Bayesian Updates

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Chapter 1

Preliminaries

1.1 Circular Distributions

I'm anticipating that I might need to put more words into this later on, so am leaving space for them here.

1.1.1 Von Mises distribution

The Von-Mises distribution is given by:

$$f(x, \mu, \kappa) = \frac{1}{2\pi I_0 \kappa} \exp(\kappa \cos(x - \mu)), \quad -\pi \le x \le \pi,$$

where $I_0(\cdot)$ is the 0th modified Bessel function, where the nth modified Bessel function is given by

$$I_n(\kappa) = \frac{1}{\pi} \int_0^{\pi} \cos(n\theta) \exp(\kappa \cos \theta) d\theta.$$

The circular mean of the Von-Mises distribution is given by:

$$\mathbb{E}\left[\exp i\theta\right] = \frac{I_1\left(\kappa\right)}{I_0\left(\kappa\right)}e^{i\mu}.$$

In general, this can be seen via

$$\mathbb{E}[e^{in\theta}] = \frac{1}{2\pi I_0(\kappa)} \int_{-\pi}^{\pi} \exp(in\theta) \exp(\kappa \cos(\theta - \mu)) d\theta$$

$$= \frac{1}{2\pi I_0(\kappa)} \int_{-\pi - \mu}^{\pi - \mu} \exp(in(\psi + \mu)) \exp(\kappa \cos \psi) d\psi$$

$$= \frac{e^{in\mu}}{2\pi I_0(\kappa)} \int_{-\pi}^{\pi} \exp(in\theta) \exp(\kappa \cos \theta) d\theta$$

$$= \frac{e^{in\mu}}{2\pi I_0(\kappa)} \int_{-\pi}^{\pi} (\cos(n\theta) + i \sin(n\theta)) \exp(\kappa \cos \theta) d\theta$$

$$= \frac{e^{in\mu}}{2\pi I_0(\kappa)} \int_{-\pi}^{\pi} \cos(n\theta) \exp(\kappa \cos \theta) d\theta$$

$$= \frac{I_{|n|}(\kappa)}{I_0(\kappa)} e^{in\mu}.$$

Note that we remove the sin integral by using the fact that the integral of an odd function over a symmetric, periodic interval is 0.

Chapter 2

Problem Statement

2.1 Setup

Goal: Given a single measurement of a Bernoulli random variable and a Von-Mises prior distribution, calculate the posterior distribution and approximate to a Von-Mises distribution.

- \bullet t time step
- ullet d_t Grover depth of quantum circuit at time t
- \bullet Y_t random variable representing a single shot measurement y_t of the quantum circuit at time t
- $\Pi(\theta|Y_1 = y_1, \dots, Y_t = y_t) = \Pi(\theta|\mathbf{Y}_t)$ 'true' posterior at time t (though values for t' < t have been used to approximate the earlier distributions)
- $\hat{\Pi}(\theta|Y_1 = y_1, \dots, Y_t = y_t) = \hat{\Pi}(\theta|\mathbf{Y}_t)$ approximate posterior at time t.

According to Bayes rule:

$$\Pi(\theta|Y_t = y_t, \mathbf{Y}_{t-1}) = \frac{\Pi(Y_t = y_t|\theta)\Pi(\theta|\mathbf{Y}_{t-1})}{\Pi(Y_t = y_t)},$$

so we need to compute each of the quantities on the RHS.

At time t, we make a measurement y_t of $Y_t \sim \text{Ber}(p_t)$ at a Grover depth of d_t where

$$p_t = \frac{1}{2}(1 - \cos((4d_t + 2)\hat{\mu}_{t-1}).$$

Thus,

$$\Pi(Y_t = y_t | \theta) = \frac{1}{2} (1 + (-1)^{y_t} \cos((4d_t + 2)\hat{\mu}_{t-1})).$$

For convenience, let $\lambda_t = 4d_t + 2$.

