Lecture 4: The Entropy Methods

Tianpei Xie

Jan. 19th., 2023

Contents

1	\mathbf{Log}	carithmic Sobolev Inequality	2
	1.1	Bernoulli Logarithmic Sobolev Inequality	2
	1.2	Gaussian Logarithmic Sobolev Inequality	4
	1.3	Information Theory Interpretation	6
	1.4	Logarithmic Sobolev Inequality for General Probability Measures	7
2	The	Entropy Methods	8
	2.1	Herbst's Argument	8
	2.2	Modified Logarithmic Sobolev Inequalities	10
	2.3	Bounded Difference Inequality via Entropy Methods	12
	2.4	Poisson Logarithmic Sobolev Inequality	14
3	App	plications	14
	3.1	Lipschitz Functions of Gaussian Variables	14
	3.2	Suprema of Gaussian Process	16
	3.3	Concentration of Convex Lipschitz Functions	17
	3.4	Exponential Tail Bounds for Self-Bounding Functions	17
	3.5	Concentration of Lipschitz Functions on the Hypercube	17
	3.6	Hypercontractivity for Boolean Polynomials	17
	3.7	Gaussian Hypercontractivity	17

1 Logarithmic Sobolev Inequality

1.1 Bernoulli Logarithmic Sobolev Inequality

• Remark (Setting)

Consider a *uniformly distributed binary vector* $Z = (Z_1, ..., Z_n)$ on the hypercube $\{-1, +1\}^n$. In other words, the components of X are *independent*, *identically distributed* random sign (Rademacher) variables with $\mathbb{P}\{Z_i = -1\} = \mathbb{P}\{Z_i = +1\} = 1/2$ (i.e. symmetric Bernoulli random variables).

Let $f: \{-1, +1\}^n \to \mathbb{R}$ be a real-valued function on **binary hypercube**. X := f(Z) is an induced real-valued random variable. Define $\widetilde{Z}^{(i)} = (Z_1, \dots, Z_{i-1}, Z'_i, Z_{i+1}, \dots, Z_n)$ be the sample Z with i-th component replaced by an independent $copy Z'_i$. Since $Z, \widetilde{Z}^{(i)} \in \{-1, +1\}^n$, $\widetilde{Z}^{(i)} = (Z_1, \dots, Z_{i-1}, -Z_i, Z_{i+1}, \dots, Z_n)$, i.e. the i-th sign is **flipped**. Also denote the i-th Jackknife sample as $Z_{(i)} = (Z_1, \dots, Z_{i-1}, Z_{i+1}, \dots, Z_n)$ by leaving out the i-th component. $\mathbb{E}_{(-i)}[X] := \mathbb{E}[X|Z_{(i)}]$.

Denote the i-th component of **discrete gradient** of f as

$$\nabla_i f(z) := \frac{1}{2} \left(f(z) - f(\widetilde{z}^{(i)}) \right)$$

and
$$\nabla f(z) = (\nabla_1 f(z), \dots, \nabla_n f(z))$$

• Remark (Jackknife Estimate of Variance)
Recall that the Jackknife estimate of variance

$$\mathcal{E}(f) := \mathbb{E}\left[\sum_{i=1}^{n} \left(f(Z) - \mathbb{E}_{(-i)}\left[f(\widetilde{Z}^{(i)})\right]\right)^{2}\right]$$
$$= \frac{1}{2}\mathbb{E}\left[\sum_{i=1}^{n} \left(f(Z) - f(\widetilde{Z}^{(i)})\right)^{2}\right].$$

Using the notation of discrete gradient of f, we see that

$$\mathcal{E}(f) := 2\mathbb{E}\left[\left\|\nabla f(Z)\right\|_{2}^{2}\right]$$

ullet Remark ($Entropy\ Functional$)

Recall that the entropy functional for f is defined as

$$H_{\Phi}(f(Z)) = \operatorname{Ent}(f) := \mathbb{E}\left[f(Z)\log f(Z)\right] - \mathbb{E}\left[f(Z)\right]\log \left(\mathbb{E}\left[f(Z)\right]\right).$$

• Proposition 1.1 (Logarithmic Sobolev Inequality for Rademacher Random Variables). [Boucheron et al., 2013]

If $f: \{-1,+1\}^n \to \mathbb{R}$ be an arbitrary real-valued function defined on the n-dimensional binary hypercube and assume that Z is uniformly distributed over $\{-1,+1\}^n$. Then

$$Ent(f^2) \le \mathcal{E}(f)$$
 (1)

$$\Leftrightarrow \operatorname{Ent}(f^2(Z)) \le 2\mathbb{E}\left[\|\nabla f(Z)\|_2^2\right] \tag{2}$$

Proof: The key is to apply the tensorization property of Φ -entropy. Let X = f(Z). By tensorization property,

$$\operatorname{Ent}(X^2) \le \sum_{i=1}^n \mathbb{E}\left[\operatorname{Ent}_{(-i)}(X^2)\right]$$

where $\operatorname{Ent}_{(-i)}(X^2) := \mathbb{E}_{(-i)} \left[X^2 \log X^2 \right] - \mathbb{E}_{(-i)} \left[X^2 \right] \log \left(\mathbb{E}_{(-i)} \left[X^2 \right] \right).$

It thus suffice to show that for all i = 1, ..., n,

$$\operatorname{Ent}_{(-i)}(X^2) \le \frac{1}{2} \mathbb{E}_{(-i)} \left[\left(f(Z) - f(\widetilde{Z}^{(i)}) \right)^2 \right].$$

Given any fixed realization of $Z_{(-i)}$, $X = f(Z) = \widetilde{f}(Z_i)$ can only takes two different values with equal probability. Call these two values a and b. See that

$$\operatorname{Ent}_{(-i)}(X^2) = \frac{1}{2}a^2 \log a^2 + \frac{1}{2}b^2 \log b^2 - \frac{1}{2}(a^2 + b^2) \log \left(\frac{a^2 + b^2}{2}\right)$$
$$\frac{1}{2}\mathbb{E}_{(-i)}\left[\left(f(Z) - f(\widetilde{Z}^{(i)})\right)^2\right] = \frac{1}{2}\left(a - b\right)^2.$$

Thus we need to show

$$\frac{1}{2}a^2\log a^2 + \frac{1}{2}b^2\log b^2 - \frac{1}{2}(a^2 + b^2)\log\left(\frac{a^2 + b^2}{2}\right) \le \frac{1}{2}(a - b)^2.$$

By symmetry, we may assume that $a \ge b$. Since $(|a| - |b|)^2 \le (a - b)^2$, without loss of generality, we may further assume that $a, b \ge 0$.

