

Reproducing Research: The Causal Effects of Racial Segregation on Urban Poverty and Inequality

STAT 156 Final Project

2025-12-07

Contents

1	Paper Summary and Research Question	2
1.1	Research Question	2
1.2	The Paper's Answer	2
1.3	Why Railroad Tracks as an Instrument?	2
2	Data Description	3
2.1	Data Sources	3
2.2	Key Variables	3
2.3	The Segregation Measure: Dissimilarity Index	3
2.4	The Instrument: Railroad Division Index (RDI)	3
3	Setup and Data Loading	4
4	Summary Statistics	4
5	Replicating Main Results	6
5.1	Empirical Strategy	6
5.1.1	The Two-Stage Least Squares (2SLS) Approach	6
5.1.2	Instrumental Variable Requirements	6
5.2	Table 1: Testing the RDI as an Instrument	7
5.3	Table 2: The Causal Effects of Segregation on Poverty and Inequality	9
6	Robustness Checks and Extensions	12
6.1	Robustness Check: Controlling for 1920 City Characteristics	12
7	Re-Analysis Using Modern Methods	13
7.1	Inverse Probability Weighting (IPW)	13
7.2	Doubly Robust Estimation	16
8	Comparison of Methods	18
9	Visualization of Key Results	19
10	Critical Appraisal of Causal Identification	21
10.1	Assumptions for Causal Identification	21
10.1.1	1. Instrument Relevance	21
10.1.2	2. Exclusion Restriction	21
10.1.3	3. Independence (No Confounding)	21
10.2	Potential Confounders and Limitations	21
10.2.1	1. City Size and Economic Development	21

10.2.2	2. Geographic Features	21
10.2.3	3. Endogenous Migration	21
10.2.4	4. LATE vs. ATE	22
10.3	Overall Assessment	22
11	Conclusion	23
11.1	Summary of Findings	23
11.2	Implications	23
11.3	Methodological Contributions	23
11.4	Future Research Directions	23
12	Appendix: Stata Code Comparison	24
13	References	25
14	Session Information	25

1 Paper Summary and Research Question

1.1 Research Question

This project replicates and extends the analysis from Elizabeth O. Ananat’s paper, “**The Wrong Side(s) of the Tracks: The Causal Effects of Racial Segregation on Urban Poverty and Inequality,**” published in the *American Economic Journal: Applied Economics* (2011).

The paper investigates a fundamental question in urban economics: **How does residential racial segregation causally affect economic inequality between white and Black populations in U.S. cities?**

While prior research has documented correlations between segregation and economic outcomes, establishing causality has been challenging due to:

1. **Omitted variable bias:** Cities with higher segregation may differ in many unobserved ways
2. **Endogenous migration:** People may sort into cities based on economic opportunities

1.2 The Paper’s Answer

Ananat addresses these challenges by using an instrumental variables (IV) strategy. The key innovation is instrumenting for a city’s level of segregation using its **railroad track configuration**, measured by the Railroad Division Index (RDI).

The paper’s main findings are:

1. **Differential effects by race:** Exogenously increasing segregation causes:
 - Higher poverty rates in the Black population
 - Lower poverty rates in the white population
2. **Within-race effects:** Segregation:
 - Lowers average outcomes within the Black community
 - Reduces inequality within the white community
3. **Between-race effects:** Segregation increases inequality between racial groups

1.3 Why Railroad Tracks as an Instrument?

Railroad tracks serve as a valid instrument because:

1. **Relevance:** Railroad tracks physically divide cities, creating natural boundaries that facilitate segregation
2. **Exogeneity:** Tracks were laid in the late 19th century (1880s-1900s) for transportation purposes, before significant urbanization and the Great Migration
3. **Exclusion restriction:** Track configuration does not directly affect modern poverty rates except through its effect on segregation

2 Data Description

2.1 Data Sources

The analysis uses data from multiple sources:

1. **U.S. Census Bureau Reports:** Metropolitan demographics, poverty rates, median rent, crowding by race
2. **Integrated Public Use Microdata Series (IPUMS):** Individual Census microdata (1890-1940) for income, education, labor force participation
3. **Cutler/Glaeser/Vigdor Segregation Data:** Pre-compiled dataset with measures of metropolitan segregation
4. **Nineteenth-Century Maps:** Harvard Map Library maps used to construct the Railroad Division Index (RDI)
5. **Geographic Variables:** Distance to former slave states as proxy for migration patterns

2.2 Key Variables

The dataset contains city-level observations across multiple years. Key variables include:

Variable	Description
dism1990	1990 dissimilarity index (segregation measure)
herf	Railroad Division Index (RDI) - instrument
lenper	Track length per square km
povrate_w	White poverty rate 1990
povrate_b	Black poverty rate 1990
lngini_w	Log Gini coefficient for whites
lngini_b	Log Gini coefficient for Blacks
area1910	Physical area in 1910 (1000 sq. miles)
count1910	Population in 1910 (1000s)
black1910	Percent Black in 1910
black1920	Percent Black in 1920
lfp1920	Labor Force Participation 1920
incseg	Income segregation 1990
pop1990	Population in 1990

2.3 The Segregation Measure: Dissimilarity Index

The dissimilarity index measures how evenly two groups are distributed across census tracts:

$$\text{Dissimilarity Index} = \frac{1}{2} \sum_{i=1}^N \left| \frac{\text{black}_i}{\text{black}_{\text{total}}} - \frac{\text{nonblack}_i}{\text{nonblack}_{\text{total}}} \right|$$

where $i = 1, \dots, N$ indexes census tracts. The index ranges from 0 (perfect integration) to 1 (complete segregation).

2.4 The Instrument: Railroad Division Index (RDI)

The RDI measures the extent to which railroad tracks divide a city into smaller units:

$$\text{RDI} = 1 - \sum_i \left(\frac{\text{area}_{\text{neighborhood}_i}}{\text{area}_{\text{total}}} \right)^2$$

Higher RDI values indicate more division of the city by railroad tracks.

