

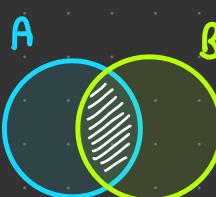
## REMEMBER

$$P(A^c) = 1 - P(A)$$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$



$$P(A \cap B) \stackrel{?}{=} P(A) \cdot P(B) \quad [\text{TO CHECK INDEPENDENCE}]$$



$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

LAW OF TOTAL PROBABILITY:  $P(A \cap B) = P(A|B) \cdot P(B) = P(B|A) P(A)$

$$P(A^c \cap B) = P(B) - P(A \cap B)$$

$$P(A^c | B) = \frac{P(A^c \cap B)}{P(B)} = \frac{P(B) - P(A \cap B)}{P(B)} = 1 - \frac{P(A \cap B)}{P(B)} = 1 - P(A|B)$$

$$\binom{N}{K} \quad N \geq K \quad \binom{N}{K} = \frac{N!}{K!(N-K)!}$$

$$P(B|A) = \frac{P(A \cap B)}{P(B)} = \frac{P(A|B) P(B)}{P(A|B) P(B) + P(A|B^c) P(B^c)}$$

## REMEMBER

$$\text{VAR}(3X) = 3 \text{ VAR}(X)$$

$$\text{COV}(x, y) = \text{COV}(y, x)$$

$$\text{COV}(x, x) = \text{VAR}(x)$$

$$\text{COV}(x, y) = E[xy] - E[x]E[y]$$

IF INDEPENDENT:  $E[xy] = E[x]E[y]$

$$\rho = \frac{\text{COV}(x, y)}{\sqrt{\sigma_x^2 \sigma_y^2}} = \frac{\text{COV}(x, y)}{\sigma_x \sigma_y}$$

$$\text{VAR}(xy) = E[(xy)^2] - (E[xy])^2$$

• VARIANCE IS ALWAYS

• COVARIANCE CAN BE POSITIVE OR NEGATIVE

• (1) OPPOSITE RELATION

• (2) DIRECTLY PROPORTIONAL

• (3) NO CORRELATION

$$\begin{aligned} & X \text{ AND } Y \text{ INDEP.} \Rightarrow \text{COV}(x, y) = 0 \\ & \text{BUT} \\ & \text{COV}(x, y) = 0 \not\Rightarrow X \text{ AND } Y \text{ INDEP.} \end{aligned}$$

## IF INDEPENDENT

$$\bullet E[3x - 2y + 7] = 3E[x] - 2E[y] + 7$$

$$\bullet \text{VAR}(3x - 2y + 7) = \text{VAR}(3x) + \text{VAR}(-2y) \quad \text{THE CONSTANT DISAPPEARS}$$

$$\bullet \text{VAR}(xy) = E[(xy)^2] - (E[xy])^2 = E[x^2]E[y^2] - (E[x])^2(E[y])^2$$

$$\text{VAR}(x) = E[x^2] - (E[x])^2$$

USE THE INVERSE FUNCTION

## DISCRETE

• BERNULLI:  $X \sim \text{BERN}(p)$        $\begin{cases} p & \text{SUCCESS} \\ 1-p & \text{FAILURE} \end{cases}$

• BINOMIAL:  $X \sim \text{BIN}(n, p)$        $P(S_N=k) = \binom{n}{k} p^k (1-p)^{n-k}$

• HYPERGEOMETRIC:  $X \sim \text{HYPERGEO}(N, n, M)$        $P(x=k) = \frac{\binom{N}{k} \binom{M-n}{N-k}}{\binom{M}{n}}$

• POISSON:  $X \sim \text{POISSON}(\lambda)$        $P(x=k) = e^{-\lambda} \frac{\lambda^k}{k!}$

• GEOMETRIC:  $X \sim \text{GEOM}(p)$        $P(x=k) = (1-p)^{k-1} p$

• UNIFORM  $X \sim \text{UNIF}(\alpha, b)$        $P = \frac{1}{b-\alpha} = \frac{1}{b-\alpha}$

• EXPONENTIAL  $X \sim \text{EXP}(\lambda)$       PROB. DENSITY FUNCTION       $\lambda e^{-\lambda t}$

