

Lecture 4: The Entropy Methods

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Jan. 19th., 2023

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1 Logarithmic Sobolev Inequality

1.1 Bernoulli Logarithmic Sobolev Inequality

- **Remark (Setting)**

Consider a **uniformly distributed binary vector** $Z = (Z_1, \dots, Z_n)$ on the hypercube $\{-1, +1\}^n$. In other words, the components of X are *independent, identically distributed random sign (Rademacher) variables* with $\mathbb{P}\{Z_i = -1\} = \mathbb{P}\{Z_i = +1\} = 1/2$ (i.e. *symmetric Bernoulli random variables*).

Let $f : \{-1, +1\}^n \rightarrow \mathbb{R}$ be a real-valued function on **binary hypercube**. $X := f(Z)$ is an induced real-valued random variable. Define $\tilde{Z}^{(i)} = (Z_1, \dots, Z_{i-1}, Z'_i, Z_{i+1}, \dots, Z_n)$ be the sample Z with i -th component replaced by an *independent copy* Z'_i . Since $Z, \tilde{Z}^{(i)} \in \{-1, +1\}^n$, $\tilde{Z}^{(i)} = (Z_1, \dots, Z_{i-1}, -Z_i, Z_{i+1}, \dots, Z_n)$, i.e. *the i -th sign is flipped*. Also denote the i -th *Jackknife sample* as $Z_{(-i)} = (Z_1, \dots, Z_{i-1}, Z_{i+1}, \dots, Z_n)$ by *leaving out the i -th component*. $\mathbb{E}_{(-i)}[X] := \mathbb{E}[X|Z_{(-i)}]$.

Denote the i -th component of **discrete gradient** of f as

$$\nabla_i f(z) := \frac{1}{2} \left(f(z) - f(\tilde{z}^{(i)}) \right)$$

and $\nabla f(z) = (\nabla_1 f(z), \dots, \nabla_n f(z))$

- **Remark (Jackknife Estimate of Variance)**

Recall that the **Jackknife estimate of variance**

$$\begin{aligned} \mathcal{E}(f) &:= \mathbb{E} \left[\sum_{i=1}^n \left(f(Z) - \mathbb{E}_{(-i)} \left[f(\tilde{Z}^{(i)}) \right] \right)^2 \right] \\ &= \frac{1}{2} \mathbb{E} \left[\sum_{i=1}^n \left(f(Z) - f(\tilde{Z}^{(i)}) \right)^2 \right]. \end{aligned}$$

Using the notation of discrete gradient of f , we see that

$$\mathcal{E}(f) := 2\mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right]$$

- **Remark (Entropy Functional)**

Recall that the entropy functional for f is defined as

$$H_\Phi(f(Z)) = \text{Ent}(f) := \mathbb{E} [f(Z) \log f(Z)] - \mathbb{E} [f(Z)] \log (\mathbb{E} [f(Z)]).$$

- **Proposition 1.1 (Logarithmic Sobolev Inequality for Rademacher Random Variables).** [Boucheron et al., 2013]

If $f : \{-1, +1\}^n \rightarrow \mathbb{R}$ be an arbitrary real-valued function defined on the n -dimensional **binary hypercube** and assume that Z is **uniformly distributed** over $\{-1, +1\}^n$. Then

$$\text{Ent}(f^2) \leq \mathcal{E}(f) \tag{1}$$

$$\Leftrightarrow \text{Ent}(f^2(Z)) \leq 2\mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right] \tag{2}$$

Proof: The key is to apply the tensorization property of Φ -entropy. Let $X = f(Z)$. By tensorization property,

$$\text{Ent}(X^2) \leq \sum_{i=1}^n \mathbb{E} [\text{Ent}_{(-i)}(X^2)]$$

where $\text{Ent}_{(-i)}(X^2) := \mathbb{E}_{(-i)} [X^2 \log X^2] - \mathbb{E}_{(-i)} [X^2] \log (\mathbb{E}_{(-i)} [X^2])$.

It thus suffice to show that for all $i = 1, \dots, n$,

$$\text{Ent}_{(-i)}(X^2) \leq \frac{1}{2} \mathbb{E}_{(-i)} \left[\left(f(Z) - f(\tilde{Z}^{(i)}) \right)^2 \right].$$

Given any fixed realization of $Z_{(-i)}$, $X = f(Z) = \tilde{f}(Z_i)$ can only takes two different values with equal probability. Call these two values a and b . See that

$$\begin{aligned} \text{Ent}_{(-i)}(X^2) &= \frac{1}{2} a^2 \log a^2 + \frac{1}{2} b^2 \log b^2 - \frac{1}{2} (a^2 + b^2) \log \left(\frac{a^2 + b^2}{2} \right) \\ \frac{1}{2} \mathbb{E}_{(-i)} \left[\left(f(Z) - f(\tilde{Z}^{(i)}) \right)^2 \right] &= \frac{1}{2} (a - b)^2. \end{aligned}$$

Thus we need to show

$$\frac{1}{2} a^2 \log a^2 + \frac{1}{2} b^2 \log b^2 - \frac{1}{2} (a^2 + b^2) \log \left(\frac{a^2 + b^2}{2} \right) \leq \frac{1}{2} (a - b)^2.$$

By symmetry, we may assume that $a \geq b$. Since $(|a| - |b|)^2 \leq (a - b)^2$, without loss of generality, we may further assume that $a, b \geq 0$.