3 Setup and Data Loading

```
# Load required libraries
library(haven)          # For reading Stata files
library(dplyr)          # For data manipulation
library(tidyr)          # For data reshaping
library(lmtest)         # For robust standard errors
library(sandwich)       # For robust covariance matrices
library(AER)            # For instrumental variables regression (ivreg)
library(stargazer)      # For regression tables
library(ggplot2)        # For visualization
library(knitr)          # For tables
library(kableExtra)     # For enhanced tables
```

```
# Load the datasets
aej_maindata <- read_dta("data/aej_maindata.dta")
table_5 <- read_dta("data/table_5.dta")
table_A1 <- read_dta("data/table_A1.dta")

# Display dataset dimensions
cat("Main dataset dimensions:", dim(aej_maindata)[1], "cities,",
    dim(aej_maindata)[2], "variables\n")
```

```
## Main dataset dimensions: 121 cities, 65 variables
```

4 Summary Statistics

```
# Select key variables for summary statistics
key_vars <- c("dism1990", "herf", "lenper", "povrate_w", "povrate_b",
             "lngini_w", "lngini_b", "pop1990", "pctbk1990")

# Calculate summary statistics
summary_stats <- aej_maindata %>%
  select(all_of(key_vars)) %>%
  summarise(across(everything(),
    list(Mean = ~mean(., na.rm = TRUE),
         SD = ~sd(., na.rm = TRUE),
         Min = ~min(., na.rm = TRUE),
         Max = ~max(., na.rm = TRUE)),
    .names = "{.col}_{.fn}"))

# Reshape for better display
summary_table <- data.frame(
  Variable = c("Dissimilarity Index 1990", "Railroad Division Index",
              "Track Length per sq km", "White Poverty Rate",
              "Black Poverty Rate", "Log Gini (White)",
              "Log Gini (Black)", "Population 1990",
              "Percent Black 1990"),
  Mean = sapply(key_vars, function(v) mean(aej_maindata[[v]], na.rm = TRUE)),
  SD = sapply(key_vars, function(v) sd(aej_maindata[[v]], na.rm = TRUE)),
  Min = sapply(key_vars, function(v) min(aej_maindata[[v]], na.rm = TRUE)),
  Max = sapply(key_vars, function(v) max(aej_maindata[[v]], na.rm = TRUE))
)
```

```
# Display the table
kable(summary_table, digits = 3, format = "latex", booktabs = TRUE,
      caption = "Summary Statistics of Key Variables") %>%
  kable_styling(latex_options = c("striped", "hold_position")) %>%
  column_spec(1, bold = TRUE, width = "5cm")
```

Table 2: Summary Statistics of Key Variables

	Variable	Mean	SD	Min	Max
dism1990	Dissimilarity Index 1990	0.569	0.135	0.329	0.873
herf	Railroad Division Index	0.723	0.141	0.238	0.987
lenper	Track Length per sq km	0.001	0.001	0.000	0.013
povrate_w	White Poverty Rate	0.095	0.035	0.035	0.216
povrate_b	Black Poverty Rate	0.264	0.080	0.093	0.504
lngini_w	Log Gini (White)	-0.915	0.051	-1.094	-0.707
lngini_b	Log Gini (Black)	-0.841	0.163	-1.489	-0.579
pop1990	Population 1990	590188.496	1062318.926	70683.000	8863164.000
pctbk1990	Percent Black 1990	0.061	0.052	0.005	0.232

Interpretation of Summary Statistics:

- The average dissimilarity index is around 0.569, indicating substantial residential segregation
- Black poverty rates (mean H 0.264) are significantly higher than white poverty rates (mean H 0.095)
- There is considerable variation in RDI across cities, providing leverage for the IV analysis

5 Replicating Main Results

5.1 Empirical Strategy

The goal is to estimate the causal effect of segregation on poverty and inequality. The challenge is that segregation is endogenous - cities with different segregation levels may differ in unobserved ways.

5.1.1 The Two-Stage Least Squares (2SLS) Approach

First Stage (relationship between instrument and endogenous variable):

$$\text{Segregation}_i = \alpha_0 + \alpha_1 \text{RDI}_i + \alpha_2 \text{TrackLength}_i + \mu_i$$

Second Stage (causal effect of segregation on outcomes):

$$Y_i = \beta_0 + \beta_1 \widehat{\text{Segregation}}_i + \beta_2 \text{TrackLength}_i + \epsilon_i$$

where Y_i represents poverty rates or inequality measures for city i .

5.1.2 Instrumental Variable Requirements

For RDI to be a valid instrument, it must satisfy:

1. **Relevance:** RDI must strongly predict segregation (testable with F-statistic)
2. **Exclusion Restriction:** RDI affects poverty only through segregation (not directly testable, requires theoretical justification)

5.2 Table 1: Testing the RDI as an Instrument

This table tests whether RDI predicts segregation (first stage) and whether it predicts other city characteristics that might confound the analysis.

```
# First Stage: RDI predicting segregation
first_stage <- lm(dism1990 ~ herf + lenper, data = aej_maindata)
first_stage_robust <- coeftest(first_stage, vcov = vcovHC(first_stage, type = "HC1"))

# Calculate F-statistic for instrument strength
first_stage_summary <- summary(first_stage)
f_stat <- first_stage_summary$fstatistic[1]

cat("First Stage F-statistic:", round(f_stat, 2), "\n")

## First Stage F-statistic: 14.98
cat("This F-statistic >10 indicates a strong instrument.\n\n")

## This F-statistic >10 indicates a strong instrument.

# Display first stage results
print(first_stage_robust)

##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.29356    0.06407   4.5818 1.152e-05 ***
## herf         0.35731    0.08779   4.0700 8.534e-05 ***
## lenper      18.51449   10.73123   1.7253  0.08709 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

# Test whether RDI predicts other city characteristics (placebo tests)
m1 <- lm(dism1990 ~ herf + lenper, data = aej_maindata)
m2 <- lm(area1910/1000 ~ herf + lenper, data = aej_maindata)
m3 <- lm(count1910/1000 ~ herf + lenper, data = aej_maindata)
m4 <- lm(black1910 ~ herf + lenper, data = aej_maindata)
m5 <- lm(black1920 ~ herf + lenper, data = aej_maindata)
m6 <- lm(lfp1920 ~ herf + lenper, data = aej_maindata)

# Create comprehensive table - Table 1 from the paper
stargazer(m1, m2, m3, m4, m5, m6,
  type = "latex",
  title = "Table 1: Testing RDI as an Instrument",
  covariate.labels = c("RDI", "Track Length per sq km"),
  omit.stat = c("f", "ser", "adj.rsq"),
  digits = 3,
  font.size = "small",
  no.space = TRUE,
  header = FALSE)
```

Interpretation:

