

1) THEORETICAL BACKGROUND

Multivariate Gaussian model (this is a key assumption for the DAGs)

$$\underline{x}_1, \dots, \underline{x}_m \stackrel{iid}{\sim} N_q(\underline{\Omega}, \Omega^{-1}) \quad \underline{x}_i \in \mathbb{R}^q$$

We'll do inference on this.

Ω s.p.d. = Precision matrix

$$\Omega \stackrel{iid}{\sim} W_{\text{Wishart}}_q(a, U) \quad \text{where } a > q-1$$

• U s.p.d.

$$p(\Omega) = C(a, U) \cdot |\Omega|^{\frac{a-q-1}{2}} \cdot \exp \left\{ -\frac{1}{2} \text{tr}(U\Omega) \right\}$$

$$p(\underline{x}_1, \dots, \underline{x}_m | \Omega) = \dots = (2\pi)^{-\frac{mq}{2}} \cdot |\Omega|^{\frac{m}{2}} \cdot \exp \left\{ -\frac{1}{2} \text{tr}(S\Omega) \right\}$$

$$\text{where } S = \sum_{i=1}^m \underline{x}_i \cdot \underline{x}_i^T = \mathbf{X}^T \mathbf{X} \in \mathbb{R}^{q \times q}$$

where \mathbf{X} is our data matrix $m \times q$

$$\begin{aligned} \text{posterior} & \quad \text{prior} \\ p(\Omega | \underline{x}_1, \dots, \underline{x}_m) & \propto p(\underline{x}_1, \dots, \underline{x}_m | \Omega) \cdot p(\Omega) \\ & \propto |\Omega|^{\frac{a+m-q-1}{2}} \cdot \exp \left\{ -\frac{1}{2} \text{tr}[(U+S)\Omega] \right\} \end{aligned}$$

$$\Rightarrow p(\Omega | \underline{x}_1, \dots, \underline{x}_m) \sim W_q(a+m, U+S)$$

the posterior is just another Wishart with modified parameters:

- $a + \text{sample size}$
- $U + \mathbf{X}^T \mathbf{X}$

$$C(a, U) = \frac{|U|^{\frac{a}{2}}}{\frac{a}{2}^{\frac{a}{2}} \cdot \Gamma_q(\frac{a}{2})} \quad \text{prior normalizing constant}$$

$$C(a+m, U+S) = \frac{|U+S|^{\frac{a+m}{2}}}{2^{\frac{(a+m)q}{2}} \cdot \Gamma_q(\frac{a+m}{2})} \quad \text{posterior normalizing constant}$$

Model selection through marginal likelihoods:

$$P(\underline{x}_1, \dots, \underline{x}_m) = \int p(\underline{x}_1, \dots, \underline{x}_m | \Omega) \cdot p(\Omega) d\Omega = \dots = \\ = (2\pi)^{-\frac{m\eta}{2}} \cdot \frac{C(a, U)}{C(a+m, U+S)} = p(x)$$

Marginal distributions:

$A \subseteq \{1, \dots, q\}$ A any subset of index of the var.

I would like to have the marginal likelihood of X_A data matrix with only the variables of columns A

$$p(X_A) ?$$

$$\underline{x}_1^A, \dots, \underline{x}_m^A \sim ?$$

$$\underline{x}_1^A, \dots, \underline{x}_m^A \mid ? \sim N_{IA_1}(\underline{\Omega}, (\underline{\Omega}_{A\bar{A}})^{-1}) \quad (\star)$$

Schur complement of Ω :

$$\underline{\Omega}_{A\bar{A}} = \underline{\Omega}_{AA} - \underline{\Omega}_{A\bar{A}} (\underline{\Omega}_{\bar{A}\bar{A}})^{-1} \underline{\Omega}_{\bar{A}A} = (\underline{\Sigma}_{AA})^{-1}$$

where: $\bar{A} = \{1, \dots, q\} \setminus A$ (all columns not in A)

$\underline{\Omega}_{AA}$ is block of Ω with rows of A
and columns of A

$$\underline{\Omega}_{A\bar{A}} \sim W_{IA_1}(a - |\bar{A}|, U_{AA}) \quad (\star\star)$$