Define

$$h(a) := \frac{1}{2} a^2 \log a^2 + \frac{1}{2} b^2 \log b^2 - \frac{1}{2} (a^2 + b^2) \log \left(\frac{a^2 + b^2}{2} \right) - \frac{1}{2} (a - b)^2$$

for $a \in [b, \infty)$. $h(b) = 0$. It suffice to check that $h'(b) = 0$ and that h is concave on $[b, \infty)$. Note that

$$\begin{aligned} h'(a) &= a \log a^2 + 1 - a \log \left(\frac{a^2 + b^2}{2} \right) - 1 - (a - b) \\ &= a \log \frac{2a^2}{(a^2 + b^2)} - (a - b). \end{aligned}$$

So $h'(b) = 0$. Moreover,

$$h''(a) = \log \frac{2a^2}{(a^2 + b^2)} + 1 - \frac{2a^2}{(a^2 + b^2)} \leq 0$$

due to inequality $\log(x) + 1 \leq x$. ■

- **Remark** (*Logarithmic Sobolev Inequality* \Rightarrow *Efron-Stein Inequality*). [Boucheron et al., 2013]

Note that for f non-negative,

$$\text{Var}(f(Z)) \leq \text{Ent}(f^2(Z)).$$

Thus *logarithmic Sobolev inequality* (1) implies

$$\text{Var}(f(Z)) \leq \mathcal{E}(f)$$

which is the *Efron-Stein inequality*.

- **Corollary 1.2** (*Logarithmic Sobolev Inequality for Asymmetric Bernoulli Random Variables*). [Boucheron et al., 2013]

If $f : \{-1, +1\}^n \rightarrow \mathbb{R}$ be an arbitrary real-valued function and $Z = (Z_1, \dots, Z_n) \in \{-1, +1\}^n$ with $p = \mathbb{P}\{Z_i = +1\}$. Then

$$\text{Ent}(f^2) \leq c(p) \mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right] \quad (3)$$

where

$$c(p) = \frac{1}{1-2p} \log \frac{1-p}{p}$$

Note that $\lim_{p \rightarrow 1/2} c(p) = 2$.

1.2 Gaussian Logarithmic Sobolev Inequality

- **Proposition 1.3** (*Gaussian Logarithmic Sobolev Inequality*). [Boucheron et al., 2013]

Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ be a **continuous differentiable** function and let $Z = (Z_1, \dots, Z_n)$ be a vector of n **independent standard Gaussian** random variables. Then

$$\text{Ent}(f^2(Z)) \leq 2 \mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right]. \quad (4)$$

Proof: We first prove for $n = 1$, where $f : \mathbb{R} \rightarrow \mathbb{R}$ is continuous differentiable and Z is standard Gaussian distribution. Without loss of generality, assume that $\mathbb{E}[f'(Z)] < \infty$ since it is trivial when $\mathbb{E}[f'(Z)] = \infty$. By density argument, it suffice to prove the proposition when f is *twice differentiable with bounded support*.

Now let $\epsilon_1, \dots, \epsilon_n$ be *independent Rademacher random variables* and introduce

$$S_n := \frac{1}{\sqrt{n}} \sum_{j=1}^n \epsilon_j.$$

Note that $\epsilon_i \in \{-1, +1\}$ with equal probability, thus

$$\begin{aligned} \mathbb{E}_{(-i)}[S_n] &= \frac{1}{2} \left[\left(\frac{1}{\sqrt{n}} \sum_{j \neq i} \epsilon_j + \frac{1}{\sqrt{n}} \right) + \left(\frac{1}{\sqrt{n}} \sum_{j \neq i} \epsilon_j - \frac{1}{\sqrt{n}} \right) \right] \\ &= \frac{1}{2} \left[\left(S_n + \frac{1 - \epsilon_i}{\sqrt{n}} \right) + \left(S_n - \frac{1 + \epsilon_i}{\sqrt{n}} \right) \right]. \end{aligned}$$

In the proof of Gaussian Poincaré inequality, we show that by *central limit theorem*,

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[\sum_{i=1}^n \left| f(S_n) - f\left(S_n - \frac{2\epsilon_i}{\sqrt{n}}\right) \right|^2 \right] = 4 \mathbb{E}[(f'(Z))^2].$$

On the other hands, for any *continuous uniformly bounded function* f , by *central limit theorem*,

$$\lim_{n \rightarrow \infty} \text{Ent} (f^2(S_n)) = \text{Ent}(f^2(Z))$$

The proof is then completed by invoking *the logarithmic Sobolev inequality* for *Rademacher random variables*

$$\begin{aligned} \text{Ent} (f^2(S_n)) &\leq \frac{1}{2} \mathbb{E} \left[\sum_{i=1}^n \left| f(S_n) - f \left(S_n - \frac{2\epsilon_i}{\sqrt{n}} \right) \right|^2 \right] \\ \Rightarrow \lim_{n \rightarrow \infty} \text{Ent} (f^2(S_n)) &\leq \frac{1}{2} \lim_{n \rightarrow \infty} \mathbb{E} \left[\sum_{i=1}^n \left| f(S_n) - f \left(S_n - \frac{2\epsilon_i}{\sqrt{n}} \right) \right|^2 \right] \\ &\Rightarrow \text{Ent}(f^2(Z)) \leq 2\mathbb{E} [(f'(Z))^2]. \end{aligned}$$

The extension of the result to dimension $n \geq 1$ follows easily from *the sub-additivity of entropy* which states that

$$\text{Ent}(f^2) \leq \sum_{i=1}^n \mathbb{E} [\mathbb{E}_{(-i)} [f^2(Z) \log f^2(Z)] - \mathbb{E}_{(-i)} [f^2(Z)] \log \mathbb{E}_{(-i)} [f^2(Z)]]$$

where $\mathbb{E}_{(-i)} [\cdot]$ denotes the integration with respect to i -th variable Z_i only. Thus by induction, for all i

$$\mathbb{E}_{(-i)} [f^2(Z) \log f^2(Z)] - \mathbb{E}_{(-i)} [f^2(Z)] \log \mathbb{E}_{(-i)} [f^2(Z)] \leq 2\mathbb{E}_{(-i)} [(\partial_i f(Z))^2].$$