• NORMAL  $X \sim \text{NORM}(\mu, \sigma^2)$       PROB. DENSITY FUNCTION       $\frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$

## CONTINUOUS

## X RANDOM VARIABLE: FIND MEAN

X DISCRETE

x	1	2	3	4	5
$p_X(x)$	1/7	1/14	3/14	2/7	2/7

X CONTINUOUS

$$f(x) = \begin{cases} 3e^{-3x}, & x > 0 \\ 0, & \text{else.} \end{cases}$$

$$E[x] = 1 \cdot \frac{1}{7} + 2 \cdot \frac{1}{14} + 3 \cdot \frac{3}{14} + 4 \cdot \frac{2}{7} + 5 \cdot \frac{2}{7}$$

$$E[x] = \int_{-\infty}^{+\infty} x f(x) dx$$

⚠ THE INTEGRAL "DOMAIN" CAN BE USUALLY REDUCED

## X RANDOM VARIABLE: FIND $E[\text{something} \rightarrow X]$

X DISCRETE

ES FIND  $E[|x-2|]$

x	1	2	3	4	5
$p_X(x)$	1/7	1/14	3/14	2/7	2/7

$$E[|x-2|] = |1-2| \cdot \frac{1}{7} + |2-2| \cdot \frac{1}{14} + |3-2| \cdot \frac{3}{14} + |4-2| \cdot \frac{2}{7} + |5-2| \cdot \frac{2}{7}$$

X CONTINUOUS

ES: FIND  $E[e^{2x}]$

$$f(x) = \begin{cases} 3e^{-3x}, & x > 0 \\ 0, & \text{else.} \end{cases}$$

$$E[e^{2x}] = \int_{-\infty}^{+\infty} e^{2x} f(x) dx = \int_0^{+\infty} e^{2x} e^{-3x} dx$$

## X RANDOM VARIABLE: FIND VARIANCE

X DISCRETE

x	1	2	3	4	5
$p_X(x)$	1/7	1/14	3/14	2/7	2/7

$$\text{VAR}(x) = E[x^2] - (E[x])^2$$

$$E[x^2] = 1 \cdot \frac{1}{7} + 2 \cdot \frac{2}{14} + 3 \cdot \frac{3}{14} + 4 \cdot \frac{2}{7} + 5 \cdot \frac{2}{7}$$

$$E[x] = 1 \cdot \frac{1}{7} + 2 \cdot \frac{2}{14} + 3 \cdot \frac{3}{14} + 4 \cdot \frac{2}{7} + 5 \cdot \frac{2}{7}$$

⚠ REMEMBER TO SQUARE THE RESULT OF  $E[x]$  TO HAVE  $(E[x])^2$  BEFORE DOING THE SUBTRACTION

X CONTINUOUS

$$f(x) = \begin{cases} 3e^{-3x}, & x > 0 \\ 0, & \text{else.} \end{cases}$$

$$\text{VAR}(x) = E[x^2] - (E[x])^2$$

$$E[x^2] = \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_0^{+\infty} x^2 3e^{-3x} dx$$

$$E[x] = \int_{-\infty}^{+\infty} x f(x) dx = \int_0^{+\infty} x 3e^{-3x} dx$$

## EXERCISE WITH A POPULATION SIZE, A %, AND $\epsilon$

BEFORE STARTING YOU NEED TO FIND

- $\epsilon$ : A SMALL NUMBER, IS THE CONFIDENCE INTERVAL
- %. THE %. OF HOW MUCH SURE YOU WANT TO BE

ON: THE SIZE OF THE SAMPLE

YOU HAVE TO FIND ONE OF THESE

## PROCEDURE

$$P(|\hat{p} - p| < \epsilon) = P(-\epsilon < \hat{p} - p < \epsilon) = P\left(\frac{-\epsilon\sqrt{N}}{\sqrt{p(1-p)}} < \bar{S}_N < \frac{\epsilon\sqrt{N}}{\sqrt{p(1-p)}}\right) = 2\Phi\left(\frac{2\epsilon\sqrt{N}}{\sqrt{p(1-p)}}\right) - 1 \geq 0.95 \Leftrightarrow \text{IS THE %.}$$

$$\epsilon = 0.02$$

$$\%. = 95$$

$$N = ? \quad \& \quad \text{FIND THIS}$$

⚠ DELETE THE DENOMINATOR  $\sqrt{p(1-p)}$  AND ADD 2 IN FRONT  
 !!! DON'T FORGET IT!!!

