

# STAT 346/446 Lecture 10

## Methods of evaluating tests

CB Section 8.3.1, DS Section 9.1

- 1 Power Function
- 2 Type I and Type II errors
- 3 Examples
- 4 Level and size of tests

# Evaluating Hypothesis tests

Hypotheses:

$$H_0 : \theta \in \Theta_0 \text{ vs. } H_1 : \theta \in \Theta_0^c$$

Test procedure:

Reject  $H_0$  if  $\mathbf{X} \in R$

- Can make two types of mistakes

- Reject  $H_0 : \theta \in \Theta_0$  when in fact  $\theta \in \Theta_0$  (Type I)
- Don't reject  $H_0 : \theta \in \Theta_0$  when in fact  $\theta \notin \Theta_0$  (Type II)

- Want the *probability* of mistakes to be small

- Recall LRT: Reject  $H_0$  if  $\lambda(\mathbf{X}) \leq c$

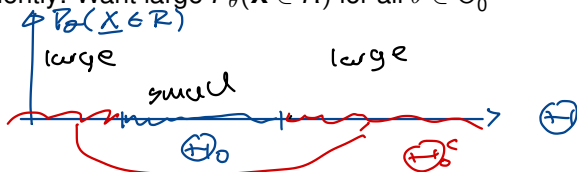
- $c$  determines the probability of these mistakes

$P_\theta(\mathbf{X} \in R) = \text{prob of rejecting } H_0$   
depends on the unknown  $\theta$

# The challenge of finding a good test procedure

Choose  $R$  (i.e.  $c$  in the LRT procedure) such that :

- $P_{\theta}(\mathbf{X} \in R)$  is small for all  $\theta \in \Theta_0$ 
  - Low probability of Type I error
  - Don't want to reject  $H_0$  when it is true
- and  $P_{\theta}(\mathbf{X} \in R^c)$  is small for all  $\theta \in \Theta_0^c$ 
  - Low probability of Type II error
  - Don't want to accept  $H_0$  when it is false
  - Equivalently: Want large  $P_{\theta}(\mathbf{X} \in R)$  for all  $\theta \in \Theta_0^c$



# Power function

Describes the properties of a test procedure

Hypotheses:

$$H_0 : \theta \in \Theta_0 \text{ vs. } H_1 : \theta \in \Theta_0^c$$

Test procedure:

Reject  $H_0$  if  $\mathbf{X} \in R$

## Def: Power function

The **power function** of a test procedure is the probability of rejecting  $H_0$  (function of  $\theta$ )

$$\beta(\theta) = P_{\theta}(\mathbf{X} \in R) \quad \text{for all } \theta \in \Theta$$

*depends on the unknown parameter  $\theta$*

where  $R$  is the *rejection region* of the test

*$\Rightarrow$  Is a function of  $\theta$ .*

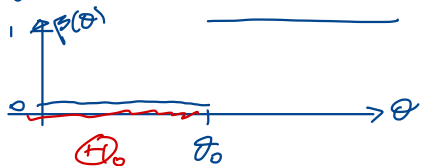
- Ideal power function:

$$\beta(\theta) = 1 \quad \text{for } \theta \in \Theta_0^c \quad \text{want to reject}$$

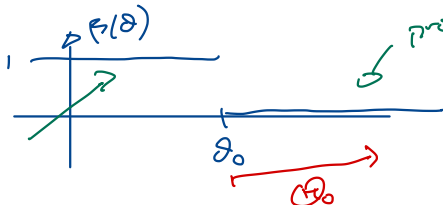
$$\beta(\theta) = 0 \quad \text{for } \theta \in \Theta_0 \quad \text{don't want to reject.}$$

Ideal power function

$$H_0: \theta \leq \theta_0 \text{ vs } H_1: \theta > \theta_0$$

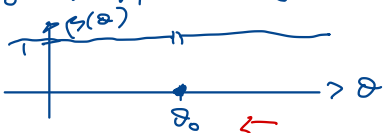


$$H_0: \theta \geq \theta_0 \text{ vs } H_1: \theta < \theta_0$$



1 - prob of  
type II  
error.

$$H_0: \theta = \theta_0 \text{ vs } H_1: \theta \neq \theta_0$$



# Type I and Type II errors

- **Type I error:** Wrongly deciding to reject  $H_0$ 
  - Rejecting  $H_0 : \theta \in \Theta_0$  when in fact  $\theta \in \Theta_0$
- **Type II error:** Wrongly deciding not to reject  $H_0$ 
  - Don't reject  $H_0 : \theta \in \Theta_0$  when in fact  $\theta \notin \Theta_0$