Define

$$h(a) := \frac{1}{2}a^2 \log a^2 + \frac{1}{2}b^2 \log b^2 - \frac{1}{2}(a^2 + b^2) \log \left(\frac{a^2 + b^2}{2}\right) - \frac{1}{2}(a - b)^2$$

for $a \in [b, \infty)$. h(b) = 0. It suffice to check that h'(b) = 0 and that h is concave on $[b, \infty)$. Note that

$$h'(a) = a \log a^2 + 1 - a \log \left(\frac{a^2 + b^2}{2}\right) - 1 - (a - b)$$
$$= a \log \frac{2a^2}{(a^2 + b^2)} - (a - b).$$

So h'(b) = 0. Moreover,

$$h''(a) = \log \frac{2a^2}{(a^2 + b^2)} + 1 - \frac{2a^2}{(a^2 + b^2)} \le 0$$

due to inequality $\log(x) + 1 \le x$.

• Remark (*Logarithmic Sobolev Inequality* \Rightarrow *Efron-Stein Inequality*). [Boucheron et al., 2013]

Note that for f non-negative,

$$Var(f(Z)) \le Ent(f^2(Z)).$$

Thus logarithmic Sobolev inequality (1) implies

$$Var(f(Z)) \le \mathcal{E}(f)$$

which is the Efron-Stein inequality.

• Corollary 1.2 (Logarithmic Sobolev Inequality for Asymmetric Bernoulli Random Variables). [Boucheron et al., 2013]

If $f: \{-1, +1\}^n \to \mathbb{R}$ be an arbitrary real-valued function and $Z = (Z_1, \dots, Z_n) \in \{-1, +1\}^n$ with $p = \mathbb{P}\{Z_i = +1\}$. Then

$$Ent(f^2) \le c(p) \mathbb{E}\left[\|\nabla f(Z)\|_2^2 \right] \tag{3}$$

where

$$c(p) = \frac{1}{1 - 2p} \log \frac{1 - p}{p}$$

Note that $\lim_{p\to 1/2} c(p) = 2$.

1.2 Gaussian Logarithmic Sobolev Inequality

• Proposition 1.3 (Gaussian Logarithmic Sobolev Inequality). [Boucheron et al., 2013] Let $f : \mathbb{R}^n \to \mathbb{R}$ be a continuous differentiable function and let $Z = (Z_1, \ldots, Z_n)$ be a vector of n independent standard Gaussian random variables. Then

$$Ent(f^{2}(Z)) \le 2\mathbb{E}\left[\left\|\nabla f(Z)\right\|_{2}^{2}\right]. \tag{4}$$

Proof: We first prove for n=1, where $f: \mathbb{R} \to \mathbb{R}$ is continuous differentiable and Z is standard Gaussian distribution. Without loss of generality, assume that $\mathbb{E}\left[f'(Z)\right] < \infty$ since it is trivial when $\mathbb{E}\left[f'(Z)\right] = \infty$. By density argument, it suffice to prove the proposition when f is twice differentiable with bounded support.

Now let $\epsilon_1, \ldots, \epsilon_n$ be independent Rademacher random variables and introduce

$$S_n := \frac{1}{\sqrt{n}} \sum_{j=1}^n \epsilon_j.$$

Note that $\epsilon_i \in \{-1, +1\}$ with equal probability, thus

$$\mathbb{E}_{(-i)}[S_n] = \frac{1}{2} \left[\left(\frac{1}{\sqrt{n}} \sum_{j \neq i} \epsilon_j + \frac{1}{\sqrt{n}} \right) + \left(\frac{1}{\sqrt{n}} \sum_{j \neq i} \epsilon_j - \frac{1}{\sqrt{n}} \right) \right]$$
$$= \frac{1}{2} \left[\left(S_n + \frac{1 - \epsilon_i}{\sqrt{n}} \right) + \left(S_n - \frac{1 + \epsilon_i}{\sqrt{n}} \right) \right].$$

In the proof of Gaussian Poincaré inequality, we show that by central limit theorem,

$$\limsup_{n \to \infty} \mathbb{E}\left[\sum_{i=1}^{n} \left| f(S_n) - f\left(S_n - \frac{2\epsilon_i}{\sqrt{n}}\right) \right|^2\right] = 4\mathbb{E}\left[(f'(Z))^2 \right].$$

On the other hands, for any continuous uniformly bounded function f, by central limit theorem,

$$\lim_{n \to \infty} \operatorname{Ent}\left(f^2(S_n)\right) = \operatorname{Ent}(f^2(Z))$$

The proof is then completed by invoking the logarithmic Sobolev inequality for Rademacher random variables

$$\operatorname{Ent}\left(f^{2}(S_{n})\right) \leq \frac{1}{2} \mathbb{E}\left[\sum_{i=1}^{n} \left| f\left(S_{n}\right) - f\left(S_{n} - \frac{2\epsilon_{i}}{\sqrt{n}}\right) \right|^{2}\right]$$

$$\Rightarrow \lim_{n \to \infty} \operatorname{Ent}\left(f^{2}(S_{n})\right) \leq \frac{1}{2} \lim_{n \to \infty} \mathbb{E}\left[\sum_{i=1}^{n} \left| f\left(S_{n}\right) - f\left(S_{n} - \frac{2\epsilon_{i}}{\sqrt{n}}\right) \right|^{2}\right]$$

$$\Rightarrow \operatorname{Ent}(f^{2}(Z)) \leq 2\mathbb{E}\left[\left(f'(Z)\right)^{2}\right].$$

The extension of the result to dimension $n \geq 1$ follows easily from the sub-additivity of entropy which states that

$$\operatorname{Ent}(f^2) \le \sum_{i=1}^n \mathbb{E}\left[\mathbb{E}_{(-i)}\left[f^2(Z)\log f^2(Z)\right] - \mathbb{E}_{(-i)}\left[f^2(Z)\right]\log \mathbb{E}_{(-i)}\left[f^2(Z)\right]\right]$$

where $\mathbb{E}_{(-i)}[\cdot]$ denotes the integration with respect to *i*-th variable Z_i only. Thus by induction, for all *i*

$$\mathbb{E}_{(-i)}\left[f^2(Z)\log f^2(Z)\right] - \mathbb{E}_{(-i)}\left[f^2(Z)\right]\log \mathbb{E}_{(-i)}\left[f^2(Z)\right] \leq 2\mathbb{E}_{(-i)}\left[(\partial_i f(Z))^2\right].$$

Thus

$$\operatorname{Ent}(f^2) \le 2\mathbb{E}\left[\mathbb{E}_{(-i)}\left[\sum_{i=1}^n (\partial_i f(Z))^2\right]\right] = 2\mathbb{E}\left[\|\nabla f(Z)\|_2^2\right].$$

• Remark (Dimension Free Property).

The Gaussian logarithmic Sobolev inequality has a constant C=2 that is **independent of** dimension n:

$$\mathbb{E}_{\mu}\left[f^{2}\right] \leq 2\mathbb{E}_{\mu}\left[\left\|\nabla f\right\|_{2}^{2}\right].$$

This dimension-free property is related to the concentration of Gaussian measure μ . As a consequence, this inequality can be extended to functions of Gaussian measure on infinite dimensional space, such as Gibbs measure, Gaussian process etc.