Thus

$$\text{Ent}(f^2) \leq 2\mathbb{E} \left[\mathbb{E}_{(-i)} \left[\sum_{i=1}^n (\partial_i f(Z))^2 \right] \right] = 2\mathbb{E} [\|\nabla f(Z)\|_2^2]. \quad \blacksquare$$

- **Remark (*Dimension Free Property*).**

The *Gaussian logarithmic Sobolev inequality* has a constant $C = 2$ that is ***independent of dimension*** n :

$$\mathbb{E}_\mu [f^2] \leq 2\mathbb{E}_\mu [\|\nabla f\|_2^2].$$

This *dimension-free property* is related to the ***concentration of Gaussian measure*** μ . As a consequence, this inequality can be extended to functions of *Gaussian measure* on ***infinite dimensional space***, such as Gibbs measure, *Gaussian process* etc.

- **Remark (*Equivalent Form of Gaussian Logarithmic Sobolev Inequality*)**

Assume $f : \mathbb{R}^n \rightarrow (0, \infty)$ and $\int_{\mathbb{R}^n} f d\mu = 1$ under Gaussian measure μ . Substituting $f \rightarrow \sqrt{f}$, the *logarithmic Sobolev inequality* becomes

$$\text{Ent}_\mu(f) = \int f \log f d\mu \leq \frac{1}{2} \int \frac{\|\nabla f\|_2^2}{f} d\mu \quad (5)$$

- **Remark** (*Gaussian Logarithmic Sobolev Inequality \Rightarrow Gaussian Poincaré Inequality*). [Boucheron et al., 2013]
Recall that *the Gaussian Poincaré inequality*

$$\text{Var}(f(Z)) \leq \mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right]$$

Since

$$(1+t) \log(1+t) = t + \frac{t^2}{2} + o(t^2)$$

as $t \rightarrow 0$, we can get for Gaussian measures,

$$\text{Ent}_\mu(1 + \epsilon h) = \frac{\epsilon^2}{2} \text{Var}_\mu(h) + o(\epsilon^2).$$

Similarly,

$$\int \frac{\|\nabla(1 + \epsilon h)\|_2^2}{1 + \epsilon h} d\mu = \epsilon^2 \int \|\nabla h\|_2^2 d\mu + o(\epsilon^2).$$

Thus from *the Gaussian logarithmic Sobolev inequality*,

$$\begin{aligned} \text{Ent}_\mu(1 + \epsilon h) &\leq \frac{1}{2} \int \frac{\|\nabla(1 + \epsilon h)\|_2^2}{1 + \epsilon h} d\mu \\ \Leftrightarrow \frac{\epsilon^2}{2} \text{Var}_\mu(h) + o(\epsilon^2) &\leq \frac{\epsilon^2}{2} \int \|\nabla h\|_2^2 d\mu + o(\epsilon^2) \\ \Leftrightarrow \text{Var}(f(Z)) &\leq \mathbb{E} \left[\|\nabla f(Z)\|_2^2 \right] \quad \text{as } \epsilon \rightarrow 0. \end{aligned}$$

Thus *the Gaussian logarithmic Sobolev inequality* implies *the Gaussian Poincaré inequality*.

1.3 Information Theory Interpretation

- **Remark** (*Information Interpretation of Gaussian Logarithmic Sobolev Inequality*)

Let ν, μ be two probability measures on $(\mathcal{X}^n, \mathcal{F})$, $\mu = \mu_1 \otimes \dots \otimes \mu_n$ and $\nu \ll \mu$. Define $f := \frac{d\nu}{d\mu}$ be the Radon-Nikodym derivative of ν with respect to μ (i.e f is the probability density function of ν with respect to μ). Then the entropy becomes ***the relative entropy***

$$\text{Ent}_\mu(f) := \mathbb{E}_\mu [f \log f] = \text{KL}(\nu \parallel \mu)$$

since $\mathbb{E}_\mu [f] = \int_{\mathcal{X}^n} f d\mu = 1$.

On the other hand, ***the (relative) Fisher information*** is defined as

$$\begin{aligned} I(\nu \parallel \mu) &:= \mathbb{E}_\nu \left[\|\nabla \log f\|_2^2 \right] \\ &= \int \left\| \frac{\nabla f}{f} \right\|_2^2 d\nu = \int \frac{\|\nabla f\|_2^2}{f^2} d\nu \\ &= \int \frac{\|\nabla f\|_2^2}{f} d\mu \end{aligned}$$

Thus the information interpretation of the Gaussian logarithmic Sobolev inequality is

$$\mathbb{KL}(\nu \parallel \mu) \leq \frac{1}{2} I(\nu \parallel \mu) \quad (6)$$

where μ is a Gaussian measure and $\nu \ll \mu$ with density function f . Note that the Fisher information metric is **the Riemannian metric** induced by the relative entropy.

1.4 Logarithmic Sobolev Inequality for General Probability Measures

- From functional analysis, we have the Sobolev inequality,

Remark (The Sobolev Inequality) [Evans, 2010]

The Sobolev inequality states for smooth function $f : \mathbb{R}^n \rightarrow \mathbb{R}$ in Sobolev space where $n \geq 3$ and $p = \frac{2n}{n-2} > 2$

$$\|f\|_p^2 \leq C_n \int_{\mathbb{R}^n} |\nabla f|^2 dx.$$

The inequality is sharp when the constant

$$C_n := \frac{1}{\pi n(n-2)} \left(\frac{\Gamma(n)}{\Gamma(n/2)} \right)^{2/n}$$