To simplify some of the computations, we're going to assert that the posterior follows a Von-Mises distribution after every update, so we calculate the new values $\hat{\mu}_t$, $\hat{\kappa}_t$ and generate our approximate posterior

$$\hat{\Pi}(\theta|\mathbf{Y}_t) \sim VM(\hat{\mu}_t, \hat{\kappa}_t).$$

2.2 Single shot updates

For simplicity, we're going to consider the first step of the update, which makes things a lot nicer. In this case, we want to know what the circular mean of the posterior distribution is after updating.

- $\Pi(\theta) \sim VM(\mu, \kappa)$ prior
- $\Pi(Y|\theta) \sim \operatorname{Ber}(\frac{1}{2}(1-\cos(\lambda\theta)))$

This gives us:

$$\begin{split} \Pi(Y=y) &= \int_{-\pi}^{\pi} \Pi(Y=y|\theta)\Pi\left(\theta\right) \mathrm{d}\theta \\ &= \frac{1}{2\pi I_0(\kappa)} \int_{-\pi}^{\pi} \frac{1}{2} (1 + (-1)^y \cos(\lambda \theta)) \exp(\kappa \cos(\theta - \mu)) \mathrm{d}\theta \\ &= \frac{1}{2\pi I_0(\kappa)} \left(\int_{-\pi}^{\pi} \frac{1}{2} \exp(\kappa \cos(\theta - \mu)) \mathrm{d}\theta \right. \\ &+ (-1)^y \int_{-\pi}^{\pi} \frac{1}{2} \cos(\lambda \theta) \exp(\kappa \cos(\theta - \mu)) \mathrm{d}\theta \right) \\ &= \frac{1}{4\pi I_0(\kappa)} \left(2\pi I_0(\kappa) + (-1)^y \int_{-\pi}^{\pi} \frac{e^{i\lambda \theta} + e^{-i\lambda \theta}}{2} \exp\left(\kappa \cos\left(\theta - \mu\right)\right) \mathrm{d}\theta \right) \\ &= \frac{1}{2} \left(1 + (-1)^y \cos(\lambda \mu) \frac{I_{\lambda}(\kappa)}{I_0(\kappa)} \right) \end{split}$$

where in the penultimate line, we use the expression for the nth circular moment. Putting this all together, and letting

$$C(y,\lambda,\mu,\kappa) = \frac{\frac{1}{2} \frac{1}{2\pi I_0(\kappa)}}{\frac{1}{2} (1 + (-1)^y \cos(\lambda \mu) \frac{I_{\lambda}(\kappa)}{I_0(\kappa)})} = \frac{1}{2\pi (I_0(\kappa) + (-1)^y \cos(\lambda \mu) I_{\lambda}(\kappa))}$$

gives

$$\begin{split} \mathbb{E}[e^{i\theta}|Y=y] &= C(y,\lambda,\mu,\kappa) \int_{-\pi}^{\pi} e^{i\theta} (1+(-1)^y \cos(\lambda\theta)) \exp(\kappa \cos(\theta-\mu)) \mathrm{d}\theta \\ &= C(y,\lambda,\mu,\kappa) \left(\int_{-\pi}^{\pi} e^{i\theta} \exp(\kappa \cos(\theta-\mu)) \mathrm{d}\theta \right. \\ &\quad + (-1)^y \int_{-\pi}^{\pi} e^{i\theta} \cos(\lambda\theta) \exp(\kappa \cos(\theta-\mu)) \mathrm{d}\theta \right) \\ &= C(y,\lambda,\mu,\kappa) \left(2\pi I_1(\kappa) e^{i\mu} \right. \\ &\quad + (-1)^y \int_{-\pi}^{\pi} e^{i\theta} \left(\frac{e^{i\lambda\theta} + e^{-i\lambda\theta}}{2} \right) \exp(\kappa \cos(\theta-\mu)) \mathrm{d}\theta \right) \\ &= 2\pi C(y,\lambda,\mu,\kappa) \left(I_1(\kappa) e^{i\mu} + \frac{(-1)^y}{2} \left(I_{\lambda+1}(\kappa) e^{i(\lambda+1)\mu} + I_{\lambda-1}(\kappa) e^{-i(\lambda-1)\mu} \right) \right) \end{split}$$