Note: since we have the prior on Ω , I'm doing all the computation using Ω instead of $\Sigma = \Omega^{-1}$

(★) + (★★) \implies we can proceed as before:

$$\Omega_{A|\bar{A}} \mid \underline{x}_1^A, \dots, \underline{x}_m^A \sim W_{|A|}(\alpha+m-|\bar{A}|, U_{AA} + S_{AA})$$

$$\Rightarrow p(\underline{x}_1^A, \dots, \underline{x}_m^A) = p(x^A) =$$

$$= \underbrace{\int p(\underline{x}_1^A, \dots, \underline{x}_m^A \mid \Omega_{A|\bar{A}}) \cdot p(\Omega_{A|\bar{A}}) d\Omega_{A|\bar{A}}}_{\text{likelihood } N_{|A|}} =$$

$$\sim W_{|A|}$$

$$= (2\pi)^{-\frac{m|\bar{A}|}{2}} \cdot \frac{c(\alpha - |\bar{A}|, U_{AA})}{c(\alpha + m - |\bar{A}|, U_{AA} + S_{AA})} =: p(x_A)$$



$$\text{where } c(\alpha - |\bar{A}|, U_{AA}) = \frac{|U_{AA}|^{\frac{|\bar{A}|}{2}}}{2^{\frac{(\alpha - |\bar{A}|) \cdot |\bar{A}|}{2}} \cdot \prod_{|\bar{A}|} \left(\frac{\alpha - |\bar{A}|}{2} \right)},$$

$$c(\alpha + m - |\bar{A}|, U_{AA} + S_{AA}) = \frac{|U_{AA} + S_{AA}|^{\frac{\alpha + m - |\bar{A}|}{2}}}{2^{\frac{(\alpha + m - |\bar{A}|) \cdot |\bar{A}|}{2}} \cdot \prod_{|\bar{A}|} \left(\frac{\alpha + m - |\bar{A}|}{2} \right)}$$

We want to implement the following algorithm:

Input:

α, U, X



Output:

$p(x_A)$

2) DAG

$$D = (V, E)$$

$V = \{1, \dots, q\} = \text{set of nodes} = \text{variables}$

$E \subseteq V \times V = \text{edges}$

$(u, v) \in E \implies \exists u \rightarrow v \text{ in } D$

$\text{par}_D(j) = \{u : (u, j) \in E\} = \text{parents of } j$

$\text{fan}_D(j) = j \cup \text{par}_D(j) = \text{family of } j$



Under D the joint p.d.f. (since we are in Gauss setting) of x_1, \dots, x_q is:

$$p(x_1, \dots, x_q | D) = \prod_{j=1}^q p(x_j | \text{par}_D(j))$$

\underline{x}_A is sub-vector of \underline{x} with columns in A .

In a Gauss DAG model:

$$x_1, \dots, x_m | \Omega_D \stackrel{\text{iid}}{\sim} N_q(\underline{0}, \Omega_D^{-1})$$

where Ω_D is: - \rightarrow pd (as before)

- Markov wrt D (satisfies the conditional independence of the DAG)

prior of Ω_D ? I need Wishart + conditional dependencies of D through Markov

Multiple methods to find it:

i) DAG - Wishart prior (Car et al., Annals of Statistics)

prior of the modified Choleski decomposition

$$\Omega \rightarrow (D, L)$$

ii) if the goal is to compute $p(x|D)$, i.e. the marginal likelihood of D :

Start from a prior on the parameter of a complete DAG model \Rightarrow (in the Gauss setting (our))

Ω is Wishart: $\Omega \sim W_q(a, U)$,

then:

$$p(x|D) = \prod_{j=1}^q \frac{p(x_{\text{fan}(j)})}{p(x_{\text{par}(j)})} \quad (\star\star)$$

$p(x_{\text{fan}(j)})$ and $p(x_{\text{par}(j)})$ are computed as Δ (Gengen and Heckerman (2002), Annals of Statistics) (Comboni and La Rocca (2012), Scandinavian Journal of Statistics)

Finding parents and family of j :
 (Cartelletti and Mascaro)

