

MM 225: AI AND DATA SCIENCE

CONTINUOUS CONDITIONAL PROBABILITY, EXPECTED VALUE, AND, VARIANCE

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LECTURE 8: CONTINUOUS CONDITIONAL PROBABILITY, EXPECTED VALUE, AND, VARIANCE



Outline

- 1 Continuous conditional probability
 - Independent events
 - Joint density and cumulative distribution function
 - Independent random variables
 - Independent trials
- 2 Examples
 - Drug effectiveness: beta density
 - Two-armed bandit problem
- 3 Expected value
 - Martingales
- 4 Variance
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CONTINUOUS CONDITIONAL PROBABILITY



Discrete to continuous

- All our discrete conditional probability results: can be carried over to continuous with appropriate modifications
- Probability mass function \rightarrow Probability density function
- If X is a continuous random variable with density function $f(x)$, and if E is an event with positive probability, conditional density $f(x|E)$ is defined as follows:

$$f(x|E) = \begin{cases} \frac{f(x)}{P(E)} & \text{if } x \in E \\ 0 & \text{if } x \notin E \end{cases}$$

- For any event F , $P(F|E) = \int_F f(x|E)dx$ is the conditional probability of F given E



Independent events

- E and F : two events with positive probability in a continuous sample space
- E and F are independent if $P(E|F) = P(E)$; $P(F|E) = P(F)$; $P(E \cap F) = P(E)P(F)$



Definition

Definition

$\mathbf{X} = (X_1, X_2, \dots, X_n)$ where X_i are continuous random variables associated with an experiment. The joint cumulative distribution function of \mathbf{X} is defined by

$$F(x_1, x_2, \dots, x_n) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n)$$

Remark

$$f(x_1, x_2, \dots, x_n) = \frac{\partial^n F(x_1, x_2, \dots, x_n)}{\partial x_1 \partial x_2 \dots \partial x_n}$$

Or,

Remark

$$F(x_1, x_2, \dots, x_n) = \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} \dots \int_{-\infty}^{x_n} f(t_1, t_2, \dots, t_n) dt_n dt_{n-1} \dots dt_1$$



Definition

Definition

Let X_1, X_2, \dots, X_n be continuous random variables with cumulative distribution functions $F_1(x), F_2(x), \dots, F_n(x)$. Then, if for any choice of x_1, x_2, \dots, x_n , $F(x_1, x_2, \dots, x_n) = F(x_1)F(x_2) \cdots F(x_n)$, then these random variables are mutually independent.



Theorems

Theorem

Let X_1, X_2, \dots, X_n be continuous random variables with density functions $f_1(x), f_2(x), \dots, f_n(x)$. Then, for any choice of x_1, x_2, \dots, x_n , $f(x_1, x_2, \dots, x_n) = f(x_1)f(x_2) \cdots f(x_n)$ is the necessary and sufficient condition for these random variables to be mutually independent.

Theorem

Let X_1, X_2, \dots, X_n be mutually independent continuous random variables and let $\phi_1(x), \phi_2(x), \dots, \phi_n(x)$ be continuous functions. Then, $\phi_1(X_1), \phi_2(X_2), \dots, \phi_n(X_n)$ are mutually independent.

A sequence X_1, X_2, \dots, X_n of random variables X_i that are mutually independent and have the same density is called an independent trials process.

Recall *iid*: independent, identical distributions!

EXAMPLES



Problem of drug effectiveness

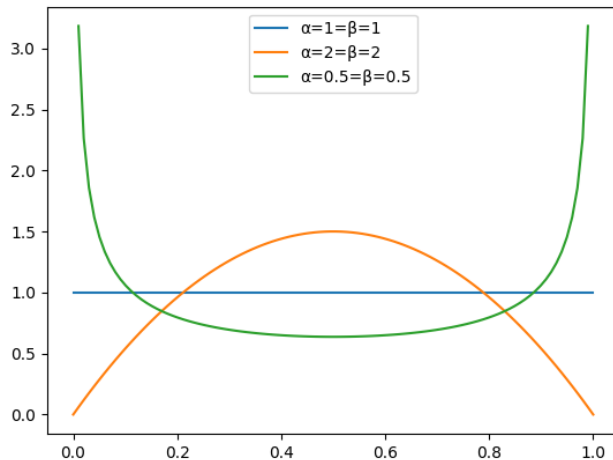
- Assumption: a drug is effective with probability x each time it is used
- Assumption: various trials are independent
- x : a continuous random variable between zero and unity
- What density function describes the distribution for x ?
- Assume beta density for x before experiment
- Beta density: sample space a combination of discrete and continuous coordinates
- $B(\alpha, \beta) = \int_0^1 x^{\alpha-1}(1-x)^{\beta-1}dx$ where α, β are integers