- **Proposition 1.4 (Euclidean Logarithmic Sobolev Inequality).**

Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ be a smooth function and m be Lebesgue measure on \mathbb{R}^n , then

$$\begin{aligned} \text{Ent}_m(f^2) &\leq \frac{n}{2} \log \left(\frac{2}{n\pi e} \mathbb{E}_m \left[\|\nabla f\|_2^2 \right] \right) \\ \Leftrightarrow \int f^2 \log \left(\frac{f^2}{\int f^2 dx} \right) dx &\leq \frac{n}{2} \log \left(\frac{2}{n\pi e} \int |\nabla f|^2 dx \right) \end{aligned} \quad (7)$$

- **Definition (Logarithmic Sobolev Inequality for General Probability Measure).**

A probability measure μ on \mathbb{R}^n is said to satisfy the logarithmic Sobolev inequality for some constant $C > 0$ if for any smooth function f

$$\text{Ent}_\mu(f^2) \leq C \mathbb{E}_\mu \left[\|\nabla f\|_2^2 \right] \quad (8)$$

holds for any **continuous differentiable** function $f : \mathbb{R}^n \rightarrow \mathbb{R}$. The left-hand side is called **the entropy functional**, which is defined as

$$\begin{aligned} \text{Ent}(f^2) &:= \mathbb{E}_\mu [f^2 \log f^2] - \mathbb{E}_\mu [f^2] \log \mathbb{E}_\mu [f^2] \\ &= \int f^2 \log \left(\frac{f^2}{\int f^2 d\mu} \right) d\mu. \end{aligned}$$

The right-hand side is defined as

$$\mathbb{E}_\mu \left[\|\nabla f\|_2^2 \right] = \int \|\nabla f\|_2^2 d\mu.$$

Thus we can rewrite the logarithmic Sobolev inequality in functional form

$$\int f^2 \log \left(\frac{f^2}{\int f^2 d\mu} \right) d\mu \leq C \int \|\nabla f\|_2^2 d\mu \quad (9)$$

- **Remark** (*Logarithmic Sobolev Inequality*)

For non-negative function f , we can replace $f \rightarrow \sqrt{f}$, so that the logarithmic Sobolev inequality becomes

$$\text{Ent}_\mu(f) \leq C \int \frac{\|\nabla f\|_2^2}{f} d\mu \quad (10)$$

- **Remark** (*Modified Logarithmic Sobolev Inequality via Convex Cost and Duality*)

For some *convex non-negative cost* $c : \mathbb{R}^n \rightarrow \mathbb{R}_+$, the *convex conjugate* of c (Legendre transform of c) is defined as

$$c^*(x) := \sup_y \{ \langle x, y \rangle - c(y) \}$$

Then we can obtain the *modified logarithmic Sobolev inequality*

$$\text{Ent}_\mu(f) \leq \int f^2 c^* \left(\frac{\nabla f}{f} \right) d\mu \quad (11)$$

2 The Entropy Methods

2.1 Herbst's Argument

- **Remark** Recall that the Φ -entropy for $\Phi(x) = x \log(x)$ as

$$H_\Phi(X) = \text{Ent}(X) := \mathbb{E}[X \log X] - \mathbb{E}[X] \log(\mathbb{E}[X]).$$

The variational formulation of $H_\Phi(X)$ is

$$\text{Ent}(X) = \sup_T \{ X (\log(T) - \log(\mathbb{E}[T])) \}$$

- **Remark** (*Tensorization Property of Entropy Functional*)

Let $\mu = \mu_1 \otimes \dots \otimes \mu_n$ be the probability distribution for $Z = (Z_1, \dots, Z_n)$ on $(\mathcal{X}^n, \mathcal{F})$. For any measurable function $f : \mathcal{X}^n \rightarrow \mathbb{R}$, let $X = f(Z_1, \dots, Z_n)$ so that $\mathbb{E}[X \log X] < \infty$. The *sub-additivity of entropy function (i.e. the tensorization property)* states that

$$\text{Ent}_{\mu_1 \otimes \dots \otimes \mu_n}(f) \leq \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} \left[\sum_{i=1}^n \text{Ent}_{\mu_i}(f) \right]$$

where the subscript μ_i indicates that the integration concerns the i -th variable only.

- **Remark** (*Entropy Functional for Moment Generating Function*)

Let $X = e^{\lambda Z}$ where Z is a random variable. The entropy function of X becomes

$$\text{Ent}(e^{\lambda Z}) = \mathbb{E}[\lambda Z e^{\lambda Z}] - \mathbb{E}[e^{\lambda Z}] \log(\mathbb{E}[e^{\lambda Z}])$$

Denote $\psi_{Z-\mathbb{E}[Z]}(\lambda) := \log \mathbb{E} [e^{\lambda(Z-\mathbb{E}[Z])}]$. Then

