## Fall 2016, EECE 5644, Homework #2

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For your convenience, these problems from the textbook are provided in the following pages. Some sub-problems have been removed. Use the provided problem statements.

- 2.1 (10 pts) Problem 2.12 from Duda's book.
- 2.2 (10 pts) Problem 3.1 from Duda's book. Skip (c) (removed from the problem statements below)
- 2.3 (10 pts) Problem 3.2 from Duda's book. Skip (b) (removed from the problem statements below)
- 2.4 (20 pts) Problem 3.3 from Duda's book.
- **2.5** (10 pts) Problem 3.4 from Duda's book.
- 2.6 (40 pts) Problem 3.17 from Duda's book.

- 12. Let  $\omega_{max}(\mathbf{x})$  be the state of nature for which  $P(\omega_{max}|\mathbf{x}) \geq P(\omega_i|\mathbf{x})$  for all i,  $i = 1, \ldots, c$ .
  - (a) Show that  $P(\omega_{max}|\mathbf{x}) \geq 1/c$ .
  - (b) Show that for the minimum-error-rate decision rule the average probability of error is given by

$$P(error) = 1 - \int P(\omega_{max}|\mathbf{x})p(\mathbf{x}) d\mathbf{x}.$$

- (c) Use these two results to show that  $P(error) \le (c-1)/c$ .
- (d) Describe a situation for which P(error) = (c 1)/c.
- 1. Let x have an exponential density

$$p(x|\theta) = \begin{cases} \theta e^{-\theta x} & x \ge 0\\ 0 & \text{otherwise.} \end{cases}$$

- (a) Plot  $p(x|\theta)$  versus x for  $\theta = 1$ . Plot  $p(x|\theta)$  versus  $\theta$ ,  $(0 \le \theta \le 5)$ , for x = 2.
- (b) Suppose that n samples  $x_1, \ldots, x_n$  are drawn independently according to  $p(x|\theta)$ . Show that the maximum-likelihood estimate for  $\theta$  is given by

$$\hat{\theta} = \frac{1}{\frac{1}{n} \sum_{k=1}^{n} x_k}.$$

2. Let x have a uniform density

$$p(x|\theta) \sim U(0, \theta) = \begin{cases} 1/\theta & 0 \le x \le \theta \\ 0 & \text{otherwise.} \end{cases}$$

(a) Suppose that n samples  $\mathcal{D} = \{x_1, \dots, x_n\}$  are drawn independently according to  $p(x|\theta)$ . Show that the maximum-likelihood estimate for  $\theta$  is  $\max[\mathcal{D}]$ —that is, the value of the maximum element in  $\mathcal{D}$ .

- 3. Maximum-likelihood methods apply to estimates of prior probabilities as well. Let samples be drawn by successive, independent selections of a state of nature  $\omega_i$  with unknown probability  $P(\omega_i)$ . Let  $z_{ik} = 1$  if the state of nature for the kth sample is  $\omega_i$  and  $z_{ik} = 0$  otherwise.
  - (a) Show that

$$P(z_{i1},\ldots,z_{in}|P(\omega_i)) = \prod_{k=1}^n P(\omega_i)^{z_{ik}} (1-P(\omega_i))^{1-z_{ik}}.$$

(b) Show that the maximum-likelihood estimate for  $P(\omega_i)$  is

$$\hat{P}(\omega_i) = \frac{1}{n} \sum_{k=1}^n z_{ik}.$$

Interpret your result in words.

**4.** Let  $\mathbf{x}$  be a d-dimensional binary (0 or 1) vector with a multivariate Bernoulli distribution

$$P(\mathbf{x}|\boldsymbol{\theta}) = \prod_{i=1}^{d} \theta_i^{x_i} (1 - \theta_i)^{1 - x_i},$$

where  $\boldsymbol{\theta} = (\theta_1, \dots, \theta_d)^t$  is an unknown parameter vector,  $\theta_i$  being the probability that  $x_i = 1$ . Show that the maximum-likelihood estimate for  $\boldsymbol{\theta}$  is

$$\hat{\boldsymbol{\theta}} = \frac{1}{n} \sum_{k=1}^{n} \mathbf{x}_{k}.$$

17. The purpose of this problem is to derive the Bayesian classifier for the *d*-dimensional multivariate Bernoulli case. As usual, work with each class separately, interpreting  $P(\mathbf{x}|\mathcal{D})$  to mean  $P(\mathbf{x}|\mathcal{D}_i, \omega_i)$ . Let the conditional probability for a given category be given by

$$P(\mathbf{x}|\boldsymbol{\theta}) = \prod_{i=1}^{d} \theta_i^{x_i} (1 - \theta_i)^{1 - x_i},$$

and let  $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$  be a set of *n* samples independently drawn according to this probability density.

(a) If  $\mathbf{s} = (s_1, \dots, s_d)^t$  is the sum of the *n* samples, show that

$$P(\mathcal{D}|\boldsymbol{\theta}) = \prod_{i=1}^{d} \theta_i^{s_i} (1 - \theta_i)^{n - s_i}.$$

(b) Assuming a uniform prior distribution for  $\theta$  and using the identity

$$\int_{0}^{1} \theta^{m} (1 - \theta)^{n} d\theta = \frac{m! n!}{(m + n + 1)!},$$

show that

$$p(\boldsymbol{\theta}|\mathcal{D}) = \prod_{i=1}^{d} \frac{(n+1)!}{s_i!(n-s_i)!} \theta_i^{s_i} (1-\theta_i)^{n-s_i}.$$

- (c) Plot this density for the case d = 1, n = 1 and for the two resulting possibilities for  $s_1$ .
- (d) Integrate the product  $P(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta}|\mathcal{D})$  over  $\boldsymbol{\theta}$  to obtain the desired conditional probability

$$P(\mathbf{x}|\mathcal{D}) = \prod_{i=1}^{d} \left(\frac{s_i+1}{n+2}\right)^{x_i} \left(1 - \frac{s_i+1}{n+2}\right)^{1-x_i}.$$

(e) If we think of obtaining  $P(\mathbf{x}|\mathcal{D})$  by substituting an estimate  $\hat{\boldsymbol{\theta}}$  for  $\boldsymbol{\theta}$  in  $P(\mathbf{x}|\boldsymbol{\theta})$ , what is the effective Bayesian estimate for  $\boldsymbol{\theta}$ ?