1}(y) = -\log(1-y)$$

$$\frac{d}{dy} g^{-1}(y) = \frac{-1}{1-y} (-1)$$

$$= \frac{1}{1-y}$$

$$y = g(x) = 1 - e^{-x} \Rightarrow y \in [0, 1)$$

$$x \in [0, \infty)$$

$$g(0) = 1 - 1 = 0$$

$$\lim_{x \rightarrow \infty} g(x) = 1 - 0 = 1$$

$$e^{-x} = 1 - y \Rightarrow -x = \log(1-y)$$

$$\Rightarrow x = -\log(1-y)$$

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$$

$$= e^{-(-\log(1-y))} \frac{1}{1-y}$$

$$\Rightarrow f_Y(y) = \begin{cases} 1 & \text{for } y \in [0, 1) \\ 0 & \text{otherwise.} \end{cases}$$

Note: $F_X(x) = \begin{cases} 1 - e^{-x} & \text{for } x \geq 0 \\ 0 & \text{otherwise.} \end{cases}$

i.e. $g(x) = f_x(x)$

Probability integral transformation

Theorem

Let X have a continuous cdf $F_X(x)$ and let $Y = F_X(X)$. Then Y is uniformly distributed on $(0, 1)$, i.e.

$$F_Y(y) = \begin{cases} 0 & , y \leq 0 \\ y & , 0 < y < 1 \\ 1 & , y \geq 1 \end{cases}$$

$$X \sim F_X(x) \Rightarrow Y = F_X(X) \sim \text{Uniform}(0, 1)$$

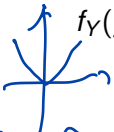
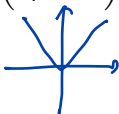
$$\text{and } Y \sim \text{Uniform} \Rightarrow X = F_X^{-1}(Y) \sim F_X$$

When g is monotone only on certain intervals

- See Theorem 2.1.8 for more detail
- If \mathcal{X} can be split into sets A_1, \dots, A_k and g can be split into $g_1(x), \dots, g_k(x)$ such that
 - $g(x) = g_i(x)$ for $x \in A_i$
 - g_i is a monotone function from A_i *onto* \mathcal{Y}

then

$$f_Y(y) = \begin{cases} \sum_{i=1}^k f_X(g_i^{-1}(y)) \left| \frac{d}{dy} g_i^{-1}(y) \right| & , y \in \mathcal{Y} \\ 0 & , \text{otherwise} \end{cases}$$

e.g. $g(x) = x^2$ or $g(x) = |x|$

$$g(x) = x^2, \quad x \in N(0,1)$$

$$g_1(x) = x^2, \quad \text{for } x \in [0, \infty) = A_1$$

$$g_2(x) = x^2, \quad \text{for } x \in (-\infty, 0) = A_2$$

$$y = x^2 \Rightarrow x = \sqrt{y} \text{ if } x \in A_1, \quad x = -\sqrt{y} \text{ if } x \in A_2$$

$$g_1(y) = \sqrt{y}, \quad \frac{d}{dy} = \frac{1}{2\sqrt{y}}$$

$$g_2(y) = -\sqrt{y}, \quad \frac{d}{dy} = -\frac{1}{2\sqrt{y}}$$

Example: Standard normal and χ^2 distribution

Let $X \sim N(0, 1)$, i.e.

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \quad \text{for } x \in \mathbb{R}$$

Let $Y = X^2$. What is the distribution of Y ?

$$\begin{aligned}
 f_Y(y) &= \frac{1}{\sqrt{2\pi}} e^{-\frac{(\sqrt{y})^2}{2}} \frac{1}{2\sqrt{y}} + \frac{1}{\sqrt{2\pi}} e^{-\frac{(-\sqrt{y})^2}{2}} \frac{1}{2\sqrt{y}} \quad \text{for } y \geq 0 \\
 &= \frac{1}{\sqrt{2\pi}} e^{-y/2} \frac{1}{\sqrt{y}}, \quad \text{for } y \geq 0 \\
 &= \text{pdf for chi-square distribution with 1 d.f.}
 \end{aligned}$$

Chi-Square
↓
 χ^2