$$\begin{aligned}
\psi'_{Z-\mathbb{E}[Z]}(\lambda) &= \frac{d}{d\lambda} \log \mathbb{E} [e^{\lambda(Z-\mathbb{E}[Z])}] \\
&= \frac{1}{\mathbb{E} [e^{\lambda(Z-\mathbb{E}[Z])}]} \mathbb{E} [(Z - \mathbb{E}[Z]) e^{\lambda(Z-\mathbb{E}[Z])}] \\
&= \frac{1}{\mathbb{E} [e^{\lambda Z}]} e^{\lambda \mathbb{E}[Z]} \mathbb{E} [(Z - \mathbb{E}[Z]) e^{\lambda(Z-\mathbb{E}[Z])}] \\
&= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \mathbb{E} [(Z - \mathbb{E}[Z]) e^{\lambda Z}] \\
&= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \mathbb{E} [Z e^{\lambda Z}] - \mathbb{E}[Z] \\
\Rightarrow \lambda \psi'_{Z-\mathbb{E}[Z]}(\lambda) &= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \left(\mathbb{E} [\lambda Z e^{\lambda Z}] - \mathbb{E} [\lambda Z] \mathbb{E} [e^{\lambda Z}] \right) \\
\Rightarrow \lambda \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \psi_{Z-\mathbb{E}[Z]}(\lambda) &= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \left\{ \mathbb{E} [\lambda Z e^{\lambda Z}] - \mathbb{E} [\lambda Z] \mathbb{E} [e^{\lambda Z}] - \mathbb{E} [e^{\lambda Z}] \log \mathbb{E} [e^{\lambda(Z-\mathbb{E}[Z])}] \right\} \\
&= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \left\{ \mathbb{E} [\lambda Z e^{\lambda Z}] - \mathbb{E} [\lambda Z] \mathbb{E} [e^{\lambda Z}] \right. \\
&\quad \left. + \mathbb{E} [e^{\lambda Z}] \mathbb{E} [\lambda Z] - \mathbb{E} [e^{\lambda Z}] \log \mathbb{E} [e^{\lambda Z}] \right\} \\
&= \frac{1}{\mathbb{E} [e^{\lambda Z}]} \left\{ \mathbb{E} [\lambda Z e^{\lambda Z}] - \mathbb{E} [e^{\lambda Z}] \log \mathbb{E} [e^{\lambda Z}] \right\} \\
&= \frac{\text{Ent}(e^{\lambda Z})}{\mathbb{E} [e^{\lambda Z}]}
\end{aligned}$$

Thus we have

$$\frac{\text{Ent}(e^{\lambda Z})}{\mathbb{E} [e^{\lambda Z}]} = \lambda \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \psi_{Z-\mathbb{E}[Z]}(\lambda). \quad (12)$$

Our strategy is based on using (12) *the sub-additivity of entropy* and then univariate calculus to derive **upper bounds** for the **derivative** of $\psi(\lambda)$. By solving the obtained **differential inequality**, we obtain tail bounds via *Chernoff's bounding*.

• **Proposition 2.1 (Herbst's Argument)** [Boucheron et al., 2013, Wainwright, 2019]

Let Z be an integrable random variable such that for some $\nu > 0$, we have, for every $\lambda > 0$,

$$\frac{\text{Ent}(e^{\lambda Z})}{\mathbb{E} [e^{\lambda Z}]} \leq \frac{\nu \lambda^2}{2} \quad (13)$$

Then, for every $\lambda > 0$, the logarithmic moment generating function of centered random variable $(Z - \mathbb{E}[Z])$ satisfies

$$\psi_{Z-\mathbb{E}[Z]}(\lambda) := \log \mathbb{E} [e^{\lambda(Z-\mathbb{E}[Z])}] \leq \frac{\nu \lambda^2}{2}.$$

Proof: The condition of the proposition means, via (12), that

$$\lambda \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \psi_{Z-\mathbb{E}[Z]}(\lambda) \leq \frac{\nu \lambda^2}{2},$$

or equivalently,

$$\frac{1}{\lambda} \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \frac{1}{\lambda^2} \psi_{Z-\mathbb{E}[Z]}(\lambda) \leq \frac{\nu}{2}.$$

Setting $G(\lambda) = \lambda^{-1} \psi_{Z-\mathbb{E}[Z]}(\lambda)$, we see that the differential inequality becomes

$$G'(\lambda) \leq \frac{\nu}{2}.$$

Since $G(\lambda) \rightarrow 0$ as $\lambda \rightarrow 0$, which implies that

$$G(\lambda) \leq \frac{\nu\lambda}{2},$$

and the result follows. \blacksquare

• **Remark (*Entropy Methods*)**

The **key strategy of entropy methods** to prove the concentration of function $f(Z)$ of independent variables Z is as follows

1. Apply **the tensorization of property of the entropy functional** on $f(Z)$

$$\text{Ent}_{\mu_1 \otimes \dots \otimes \mu_n}(f) \leq \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} \left[\sum_{i=1}^n \text{Ent}_{\mu_i}(f) \right]$$

where $\mu := \mu_1 \otimes \dots \otimes \mu_n$ is the distribution of Z and the subscript μ_i indicates that the integration concerns the i -th variable only.

2. Then we **bound the entropy** for each individual variables Z_i conditioning on the rest of them, i.e. $Z_{(-i)} := (Z_1, \dots, Z_{i-1}, Z_{i+1}, \dots, Z_n)$. Typically, we use **the Herbst's argument**, when $f(Z) = e^{\lambda Z}$

$$\text{Ent}_{\mu_i}(e^{\lambda Z}) \leq \phi(\lambda) \mathbb{E}_{\mu_i} [e^{\lambda Z}].$$

Or, we can use **the logarithmic Sobolev inequality** if Z is *Gaussian* or *Bernoulli distribution*,

$$\text{Ent}_{\mu_i}(e^{\lambda Z}) \leq 2\mathbb{E}_{\mu_i} \left[\left\| \nabla e^{\lambda Z/2} \right\|_2^2 \right] = \frac{\lambda^2}{2} \mathbb{E}_{\mu_i} \left[e^{\lambda Z} \|\nabla Z\|_2^2 \right].$$

3. Then we **bound the entire entropy**

$$\text{Ent}_{\mu_1 \otimes \dots \otimes \mu_n}(e^{\lambda Z}) \leq \phi(\lambda) \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} \left[\sum_{i=1}^n \mathbb{E}_{\mu_i} [e^{\lambda Z}] \right] = n\phi(\lambda) \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} [e^{\lambda Z}]$$

or

$$\text{Ent}_{\mu_1 \otimes \dots \otimes \mu_n}(f) \leq \frac{\lambda^2}{2} \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} \left[\sum_{i=1}^n \mathbb{E}_{\mu_i} [e^{\lambda Z} \|\nabla Z\|_2^2] \right].$$

4. Apply **the Herbst's argument** to obtain **differential inequality** for **the logarithmic moment generating function** $\psi_{f(Z)}$
5. Obtain concentration results based on **Chernoff's inequality**.