"confusion matrix"  
↙

Truth (true state)

		Decision	
		Choose $H_0$	Choose $H_1$
$H_0$ is true ( $\theta \in \Theta_0$ )		no error	Type I error
$H_1$ is true ( $\theta \in \Theta_0^c$ )		Type II error	no error

Relation to power function:

- If  $\theta \in \Theta_0$ :  $\beta(\theta)$  = probability of type I error
- If  $\theta \in \Theta_0^c$ :  $1 - \beta(\theta)$  = probability of type II error  
 $\beta(\theta)$  = power of test when  $\theta \in \Theta_0^c$

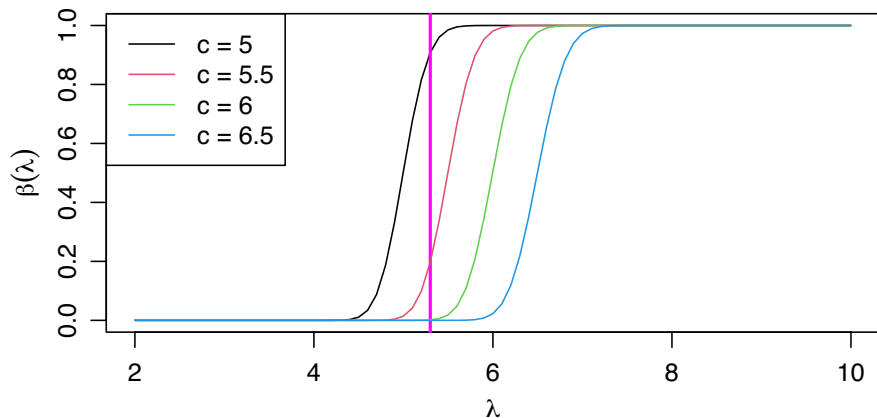
## Example: Power function for a Poisson likelihood

- Let  $X_1, \dots, X_n$  be a random sample from  $\text{Poisson}(\lambda)$
- We want to test the hypotheses

$$H_0 : \lambda \leq 5.3 \quad \text{versus} \quad H_1 : \lambda > 5.3$$

- We will reject  $H_0$  if  $\bar{X} \geq 6$
- What is the power function for this test?
- How does the power function change if instead we reject if  $\bar{X} \geq 6.5$ ? *The function shifts to the right but keeps the same shape.*
- How does the power function change if sample size is increased? *The function becomes steeper and closer to the ideal form of 0 in  $\mathbb{R}_0$  and 1 in  $\mathbb{R}_0^c$*

# Power function for Poisson Example





## Example: After-School

- The Chapel Hill & Carrboro City Schools (CHCCS) operate After-School programs in all the Elementary schools in the district.
- Each month they send invoices to families that have children enrolled in these programs for both the monthly fee and any extra charges, such as care during teacher workdays (eg. election day)
- Customers are expected to use their own envelopes and stamps to return their payments
- Currently, the time it takes to pay bills has a mean of 24 days and a standard deviation of 6 days.

## Example: After-School, continued

- Suppose the chief financial officer (CFO) believes that including a stamped self-addressed envelope would decrease the amount of time it takes to pay the bills.
- She calculates that the improved cash flow from a 2-day decrease in the payment period would pay for the costs for envelopes and stamps. Further decrease would generate profit.
- To test this she randomly selects 220 customers and includes a stamped self-addressed envelope with their invoice
- She assumes that the time to pay a bill follows the normal distribution  $N(\mu, \sigma^2)$ , where  $\sigma^2$  is known.

Using the data from her experiment, how can she conclude whether this plan will be profitable?

# Power function for after-school

- We have  $X_1, \dots, X_n$  i.i.d.  $N(\mu, \sigma^2)$  and we are interested in testing   
 ~~24-2~~

$$H_0 : \mu \geq 22 \quad \text{and} \quad H_1 : \mu < 22$$

We assume that  $\sigma^2$  is known,  $\sigma^2 = 6^2$

- Consider the LRT procedure:
  - Reject  $H_0$  if

$$\frac{\bar{X}_n - \mu_0}{\sigma/\sqrt{n}} \leq -\sqrt{-2 \log(c)} \equiv c^*$$

that is if

$$\frac{\bar{X}_n - 22}{6/\sqrt{220}} \leq c^*$$

- What is the power function for this test?