Search-and-score method:

Starting D \longrightarrow construct D^* , then
 accept/reject D^* based on:

$$p(\text{accept } D^*) \propto \frac{p(x|D^*)}{p(x|D)} \cdot \frac{p(D^*)}{p(D)} \cdot \frac{q(D|D^*)}{q(D^*|D)}$$

$$\alpha_{D^*, D} = \min \left\{ 1; \frac{p(x|D^*)}{p(x|D)} \cdot \frac{p(D^*)}{p(D)} \cdot \frac{q(D|D^*)}{q(D^*|D)} \right\}$$

D^* accepted with prob. $\alpha_{D^*, D}$



3) PC-algorithm:

$$X_u \perp\!\!\!\perp X_v \mid X_S \quad S = \{1, \dots, q\} \setminus \{u, v\}$$

Until now we know how to compute only through frequentist method. With Gauss assumpt. \Rightarrow test on partial correlation coefficient.

$$X = (x_1, \dots, x_q)$$

$$f_{u,v|S} = \text{corr}(X_u, X_v \mid X_S)$$

$A = \{u, v, S\} \xrightarrow{\text{we compute}} \text{marginal distribution of } X_A$

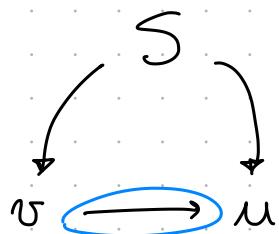
$$X_u, X_v \mid X_S \sim N_2(\dots, \dots)$$

↑
we should use these
to compute $f_{u,v|S}$

But! Testing $f_{u,v|S} = 0$ is the same as

testing $L_{u|v,S} = 0 \leftarrow \text{regression coefficient}$
of u on $\{v, S\}$

Interpretation: $D^{(1)}$:



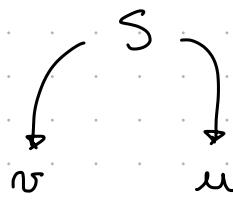
$L_{u|v,S}$ measure how much important the $v \rightarrow u$ arrow.

$$L_{u|v,S} = 0 \implies \text{no arrow}$$

and it's the same as $p(X_u \mid X_{pa(u)})$

Remember! we are in the skeleton estimation phase, therefore I don't need to know the direction of the arrow, I just want to know if $u \perp\!\!\!\perp v | S$ for now.

Let $D^{(0)}$:



Compare, in a Bayesian way, $D^{(1)}$ and $D^{(0)}$ through Bayes-factor:

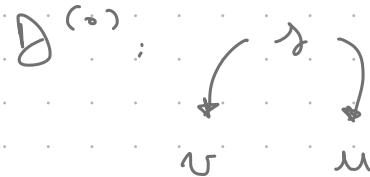
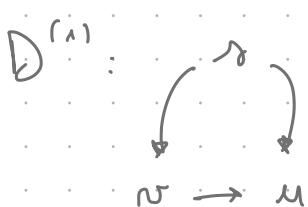
$$\text{BF}_{01} = \frac{m(x | D^{(0)})}{m(x | D^{(1)})}$$

← marginals that I find with the previous formulas

Based on BF_{01} I decide if I include the $v \rightarrow u$ arrow.

$$p(x | D) = \prod_{j=1}^q \frac{p(x_{ja_D(j)})}{p(x_{pa_D(j)})}$$

Notice that since we are in Gauss setting, this should be exactly equal to $p(u \rightarrow v)$



$$\frac{p(x | D^{(0)})}{p(x | D^{(1)})} = \frac{\frac{p(x_{ja_0(u)})}{p(x_{pa_0(u)})}}{\frac{p(x_{ja_1(u)})}{p(x_{pa_1(u)})}}$$



I need only mode u , difference between u, v and $u \rightarrow v$ will emerge by itself

Next step

- i) R function to evaluate $p(x_A)$ (see 1))
- ii) DAG D \rightsquigarrow function to evaluate $p(x|D)$
(using $p(x_A)$) (see 2))
- iii) Modify the pcalg "skeleton" function to use
BF instead of the already implemented pc algorithm
function