NOW REWRITE S.T.  $\Phi$  IS ALONE:

$$\Phi(2\epsilon\sqrt{N}) \geq 0.975 \Rightarrow 2\epsilon\sqrt{N} = 1.96 \Rightarrow N = \left(\frac{1.96}{2\epsilon}\right)^2 \Rightarrow N = 2401 \quad \text{Ü}$$

Exercise 4.6. A pollster would like to estimate the fraction  $p$  of people in a population who intend to vote for a particular candidate. How large must a random sample be in order to be at least 95% certain that the fraction  $\hat{p}$  of positive answers in the sample is within 0.02 of the true  $p$ ?

$$\epsilon = 0.02$$

$$\%. = 95$$

$$N = ? \quad \& \quad \text{FIND THIS}$$

⚠ FIND ON THE  $\Phi$  THE VALUE CORRESPONDING, IF THERE IS NONE CHOOSE THE ONE BEFORE IT

# DISCRETE

CHECK IF  $P(x,y)$  IS A GENUINE JOINT MASS FUNCTION

		Y			
		0	1	2	3
X	1	$\frac{1}{15}$	$\frac{1}{15}$	$\frac{2}{15}$	$\frac{1}{15}$
	2	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{5}$	$\frac{1}{10}$
3	$\frac{1}{30}$	$\frac{1}{30}$	0	$\frac{1}{10}$	$\frac{1}{30}$

YOU NEED TO CHECK THAT BOTH MARGINALS IF SUMMED ARE BOTH EQUAL TO 1  
 $\frac{1}{15} + \frac{1}{15} + \frac{2}{15} + \frac{1}{15} = 1$   
 $\frac{1}{10} + \frac{1}{10} + \frac{1}{5} + \frac{1}{10} = 1$   
 $\frac{1}{30} + \frac{1}{30} + 0 + \frac{1}{10} = 1$

MARGINAL MASS FUNCTION

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

$\rightarrow$

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

$0.35 \quad 0.20 \quad 0.15 \quad 0.30 \quad \boxed{1}$

CALCULATE THE PROBABILITY

		Y			
		0	1	2	3
X	1	$\frac{1}{15}$	$\frac{1}{15}$	$\frac{2}{15}$	$\frac{1}{15}$
	2	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{5}$	$\frac{1}{10}$
3	$\frac{1}{30}$	$\frac{1}{30}$	0	$\frac{1}{10}$	

$$\text{ES: } P(X+Y^2 \leq 2)$$

JUST FIND THE VALUE OF X AND Y THAT SATISFY THE REQUIREMENT  $(1,0), (2,0), (1,1)$  AND SUM THEIR RESPECTIVE PROBABILITY.

$$P(X+Y^2 \leq 2) = \frac{1}{15} + \frac{1}{10} + \frac{1}{15}$$

CALCULATE THE EXPECTATION, VARIANCE AND COVARIANCE

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

$$\text{ES: } E[Xe^Y]$$

PRODUCT OF THE PROBABILITY WITH THEIR RESPECTIVE X AND Y FROM THE EXPECTATION FORMULA. THEN SUM ALL  $0.1 \cdot 1e^0 + 0.15 \cdot 1e^1 + \dots$

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
3	0.05	0	0.10	0.05	

		Y			
		0	1	2	3
X					

### CUMULATIVE DISTRIBUTION FUNCTION

$x$	1	2	3	4	5
$p_X(x)$	1/7	1/14	3/14	2/7	2/7

$$f(s) = P(X \leq s) = \begin{cases} 0 & s < 1 \\ 1/7 & 1 \leq s < 2 \\ 3/14 & 2 \leq s < 3 \\ 6/14 & 3 \leq s < 4 \\ 10/14 & 4 \leq s < 5 \\ 1 & s \geq 5 \end{cases}$$