$$\mathbb{P} \{f(Z) - \mathbb{E}[f(Z)] > t\} \leq \inf_{\lambda > 0} \exp(\psi_{f(Z)}(\lambda) - \lambda t) \leq \inf_{\lambda > 0} \exp(\phi(\lambda) - \lambda t)$$

2.2 Bounded Difference Inequality via Entropy Methods

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2.3 Modified Logarithmic Sobolev Inequalities

- **Proposition 2.2 (A Modified Logarithmic Sobolev Inequalities for Moment Generating Function)** [Boucheron et al., 2013]

Consider independent random variables Z_1, \dots, Z_n taking values in \mathcal{X} , a real-valued function $f : \mathcal{X}^n \rightarrow \mathbb{R}$ and the random variable $X = f(Z_1, \dots, Z_n)$. Also denote $Z_{(-i)} = (Z_1, \dots, Z_{i-1}, Z_{i+1}, \dots, Z_n)$ and $X_{(-i)} = f_i(Z_{(-i)})$ where $f_i : \mathcal{X}^{n-1} \rightarrow \mathbb{R}$ is an arbitrary function. Let $\phi(x) = e^x - x - 1$. Then for all $\lambda \in \mathbb{R}$,

$$\lambda \mathbb{E} [X e^{\lambda X}] - \mathbb{E} [e^{\lambda X}] \log \mathbb{E} [e^{\lambda X}] \leq \sum_{i=1}^n \mathbb{E} [e^{\lambda X} \phi(-\lambda(X - X_{(-i)}))] \quad (14)$$

Proof: Recall the tensorization of entropy

$$\text{Ent}_{\mu_1 \otimes \dots \otimes \mu_n}(Y) \leq \mathbb{E}_{\mu_1 \otimes \dots \otimes \mu_n} \left[\sum_{i=1}^n \text{Ent}_{\mu_i}(Y) \right].$$

We bound each term on the right-hand side by the variational formulation of entropy

$$\text{Ent}_{\mu_i}(Y) \leq \mathbb{E}_{\mu_i} [Y(\log Y - \log u) - (Y - u)]$$

for any $u > 0$. Let $u = Y_{(-i)} = g_i(Z_{(-i)})$. We have

$$\text{Ent}_{\mu_i}(Y) \leq \mathbb{E}_{\mu_i} [Y(\log Y - \log Y_{(-i)}) - (Y - Y_{(-i)})].$$

Applying above inequality to the variable $Y = e^{\lambda X}$ and $Y_{(-i)} = e^{\lambda X_{(-i)}}$, one obtain

$$\begin{aligned} \text{Ent}_{\mu_i}(e^{\lambda X}) &\leq \mathbb{E}_{\mu_i} [e^{\lambda X} (\log e^{\lambda X} - \log e^{\lambda X_{(-i)}}) - (e^{\lambda X} - e^{\lambda X_{(-i)}})] \\ &= \mathbb{E}_{\mu_i} [e^{\lambda X} (\lambda(X - X_{(-i)}) - (e^{\lambda X} - e^{\lambda X_{(-i)}}))] \\ &= \mathbb{E}_{\mu_i} [e^{\lambda X} (\lambda(X - X_{(-i)}) - e^{-\lambda X} (e^{\lambda X} - e^{\lambda X_{(-i)}}))] \\ &= \mathbb{E}_{\mu_i} [e^{\lambda X} (\lambda(X - X_{(-i)}) + e^{-\lambda(X - X_{(-i)})} - 1)] \\ &= \mathbb{E}_{\mu_i} [e^{\lambda X} \phi(-\lambda(X - X_{(-i)}))] \end{aligned}$$

where $\phi(x) = e^x - x - 1$. Thus the proof is completed. \blacksquare

- **Proposition 2.3 (Symmetrized Modified Logarithmic Sobolev Inequalities)** [Boucheron et al., 2013]

Consider independent random variables Z_1, \dots, Z_n taking values in \mathcal{X} , a real-valued function $f : \mathcal{X}^n \rightarrow \mathbb{R}$ and the random variable $X = f(Z_1, \dots, Z_n)$. Also denote $\tilde{X}^{(i)} = f(Z_1, \dots, Z_{i-1}, Z'_i, Z_{i+1}, \dots, Z_n)$. Let $\phi(x) = e^x - x - 1$. Then for all $\lambda \in \mathbb{R}$,

$$\lambda \mathbb{E} [X e^{\lambda X}] - \mathbb{E} [e^{\lambda X}] \log \mathbb{E} [e^{\lambda X}] \leq \sum_{i=1}^n \mathbb{E} [e^{\lambda X} \phi(-\lambda(X - \tilde{X}^{(i)}))] \quad (15)$$

Moreover, denoting $\tau(x) = x(e^x - 1)$, for all $\lambda \in \mathbb{R}$,

$$\begin{aligned}\lambda \mathbb{E} \left[X e^{\lambda X} \right] - \mathbb{E} \left[e^{\lambda X} \right] \log \mathbb{E} \left[e^{\lambda X} \right] &\leq \sum_{i=1}^n \mathbb{E} \left[e^{\lambda X} \tau(-\lambda(X - \tilde{X}^{(i)})_+) \right], \\ \lambda \mathbb{E} \left[X e^{\lambda X} \right] - \mathbb{E} \left[e^{\lambda X} \right] \log \mathbb{E} \left[e^{\lambda X} \right] &\leq \sum_{i=1}^n \mathbb{E} \left[e^{\lambda X} \tau(\lambda(\tilde{X}^{(i)} - X)_-) \right].\end{aligned}$$

