

$$\textcircled{1}. Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

$$\begin{aligned} a) E[\hat{\beta}_0] &= E[\bar{Y} - \hat{\beta}_1 \bar{x}] = E[\bar{Y}] - \bar{x} E[\hat{\beta}_1] \\ &= \beta_0 + \beta_1 \bar{x} - \beta_1 \bar{x} \\ &= \beta_0 \end{aligned}$$

$$b) L(\beta_0, \beta_1, \sigma^2 | (x_i, y_i)_{i=1}^n) = \prod \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2} (y_i - x_i \beta_1 - \beta_0)^2\right\}$$

$$\ell(\sigma^2) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2$$

$$\frac{\partial \ell}{\partial \sigma^2} = -\frac{n}{2} \cdot \frac{1}{\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum (y_i - \beta_0 - \beta_1 x_i)^2$$

⋮

$$\hat{\sigma}^2 = \frac{\sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2}{n} = \frac{SSE}{n}$$

$$\begin{aligned} c.) E[SSTot] &= E\left[\sum_i (y_i - \bar{y})^2\right] \\ &= \sum_i E(y_i - \bar{y})^2 \\ &= \sum_i \left[\text{Var}(y_i - \bar{y}) + (E(y_i - \bar{y}))^2 \right] \\ &= \sum_i \left[\sigma^2 \left(\frac{n-1}{n}\right) + \beta_1^2 (x_i - \bar{x})^2 \right] \end{aligned}$$

$$E[SSTot] = (n-1)\sigma^2 + \beta_1^2 \sum_i (x_i - \bar{x})^2$$

$$\begin{aligned} E[SSR] &= \sum_i E[(\hat{\beta}_0 + \hat{\beta}_1 x_i - \bar{y})^2] = \sum_i E[\hat{\beta}_1^2 (x_i - \bar{x})^2] \\ &= E[\hat{\beta}_1^2] \sum (x_i - \bar{x})^2 = (\text{Var}(\hat{\beta}_1) + (E(\hat{\beta}_1))^2) \sum (x_i - \bar{x})^2 \\ &= \left(\frac{\sigma^2}{\sum (x_i - \bar{x})^2} + \beta_1^2 \right) \sum (x_i - \bar{x})^2 = \sigma^2 + \beta_1^2 \sum (x_i - \bar{x})^2 \end{aligned}$$

$$\Rightarrow E[SSE] = E[SSTot] - E[SSR] = (n-2)\sigma^2.$$

(1. d.) An unbiased est. for σ^2 is $s^2 = \frac{SSE}{n-2}$

e.) Shown in (c).

$$f.) \hat{\mu}(\bar{x}) = \hat{\beta}_0 + \hat{\beta}_1 \bar{x} = \bar{y} - \hat{\beta}_1 \bar{x} + \hat{\beta}_1 \bar{x} \\ = \bar{y}$$

g.) $\hat{\beta}_0$ and $\hat{\beta}_1$ are linear transformations of the y_i and are therefore normally distributed.

(2. a.)

$$L(\beta_1 | (x_i, y_i)) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(y_i - \beta_1 x_i)^2\right\}$$
$$l(\beta_1) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_i (y_i - \beta_1 x_i)^2$$
$$\frac{\partial l}{\partial \beta_1} = -\frac{1}{2\sigma^2} \cdot \left(-2 \sum_i x_i (y_i - \beta_1 x_i)\right) = 0$$
$$\hat{\beta}_1 = \frac{\sum x_i y_i}{\sum x_i^2}$$

b.) The model without β_0 forces the fitted line to have y -intercept = 0.