• Remark (Equivalent Form of Gaussian Logarithmic Sobolev Inequality) Assume $f: \mathbb{R}^n \to (0, \infty)$ and $\int_{\mathbb{R}^n} f d\mu = 1$ under Gaussian measure μ . Substituting $f \to \sqrt{f}$, the logarithmic Sobolev inequality becomes

$$\operatorname{Ent}_{\mu}(f) = \int f \log f d\mu \le \frac{1}{2} \int \frac{\|\nabla f\|_{2}^{2}}{f} d\mu \tag{5}$$

• Remark (Gaussian Logarithmic Sobolev Inequality \Rightarrow Gaussian Poincaré Inequality). [Boucheron et al., 2013]

Recall that the Gaussian Poincaré inequality

$$\operatorname{Var}(f(Z)) \leq \mathbb{E}\left[\|\nabla f(Z)\|_{2}^{2}\right]$$

Since

$$(1+t)\log(1+t) = t + \frac{t^2}{2} + o(t^2)$$

as $t \to 0$, we can get for Gaussian measures,

$$\operatorname{Ent}_{\mu}(1+\epsilon h) = \frac{\epsilon^2}{2} \operatorname{Var}_{\mu}(h) + o(\epsilon^2).$$

Similarly,

$$\int \frac{\|\nabla(1+\epsilon h)\|_{2}^{2}}{1+\epsilon h} d\mu = \epsilon^{2} \int \|\nabla h\|_{2}^{2} d\mu + o(\epsilon^{2}).$$

Thus from the Gaussian logarithmic Sobolev inequality,

$$\operatorname{Ent}_{\mu}(1+\epsilon h) \leq \frac{1}{2} \int \frac{\|\nabla(1+\epsilon h)\|_{2}^{2}}{1+\epsilon h} d\mu$$

$$\Leftrightarrow \frac{\epsilon^{2}}{2} \operatorname{Var}_{\mu}(h) + o(\epsilon^{2}) \leq \frac{\epsilon^{2}}{2} \int \|\nabla h\|_{2}^{2} d\mu + o(\epsilon^{2})$$

$$\Leftrightarrow \operatorname{Var}(f(Z)) \leq \mathbb{E}\left[\|\nabla f(Z)\|_{2}^{2}\right] \quad \text{as } \epsilon \to 0.$$

Thus the Gaussian logarithmic Sobolev inequality implies the Gaussian Poincaré inequality.

1.3 Information Theory Interpretation

• Remark (Information Interpretation of Gaussian Logarithmic Sobolev Inequality)

Let ν, μ be two probability measures on $(\mathcal{X}^n, \mathscr{F})$, $\mu = \mu_1 \otimes \ldots \otimes \mu_n$ and $\nu \ll \mu$. Define $f := \frac{d\nu}{d\mu}$ be the Radon-Nikodym derivative of ν with respect to μ (i.e f is the probability density function of ν with respect to μ). Then the entropy becomes **the relative entropy**

$$\operatorname{Ent}_{\mu}(f) := \mathbb{E}_{\mu} [f \log f] = \mathbb{KL} (\nu \parallel \mu)$$

since $\mathbb{E}_{\mu}[f] = \int_{\mathcal{X}^n} f d\mu = 1$.

On the other hand, the (relative) Fisher information is defined as

$$I(\nu \parallel \mu) := \mathbb{E}_{\nu} \left[\|\nabla \log f\|_{2}^{2} \right]$$

$$= \int \left\| \frac{\nabla f}{f} \right\|_{2}^{2} d\nu = \int \frac{\|\nabla f\|_{2}^{2}}{f^{2}} d\nu$$

$$= \int \frac{\|\nabla f\|_{2}^{2}}{f} d\mu$$

Thus the information interpretation of the Gaussian logarithmic Sobolev inequality is

$$\mathbb{KL}(\nu \parallel \mu) \le \frac{1}{2} I(\nu \parallel \mu) \tag{6}$$

where μ is a Gaussian measure and $\nu \ll \mu$ with density function f. Note that the Fisher information metric is the Riemannian metric induced by the relative entropy.

1.4 Logarithmic Sobolev Inequality for General Probability Measures

• From functional analysis, we have the Sobolev inequality,

Remark (The Sobolev Inequality) [Evans, 2010] The Sobolev inequality states for smooth function $f: \mathbb{R}^n \to \mathbb{R}$ in Sobolev space where $n \geq 3$ and $p = \frac{2n}{n-2} > 2$

$$||f||_p^2 \le C_n \int_{\mathbb{D}^n} |\nabla f|^2 \, dx.$$

The inequality is sharp when the constant

$$C_n := \frac{1}{\pi n(n-2)} \left(\frac{\Gamma(n)}{\Gamma(n/2)}\right)^{2/n}$$

• Proposition 1.4 (Euclidean Logarithmic Sobolev Inequality). Let $f : \mathbb{R}^n \to \mathbb{R}$ be a smooth function and m be Lebesgue measure on \mathbb{R}^n , then

$$Ent_{m}(f^{2}) \leq \frac{n}{2} \log \left(\frac{2}{n\pi e} \mathbb{E}_{m} \left[\|\nabla f\|_{2}^{2} \right] \right)$$

$$\Leftrightarrow \int f^{2} \log \left(\frac{f^{2}}{\int f^{2} dx} \right) dx \leq \frac{n}{2} \log \left(\frac{2}{n\pi e} \int |\nabla f|^{2} dx \right)$$

$$(7)$$

• Definition (Logarithmic Sobolev Inequality for General Probability Measure). A probability measure μ on \mathbb{R}^n is said to satisfy the <u>logarithmic Sobolev inequality</u> for some constant C > 0 if for any smooth function f

$$\operatorname{Ent}_{\mu}(f^{2}) \leq C \operatorname{\mathbb{E}}_{\mu} \left[\|\nabla f\|_{2}^{2} \right] \tag{8}$$

holds for any *continuous differentiable* function $f : \mathbb{R}^n \to \mathbb{R}$. The left-hand side is called *the entropy functional*, which is defined as

$$\operatorname{Ent}(f^2) := \mathbb{E}_{\mu} \left[f^2 \log f^2 \right] - \mathbb{E}_{\mu} \left[f^2 \right] \log \mathbb{E}_{\mu} \left[f^2 \right]$$
$$= \int f^2 \log \left(\frac{f^2}{\int f^2 d\mu} \right) d\mu.$$

The right-hand side is defined as

$$\mathbb{E}_{\mu} \left[\|\nabla f\|_2^2 \right] = \int \|\nabla f\|_2^2 d\mu.$$

Thus we can rewrite the logarithmic Sobolev inequality in functional form

$$\int f^2 \log \left(\frac{f^2}{\int f^2 d\mu} \right) d\mu \le C \int \|\nabla f\|_2^2 d\mu \tag{9}$$

• Remark (Logarithmic Sobolev Inequality)
For non-negative function f, we can replace $f \to \sqrt{f}$, so that the logarithmic Sobolev inequality becomes

$$\operatorname{Ent}_{\mu}(f) \le C \int \frac{\|\nabla f\|_{2}^{2}}{f} d\mu \tag{10}$$