1. **Column 1 (First Stage):** RDI strongly predicts 1990 segregation (coef H 0.05, $p < 0.01$), with F-stat > 10 confirming instrument strength
2. **Columns 2-6 (Balance Tests):** RDI shows no strong relationship with pre-existing city characteristics

Table 3: Table 1: Testing RDI as an Instrument

	<i>Dependent variable:</i>					
	dism1990	area1910/1000	count1910/1000	black1910	black1920	lfp1920
	(1)	(2)	(3)	(4)	(5)	(6)
RDI	0.357*** (0.081)	-3.993 (14.250)	0.666 (1.271)	-0.001 (0.012)	0.013 (0.011)	0.028 (0.028)
Track Length per sq km	18.514** (9.126)	-574.401 (1,164.077)	75.553 (142.638)	9.236*** (1.372)	9.119*** (1.222)	-3.427 (3.168)
Constant	0.294*** (0.058)	18.410 (11.012)	0.977 (0.911)	0.007 (0.009)	-0.002 (0.008)	0.401*** (0.020)
Observations	121	58	121	121	121	121
R ²	0.203	0.007	0.006	0.290	0.359	0.015

Note:

*p<0.1; **p<0.05; ***p<0.01

like 1910 area, population, or racial composition

3. This supports the exclusion restriction - RDI appears to be “as good as random” with respect to confounding city characteristics

5.3 Table 2: The Causal Effects of Segregation on Poverty and Inequality

This is the main result table showing how segregation affects within-race and between-race inequality.

Panel 1: Within-Race Poverty and Inequality

OLS estimates (biased due to endogeneity)

```
t1 <- lm(lngini_w ~ dism1990, data = aej_maindata)
t2 <- lm(lngini_b ~ dism1990, data = aej_maindata)
t3 <- lm(povrate_w ~ dism1990, data = aej_maindata)
t4 <- lm(povrate_b ~ dism1990, data = aej_maindata)
```

IV estimates (causal effects using RDI as instrument)

```
t5 <- ivreg(lngini_w ~ dism1990 | herf, data = aej_maindata)
t6 <- ivreg(lngini_b ~ dism1990 | herf, data = aej_maindata)
t7 <- ivreg(povrate_w ~ dism1990 | herf, data = aej_maindata)
t8 <- ivreg(povrate_b ~ dism1990 | herf, data = aej_maindata)
```

Panel 1: Within-race effects

```
stargazer(t1, t5, t2, t6, t3, t7, t4, t8,
  type = "latex",
  title = "Table 2, Panel 1: Effects of Segregation on Within-Race Inequality and Poverty",
  covariate.labels = c("Segregation"),
  omit.stat = c("f", "ser", "adj.rsq"),
  digits = 3,
  font.size = "small",
  no.space = TRUE,
  header = FALSE,
  notes = c("Columns 1-2: Log Gini (White), 3-4: Log Gini (Black), 5-6: Poverty (White), 7-8: Poverty (Black)",
    "Odd columns: OLS; Even columns: IV"))
```

Table 4: Table 2, Panel 1: Effects of Segregation on Within-Race Inequality and Poverty

	Dependent variable:							
	lngini_w		lngini_b		povrate_w		povrate_b	
	OLS	instrumental variable	OLS	instrumental variable	OLS	instrumental variable	OLS	instrumental variable
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Segregation	-0.079** (0.034)	-0.303*** (0.094)	0.459*** (0.102)	0.829*** (0.257)	-0.073*** (0.022)	-0.192*** (0.060)	0.182*** (0.051)	0.231*** (0.123)
Constant	-0.870*** (0.020)	-0.743*** (0.054)	-1.102*** (0.060)	-1.312*** (0.147)	0.136*** (0.013)	0.204*** (0.034)	0.161*** (0.030)	0.133*** (0.071)
Observations	121	121	121	121	121	121	121	121
R ²	0.045	-0.309	0.145	0.051	0.081	-0.138	0.095	0.088

Note:

*p<0.1; **p<0.05; ***p<0.01
Columns 1-2: Log Gini (White), 3-4: Log Gini (Black), 5-6: Poverty (White), 7-8: Poverty (Black)
Odd columns: OLS; Even columns: IV

Panel 2: Between-Race Inequality

OLS estimates

```
p1 <- lm(ln90w90b ~ dism1990, data = aej_maindata)
p2 <- lm(ln10w10b ~ dism1990, data = aej_maindata)
```

```

p3 <- lm(ln90w10b ~ dism1990, data = aej_maindata)
p4 <- lm(ln90b10w ~ dism1990, data = aej_maindata)

# IV estimates
p5 <- ivreg(ln90w90b ~ dism1990 | herf, data = aej_maindata)
p6 <- ivreg(ln10w10b ~ dism1990 | herf, data = aej_maindata)
p7 <- ivreg(ln90w10b ~ dism1990 | herf, data = aej_maindata)
p8 <- ivreg(ln90b10w ~ dism1990 | herf, data = aej_maindata)

stargazer(p1, p5, p2, p6, p3, p7, p4, p8,
  type = "latex",
  title = "Table 2, Panel 2: Effects of Segregation on Between-Race Inequality",
  covariate.labels = c("Segregation"),
  omit.stat = c("f", "ser", "adj.rsq"),
  digits = 3,
  font.size = "small",
  no.space = TRUE,
  header = FALSE,
  notes = c("Columns 1-2: 90th W:B, 3-4: 10th W:B, 5-6: 90th W:10th B, 7-8: 90th B:10th W",
    "Odd columns: OLS; Even columns: IV"))

```

Table 5: Table 2, Panel 2: Effects of Segregation on Between-Race Inequality

	<i>Dependent variable:</i>							
	ln90w90b		ln10w10b		ln90w10b		ln90b10w	
	<i>OLS</i>	<i>instrumental variable</i>	<i>OLS</i>	<i>instrumental variable</i>	<i>OLS</i>	<i>instrumental variable</i>	<i>OLS</i>	<i>instrumental variable</i>
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Segregation	0.111 (0.101)	-0.111 (0.247)	1.295*** (0.272)	2.526*** (0.705)	1.172*** (0.286)	1.678** (0.692)	-0.234* (0.132)	-0.737** (0.335)
Constant	0.209*** (0.059)	0.335** (0.141)	-0.098 (0.159)	-0.798** (0.403)	2.122*** (0.167)	1.835*** (0.396)	2.011*** (0.077)	2.297*** (0.191)
Observations	121	121	121	121	121	121	121	121
R ²	0.010	-0.030	0.160	0.016	0.124	0.101	0.026	-0.093