Proof: Note that $X_{(-i)}$ and $\tilde{X}^{(i)}$ are both independent from Z_i . The first inequality is the same as the proposition above. For the second inequality, use the fact that

$$\begin{aligned}\mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(\lambda(\tilde{X}^{(i)} - X)_+ \right) \right] &= \mathbb{E}_{\mu_i} \left[e^{\lambda \tilde{X}^{(i)}} \phi \left(\lambda(X - \tilde{X}^{(i)})_+ \right) \right] \\ &= \mathbb{E}_{\mu_i} \left[e^{\lambda X} e^{-\lambda(X - \tilde{X}^{(i)})} \phi \left(\lambda(X - \tilde{X}^{(i)})_+ \right) \right].\end{aligned}$$

and

$$\begin{aligned}\mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(-\lambda(\tilde{X}^{(i)} - X) \right) \right] &= \mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(-\lambda(\tilde{X}^{(i)} - X)_+ \right) \right] + \mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(\lambda(\tilde{X}^{(i)} - X)_+ \right) \right] \\ &= \mathbb{E}_{\mu_i} \left[e^{\lambda X} \left\{ \phi \left(-\lambda(\tilde{X}^{(i)} - X) \right) + e^{-\lambda(X - \tilde{X}^{(i)})} \phi \left(\lambda(X - \tilde{X}^{(i)})_+ \right) \right\} \right].\end{aligned}$$

Finally note that $\phi(x) + e^x \phi(-x) = \tau(x) = x(e^x - 1)$. ■

2.4 Poisson Logarithmic Sobolev Inequality

- **Proposition 2.4 (Modified Logarithmic Sobolev Inequality for Bernoulli Random Variable).** [Boucheron et al., 2013]

Let $f : \{0, 1\} \rightarrow (0, \infty)$ be a **non-negative** real-valued function defined on the binary set $\{0, 1\}$. Define **the discrete derivative** of f at $x \in \{0, 1\}$ by

$$\nabla f := f(1 - x) - f(x).$$

Let X be a Bernoulli random variable with parameter $p \in (0, 1)$ (i.e. $\mathbb{P}\{X = 1\} = p$). Then

$$\text{Ent}(f(X)) \leq (p(1 - p)) \mathbb{E} [\nabla f(X) \nabla \log f(X)]. \quad (16)$$

and

$$\text{Ent}(f(X)) \leq (p(1 - p)) \mathbb{E} \left[\frac{|\nabla f(X)|^2}{f(X)} \right]. \quad (17)$$

- **Proposition 2.5 (Poisson Logarithmic Sobolev Inequality).** [Boucheron et al., 2013]
Let $f : \mathbb{N} \rightarrow (0, \infty)$ be a **non-negative** real-valued function defined on the set of non-negative integers \mathbb{N} . Define **the discrete derivative** of f at $x \in \mathbb{N}$ by

$$\nabla f := f(x + 1) - f(x).$$

Let X be a Poisson random variable. Then

$$\text{Ent}(f(X)) \leq (\mathbb{E}[X]) \mathbb{E} [\nabla f(X) \nabla \log f(X)]. \quad (18)$$

and

$$\text{Ent}(f(X)) \leq (\mathbb{E}[X]) \mathbb{E} \left[\frac{|\nabla f(X)|^2}{f(X)} \right]. \quad (19)$$

3 Applications

3.1 Lipschitz Functions of Gaussian Variables

- **Theorem 3.1 (Rademacher Theorem).**

If $f : U \rightarrow \mathbb{R}$ is a L -Lipschitz function where $U \subseteq \mathbb{R}^n$, then f is **differentiable almost everywhere** in U and **the essential supremum of the norm of its derivative is bounded by its Lipschitz constant**.

- **Theorem 3.2 (Lipschitz Functions of Gaussian Variables)** [Boucheron et al., 2013]
Let $Z = (Z_1, \dots, Z_n)$ be a vector of n **independent standard normal** random variables. Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ denote an **L -Lipschitz function**, that is, there exists a constant $L > 0$ such that for all $x, y \in \mathbb{R}^n$,

$$|f(x) - f(y)| \leq L \|x - y\|.$$

Then, for all $\lambda \in \mathbb{R}$,

$$\psi_{f(Z) - \mathbb{E}[f(Z)]}(\lambda) := \log \mathbb{E} \left[e^{\lambda(f(Z) - \mathbb{E}[f(Z)])} \right] \leq \frac{L^2 \lambda^2}{2} \quad (20)$$

Proof: By a standard density argument we may assume that f is *differentiable* with *gradient uniformly bounded by L* according to *Rademacher theorem*. We may also assume, without loss of generality, that $\mathbb{E}[f(Z)] = 0$. Using the *Gaussian logarithmic Sobolev inequality* for the function $e^{\lambda f/2}$, we obtain

$$\begin{aligned} \text{Ent}(e^{\lambda f}) &\leq 2 \mathbb{E} \left[\left\| \nabla e^{\lambda f/2} \right\|_2^2 \right] \\ &= 2 \mathbb{E} \left[\left\| \frac{\lambda}{2} e^{\lambda f/2} \nabla f \right\|_2^2 \right] \\ &= \frac{\lambda^2}{2} \mathbb{E} \left[e^{\lambda f} \|\nabla f\|_2^2 \right] \\ &\leq \frac{\lambda^2}{2} \mathbb{E} \left[e^{\lambda f} \right] L^2 \\ \Rightarrow \frac{\text{Ent}(e^{\lambda f})}{\mathbb{E} [e^{\lambda f}]} &\leq \frac{L^2 \lambda^2}{2} \end{aligned}$$

By Herbst's argument, this implies that $\lambda \psi'(\lambda) - \psi(\lambda) \leq \frac{L^2 \lambda^2}{2}$. Let $G(\lambda) := \lambda^{-1} \psi(\lambda)$, we have $G'(\lambda) \leq L^2/2$, so we have

$$\psi_{f(Z)}(\lambda) \leq \frac{L^2 \lambda^2}{2} \quad \blacksquare$$

- **Theorem 3.3 (Gaussian Concentration Inequality / The Tsirelson-Ibragimov-Sudakov Inequality)** [Boucheron et al., 2013, Wainwright, 2019]