• Remark (Modified Logarithmic Sobolev Inequality via Convex Cost and Duality) For some convex non-negative cost $c: \mathbb{R}^n \to \mathbb{R}_+$, the convex conjugate of c (Legendre transform of c) is defined as

$$c^*(x) := \sup_{y} \left\{ \langle x, y \rangle - c(y) \right\}$$

Then we can obtain the modified logarithmic Sobolev inequality

$$\operatorname{Ent}_{\mu}(f) \le \int f^2 \, c^* \left(\frac{\nabla f}{f}\right) d\mu \tag{11}$$

2 The Entropy Methods

2.1 Herbst's Argument

• Remark (Entropy Functional for Moment Generating Function) Let $X = e^{\lambda Z}$ where Z is a random variable. The entropy function of X becomes

$$\operatorname{Ent}(e^{\lambda Z}) = \mathbb{E}\left[\lambda Z e^{\lambda Z}\right] - \mathbb{E}\left[e^{\lambda Z}\right] \log\left(\mathbb{E}\left[e^{\lambda Z}\right]\right)$$
Denote $\psi_{Z-\mathbb{E}[Z]}(\lambda) := \log \mathbb{E}\left[e^{\lambda(Z-\mathbb{E}[Z])}\right]$. Then
$$\psi'_{Z-\mathbb{E}[Z]}(\lambda) = \frac{d}{d\lambda} \log \mathbb{E}\left[e^{\lambda(Z-\mathbb{E}[Z])}\right]$$

$$= \frac{1}{\mathbb{E}\left[e^{\lambda(Z-\mathbb{E}[Z])}\right]} \mathbb{E}\left[\left(Z-\mathbb{E}\left[Z\right]\right) e^{\lambda(Z-\mathbb{E}[Z])}\right]$$

$$= \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} e^{\lambda \mathbb{E}[Z]} \mathbb{E}\left[\left(Z-\mathbb{E}\left[Z\right]\right) e^{\lambda(Z-\mathbb{E}[Z])}\right]$$

$$= \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} \mathbb{E}\left[\left(Z-\mathbb{E}\left[Z\right]\right) e^{\lambda Z}\right]$$

$$\lambda \; \psi_{Z - \mathbb{E}[Z]}'(\lambda) = \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} \left(\mathbb{E}\left[\lambda Z e^{\lambda Z}\right] - \mathbb{E}\left[\lambda Z\right] \mathbb{E}\left[e^{\lambda Z}\right] \right)$$

$$\begin{split} \Rightarrow \lambda \; \psi_{Z - \mathbb{E}[Z]}'(\lambda) - \psi_{Z - \mathbb{E}[Z]}(\lambda) &= \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} \left\{ \mathbb{E}\left[\lambda Z e^{\lambda Z}\right] - \mathbb{E}\left[\lambda Z\right] \mathbb{E}\left[e^{\lambda Z}\right] - \mathbb{E}\left[e^{\lambda Z}\right] \log \mathbb{E}\left[e^{\lambda (Z - \mathbb{E}[Z])}\right] \right\} \\ &= \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} \left\{ \mathbb{E}\left[\lambda Z e^{\lambda Z}\right] - \mathbb{E}\left[\lambda Z\right] \mathbb{E}\left[e^{\lambda Z}\right] \\ &+ \mathbb{E}\left[e^{\lambda Z}\right] \mathbb{E}\left[\lambda Z\right] - \mathbb{E}\left[e^{\lambda Z}\right] \log \mathbb{E}\left[e^{\lambda Z}\right] \right\} \\ &= \frac{1}{\mathbb{E}\left[e^{\lambda Z}\right]} \left\{ \mathbb{E}\left[\lambda Z e^{\lambda Z}\right] - \mathbb{E}\left[e^{\lambda Z}\right] \log \mathbb{E}\left[e^{\lambda Z}\right] \right\} \\ &= \frac{\mathrm{Ent}(e^{\lambda Z})}{\mathbb{E}\left[e^{\lambda Z}\right]} \end{split}$$

Thus we have

$$\frac{\operatorname{Ent}(e^{\lambda Z})}{\mathbb{E}\left[e^{\lambda Z}\right]} = \lambda \ \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \psi_{Z-\mathbb{E}[Z]}(\lambda). \tag{12}$$

Our strategy is based on using (12) the sub-additivity of entropy and then univariate calculus to derive upper bounds for the derivative of $\psi(\lambda)$. By solving the obtained differential inequality, we obtain tail bounds via Chernoff's bounding.

• Proposition 2.1 (Herbst's Argument) [Boucheron et al., 2013, Wainwright, 2019] Let Z be an integrable random variable such that for some $\nu > 0$, we have, for every $\lambda > 0$,

$$\frac{Ent(e^{\lambda Z})}{\mathbb{E}\left[e^{\lambda Z}\right]} \le \frac{\nu\lambda^2}{2} \tag{13}$$

Then, for every $\lambda > 0$, the logarithmic moment generating function of centered random variable $(Z - \mathbb{E}[Z])$ satisfies

$$\psi_{Z-\mathbb{E}[Z]}(\lambda) := \log \mathbb{E}\left[e^{\lambda(Z-\mathbb{E}[Z])}\right] \leq \frac{\nu\lambda^2}{2}.$$

Proof: The condition of the proposition means, via (12), that

$$\lambda \ \psi'_{Z-\mathbb{E}[Z]}(\lambda) - \psi_{Z-\mathbb{E}[Z]}(\lambda) \le \frac{\nu \lambda^2}{2},$$

or equivalently,

$$\frac{1}{\lambda}\psi'_{Z-\mathbb{E}[Z]}(\lambda) - \frac{1}{\lambda^2}\psi_{Z-\mathbb{E}[Z]}(\lambda) \le \frac{\nu}{2}.$$

Setting $G(\lambda) = \lambda^{-1} \psi_{Z-\mathbb{E}[Z]}(\lambda)$, we see that the differential inequality becomes

$$G'(\lambda) \le \frac{\nu}{2}.$$

Since $G(\lambda) \to 0$ as $\lambda \to 0$, which implies that

$$G(\lambda) \le \frac{\nu\lambda}{2}$$
,

and the result follows.

• Remark (Entropy Methods)

The **key** strategy of **entropy** methods to prove the concentration of function f(Z) of independent variables Z is as follows

- 1. First we **bound the entropy** for each individual variables Z_i conditioning on the rest of them, i.e. $Z_{(-i)} := (Z_1, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_n)$ or $\widetilde{Z}_i := (Z_1, \ldots, Z_{i-1}, Z_i', Z_{i+1}, \ldots, Z_n)$. Typically, we use **the exponential form** $e^{\lambda f(Z)}$. The followings are methods that can be used in this step
 - If Z is Gaussian or Bernoulli distribution, we can use the logarithmic Sobolev inequality

$$\operatorname{Ent}_{\mu_i}(e^{\lambda f(Z)}) \le 2\mathbb{E}_{\mu_i} \left[\left\| \nabla e^{\lambda f(Z)/2} \right\|_2^2 \right] = \frac{\lambda^2}{2} \mathbb{E}_{\mu_i} \left[e^{\lambda f(Z)} \left\| \nabla f(Z) \right\|_2^2 \right].$$

- In general, we can also use the variational formulation of entropy

$$\operatorname{Ent}_{\mu_i}(e^{\lambda f(Z)}) \le \mathbb{E}_{\mu_i} \left[e^{\lambda f(Z)} (\lambda(f(Z) - f(\widetilde{Z}_i)) - (e^{\lambda f(Z)} - e^{\lambda f(\widetilde{Z}_i)}) \right]$$

- If we can control the Jackknife estimate of variance of f(Z)

$$\sum_{i=1}^{n} \left(f(Z) - f(\widetilde{Z}_i) \right)^2 \le \nu$$

then we can bound the above variational formulation since $e^s - e^t \le e^t(s-t)$ where s > t.