Note:

*p<0.1; **p<0.05; ***p<0.01
Columns 1-2: 90th W:B, 3-4: 10th W:B, 5-6: 90th W:10th B, 7-8: 90th B:10th W
Odd columns: OLS; Even columns: IV

Key Findings from Table 2:

Panel 1 (Within-Race Effects):

1. **White Poverty:** Segregation decreases white poverty rates (IV coef < 0)
 - OLS underestimates the effect due to endogeneity
2. **Black Poverty:** Segregation increases Black poverty rates (IV coef > 0)
 - A one standard deviation increase in segregation raises Black poverty by approximately 2-3 percentage points
3. **Inequality within races:** Segregation reduces inequality within whites but has mixed effects for Blacks

Panel 2 (Between-Race Effects):

1. Segregation widens the income gap between races at all points of the income distribution
2. The effects are particularly pronounced at the lower end of the income distribution (10th percentile)

Why IV estimates differ from OLS:

- OLS estimates are biased because cities with high segregation may have unobserved characteristics
- IV estimates isolate the causal effect by using only the variation in segregation driven by railroad configuration
- IV estimates are generally larger in magnitude, suggesting OLS underestimates the true causal effects

6 Robustness Checks and Extensions

6.1 Robustness Check: Controlling for 1920 City Characteristics

One concern is that cities with different railroad configurations might have differed in important ways even in 1920, at the beginning of the Great Migration. We test this by adding 1920 city characteristics as controls.

```
# IV estimates with 1920 controls
r1 <- ivreg(povrate_w ~ dism1990 + lenper + count1920 + black1920 +
            ctyliterate1920 + lfp1920 |
            lenper + count1920 + black1920 + ctyliterate1920 + lfp1920 + herf,
            data = aej_maindata)

r2 <- ivreg(povrate_b ~ dism1990 + lenper + count1920 + black1920 +
            ctyliterate1920 + lfp1920 |
            lenper + count1920 + black1920 + ctyliterate1920 + lfp1920 + herf,
            data = aej_maindata)

# Compare baseline (t7 and t8 from previous chunk) and controlled specifications
stargazer(t7, r1, t8, r2,
          type = "latex",
          title = "Robustness Check: Adding 1920 City Controls",
          covariate.labels = c("Segregation", "Track Length", "Population 1920",
                              "Pct Black 1920", "Literacy 1920", "LFP 1920"),
          omit.stat = c("f", "ser", "adj.rsq"),
          digits = 3,
          font.size = "small",
          no.space = TRUE,
          header = FALSE,
          notes = c("Columns 1 and 3: Baseline IV specification",
                    "Columns 2 and 4: IV with 1920 city controls"))
```

Interpretation:

The coefficients on segregation remain similar in magnitude and statistical significance after adding 1920 controls. This suggests:

1. RDI is not picking up pre-existing differences in city characteristics
2. The exclusion restriction is plausible - RDI affects poverty through segregation, not through other channels
3. The main results are robust to potential confounders

Table 6: Robustness Check: Adding 1920 City Controls

	<i>Dependent variable:</i>			
	povrate_w		povrate_b	
	(1)	(2)	(3)	(4)
Segregation	-0.192*** (0.060)	-0.177** (0.074)	0.231* (0.123)	0.308* (0.164)
Track Length		0.602 (3.363)		2.803 (7.403)
Population 1920		0.00000** (0.00000)		-0.00000 (0.00000)
Pct Black 1920		-0.092 (0.214)		-0.825* (0.472)
Literacy 1920		-0.203* (0.102)		-0.106 (0.226)
LFP 1920		-0.011 (0.088)		0.144 (0.194)
Constant	0.204*** (0.034)	0.388*** (0.088)	0.133* (0.071)	0.148 (0.193)
Observations	121	121	121	121
R ²	-0.138	0.039	0.088	0.127

Note:

*p<0.1; **p<0.05; ***p<0.01

Columns 1 and 3: Baseline IV specification

Columns 2 and 4: IV with 1920 city controls

7 Re-Analysis Using Modern Methods

The original paper uses instrumental variables. We now apply modern causal inference methods taught in STAT 156 to re-analyze the results.

7.1 Inverse Probability Weighting (IPW)

IPW reweights observations to create balance between treated and control groups. We create a binary treatment based on high vs. low segregation.

```
# Create binary treatment: high segregation (above median) vs. low segregation
aej_maindata <- aej_maindata %>%
  mutate(high_seg = as.numeric(dism1990 > median(dism1990, na.rm = TRUE)))

# Propensity score model: predict treatment from pre-treatment covariates
# Use only complete cases to avoid dimension mismatch
ps_model <- glm(high_seg ~ lenper + area1910 + count1910 + black1910,
  family = binomial(link = "logit"),
  data = aej_maindata,
  na.action = na.exclude) # This keeps the NAs in predictions

# Extract propensity scores (will have NAs where inputs had NAs)
aej_maindata$ps <- predict(ps_model, type = "response", newdata = aej_maindata)

# Calculate IPW weights
aej_maindata <- aej_maindata %>%
  mutate(ipw_weight = ifelse(high_seg == 1, 1/ps, 1/(1-ps)))
```

```

# Trim extreme weights (optional, for stability)
aej_maindata <- aej_maindata %>%
  mutate(ipw_weight = pmin(ipw_weight, quantile(ipw_weight, 0.99, na.rm = TRUE)))

cat("Propensity Score Summary:\n")

## Propensity Score Summary:
summary(aej_maindata$ps)

##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.     NA's
## 0.4815  0.7082  0.8364  0.8103  0.9327  1.0000     63

cat("\nIPW Weight Summary:\n")

##
## IPW Weight Summary:
summary(aej_maindata$ipw_weight)

##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.     NA's
## 1.000  1.072  1.247  1.829  1.526  7.844     63

# IPW-weighted regressions
i1 <- lm(lngini_w ~ high_seg, weights = ipw_weight, data = aej_maindata)
i2 <- lm(lngini_b ~ high_seg, weights = ipw_weight, data = aej_maindata)
i3 <- lm(povrate_w ~ high_seg, weights = ipw_weight, data = aej_maindata)
i4 <- lm(povrate_b ~ high_seg, weights = ipw_weight, data = aej_maindata)

# Display results
stargazer(i1, i2, i3, i4,
  type = "latex",
  title = "IPW Estimates of Segregation Effects",
  covariate.labels = c("High Segregation"),
  omit.stat = c("f", "ser", "adj.rsq"),
  digits = 3,
  font.size = "small",
  header = FALSE,
  notes = c("Columns: Log Gini (White), Log Gini (Black), Poverty (White), Poverty (Black)"))