### CHECK THE INDEPENDENCE OF RANDOM VARIABLE X AND Y

⚠ BEFORE YOU NEED TO CALCULATE THE MARGINAL MASS FUNCTION

		Y			
		0	1	2	3
X	1	1/15	1/15	2/15	1/15
	2	1/10	1/10	1/5	1/10
	3	1/30	1/30	0	1/10

- YOU NEED TO CHECK THAT THE PRODUCT OF ALL THE MARGINALS IS EQUAL TO THE RESPECTIVE INTERSECTION  
 $\frac{1}{15} \cdot \frac{1}{10} \cdot \frac{1}{30} = \frac{1}{90}$   
 • IF YOU HAVE A ZERO  $\Rightarrow$  NOT INDEPENDENT = DEPENDENT  
 • IF IT IS TRUE  $\forall$  MARGINALS  $\Rightarrow$  INDEPENDENT

### PROBABILITY MASS FUNCTION

Exercise 3.10. Let  $X$  have probability mass function

$$P(X = -1) = \frac{1}{2}, \quad P(X = 0) = \frac{1}{3}, \quad \text{and} \quad P(X = 1) = \frac{1}{6}.$$

Calculate  $E(|X|)$  using the approaches in (a) and (b) below.

- (a) First find the probability mass function of the random variable  $Y = |X|$  and using that compute  $E(|X|)$ .

GIVEN  $Y = |X|$  FIND THE VALUES THAT  $Y$  CAN HAVE  
 •  $y=1 \Rightarrow$  FROM  $X=-1$  AND  $X=1$   
 •  $y=0 \Rightarrow$  FROM  $X=0$

$$\text{IF } y=1: \quad P(y=1) = P(x=1) + P(x=-1) = \frac{1}{2} + \frac{1}{6} = \frac{2}{3}$$

$$\text{IF } y=0: \quad P(y=0) = P(x=0) = \frac{1}{3}$$

## DISCRETE

## CONDITIONAL

### CONDITIONAL MASS FUNCTION OF X GIVEN Y

⚠ FIND THE MARGINAL BEFORE

$$P(x=1|y=0) = \frac{P(x=1, y=0)}{P(y=0)} = \frac{0.10}{0.35} = \dots$$

$$P(x=2|y=0) = \frac{P(x=2, y=0)}{P(y=0)} = \frac{0.20}{0.35} = \dots$$

$$P(x=3|y=0) = \frac{P(x=3, y=0)}{P(y=0)} = \frac{0.05}{0.35} = \dots$$

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
	3	0.05	0	0.10	0.05

0.3  
0.2  
0.1

### CONDITIONAL MASS FUNCTION

GIVEN A JOINT PROBABILITY MASS FUNCTION FIND  $f_{x|y}(x|y)$  ①

### FIND MARGINAL

		Y		
		0	1	2
X	1	0	$\frac{1}{9}$	0
	2	$\frac{1}{3}$	$\frac{2}{9}$	$\frac{1}{9}$
	3	0	$\frac{1}{9}$	$\frac{1}{9}$

1/9  
6/9  
2/9

$\frac{3}{9}, \frac{4}{9}, \frac{2}{9}$

Exercise 10.1. The joint probability mass function of the random variables  $(X, Y)$  is given by the following table:

		Y		
		0	1	2
X	1	0	$\frac{1}{9}$	0
	2	$\frac{1}{3}$	$\frac{2}{9}$	$\frac{1}{9}$
	3	0	$\frac{1}{9}$	$\frac{1}{9}$

(a) Find the conditional probability mass function of  $X$  given  $Y=y$ .

$$P(x=1|y=0) = \frac{P(x=1, y=0)}{P(y=0)} = \frac{0}{3/9} = 0 \quad P(x=2|y=0) = \frac{1/3}{3/9} = 1 \quad P(x=3|y=0) = 0$$

$$P(x=1|y=1) = \frac{P(x=1, y=1)}{P(y=1)} = \frac{1/9}{4/9} = 1/4 \quad P(x=2|y=1) = \frac{2/9}{4/9} = 1/2 \quad P(x=3|y=1) = \frac{1/9}{4/9} = 1/4$$

$$P(x=1|y=2) = \frac{P(x=1, y=2)}{P(y=2)} = \frac{0}{2/9} = 0 \quad P(x=2|y=2) = \frac{1/9}{4/9} = 1/4 \quad P(x=3|y=2) = \frac{1/9}{2/9} = 1/2$$

TWO VARIABLES

(JOINT)