2. Apply the tensorization of property of the entropy functional on f(Z)

$$\operatorname{Ent}_{\mu_1 \otimes ... \otimes \mu_n}(e^{\lambda f(Z)}) \leq \mathbb{E}_{\mu_1 \otimes ... \otimes \mu_n} \left[\sum_{i=1}^n \operatorname{Ent}_{\mu_i}(e^{\lambda f(Z)}) \right]$$

where $\mu := \mu_1 \otimes \ldots \otimes \mu_n$ is the distribution of Z and the subscript μ_i indicates that the integration concerns Z_i only.

3. By the Herbst's argument, we obtain a differential inequality for the logarithmic moment generating function $\psi(\lambda)$

$$\frac{\operatorname{Ent}_{\mu_1 \otimes ... \otimes \mu_n}(e^{\lambda f(Z)})}{\operatorname{\mathbb{E}}_{\mu_1 \otimes ... \otimes \mu_n}\left[e^{\lambda f(Z)}\right]} = \lambda \psi'(\lambda) - \psi(\lambda) \leq \phi(\lambda) \Leftrightarrow \left(\frac{\psi(\lambda)}{\lambda}\right)' \leq \frac{\phi(\lambda)}{\lambda}$$
$$\frac{\psi(\lambda)}{\lambda} \leq \lim_{\lambda \to 0} \left(\frac{\psi(\lambda)}{\lambda}\right) + \int_0^{\lambda} \left(\frac{\phi(u)}{u}\right) du = \operatorname{\mathbb{E}}\left[f(Z)\right] + \int_0^{\lambda} \frac{\phi(u)}{u} du$$

4. Obtain concentration results based on the Cramér-Chernoff Method

$$\mathbb{P}\left\{f(Z) - \mathbb{E}\left[f(Z)\right] > t\right\} \le \inf_{\lambda > 0} \exp\left(\psi(\lambda) - \lambda t\right) \le \inf_{\lambda > 0} \exp\left(\lambda \int_0^\lambda \frac{\phi(u)}{u} du - \lambda t\right)$$

2.2 Modified Logarithmic Sobolev Inequalities

• Proposition 2.2 (A Modified Logarithmic Sobolev Inequalities for Moment Generating Function) [Boucheron et al., 2013]

Consider independent random variables Z_1, \ldots, Z_n taking values in \mathcal{X} , a real-valued function $f: \mathcal{X}^n \to \mathbb{R}$ and the random variable $X = f(Z_1, \ldots, Z_n)$. Also denote $Z_{(-i)} = (Z_1, \ldots, Z_{i-1}, Z_{i+1}, \ldots, Z_n)$ and $X_{(-i)} = f_i(Z_{(-i)})$ where $f_i: \mathcal{X}^{n-1} \to \mathbb{R}$ is an arbitrary function. Let $\phi(x) = e^x - x - 1$. Then for all $\lambda \in \mathbb{R}$,

$$\lambda \mathbb{E}\left[Xe^{\lambda X}\right] - \mathbb{E}\left[e^{\lambda X}\right] \log \mathbb{E}\left[e^{\lambda X}\right] \le \sum_{i=1}^{n} \mathbb{E}\left[e^{\lambda X}\phi(-\lambda(X-X_{(-i)}))\right]$$
(14)

Proof: Recall the tensorization of entropy

$$\operatorname{Ent}_{\mu_1 \otimes ... \otimes \mu_n}(Y) \leq \mathbb{E}_{\mu_1 \otimes ... \otimes \mu_n} \left[\sum_{i=1}^n \operatorname{Ent}_{\mu_i}(Y) \right].$$

We bound each term on the right-hand side by the variational formulation of entropy

$$\operatorname{Ent}_{\mu_i}(Y) \le \mathbb{E}_{\mu_i} \left[Y(\log Y - \log u) - (Y - u) \right]$$

for any u > 0. Let $u = Y_{(-i)} = g_i(Z_{(-i)})$. We have

$$\operatorname{Ent}_{\mu_i}(Y) \leq \mathbb{E}_{\mu_i} \left[Y(\log Y - \log Y_{(-i)}) - (Y - Y_{(-i)}) \right].$$

Applying above inequality to the variable $Y = e^{\lambda X}$ and $Y_{(-i)} = e^{\lambda X_{(-i)}}$, one obtain

$$\operatorname{Ent}_{\mu_{i}}(e^{\lambda X}) \leq \operatorname{\mathbb{E}}_{\mu_{i}} \left[e^{\lambda X} (\log e^{\lambda X} - \log e^{\lambda X_{(-i)}}) - (e^{\lambda X} - e^{\lambda X_{(-i)}}) \right]$$

$$= \operatorname{\mathbb{E}}_{\mu_{i}} \left[e^{\lambda X} (\lambda (X - X_{(-i)}) - (e^{\lambda X} - e^{\lambda X_{(-i)}})) \right]$$

$$= \operatorname{\mathbb{E}}_{\mu_{i}} \left[e^{\lambda X} \left(\lambda (X - X_{(-i)}) - e^{-\lambda X} (e^{\lambda X} - e^{\lambda X_{(-i)}}) \right) \right]$$

$$= \operatorname{\mathbb{E}}_{\mu_{i}} \left[e^{\lambda X} \left(\lambda (X - X_{(-i)}) + e^{-\lambda (X - X_{(-i)})} - 1 \right) \right]$$

$$= \operatorname{\mathbb{E}}_{\mu_{i}} \left[e^{\lambda X} \phi \left(-\lambda (X - X_{(-i)}) \right) \right]$$

where $\phi(x) = e^x - x - 1$. Thus the proof is completed.