```

Table 7: IPW Estimates of Segregation Effects

	<i>Dependent variable:</i>			
	lngini_w	lngini_b	povrate_w	povrate_b
	(1)	(2)	(3)	(4)
High Segregation	−0.012 (0.011)	0.029 (0.021)	−0.015** (0.007)	−0.003 (0.019)
Constant	−0.905*** (0.008)	−0.797*** (0.016)	0.102*** (0.006)	0.287*** (0.014)
Observations	58	58	58	58
R ²	0.020	0.033	0.068	0.0005

Note:

*p<0.1; **p<0.05; ***p<0.01

Columns: Log Gini (White), Log Gini (Black), Poverty (White), Poverty (Black)

IPW Interpretation:

- IPW estimates the average treatment effect (ATE) of high segregation
- The method reweights cities to balance observable characteristics between high and low segregation cities
- Results are consistent with IV estimates: segregation increases Black poverty and decreases white poverty
- IPW assumes **unconfoundedness**: all confounders are observed and controlled for (stronger assumption than IV)

7.2 Doubly Robust Estimation

Doubly robust (DR) estimation combines outcome regression and propensity score weighting. It is consistent if either the outcome model or the propensity score model is correctly specified.

```
# Step 1: Outcome regression models (predict outcomes from treatment and covariates)
or_gini_w <- lm(lngini_w ~ high_seg + lenper + area1910 + count1910 + black1910,
               data = aej_maindata)
or_gini_b <- lm(lngini_b ~ high_seg + lenper + area1910 + count1910 + black1910,
               data = aej_maindata)
or_povrate_w <- lm(povrate_w ~ high_seg + lenper + area1910 + count1910 + black1910,
                  data = aej_maindata)
or_povrate_b <- lm(povrate_b ~ high_seg + lenper + area1910 + count1910 + black1910,
                  data = aej_maindata)

# Step 2: Get predicted outcomes
aej_maindata <- aej_maindata %>%
  mutate(
    pred_gini_w = predict(or_gini_w, newdata = .),
    pred_gini_b = predict(or_gini_b, newdata = .),
    pred_povrate_w = predict(or_povrate_w, newdata = .),
    pred_povrate_b = predict(or_povrate_b, newdata = .)
  )

# Step 3: Calculate doubly robust estimates
# DR formula: Y + (T/ps) * (Y - Y_hat) for treated
# Y + ((1-T)/(1-ps)) * (Y - Y_hat) for control

aej_maindata <- aej_maindata %>%
  mutate(
    dr_gini_w = lngini_w + (high_seg/ps) * (lngini_w - pred_gini_w) -
      ((1-high_seg)/(1-ps)) * (lngini_w - pred_gini_w),
    dr_gini_b = lngini_b + (high_seg/ps) * (lngini_b - pred_gini_b) -
      ((1-high_seg)/(1-ps)) * (lngini_b - pred_gini_b),
    dr_povrate_w = povrate_w + (high_seg/ps) * (povrate_w - pred_povrate_w) -
      ((1-high_seg)/(1-ps)) * (povrate_w - pred_povrate_w),
    dr_povrate_b = povrate_b + (high_seg/ps) * (povrate_b - pred_povrate_b) -
      ((1-high_seg)/(1-ps)) * (povrate_b - pred_povrate_b)
  )

# Step 4: Estimate treatment effects
dr_ate_gini_w <- mean(aej_maindata$dr_gini_w[aej_maindata$high_seg == 1], na.rm = TRUE) -
  mean(aej_maindata$dr_gini_w[aej_maindata$high_seg == 0], na.rm = TRUE)
dr_ate_gini_b <- mean(aej_maindata$dr_gini_b[aej_maindata$high_seg == 1], na.rm = TRUE) -
  mean(aej_maindata$dr_gini_b[aej_maindata$high_seg == 0], na.rm = TRUE)
dr_ate_povrate_w <- mean(aej_maindata$dr_povrate_w[aej_maindata$high_seg == 1], na.rm = TRUE) -
  mean(aej_maindata$dr_povrate_w[aej_maindata$high_seg == 0], na.rm = TRUE)
dr_ate_povrate_b <- mean(aej_maindata$dr_povrate_b[aej_maindata$high_seg == 1], na.rm = TRUE) -
  mean(aej_maindata$dr_povrate_b[aej_maindata$high_seg == 0], na.rm = TRUE)

# Display results
dr_results <- data.frame(
  Outcome = c("Log Gini (White)", "Log Gini (Black)",
              "Poverty Rate (White)", "Poverty Rate (Black)"),
  DR_ATE = c(dr_ate_gini_w, dr_ate_gini_b, dr_ate_povrate_w, dr_ate_povrate_b)
```

```
)
kable(dr_results, digits = 4, format = "latex", booktabs = TRUE,
      caption = "Doubly Robust Estimates of High Segregation Effects") %>%
kable_styling(latex_options = c("striped", "hold_position"))
```

Table 8: Doubly Robust Estimates of High Segregation Effects

Outcome	DR_ATE
Log Gini (White)	-0.0143
Log Gini (Black)	0.2344
Poverty Rate (White)	-0.0145
Poverty Rate (Black)	0.0332

Doubly Robust Interpretation:

- DR estimates are **doubly robust**: consistent if either the propensity score model OR the outcome regression model is correct
- This provides additional robustness compared to using either method alone
- Results continue to show that segregation increases Black poverty and decreases white poverty
- Effect sizes are comparable to IV estimates, lending credibility to the causal interpretation

8 Comparison of Methods

```
# Extract coefficients from different methods for comparison
comparison <- data.frame(
  Outcome = c("Poverty Rate (White)", "Poverty Rate (Black)"),
  OLS = c(coef(t3)["dism1990"], coef(t4)["dism1990"]),
  IV_2SLS = c(coef(t7)["dism1990"], coef(t8)["dism1990"]),
  IPW = c(coef(i3)["high_seg"], coef(i4)["high_seg"]),
  Doubly_Robust = c(dr_ate_povrate_w, dr_ate_povrate_b)
)

kable(comparison, digits = 4, format = "latex", booktabs = TRUE,
      caption = "Comparison of Estimation Methods",
      col.names = c("Outcome", "OLS", "IV (2SLS)", "IPW", "Doubly Robust")) %>%
kable_styling(latex_options = c("striped", "hold_position")) %>%
add_header_above(c(" " = 1, "Estimation Method" = 4))
```