CUMULATIVE DISTRIBUTION FUNCTION

$$f(x) = \begin{cases} 3e^{-3x}, & x > 0 \\ 0, & \text{else.} \end{cases}$$

P(X ≤ s) =

$$\int_{-\infty}^s 3e^{-3x} dx = 1 - e^{-3s}$$

⚠ USE THE LETTER S

$$\text{IF THERE ARE 2 VAR: } P(X \leq s, Y \leq t) = \int_{-\infty}^s \int_{-\infty}^t f(x, y) dy dx$$

### CHECK THE INDEPENDENCE OF RANDOM VARIABLE X AND Y

$$f(x, y) = \begin{cases} \frac{12}{7}(xy + y^2), & 0 \leq x \leq 1 \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise.} \end{cases}$$

⚠ BEFORE YOU NEED TO CALCULATE THE MARGINAL MASS FUNCTION

• YOU NEED TO CHECK IF THE PRODUCT OF THE TWO MARGINAL IS EQUAL

TO  $f(x, y)$

$$\bullet f(x, y) \stackrel{?}{=} \left( \int_0^1 f(x, y) dy \right) \left( \int_0^1 f(x, y) dx \right)$$

• IF THE EQUATION IS FALSE  $\Rightarrow$  NOT INDEPENDENT = DEPENDENT

• IF THE EQUATION IS TRUE  $\Rightarrow$  INDEPENDENT

### PROBABILITY DENSITY FUNCTION

⚠ REMEMBER THE FORMULA:  $f_y = f_x(h^{-1}(y)) \cdot \left[ \frac{d}{dy} h^{-1}(y) \right]$

Exercise 5.7. Suppose  $X \sim \text{Exp}(\lambda)$  and  $Y = \ln X$ . Find the probability density function of  $Y$ .

$$X \sim \text{EXP}(\lambda) \quad Y = \ln(X)$$

$h^{-1}$  IS THE INVERSE OF  $Y$ : SINCE  $Y = \ln(X) \Rightarrow h^{-1} = X = e^Y$

$$f_y = f_x(e^y) \cdot \left[ \frac{d}{dy} e^y \right] = \lambda e^{-\lambda e^y} e^y = \lambda e^{y-\lambda e^y}$$

THIS IS THE FUNCTION OF  $X$  THAT IS  $\lambda e^{-\lambda e^y}$   
 BECAUSE IS EXPONENTIAL WITH  $t = h^{-1} = e^y$

## DISCRETE

## CONDITIONAL

## CONTINUOUS

### CONDITIONAL MASS FUNCTION OF X GIVEN Y

⚠ FIND THE MARGINAL BEFORE

$$P(x=1|y=0) = \frac{P(x=1, y=0)}{P(y=0)} = \frac{0.10}{0.35} = \dots$$

$$P(x=2|y=0) = \frac{P(x=2, y=0)}{P(y=0)} = \frac{0.20}{0.35} = \dots$$

$$P(x=3|y=0) = \frac{P(x=3, y=0)}{P(y=0)} = \frac{0.05}{0.35} = \dots$$

		Y			
		0	1	2	3
X	1	0.10	0.15	0	0.05
	2	0.20	0.05	0.05	0.20
	3	0.05	0	0.10	0.05

0.3  
0.2  
0.1

### CONDITIONAL MASS FUNCTION

GIVEN A JOINT PROBABILITY MASS FUNCTION FIND  $f_{x|y}(x|y)$  ①

### FIND MARGINAL

		Y		
0	1	2		
X	1	0	$\frac{1}{9}$	0

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## CONDITIONAL PROBABILITIES

GIVEN A JOINT MASS FUNCTION FIND  $P(X \in A | Y = y)$

USELESS?

$$\text{FAKE-ES: } P(X \in A | Y = y) = P(X \in A | Y = 0) + P(X \in A | Y = 1) + P(X \in A | Y = 2)$$

$$\begin{aligned} P(X=1|Y=0) &= \frac{P(X=1, Y=0)}{P(Y=0)} = \frac{0}{3/9} \\ P(X=1|Y=1) &= \frac{P(X=1, Y=1)}{P(Y=1)} = \frac{1/9}{4/9} = 1/4 \\ P(X=1|Y=2) &= \frac{P(X=1, Y=2)}{P(Y=2)} = \frac{0}{2/9} = 0 \end{aligned}$$

$$P(X \in A | Y = y) = 0 + \frac{1}{4} + 0 = 1/4$$

(2)

## CONDITIONAL PROBABILITIES

GIVEN A JOINT DENSITY FUNCTION FIND  $P(X \in A | Y = y)$

(2)

Exercise 10.6. Let the joint density function of  $(X, Y)$  be

$$f(x, y) = \frac{x+y}{4}, \quad 0 < x < y < 2.$$

(a) Find  $f_{X|Y}(x|y)$ .