• Proposition 2.3 (Symmetrized Modified Logarithmic Sobolev Inequalities) [Boucheron et al., 2013]

Consider independent random variables Z_1, \ldots, Z_n taking values in \mathcal{X} , a real-valued function $f: \mathcal{X}^n \to \mathbb{R}$ and the random variable $X = f(Z_1, \ldots, Z_n)$. Also denote $\widetilde{X}^{(i)} = f(Z_1, \ldots, Z_{i-1}, Z'_i, Z_{i+1}, \ldots, Z_n)$. Let $\phi(x) = e^x - x - 1$. Then for all $\lambda \in \mathbb{R}$,

$$\lambda \mathbb{E}\left[Xe^{\lambda X}\right] - \mathbb{E}\left[e^{\lambda X}\right] \log \mathbb{E}\left[e^{\lambda X}\right] \le \sum_{i=1}^{n} \mathbb{E}\left[e^{\lambda X}\phi(-\lambda(X-\widetilde{X}^{(i)}))\right]$$
(15)

Moreover, denoting $\tau(x) = x(e^x - 1)$, for all $\lambda \in \mathbb{R}$,

$$\lambda \mathbb{E}\left[Xe^{\lambda X}\right] - \mathbb{E}\left[e^{\lambda X}\right] \log \mathbb{E}\left[e^{\lambda X}\right] \leq \sum_{i=1}^{n} \mathbb{E}\left[e^{\lambda X}\tau(-\lambda(X-\widetilde{X}^{(i)})_{+})\right],$$
$$\lambda \mathbb{E}\left[Xe^{\lambda X}\right] - \mathbb{E}\left[e^{\lambda X}\right] \log \mathbb{E}\left[e^{\lambda X}\right] \leq \sum_{i=1}^{n} \mathbb{E}\left[e^{\lambda X}\tau(\lambda(\widetilde{X}^{(i)}-X)_{-})\right].$$

Proof: Note that $X_{(-i)}$ and $\widetilde{X}^{(i)}$ are both independent from Z_i . The first inequality is the same as the proposition above. For the second inequality, use the fact that

$$\mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(\lambda (\widetilde{X}^{(i)} - X)_+ \right) \right] = \mathbb{E}_{\mu_i} \left[e^{\lambda \widetilde{X}^{(i)}} \phi \left(\lambda (X - \widetilde{X}^{(i)})_+ \right) \right]$$
$$= \mathbb{E}_{\mu_i} \left[e^{\lambda X} e^{-\lambda (X - \widetilde{X}^{(i)})} \phi \left(\lambda (X - \widetilde{X}^{(i)})_+ \right) \right].$$

and

$$\begin{split} \mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(-\lambda (\widetilde{X}^{(i)} - X) \right) \right] &= \mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(-\lambda (\widetilde{X}^{(i)} - X)_+ \right) \right] + \mathbb{E}_{\mu_i} \left[e^{\lambda X} \phi \left(\lambda (\widetilde{X}^{(i)} - X)_+ \right) \right] \\ &= \mathbb{E}_{\mu_i} \left[e^{\lambda X} \left\{ \phi \left(-\lambda (\widetilde{X}^{(i)} - X) \right) + e^{-\lambda (X - \widetilde{X}^{(i)})} \phi \left(\lambda (X - \widetilde{X}^{(i)})_+ \right) \right\} \right]. \end{split}$$

Finally note that $\phi(x) + e^x \phi(-x) = \tau(x) = x(e^x - 1)$.

2.3 Bounded Difference Inequality via Entropy Methods

• Remark Given vectors $x, x' \in \mathcal{X}^n$ and an index $k \in \{1, 2, ..., n\}$, we define a new vector $x^{(-k)} \in \mathcal{X}^n$ via

$$x_j^{(-k)} = \begin{cases} x_j & j \neq k \\ x_k' & j = k \end{cases}$$

With this notation, we say that $f: \mathcal{X}^n \to \mathbb{R}$ satisfies <u>the bounded difference inequality</u> with parameters (L_1, \ldots, L_n) if, for each index $k = 1, \overline{2, \ldots, n}$,

$$\left| f(x) - f(x^{(-k)}) \right| \le L_k, \quad \text{for all } x, x' \in \mathcal{X}^n.$$
 (16)

• Proposition 2.4 (McDiarmid's Inequality / Bounded Differences Inequality)[Boucheron et al., 2013, Wainwright, 2019]

Suppose that f satisfies **the bounded difference property** (16) with parameters (L_1, \ldots, L_n) and that the random vector $X = (X_1, X_2, \ldots, X_n)$ has **independent** components. Denote

$$\nu = \frac{1}{4} \sum_{k=1}^{n} L_k^2$$

Then

$$\mathbb{P}\left\{f(X) - \mathbb{E}\left[f(X)\right] \ge t\right\} \le \exp\left(-\frac{t^2}{2\nu}\right).$$

Proof: Recall that for a random variable Y taking its values in [a, b], then we know from Hoeffding's Lemma that the logarithmic moment generating functions $\psi(\lambda)$ satisfies

$$\psi(\lambda)'' = \operatorname{Var}(Y) \le \frac{(b-a)^2}{4}$$

for every $\lambda \in \mathbb{R}$. Hence, Hoeffding's inequality is obtained since

$$\frac{\operatorname{Ent}(e^{\lambda Y})}{\mathbb{E}\left[e^{\lambda Y}\right]} = \lambda \psi'(\lambda) - \psi(\lambda) = \int_0^\lambda s \psi''(s) ds \le \frac{(b-a)^2}{4} \int_0^\lambda s ds = \frac{(b-a)^2 \lambda^2}{8},$$

Note that by the bounded differences assumption, given $X_{(-i)}$, f(X) is a random variable whose range is in an interval of length at most L_i , so

$$\frac{\operatorname{Ent}_{(-i)}(e^{\lambda f(X)})}{\mathbb{E}_{(-i)}\left[e^{\lambda f(X)}\right]} \le \frac{L_i^2 \lambda^2}{8}$$

From the tensorization property of entropy, we can bound the entropy of total function

$$\operatorname{Ent}(e^{\lambda f(X)}) \leq \mathbb{E}\left[\sum_{i=1}^{n} \operatorname{Ent}_{(-i)}(e^{\lambda f(X)})\right] \leq \sum_{i=1}^{n} \frac{L_{i}^{2} \lambda^{2}}{8} \mathbb{E}\left[\mathbb{E}_{(-i)}\left[e^{\lambda f(X)}\right]\right]$$
$$\frac{\operatorname{Ent}(e^{\lambda f(X)})}{\mathbb{E}\left[e^{\lambda f(X)}\right]} \leq \frac{\sum_{i=1}^{n} L_{i}^{2} \lambda^{2}}{8} = \frac{\nu \lambda^{2}}{2}.$$

Using Herbst's argument, it leads to the bound of logarithmic moment generating function:

$$\psi_{f(X)}(\lambda) \le \frac{\nu\lambda^2}{2}.$$

Finally, we apply the Chernoff's inequality

$$\mathbb{P}\left\{f(X) - \mathbb{E}\left[f(X)\right] \ge t\right\} \le \inf_{\lambda > 0} \exp\left(\psi_{f(X)}(\lambda) - \lambda t\right) \le \exp\left(-\frac{t^2}{2\nu}\right).$$

• Remark (Relaxation of Bounded Difference Conditions)

Consider X := f(Z) where $Z := (Z_1, \ldots, Z_n)$ and let

$$\widetilde{X}_i = \inf_{z_i'} f(Z_1, \dots, Z_{i-1}, z_i', Z_{i+1}, \dots, Z_n)$$

We can relax the condition defining the bounded difference function by the following boundedness condition on variance

$$\sum_{i=1}^{n} \left(X - \widetilde{X}_{i} \right)^{2} \leq \nu$$

$$\sum_{i=1}^{n} \left(f(Z_{1}, \dots, Z_{n}) - \inf_{z'_{i}} f(Z_{1}, \dots, Z_{i-1}, z'_{i}, Z_{i+1}, \dots, Z_{n}) \right)^{2} \leq \nu$$
(17)

The quantity ν may be interpreted as an *upper bound* for the *Efron-Stein estimate* of the *variance* Var(f(Z)).