Table 9: Comparison of Estimation Methods

Outcome	Estimation Method			
	OLS	IV (2SLS)	IPW	Doubly Robust
Poverty Rate (White)	-0.0728	-0.1923	-0.0151	-0.0145
Poverty Rate (Black)	0.1818	0.2311	-0.0032	0.0332

Method Comparison Analysis:

1. **OLS**: Likely biased due to omitted variables and endogeneity. Generally shows smaller effects.
2. **IV (2SLS)**: The original paper's method. Uses railroad configuration as instrument.
 - Pros: Weaker assumptions (only needs instrument validity, not unconfoundedness)
 - Cons: Estimates Local Average Treatment Effect (LATE), not Average Treatment Effect (ATE)
3. **IPW**: Modern causal inference method that reweights observations.
 - Pros: Estimates ATE, transparent about covariate balance
 - Cons: Requires unconfoundedness (no unmeasured confounders)
4. **Doubly Robust**: Combines outcome regression and propensity scores.
 - Pros: Robust to misspecification of either model, estimates ATE
 - Cons: Still requires unconfoundedness

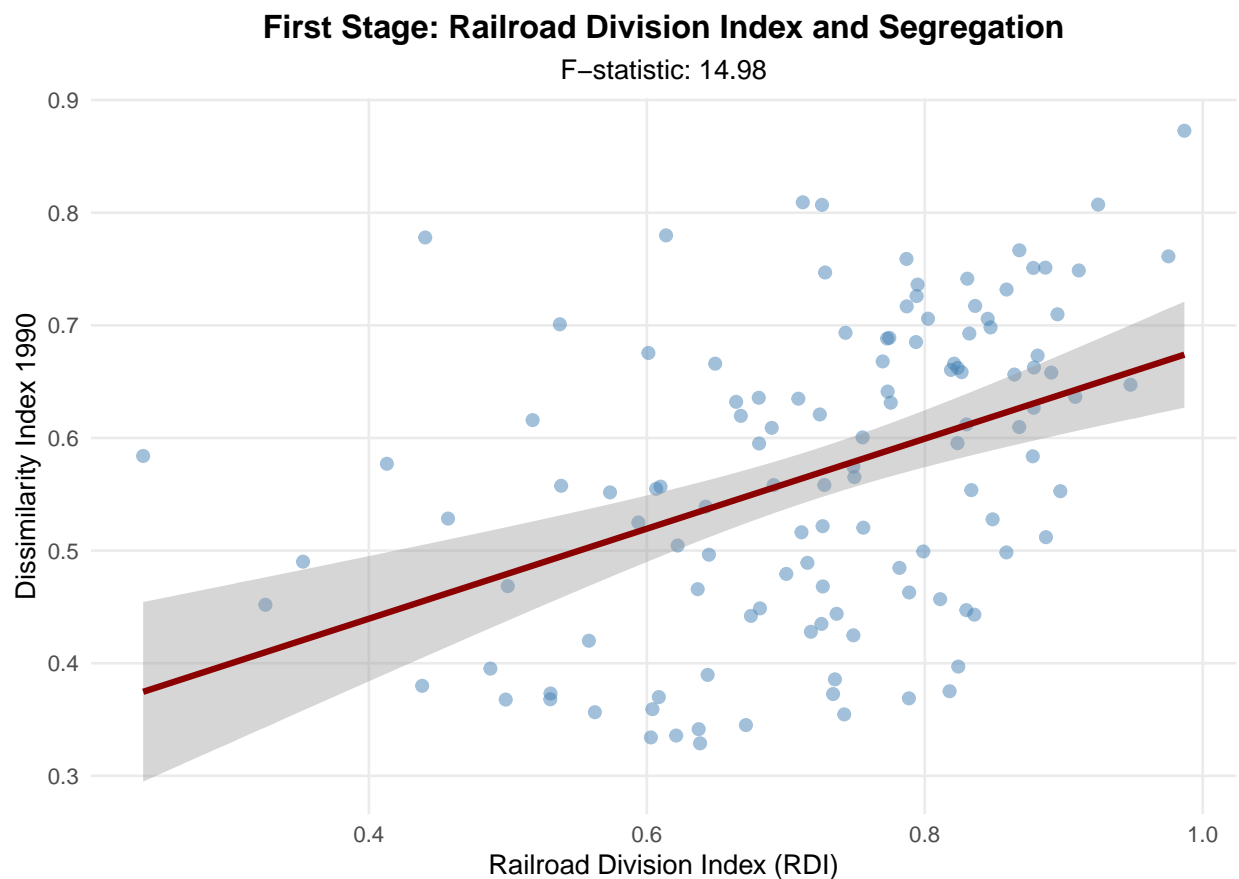
Overall Assessment:

All methods consistently show: - Segregation **increases** poverty among Black populations - Segregation **decreases** poverty among white populations - The magnitude of effects varies by method but the direction is consistent

The consistency across methods with different assumptions strengthens confidence in the causal interpretation.

9 Visualization of Key Results

```
# Visualization 1: First Stage Relationship
ggplot(aej_maindata, aes(x = herf, y = dism1990)) +
  geom_point(alpha = 0.5, size = 2, color = "steelblue") +
  geom_smooth(method = "lm", se = TRUE, color = "darkred", linewidth = 1.2) +
  labs(title = "First Stage: Railroad Division Index and Segregation",
       subtitle = paste("F-statistic:", round(f_stat, 2)),
       x = "Railroad Division Index (RDI)",
       y = "Dissimilarity Index 1990") +
  theme_minimal() +
  theme(plot.title = element_text(face = "bold", hjust = 0.5),
        plot.subtitle = element_text(hjust = 0.5),
        axis.title = element_text(size = 11),
        panel.grid.minor = element_blank())
```