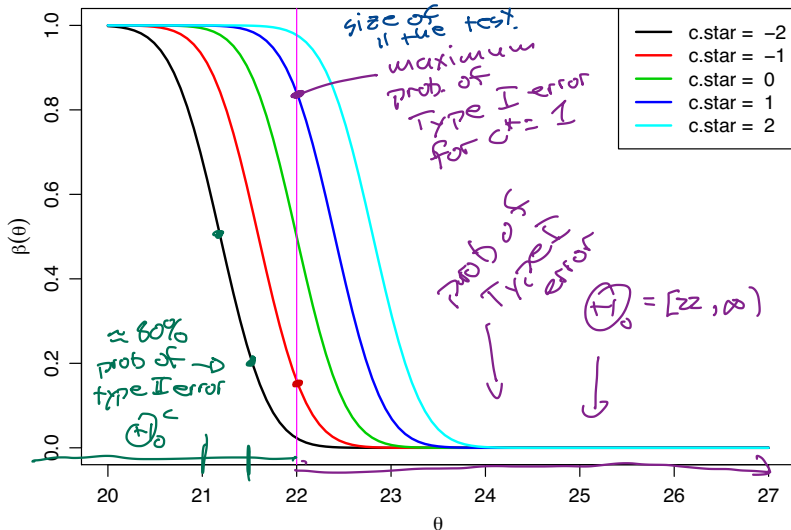
## Power function for after-school

$$n = 220$$

$$\sigma = \sqrt{6}$$

$$\mu_0 = 22$$

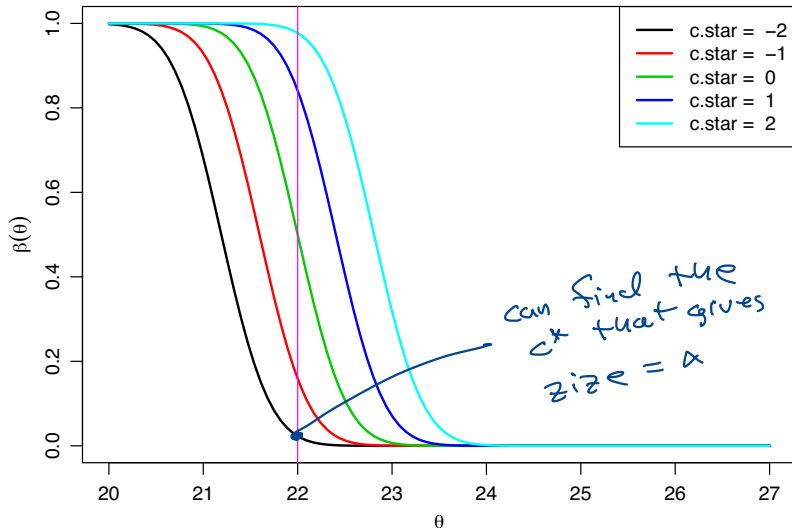
Power function for school example



# Power function for after-school

$\sum_{i=1}^n$  same for  
 $\sigma^2$  many  $c$   
 values

Power function for school example



# Level and size of tests

↳ Refer to power function only on  $\Theta_0$

## Definition: Level and size

A test with power function  $\beta(\theta)$

- is a **size**  $\alpha$  test if  $\sup_{\theta \in \Theta_0} \beta(\theta) = \alpha$
- is a **level**  $\alpha$  test if  $\sup_{\theta \in \Theta_0} \beta(\theta) \leq \alpha$

$$0 \leq \alpha \leq 1$$

- size of a test  $\leq$  level of a test
- Sometimes no distinction made between size and level.
- Choose test procedure that have a certain size (e.g. 0.1, 0.05, 0.01, 0.001)
  - Not unique, have a class  $\mathcal{C}$  of level  $\alpha$  tests
- For discrete distributions, exact size not always attainable.
  - Randomized tests

# Level and size of tests

- Want  $\beta(\theta)$  to be small for  $\theta \in \Theta_0$  and large for  $\theta \in \Theta_0^c$
- Generally there is a trade-off between these probabilities
- A common approach: Choose a number  $\alpha$  and pick a procedure (e.g.  $c^*$ ) such that

$$\beta(\theta) \leq \alpha \quad \text{for } \theta \in \Theta_0$$

That is, we put an upper bound on the probability of type I error.