Let $Z = (Z_1, \dots, Z_n)$ be a vector of n **independent standard normal** random variables. Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ denote an **L -Lipschitz function**. Then, for all $t > 0$,

$$\mathbb{P} \{ f(Z) - \mathbb{E}[f(Z)] \geq t \} \leq \exp \left(-\frac{t^2}{2L^2} \right). \quad (21)$$

Proof: The previous theorem states that $f(Z)$ is *sub-Gaussian distributed* with parameter $\nu = L^2$. By Markov's inequality,

$$\begin{aligned}\mathbb{P}\{f(Z) - \mathbb{E}[f(Z)] \geq t\} &= \mathbb{P}\left\{e^{\lambda(f(Z) - \mathbb{E}[f(Z)])} \geq e^{\lambda t}\right\} \leq \frac{\mathbb{E}\left[e^{\lambda f(Z) - \mathbb{E}[f(Z)]}\right]}{e^{\lambda t}} \\ &\leq \inf_{\lambda > 0} \exp(\psi(\lambda) - \lambda t) \\ &\leq \inf_{\lambda > 0} \exp\left(\frac{L^2 \lambda^2}{2} - \lambda t\right) \\ &= \exp\left(-\frac{t^2}{2L^2}\right)\end{aligned}$$

where $\lambda^* = \frac{t}{L^2}$. ■

- **Remark (*Dimension-Free Concentration*)**

An **important** feature of the theorem is that the right-hand side **does not depend on the dimension** n . This inequality has served as a benchmark for the development of concentration inequalities during the last three decades.

3.2 Suprema of Gaussian Process

- **Definition (*Gaussian Process*)**

Let T be a *metric space*. A stochastic process $(X_t)_{t \in T}$ is a **Gaussian process indexed by T** if for any finite collection $\{t_1, \dots, t_n\} \subset T$, the vector $(X_{t_1}, \dots, X_{t_n})$ has a *jointly Gaussian distribution*.

In addition, we assume that T is **totally bounded** (i.e. for every $t > 0$ it can be covered by *finitely many* balls of radius t) and that the *Gaussian process* is **almost surely continuous**, that is, with probability 1, X_t is a *continuous function* of t .

- **Theorem 3.4 (*Concentration of Suprema of Gaussian Process*)** [Boucheron et al., 2013, Vershynin, 2018, Wainwright, 2019, Giné and Nickl, 2021]
Let $(X_t)_{t \in T}$ be an **almost surely continuous centered Gaussian process** indexed by a **totally bounded set** T . If

$$\sigma^2 := \sup_{t \in T} \mathbb{E}[X_t^2],$$

then $Z = \sup_{t \in T} X_t$ satisfies $\text{Var}(Z) \leq \sigma^2$, and for all $u > 0$,

$$\mathbb{P}\{Z - \mathbb{E}[Z] \geq u\} \leq \exp\left(-\frac{u^2}{2\sigma^2}\right) \quad (22)$$

and

$$\mathbb{P}\{\mathbb{E}[Z] - Z \geq u\} \leq \exp\left(-\frac{u^2}{2\sigma^2}\right) \quad (23)$$

Proof: We assume that T is a **finite set**. The extension to arbitrary *totally bounded* T is based on a **separability argument** and **monotone convergence**. We may assume, for simplicity, that $T = \{1, \dots, n\}$. Let Σ be the covariance matrix of the centered Gaussian

vector $X = (X_1, \dots, X_n)$. Denote by A the *square root* of the positive semidefinite matrix Γ . If $Y = (Y_1, \dots, Y_n)$ is a vector of *independent standard normal random variables*, then

$$f(Y) = \max_{i=1, \dots, n} (AY)_i$$

has the same distribution as $Z = \max_{i=1, \dots, n} X_i$. Hence, we can apply *the Gaussian concentration inequality* by bounding the *Lipschitz constant* of f . By *the Cauchy-Schwarz inequality*, for all $u, v \in \mathbb{R}^n$ and $i = 1, \dots, n$,

$$|(Au)_i - (Av)_i| \leq \|A_{i:}\|_2 \|u - v\|_2$$

Since $\|A_{i:}\|_2^2 = \sum_j A_{i,j}^2 = \text{Var}(X_i)$, we get

$$|f(u) - f(v)| \leq \max_{i=1, \dots, n} |(Au)_i - (Av)_i| \leq \sigma \|u - v\|_2$$

Therefore, f is *Lipschitz with constant* σ and the tail bounds follow from the Gaussian concentration inequality. The variance bound follows from *the Gaussian Poincaré inequality*. ■

3.3 Concentration of Convex Lipschitz Functions

3.4 Concentration on the Hypercube

3.5 Gaussian Hypercontractivity

3.6 Hypercontractivity for Boolean Polynomials

3.7 The Johnson-Lindenstrauss Lemma

3.8 Exponential Tail Bounds for Self-Bounding Functions

References

- Stéphane Boucheron, Gábor Lugosi, and Pascal Massart. *Concentration inequalities: A nonasymptotic theory of independence*. Oxford university press, 2013.
- Lawrence C Evans. *Partial differential equations*, volume 19. American Mathematical Soc., 2010.
- Evarist Giné and Richard Nickl. *Mathematical foundations of infinite-dimensional statistical models*. Cambridge university press, 2021.
- Roman Vershynin. *High-dimensional probability: An introduction with applications in data science*, volume 47. Cambridge university press, 2018.
- Martin J Wainwright. *High-dimensional statistics: A non-asymptotic viewpoint*, volume 48. Cambridge University Press, 2019.