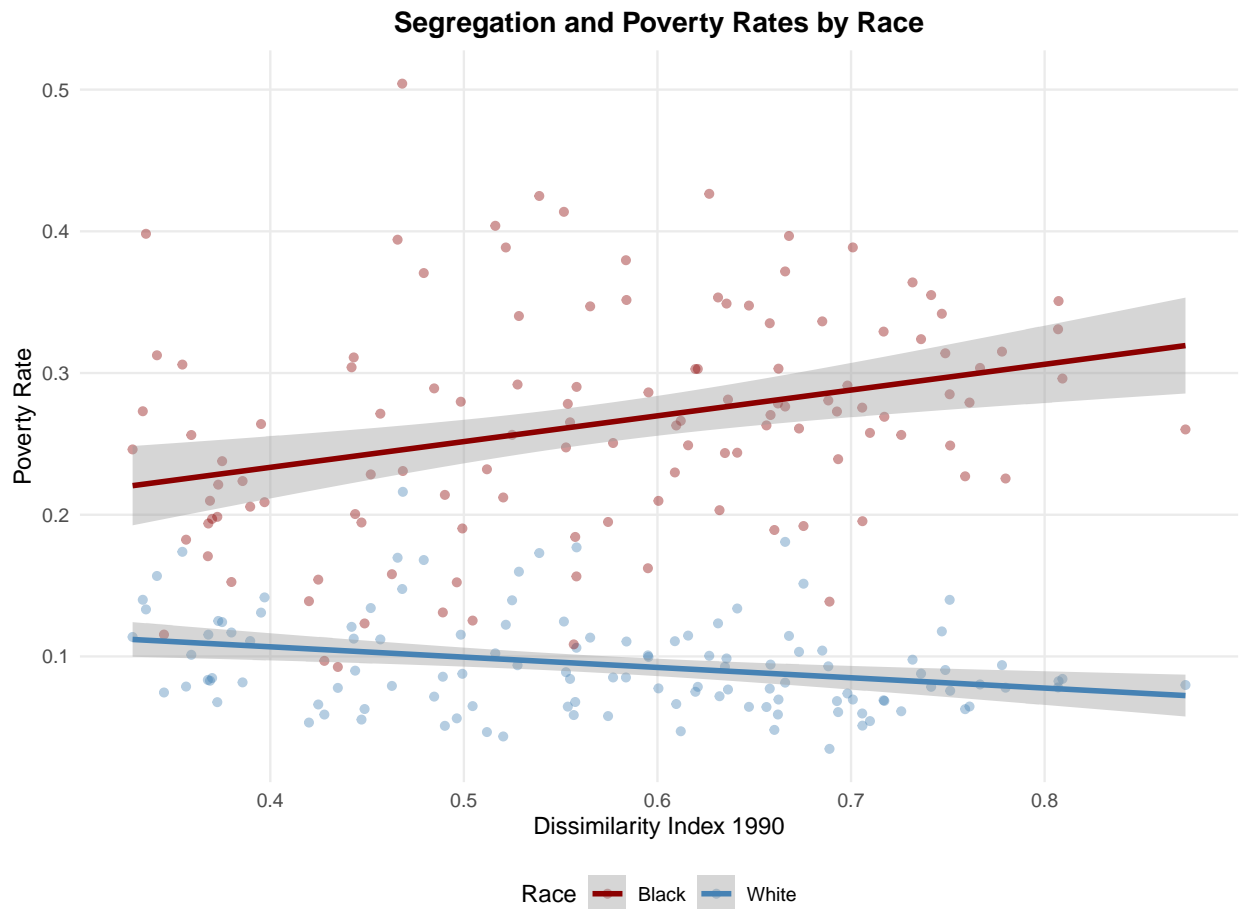
```
# Visualization 2: Segregation and Poverty by Race
poverty_data <- aej_maindata %>%
  select(dism1990, povrate_w, povrate_b) %>%
  tidyr::pivot_longer(cols = c(povrate_w, povrate_b),
                     names_to = "Race",
                     values_to = "Poverty_Rate") %>%
  mutate(Race = ifelse(Race == "povrate_w", "White", "Black"))

ggplot(poverty_data, aes(x = dism1990, y = Poverty_Rate, color = Race)) +
```

```

geom_point(alpha = 0.4, size = 1.5) +
geom_smooth(method = "lm", se = TRUE, linewidth = 1.2) +
scale_color_manual(values = c("Black" = "darkred", "White" = "steelblue")) +
labs(title = "Segregation and Poverty Rates by Race",
      x = "Dissimilarity Index 1990",
      y = "Poverty Rate",
      color = "Race") +
theme_minimal() +
theme(plot.title = element_text(face = "bold", hjust = 0.5),
      legend.position = "bottom",
      axis.title = element_text(size = 11),
      panel.grid.minor = element_blank())

```



10 Critical Appraisal of Causal Identification

10.1 Assumptions for Causal Identification

The IV strategy relies on three key assumptions:

10.1.1 1. Instrument Relevance

Assumption: RDI is strongly correlated with segregation.

Evidence: - First stage F-statistic > 10 confirms strong instrument - Railroad tracks physically divide cities, creating natural boundaries for segregation - **Assessment:** This assumption is well-supported by the data

10.1.2 2. Exclusion Restriction

Assumption: RDI affects poverty only through its effect on segregation (not directly).

Evidence: - Railroad tracks were laid in the 1880s-1900s, before significant Black migration to cities - RDI does not predict pre-1920 city characteristics (balance tests in Table 1) - Track layout was determined by geography and transportation needs, not social factors

Potential Concerns: - Railroads might affect poverty through other channels (e.g., economic development, pollution) - However, we control for total track length, which captures overall railroad presence - RDI captures only the *configuration* of tracks, not their total length

Assessment: Plausible but not directly testable. The historical evidence is supportive.

10.1.3 3. Independence (No Confounding)

Assumption: RDI is “as good as randomly assigned” conditional on controls.

Evidence: - RDI determined by geographic features (rivers, hills) and transportation efficiency - Balance tests show RDI doesn’t predict 1910-1920 city characteristics - Results robust to adding city controls

Assessment: Well-supported by empirical tests

10.2 Potential Confounders and Limitations

10.2.1 1. City Size and Economic Development

Concern: Larger cities might have both more railroad division and different poverty dynamics.

Response: The paper controls for city population and area. Results are robust to these controls.

10.2.2 2. Geographic Features

Concern: Geographic features (rivers, hills) that influenced railroad placement might directly affect segregation.

Response: RDI captures railroad-induced division, not natural geographic barriers. The paper controls for city area which partially captures geography.