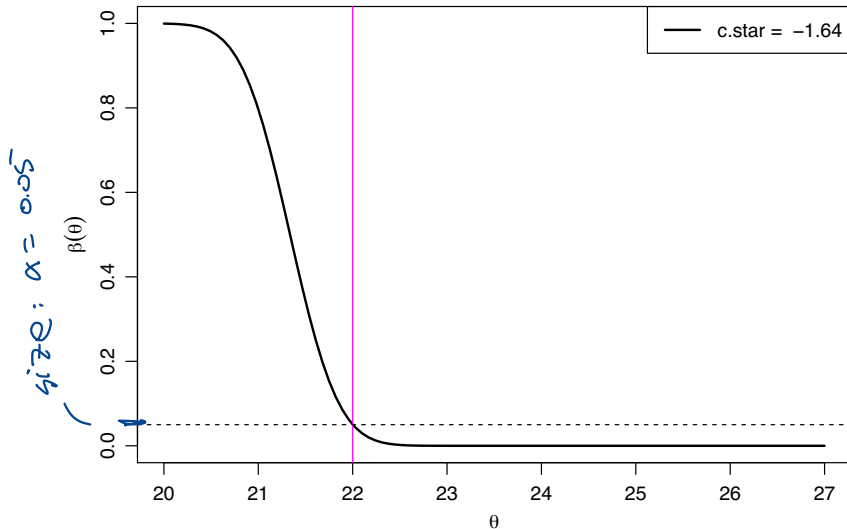
- The test is then a level  $\alpha$  test or we say that the test has **significance level**  $\alpha$

Example: Find the  $c^*$  value that makes the schools-example test a level 0.05 test.

$$c^* = -z_{0.05} = -1.64$$

# Power function for a size 0.05 test for after-school

Power function for size 0.05 test, school example





# Upper quantile notation

- A standard notation for upper quantiles.
- $z_\alpha$  is the value where

$$P(Z > z_\alpha) = \alpha \quad \text{where } Z \sim N(0, 1)$$

- $t_{n,\alpha}$  is the value where

$$P(T > t_{n,\alpha}) = \alpha \quad \text{where } T \sim t_n$$

- $\chi_{n,\alpha}^2$  is the value where

$$P(X > \chi_{n,\alpha}^2) = \alpha \quad \text{where } X \sim \chi_n^2$$

# Size / Level for the Poisson Example

- The set-up:

- $X_1, \dots, X_n$  i.i.d.  $\text{Poisson}(\lambda)$
- $H_0 : \lambda \leq 5.3$  versus  $H_1 : \lambda > 5.3$
- Reject  $H_0$  if  $\bar{X} \geq c$  for some  $c$

- Power function:

$$\beta(\lambda) = P_\lambda(T \geq nc) \quad \text{where } T = \sum_{i=1}^n X_i$$

$$= 1 - \sum_{t=0}^{\lceil nc-1 \rceil} \frac{e^{-n\lambda} (n\lambda)^t}{t!}$$

want to find the  $c$  (or  $c^* = \lceil nc-1 \rceil$ )  
such that  $\alpha = \sup_{\lambda \in (-\infty, 5.3]} P_\lambda(X) = P_{5.3}(X)$

$\frac{d}{d\lambda}$  to check

A bit complicated function of  $\lambda$  but we saw in the RMarkdown file that it is a monotone increasing function of  $\lambda$  - *Need to justify this!*

# Size / Level for the Poisson Example

- Since the power function  $\beta(\lambda)$  is a monotone increasing function of  $\lambda$ :

$$\sup_{\lambda \leq 5.3} \beta(\lambda) = \beta(5.3) = 1 - \sum_{t=0}^{\lceil nc-1 \rceil} \frac{e^{n5.3} (n5.3)^t}{t!}$$

*True! this*

- Want to find a  $c$  such that

$$1 - \sum_{t=0}^{\lceil nc-1 \rceil} \frac{e^{n5.3} (n5.3)^t}{t!} = \alpha$$

Can't "solve for  $c$ " so we try a few values - see RMarkdown file

## Size / Level for the Poisson Example

\*: Based on the fact that  $\beta(\lambda)$  is a monotone increasing function of  $\lambda$   
 Justification for our statement that the power function

$$\beta(\lambda) = 1 - \sum_{t=0}^{\lceil nc-1 \rceil} \frac{e^{-n\lambda} (n\lambda)^t}{t!}$$

is a monotone increasing function of  $\lambda$

- One way: Take arbitrary  $\lambda_1$  and  $\lambda_2$  and assume that  $\lambda_1 < \lambda_2$ . Then show that  $\beta(\lambda_1) \leq \beta(\lambda_2)$
- Another way: Show that

$$\frac{d}{d\lambda} \beta(\lambda) > 0 \quad \text{for all } \lambda > 0$$

# Unbiased tests

## Definition

A test with power function  $\beta(\theta)$  is called **unbiased** if

$$\beta(\theta) \leq \beta(\theta') \quad \forall \theta \in \Theta_0, \theta' \in \Theta_0^c$$

- For any  $\theta$  in  $\Theta_0$  we want the probability of rejecting  $H_0$  to be smaller than for any  $\theta$  in  $\Theta_0^c$

*not unbiased.*