where in the penultimate line, we use the fact that

$$\int_{-\pi}^{\pi} e^{in\theta} \exp(\kappa \cos(\theta - \mu)) d\theta = 2\pi I_0(\kappa) \mathbb{E}[e^{in\theta}] = I_{|n|}(\kappa) e^{in\mu}.$$

This gives us that

$$\mathbb{E}[e^{i\theta}|Y=y] = \frac{I_1(\kappa)e^{i\mu} + \frac{(-1)^y}{2}\left(I_{\lambda+1}(\kappa)e^{i(\lambda+1)\mu} + I_{\lambda-1}(\kappa)e^{-i(\lambda-1)\mu}\right)}{I_0(\kappa) + (-1)^y\cos(\lambda\mu)I_{\lambda}(\kappa)}.$$

If we then take expectations over Y (i.e. multiply by $\Pi(Y=y)$) and sum) this gives us

$$\mathbb{E}[e^{i\theta}] = \frac{I_1(\kappa)}{I_0(\kappa)} e^{i\mu}.$$

So, we can infer that we do not expect the angular parameter μ to move. To infer something about κ , we need to consider $\mathbb{E}[R]$. As before, let's consider $\mathbb{E}[R|Y=y]$. From the above, we can deduce that

$$\begin{split} \mathbb{E}[R|Y=y]^2 &= \frac{\left(I_1 + \frac{(-1)^y}{2} \left(e^{i\lambda\mu}I_{\lambda+1} + e^{-i\lambda\mu}I_{\lambda-1}\right)\right) \left(I_1 + \frac{(-1)^y}{2} \left(e^{-i\lambda\mu}I_{\lambda+1} + e^{i\lambda\mu}I_{\lambda-1}\right)\right)}{\left(I_0 + (-1)^y \cos(\lambda\mu)I_{\lambda}\right)^2} \\ &= \frac{I_1^2 + \frac{1}{4} \left(I_{\lambda+1}^2 + I_{\lambda-1}^2\right) + \frac{1}{2} \cos\left(2\lambda\mu\right) I_{\lambda+1}I_{\lambda-1} + (-1)^y I_1 \left(I_{\lambda+1} + I_{\lambda-1}\right) \cos\lambda\mu}{\left(I_0 + (-1)^y \cos(\lambda\mu)I_{\lambda}\right)^2} \\ &= \frac{N_y^2}{\left(I_0 + (-1)^y \cos(\lambda\mu)I_{\lambda}\right)^2}, \end{split}$$

where for brevity, we have suppressed the argument κ for each of the Bessel functions I_{ν} .

Calculating $\mathbb{E}[R]$ then, is achieved by multiplying by $\Pi(Y=y)$, square-rooting, and summing. This results in the sum of the square roots of the numerators multiplied by a constant factor of $\frac{1}{2I_0(\kappa)}$, i.e.

$$\begin{split} \mathbb{E}[R] &= \frac{N_0 + N_1}{2I_0(\kappa)} \\ &= \frac{1}{2I_0(\kappa)} \sqrt{I_1^2 + \frac{1}{4} \left(I_{\lambda+1}^2 + I_{\lambda-1}^2\right) + \frac{1}{2} \cos(2\lambda\mu) I_{\lambda+1} I_{\lambda-1} + I_1 \left(I_{\lambda+1} + I_{\lambda-1}\right) \cos(\lambda\mu)} \\ &+ \frac{1}{2I_0(\kappa)} \sqrt{I_1^2 + \frac{1}{4} \left(I_{\lambda+1}^2 + I_{\lambda-1}^2\right) + \frac{1}{2} \cos(2\lambda\mu) I_{\lambda+1} I_{\lambda-1} - I_1 \left(I_{\lambda+1} + I_{\lambda-1}\right) \cos(\lambda\mu)} \end{split}$$