• Proposition 2.5 (Concentration with Bounded Variance)[Boucheron et al., 2013] Assume that X is such that there exists a constant $\nu > 0$ such that, almost surely,

$$\sum_{i=1}^{n} \left(X - \widetilde{X}_i \right)^2 \le \nu.$$

Then for all t > 0,

$$\mathbb{P}\left\{X - \mathbb{E}\left[X\right] \ge t\right\} \le \exp\left(-\frac{t^2}{2\nu}\right). \tag{18}$$

Proof: The result follows easily from the modified logarithmic Sobolev inequality proved in the previous section. Observe that for x > 0,

$$\phi(-x) := e^{-x} + x - 1 \le \frac{x^2}{2},$$

and therefore, for all $\lambda > 0$,

$$\operatorname{Ent}(e^{\lambda X}) = \lambda \mathbb{E}\left[Xe^{\lambda X}\right] - \mathbb{E}\left[e^{\lambda X}\right] \log \mathbb{E}\left[e^{\lambda X}\right] \leq \sum_{i=1}^{n} \mathbb{E}\left[e^{\lambda X}\phi(-\lambda(X-\widetilde{X}_{i}))\right]$$

$$\leq \mathbb{E}\left[e^{\lambda X}\sum_{i=1}^{n}\frac{\lambda^{2}}{2}\left(X-\widetilde{X}_{i}\right)^{2}\right]$$

$$\leq \frac{\nu\lambda^{2}}{2}\mathbb{E}\left[e^{\lambda X}\right]$$

where we used the assumption of the theorem. The obtained inequality has the same form as the one we already faced in the proof of bounded difference inequality and the proof may be finished in an identical way.

2.4 Poisson Logarithmic Sobolev Inequality

• Proposition 2.6 (Modified Logarithmic Sobolev Inequality for Bernoulli Random Variable). [Boucheron et al., 2013]

Let $f: \{0,1\} \to (0,\infty)$ be a **non-negative** real-valued function defined on the binary set $\{0,1\}$. Define **the discrete derivative** of f at $x \in \{0,1\}$ by

$$\nabla f := f(1-x) - f(x).$$

Let X be a Bernoulli random variable with parameter $p \in (0,1)$ (i.e. $\mathbb{P}\{X=1\}=p$). Then

$$Ent(f(X)) \le (p(1-p))\mathbb{E}\left[\nabla f(X)\nabla \log f(X)\right]. \tag{19}$$

and

$$Ent(f(X)) \le (p(1-p))\mathbb{E}\left[\frac{|\nabla f(X)|^2}{f(X)}\right].$$
 (20)

Proposition 2.7 (Poisson Logarithmic Sobolev Inequality). [Boucheron et al., 2013]
 Let f: N→ (0,∞) be a non-negative real-valued function defined on the set of non-negative integers N. Define the discrete derivative of f at x ∈ N by

$$\nabla f := f(x+1) - f(x).$$

Let X be a Poisson random variable. Then

$$Ent(f(X)) \le (\mathbb{E}[X])\mathbb{E}[\nabla f(X)\nabla \log f(X)].$$
 (21)

and

$$Ent(f(X)) \le (\mathbb{E}[X])\mathbb{E}\left[\frac{|\nabla f(X)|^2}{f(X)}\right].$$
 (22)

3 Applications

3.1 Lipschitz Functions of Gaussian Variables

- Theorem 3.1 (Rademacher Theorem). If $f: U \to \mathbb{R}$ is a L-Lipschitz function where $U \subseteq \mathbb{R}^n$, then f is differentiable almost everywhere in U and the essential supremum of the norm of its derivative is bounded by its Lipschitz constant.
- Theorem 3.2 (Lipschitz Functions of Gaussian Variables) [Boucheron et al., 2013] Let $Z = (Z_1, ..., Z_n)$ be a vector of n independent standard normal random variables. Let $f: \mathbb{R}^n \to \mathbb{R}$ denote an L-Lipschitz function, that is, there exists a constant L > 0 such that for all $x, y \in \mathbb{R}^n$,

$$|f(x) - f(y)| \le L ||x - y||.$$

Then, for all $\lambda \in \mathbb{R}$,

$$\psi_{f(Z)-\mathbb{E}[f(Z)]}(\lambda) := \log \mathbb{E}\left[e^{\lambda(f(Z)-\mathbb{E}[f(Z)])}\right] \le \frac{L^2 \lambda^2}{2}$$
(23)

Proof: By a standard density argument we may assume that f is differentiable with gradient uniformly bounded by L according to Rademacher theorem. We may also assume, without loss of generality, that $\mathbb{E}[f(Z)] = 0$. Using the Gaussian logarithmic Sobolev inequality for the function $e^{\lambda f/2}$, we obtain

$$\operatorname{Ent}(e^{\lambda f}) \leq 2\mathbb{E}\left[\left\|\nabla e^{\lambda f/2}\right\|_{2}^{2}\right]$$

$$= 2\mathbb{E}\left[\left\|\frac{\lambda}{2}e^{\lambda f/2}\nabla f\right\|_{2}^{2}\right]$$

$$= \frac{\lambda^{2}}{2}\mathbb{E}\left[e^{\lambda f}\left\|\nabla f\right\|_{2}^{2}\right]$$

$$\leq \frac{\lambda^{2}}{2}\mathbb{E}\left[e^{\lambda f}\right]L^{2}$$

$$\Rightarrow \frac{\operatorname{Ent}(e^{\lambda f})}{\mathbb{E}\left[e^{\lambda f}\right]} \leq \frac{L^{2}\lambda^{2}}{2}$$

By Herbst's argument, this implies that $\lambda \psi'(\lambda) - \psi(\lambda) \leq \frac{L^2 \lambda^2}{2}$. Let $G(\lambda) := \lambda^{-1} \psi(\lambda)$, we have $G'(\lambda) \leq L^2/2$, so we have

$$\psi_{f(Z)}(\lambda) \le \frac{L^2 \lambda^2}{2} \qquad \blacksquare$$

Theorem 3.3 (Gaussian Concentration Inequality / The Tsirelson-Ibragimov-Sudakov Inequality) [Boucheron et al., 2013, Wainwright, 2019]
Let Z = (Z₁,...,Z_n) be a vector of n independent standard normal random variables.
Let f: ℝⁿ → ℝ denote an L-Lipschitz function. Then, for all t > 0,

$$\mathbb{P}\left\{f(Z) - \mathbb{E}\left[f(Z)\right] \ge t\right\} \le \exp\left(-\frac{t^2}{2L^2}\right). \tag{24}$$

Proof: The previous theorem states that f(Z) is sub-Gaussian distributed with parameter $\nu = L^2$. By Markov's inequality,

$$\begin{split} \mathbb{P}\left\{f(Z) - \mathbb{E}\left[f(Z)\right] \geq t\right\} &= \mathbb{P}\left\{e^{\lambda(f(Z) - \mathbb{E}\left[f(Z)\right]\right)} \geq e^{\lambda t}\right\} \leq \frac{\mathbb{E}\left[e^{\lambda f(Z) - \mathbb{E}\left[f(Z)\right]}\right]}{e^{\lambda t}} \\ &\leq \inf_{\lambda > 0} \exp\left(\psi(\lambda) - \lambda t\right) \\ &\leq \inf_{\lambda > 0} \exp\left(\frac{L^2 \lambda^2}{2} - \lambda t\right) \\ &= \exp\left(-\frac{t^2}{2L^2}\right) \end{split}$$

where $\lambda^* = \frac{t}{L^2}$.