(b) Compute the conditional probabilities  $P(X < 1/2 | Y = 1)$  and  $P(X < 3/2 | Y = 1)$ .

FORMULA:

$$P(X \in A | Y = y) = \int_A f_{X|Y}(x|y) dx$$

I KNOW

$$\bullet f(x, y) = \frac{x+y}{4}$$

$\bullet 0 < x < y < 2$

$$\bullet f_{X|Y}(x|y) = \frac{x+y}{y^2} \quad \begin{array}{l} \text{FROM BEFORE} \\ \text{ALWAYS CHECK X!!!} \\ \text{IT MUST BE INSIDE ITS DOMAIN} \Rightarrow \text{IF } P(X < 3/2 | Y = 1) \\ \text{BUT } 0 < x < y < 2, \\ \text{IF } P(X > 1/2 | Y = 1) \text{ BUT} \\ 0 < x < y < 2 \text{ AND } Y = 1 \text{ SO } \boxed{0 < x < 1} \end{array}$$

$\Rightarrow$  TAKE  $f_{X|Y}(x|y)$

$$\bullet P(X < 1/2 | Y = 1) = \int_0^{1/2} \frac{x+y}{y^2} dx = \frac{1}{3} \int_0^{1/2} x^2 + x dx = \frac{1}{3} \left[ \frac{x^3}{3} + \frac{x^2}{2} \right]_0^{1/2} = \frac{1}{3} \cdot \frac{5}{8} = \boxed{\frac{5}{24}}$$

$\downarrow$  You know  $y = 1$ , so  
SUBSTITUTE  $y$  WITH 1

## CONDITIONAL EXPECTATION

GIVEN A JOINT DENSITY FUNCTION COMPUTE  $E[X|Y=y]$

(3)

Exercise 10.1. The joint probability mass function of the random variables  $(X, Y)$  is given by the following table:

		Y		
		0	1	2
X	0	0	$\frac{1}{9}$	0
	1	$\frac{1}{3}$	$\frac{2}{9}$	$\frac{1}{9}$
	2	0	$\frac{1}{9}$	$\frac{1}{9}$

(a) Find the conditional probability mass function of  $X$  given  $Y = y$ .

(b) Find the conditional expectation  $E[X|Y=y]$  for each of  $y = 0, 1, 2$ .

$$E[X|Y=y] \text{ FOR EACH } y = 0, 1, 2$$

YOU NEED TO FIND BEFORE THE CONDITIONAL MASS FUNCTION (1)

- $E[X|Y=0] = 1 \cdot P(X=1|Y=0) + 2P(X=2|Y=0) + 3P(X=3|Y=0) = 1 \cdot 0 + 2 \cdot 1 + 3 \cdot 0 = \boxed{2}$
- $E[X|Y=1] = 1 \cdot P(X=1|Y=1) + 2P(X=2|Y=1) + 3P(X=3|Y=1) = 1 \cdot \frac{1}{3} + 2 \cdot \frac{1}{9} + 3 \cdot \frac{1}{9} = \boxed{2}$
- $E[X|Y=2] = 1 \cdot P(X=1|Y=2) + 2P(X=2|Y=2) + 3P(X=3|Y=2) = 1 \cdot 0 + 2 \cdot \frac{1}{9} + 3 \cdot \frac{1}{9} = \boxed{5/2}$