2.3 Single shot updates - Gaussian

Now we're going to assume a different prior and posterior:

•
$$\Pi(\theta) \sim N(\mu, \sigma^2)$$
 - prior

•
$$\Pi(Y = y|\theta) \sim \operatorname{Ber}(\lambda) = \frac{1}{2}(1 + (-1)^y \cos(\lambda \theta))$$

Bayes rule, again, states that

$$\Pi'(\theta|Y=y) = \frac{\Pi(\theta)\mathcal{L}(y,\theta)}{\Pi(Y=y)}.$$
(2.1)

We assume that

$$\Pi(\theta) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}},\tag{2.2}$$

we know that

$$\mathcal{L}(y,\theta) = \frac{1}{2}(1 + (-1)^y \cos(\lambda \theta)), \tag{2.3}$$

and by definition

$$\Pi(Y = y) = \int \mathcal{L}(y, \theta) \Pi(\theta) d\theta.$$
 (2.4)

Let us now define the bias to be

$$\Lambda(\theta) = 2\mathcal{L}(0, \theta) - 1, \tag{2.5}$$

which gives the likelihood as

$$\mathcal{L}(y,\theta) = \frac{1}{2}(1 + (-1)^y \Lambda(\theta)). \tag{2.6}$$

Recognise that in this case $\Lambda(\theta) = \cos(\lambda \theta)$.

Let us define expected bias b and the chi function χ as

$$b = \int \Pi(\theta) \Lambda(\theta) d\theta \tag{2.7}$$

$$\chi = \frac{1}{\sigma^2} \int (\theta - \mu) \Pi(\theta) \Lambda(\theta) d\theta$$
 (2.8)

Now, putting that in to equation (??) gives

$$\Pi(Y = y) = \frac{1}{2} \left[\int \Pi(\theta) d\theta + (-1)^y \int \Pi(\theta) \Lambda(\theta) d\theta \right]. \tag{2.9}$$

The first part equals 1 by definition. The second part is the expected bias, also by definition, i.e.,

$$\Pi(Y=y) = \frac{1}{2} \Big[1 + (-1)^y b \Big]. \tag{2.10}$$

The expected bias b is given by

$$b = \int \Pi(\theta) \Lambda(\theta) d\theta \tag{2.11}$$

$$= \int \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}} \cos(\lambda\theta) d\theta, \qquad (2.12)$$

which apparently (Koh paper, equation 153 on page 25) is an 'identity':

$$b = e^{-\frac{1}{2}\lambda^2\sigma^2}\cos(\lambda\mu). \tag{2.13}$$

Putting this all together gives for the posterior

$$\Pi'(\theta|Y=y) = \frac{\Pi(\theta)\mathcal{L}(y,\theta)}{\Pi(Y=y)}$$
(2.14)

$$= \frac{1}{4\sqrt{2\pi\sigma^2}} \frac{e^{-\frac{1}{2}(\frac{\theta-\mu}{\sigma})^2} (1 + (-1)^y \cos(\lambda\theta))}{1 + (-1)^y e^{-\frac{1}{2}\lambda^2\sigma^2} \cos(\lambda\mu)}$$
(2.15)

Appendix A

Integrals

A.1 Normal Distribution

First, let us note that

$$\frac{\mathrm{d}}{\mathrm{d}x}(e^{-x^2}) = -2xe^{-x^2}.$$

The integrals we are interested in computing, are either of the form

$$I_{2n}(k) = \int_{-\infty}^{\infty} \theta^{2n} \cos(k\theta) e^{-\theta^2} d\theta \text{ or } I_{2n+1} = \int_{-\infty}^{\infty} \theta^{2n+1} \sin(k\theta) e^{-\theta^2} d\theta$$