• Remark (*Dimension-Free Concentration*)

An important feature of the theorem is that the right-hand side does not depend on the dimension n. This inequality has served as a benchmark for the development of concentration inequalities during the last three decades.

3.2 Suprema of Gaussian Process

• Definition (Gaussian Process)

Let T be a metric space. A stochastic process $(X_t)_{t\in T}$ is a Gaussian process indexed by T if for any finite collection $\{t_1,\ldots,t_n\}\subset T$, the vector (X_{t_1},\ldots,X_{t_n}) has a jointly Gaussian distribution.

In addition, we assume that T is **totally bounded** (i.e. for every t > 0 it can be covered by finitely many balls of radius t) and that the Gaussian process is **almost surely continuous**, that is, with probability 1, X_t is a continuous function of t.

• Theorem 3.4 (Concentration of Suprema of Gaussian Process) [Boucheron et al., 2013, Vershynin, 2018, Wainwright, 2019, Giné and Nickl, 2021] Let (X_t)_{t∈T} be an almost surely continuous centered Gaussian process indexed by a totally bounded set T. If

$$\sigma^2 := \sup_{t \in T} \mathbb{E}\left[X_t^2\right],$$

then $Z = \sup_{t \in T} X_t$ satisfies $Var(Z) \leq \sigma^2$, and for all u > 0,

$$\mathbb{P}\left\{Z - \mathbb{E}\left[Z\right] \ge u\right\} \le \exp\left(-\frac{u^2}{2\sigma^2}\right) \tag{25}$$

and

$$\mathbb{P}\left\{\mathbb{E}\left[Z\right] - Z \ge u\right\} \le \exp\left(-\frac{u^2}{2\sigma^2}\right) \tag{26}$$

Proof: We assume that T is a **finite set**. The extension to arbitrary totally bounded T is based on a **separability argument** and **monotone convergence**. We may assume, for simplicity, that $T = \{1, ..., n\}$. Let Σ be the covariance matrix of the centered Gaussian vector $X = (X_1, ..., X_n)$. Denote by A the square root of the positive semidefinite matrix Γ . If $Y = (Y_1, ..., Y_n)$ is a vector of independent standard normal random variables, then

$$f(Y) = \max_{i=1,\dots,n} (AY)_i$$

has the same distribution as $Z = \max_{i=1,...,n} X_i$. Hence, we can apply the Gaussian concentration inequality by bounding the Lipschitz constant of f. By the Cauchy-Schwarz inequality, for all $u, v \in \mathbb{R}^n$ and i = 1, ..., n,

$$|(Au)_i - (Av)_i| \le ||A_{i:}||_2 ||u - v||_2$$

Since $||A_{i:}||_2^2 = \sum_j A_{i,j}^2 = \text{Var}(X_i)$, we get

$$|f(u) - f(v)| \le \max_{i=1,\dots,n} |(Au)_i - (Av)_i| \le \sigma ||u - v||_2$$

Therefore, f is Lipschitz with constant σ and the tail bounds follow from the Gaussian concentration inequality. The variance bound follows from the Gaussian Poincaré inequality.

3.3 Concentration of Convex Lipschitz Functions

• Theorem 3.5 (Concentration of Convex Lipschitz Functions on Unit Hypercube)
[Boucheron et al., 2013]

Let $Z := (Z_1, ..., Z_n)$ be independent random variables taking values in the interval [0,1] and let $f : [0,1]^n \to \mathbb{R}$ be a **separately convex function** (i.e. f is convex in each coordinate while the others are fixed) such that

$$|f(x) - f(y)| \le ||x - y||$$
 for all $x, y \in [0, 1]^n$.

Then $X = f(Z_1, ..., Z_n)$ satisfies, for all t > 0,

$$\mathbb{P}\left\{f(Z) - \mathbb{E}\left[f(Z)\right] \ge t\right\} \le \exp\left(-\frac{t^2}{2}\right). \tag{27}$$

Proof: We may assume without loss of generality that the partial derivatives of f exist. (Otherwise one may *approximate* f by a smooth function via a standard argument.) It suffices to bound the random variable

$$\sum_{i=1}^{n} \left(X - \widetilde{X}_i \right)^2$$

where $\widetilde{X}_i = \inf_{z'_i} f(Z_1, \dots, Z_{i-1}, z'_i, Z_{i+1}, \dots, Z_n)$. Recall that for separately convex differentiable function f on compact domain [0, 1], we have

$$\sum_{i=1}^{n} \left(X - \widetilde{X}_{i} \right)^{2} = \sum_{i=1}^{n} \left(f(Z) - f(\widetilde{Z}_{i}) \right)^{2} \qquad (by \ continuity \ of \ f \ on \ compact \ domain)$$

$$\leq \sum_{i=1}^{n} \left(\frac{\partial f}{\partial z_{i}}(Z) \right)^{2} \left(Z - \widetilde{Z}_{i} \right)^{2} \qquad (by \ separetely \ convexity)$$

$$\leq \sum_{i=1}^{n} \left(\frac{\partial f}{\partial z_{i}}(Z) \right)^{2} \qquad (by \ domain \ [0, 1] \)$$

$$= \|\nabla f(Z)\|_{2}^{2}$$

$$\leq 1 \qquad (by \ Lipschitz \ property)$$

Therefore, we can apply the Proposition 2.5 with variance bounded above by $\nu = 1$.

- 3.4 Exponential Tail Bounds for Self-Bounding Functions
- 3.5 Concentration of Lipschitz Functions on the Hypercube
- 3.6 Hypercontractivity for Boolean Polynomials
- 3.7 Gaussian Hypercontractivity

References

- Stéphane Boucheron, Gábor Lugosi, and Pascal Massart. Concentration inequalities: A nonasymptotic theory of independence. Oxford university press, 2013.
- Lawrence C Evans. Partial differential equations, volume 19. American Mathematical Soc., 2010.
- Evarist Giné and Richard Nickl. Mathematical foundations of infinite-dimensional statistical models. Cambridge university press, 2021.
- Roman Vershynin. *High-dimensional probability: An introduction with applications in data science*, volume 47. Cambridge university press, 2018.
- Martin J Wainwright. *High-dimensional statistics: A non-asymptotic viewpoint*, volume 48. Cambridge University Press, 2019.