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$$B(\alpha, \beta, x) = \begin{cases} \frac{x^{\alpha-1}(1-x)^{\beta-1}}{B(\alpha, \beta)} & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$



Beta: density function





Experiment

- x before experiment; described by beta density
- Give drugs to n subjects and record the number of successes; say i
- i : discrete random variable
- Set of possible outcomes of experiment: ordered pair (x, i)
- $m(i|x) = \binom{n}{i} x^i (1-x)^{n-i}$
- $f(x, i) = m(i|x) B(\alpha, \beta, x)$
- If i successes are observed, the probability density that the drug is effective is $B(\alpha + i, \beta + n - i)$; if the drug is now tested on the next subject, the probability that the drug is effective is $\frac{\alpha + i}{\alpha + \beta + n}$
- For large n , probability of success after experiment \approx the proportion of successes in the experiment
- Suppose there are two drugs: how to choose between them?



Two-armed bandit problem

Suppose in a Casino, there are two slot machines. Each machine either gives 1 dollar or nothing. The probabilities for the first and second machines to give 1 dollar are x and y , respectively. Suppose you play the game ten times, each time choosing one or the other machine. Suppose you want to maximize the number of times you win, what should be the strategy of play?



Strategies

- **Play the best machine**

At every stage, calculate the probability of pay off for each machine and choose the machine with the highest probability.

- Weakness: chance might make you continue playing with the machine with lower return

- **Play the winner**

Play the same machine when you win but switch machine when you lose

- **Homework:** Simulate the strategies

- **Importance:** Two armed bandit: important in clinical trials!

EXPECTED VALUE



A couple of problems

- Suppose a die is rolled. If odd number turns up, G wins an amount equal to the number. If even number turns up, G loses an amount equal to the number. Is this a reasonable game for G to play?
- Let us generate a number from the interval $[0,1]$. If you generate a large number of such random numbers, what is the average value of the generated numbers?



Definitions

Definition

Let X be a discrete random variable with sample space Ω and distribution function $m(x)$. The expected value $E(X) = \sum_{x \in \Omega} xm(x)$ provided this sum converges absolutely. We typically refer to $E(X)$ as the mean and denote it by μ . X does not have an expected value if there is no convergence.

Definition

Let X be a real valued random variable with density function $f(x)$. The expected value $\mu = E(X) = \int_{-\infty}^{\infty} xf(x)dx$ provided the integral $\int_{-\infty}^{\infty} |x|f(x)dx$ is finite.



Cauchy distribution

- Consider the random variable X with density
$$f_X(x) = \frac{a}{\pi} \frac{1}{a^2 + x^2}$$
- No expectation: integral diverges
- No variance either!



Properties

Theorem

Let X_1, X_2, \dots, X_n be random variables with finite expected values (discrete or continuous). Let c_1, c_2, \dots, c_n be constants. Then,

$$E(c_1X_1 + c_2X_2 + \dots + c_nX_n) = c_1E(X_1) + c_2E(X_2) + \dots + c_nE(X_n).$$

Remark

First fundamental mystery of probability: Expectations add irrespective of whether the summands are mutually independent or not.



Theorems

Theorem

If X is a discrete random variable with sample space Ω and distribution function $m(x)$, and if $\phi : \Omega \rightarrow R$ is a function, then, $E(\phi(X)) = \sum_{x \in \Omega} \phi(x)m(x)$ provided the series converges absolutely.