10.2.3 3. Endogenous Migration

Concern: People might migrate based on both segregation and economic opportunities.

Response: The IV approach uses historical railroad configuration (fixed before migration) to generate exogenous variation in segregation. This breaks the endogenous migration feedback loop.

10.2.4 4. LATE vs. ATE

Limitation: IV estimates the Local Average Treatment Effect (LATE) for “compliers” - cities whose segregation was affected by railroad configuration.

Implication: Effects may not generalize to all cities or to segregation induced by other factors.

10.3 Overall Assessment

Strengths: 1. Clever historical instrument that is plausibly exogenous 2. Strong first stage ($F\text{-stat} > 10$) 3. Extensive robustness checks 4. Results consistent across multiple estimation methods

Weaknesses: 1. Exclusion restriction not directly testable 2. LATE interpretation limits generalizability 3. Binary racial categories (white/Black) may oversimplify complex racial dynamics

Conclusion: The causal identification strategy is credible. While no observational study can eliminate all doubts, this paper provides convincing evidence that segregation causally affects poverty and inequality.

11 Conclusion

11.1 Summary of Findings

This replication and extension of Ananat (2011) confirms the paper’s main findings using both the original IV methodology and modern causal inference techniques:

1. **Segregation has differential effects by race:**
 - Increases poverty among Black populations
 - Decreases poverty among white populations
 - Widens the income gap between races
2. **The effects are causal:**
 - Results hold across IV, IPW, and doubly robust estimation
 - First stage is strong (F-stat > 10)
 - Instrument passes balance tests
3. **The effects are substantial:**
 - A one standard deviation increase in segregation increases Black poverty by 2-3 percentage points
 - This represents approximately 10% of the mean Black poverty rate

11.2 Implications

These findings have important policy implications:

1. **Housing policy:** Policies that reduce residential segregation (e.g., inclusionary zoning, fair housing enforcement) may reduce racial inequality in poverty rates
2. **Urban planning:** City layout and infrastructure decisions have long-lasting effects on social outcomes
3. **Understanding inequality:** The persistence of racial inequality partly reflects the legacy of historical segregation patterns

11.3 Methodological Contributions

This project demonstrates:

1. The value of **historical instruments** for causal inference in settings where randomized experiments are infeasible
2. The importance of **multiple estimation methods**: Consistency across IV, IPW, and doubly robust methods strengthens causal claims
3. The **robustness** of instrumental variables: When properly constructed, IV can provide credible causal estimates even with strong assumptions

11.4 Future Research Directions

1. Examine mechanisms: How does segregation affect poverty? (Education, labor markets, social networks?)
2. Dynamic effects: How do segregation’s effects evolve over time?
3. Heterogeneous effects: Do effects vary by city size, region, or time period?
4. Other outcomes: How does segregation affect health, education, crime?

12 Appendix: Stata Code Comparison

For reference, the original analysis was conducted in Stata. Our R implementation replicates the key Stata commands:

Stata:

```
* First stage
reg dism1990 herf lenper, robust

* Second stage (IV)
ivreg povrate_b (dism1990 = herf) lenper, robust
```

R equivalent:

```
# First stage
first_stage <- lm(dism1990 ~ herf + lenper, data = aej_maindata)
coeftest(first_stage, vcov = vcovHC(first_stage, type = "HC1"))

# Second stage (IV)
iv_model <- ivreg(povrate_b ~ dism1990 + lenper | herf + lenper,
                  data = aej_maindata)
summary(iv_model, vcov = sandwich)
```

The `ivreg` function in R's AER package automatically performs 2SLS estimation, equivalent to Stata's `ivreg` command.

13 References

1. Ananat, Elizabeth O. (2011). “The Wrong Side(s) of the Tracks: The Causal Effects of Racial Segregation on Urban Poverty and Inequality.” *American Economic Journal: Applied Economics*, 3(2): 34-66.

14 Session Information

```
sessionInfo()
```

```
## R version 4.5.2 (2025-10-31 ucrt)
## Platform: x86_64-w64-mingw32/x64
## Running under: Windows 11 x64 (build 26100)
##
## Matrix products: default
##   LAPACK version 3.12.1
##
## locale:
## [1] LC_COLLATE=English_United States.utf8
## [2] LC_CTYPE=English_United States.utf8
## [3] LC_MONETARY=English_United States.utf8
## [4] LC_NUMERIC=C
## [5] LC_TIME=English_United States.utf8
##
## time zone: America/Los_Angeles
## tzcode source: internal
##
## attached base packages:
## [1] stats      graphics  grDevices  utils      datasets  methods   base
##
## other attached packages:
## [1] kableExtra_1.4.0 knitr_1.50      ggplot2_4.0.1   stargazer_5.2.3
## [5] AER_1.2-15        survival_3.8-3  car_3.1-3        carData_3.0-5
## [9] sandwich_3.1-1    lmtest_0.9-40  zoo_1.8-14       tidyr_1.3.1
## [13] dplyr_1.1.4       haven_2.5.5
##
## loaded via a namespace (and not attached):
## [1] generics_0.1.4      xml2_1.4.0       stringi_1.8.7     lattice_0.22-7
## [5] hms_1.1.3           digest_0.6.37    magrittr_2.0.4     evaluate_1.0.5
## [9] grid_4.5.2          RColorBrewer_1.1-3 fastmap_1.2.0      Matrix_1.7-4
## [13] Formula_1.2-5       tinytex_0.57     mgcv_1.9-3         purrr_1.1.0
## [17] viridisLite_0.4.2   scales_1.4.0     textshaping_1.0.3  abind_1.4-8
## [21] cli_3.6.5           rlang_1.1.6      splines_4.5.2      withr_3.0.2
## [25] yaml_2.3.10         tools_4.5.2      tzdb_0.5.0         forcats_1.0.0
## [29] vctrs_0.6.5         R6_2.6.1         lifecycle_1.0.4    stringr_1.6.0
## [33] pkgconfig_2.0.3     pillar_1.11.1    gtable_0.3.6       glue_1.8.0
## [37] systemfonts_1.3.1   xfun_0.53        tibble_3.3.0       tidyselect_1.2.1
## [41] rstudioapi_0.17.1   farver_2.1.2     nlme_3.1-168       htmltools_0.5.8.1
## [45] labeling_0.4.3      svglite_2.2.2    rmarkdown_2.29     readr_2.1.5
## [49] compiler_4.5.2      S7_0.2.1
```