Theorem

If X is a real-valued random variable and if $\phi : \Omega \rightarrow R$ is a continuous real-valued function with domain $[a,b]$, then, $E(\phi(X)) = \int_{-\infty}^{\infty} \phi(x)f_X(x)dx$ provided the integral exists.



Independence

Theorem

If X_1, X_2, \dots, X_n are n mutually independent random variables (discrete or continuous) with expectation values $E(X_i)$, then, $E(X_1 X_2 \dots X_n) = E(X_1) E(X_2) \dots E(X_n)$

A problem: Let $Z = (X, Y)$ be a point chosen at random in the unit square. Let $A = X^2$ and $B = Y^2$. What is $E(AB)$?

- A and B are independent. Why? If random variables are independent, their functions are also independent.
- So, the answer is easy using the above theorem! $(\frac{1}{3} \cdot \frac{1}{3} = \frac{1}{9})$.
- Calculation from definition: Complicated. $f(AB) = -\frac{\log(t)}{4\sqrt{t}}$



Conditional expectation

Definition

If F is any event and X is a random variable with sample space $\Omega = \{x_1, x_2, \dots\}$, then the conditional expectation given F is $E(X|F) = \sum_j x_j P(X = x_j|F)$.

Theorem

Let X be a random variable with sample space Ω . If F_1, F_2, \dots, F_r are events such that $F_i \cap F_j = \emptyset$ if $i \neq j$ and $\Omega = \cup_j F_j$, then, $E(X) = \sum_j E(X|F_j)P(F_j)$.

MARTINGALES



Coin toss problem

- **Recall the problem:**

Suppose we toss a coin. Assume the coin is fair. If head (H) comes up, G(uru) gets Re. 1 and H(ina) loses Re. 1; if tail (T) comes up, H gets Re. 1 and G loses Re. 1. Suppose the coin is tossed 40 times.

- Let S_1, S_2, \dots, S_n be G's accumulated fortune in playing this game.
- $E(S_n | S_{n-1} = a, \dots, S_1 = r) = \frac{1}{2}(a + 1) + \frac{1}{2}(a - 1) = a$
- G's expected fortune in the next play is equal to his present fortune: game is *fair*
- Martingale: a fair game
- If the coin is not fair, $E(S_n | S_{n-1} = a, \dots, S_1 = r) = p(a + 1) + q(a - 1) = a + p - q$
- If $p < q$ ($p > q$), game is unfavourable (favourable)
- Can we make a fair game, a favourable game? No
- Martingale doubling system : tutorial applying this to stock market
- Efforts to make unfavourable game, favourable!

VARIANCE



Definitions

Definition

Let X be a discrete random variable with expected value $E(X) = \mu$. Then, the variance of X is $V(X) = E((X - \mu)^2)$. Further, $V(x) = \sum_x (x - \mu)^2 m(x) = E(X^2) - \mu^2$ where m is the distribution function of X .

Definition

Let X be a real valued random variable with density function $f(x)$. The variance $\sigma^2 = V(X) = E((X - \mu)^2)$. Further, $V(x) = \sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx = E(X^2) - \mu^2$ where $\mu = E(X)$.

Remark

Standard deviation: $\sigma = \sqrt{V(X)}$



Properties

- If X is any random variable, and c is any constant, $V(cX) = c^2 V(X)$ and $V(X + c) = V(X)$.
- If X and Y are two independent random variables, $V(X + Y) = V(X) + V(Y)$



Standardisation

Theorem

Let X_1, X_2, \dots, X_n be an independent trials process with $E(X) = \mu$ and $V(x) = \sigma^2$. Let $S_n = X_1 + X_2 + \dots + X_n$ be the sum and $A_n = \frac{S_n}{n}$ be the average. Then, $E(S_n) = n\mu$, $V(S_n) = n\sigma^2$, $E(A_n) = \mu$ and $V(A_n) = \frac{\sigma^2}{n}$.

Remark

Standard deviation of the average goes to 0 as $n \rightarrow \infty$.

Remark

If we set $S_n^* = \frac{S_n - n\mu}{\sqrt{n\sigma^2}}$, then, $E(S_n^*) = 0$ and $V(S_n^*) = 1$.

THANK YOU!!!