FIND MARGINAL

		Y		
		0	1	2
X	0	0	$\frac{1}{9}$	0
	1	$\frac{1}{3}$	$\frac{2}{9}$	$\frac{1}{9}$
	2	0	$\frac{1}{9}$	$\frac{1}{9}$

$$\frac{1}{9}$$

$$\frac{6}{9}$$

$$\frac{2}{9}$$

$$\frac{3}{9}$$

$$\frac{4}{9}$$

$$\frac{2}{9}$$

$X, Y$  INDEPENDENT RANDOM VARIABLES.  
 $Z = X+Y$  FIND  $P(Z=3)$

$$\text{ES: } X \sim \text{Poisson}(2), \quad Y \sim \text{Geom}\left(\frac{2}{3}\right)$$

$$\downarrow \quad \downarrow$$

$$P(X=k) = \frac{e^{-2}}{k!} \cdot \frac{2^k}{k!} \quad P(Y=k) = (1-p)^{k-1} \cdot p$$

FIND ALL THE COMBINATIONS THAT SATISFY THE Z FUNCTION

- $x=0, y=3 \rightarrow P(X=0)P(Y=3) = e^{-2} \cdot \left(\frac{1}{3}\right)^2 \cdot \frac{2}{3} = \frac{2}{27} e^{-2}$
- $x=1, y=2 \rightarrow P(X=1)P(Y=2) = e^{-2} \cdot 2 \cdot \frac{1}{3} \cdot \frac{2}{3} = \frac{4}{27} e^{-2}$
- $x=2, y=1 \rightarrow P(X=2)P(Y=1) = e^{-2} \cdot \frac{1}{2} \cdot \frac{2}{3} = \frac{1}{3} e^{-2}$
- $x=3, y=0 \rightarrow P(X=3)P(Y=0) = e^{-2} \cdot \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{27}$

DECEASE THE  $k$   
ENTIRE LINE, MAKES  
NO SENSE

$$P(Z=3) = \frac{2}{27} e^{-2} + \frac{4}{27} e^{-2} + \frac{1}{3} e^{-2} = \boxed{\frac{50}{27} e^{-2}}$$

RANDOM VARIABLE

FIRST KIND OF TRANSFORMATION

$$Z = g(x, y) \quad g: \mathbb{R}^2 \rightarrow \mathbb{R}^1$$

GIVEN X, Y RANDOM VECTOR

SECOND KIND OF TRANSFORMATION

$$(u, v) = h(x, y) \quad h: \mathbb{R}^2 \rightarrow \mathbb{R}^2$$

ES: GIVEN THAT  $(X, Y)$  IS A CONTINUOUS RANDOM VECTOR FIND  $f_Z$

$$F_Z(s) = P[Z \leq s], \quad s \in \mathbb{R}$$

ES: GIVEN THAT  $(X, Y)$  IS A CONTINUOUS RANDOM VECTOR FIND  $f_{(U,V)}$

$$1) \{g(x, y) \leq s\} = \{(x, y) \in \mathbb{R}^2 \mid g(x, y) \leq s\}$$

PROBLEM OF ANALYTIC GEOMETRY

ES: I KNOW THAT  $g(x, y) = x+y$

$$\{(x, y) \in \mathbb{R}^2 \mid x+y \leq s\}$$



$$\{(x, y) \in \mathbb{R}^2 \mid x+y \leq s\}$$

## TRANSFORMATION $R^1 \rightarrow R^1$

### CONTINUOUS (SAME ES AS BEFORE)

ES:  $X \sim N(1,2)$  FIND PROB. DENSITY FUNCTION OF  $Y = X^3$

$$f_Y = f_X(h^{-1}(y)) \left[ \frac{d}{dy} h^{-1}(y) \right]$$

1) CHECK THAT:

$\cup$   $\circ X^3$  IS 1-TO-1 (FROM GRAPH)

$\cup$   $\circ X^3$  IS DIFF. BLE (FROM GRAPH)

$\cup$   $\circ h^{-1}$  IS DIFF. BLE

$$2) Y = X^3 \Rightarrow h^{-1} = X = \sqrt[3]{Y}$$



APPLY FORMULA

$$f_Y = f_X(\sqrt[3]{y}) \cdot \left[ \frac{d}{dy} \sqrt[3]{y} \right] \Rightarrow f_Y = f_X(\sqrt[3]{y}) \cdot \frac{1}{3} \frac{1}{\sqrt[3]{y^2}} \Rightarrow f_Y = \frac{1}{\sqrt[3]{2\pi}} e^{-\frac{y^2}{2}} \cdot \frac{1}{3} \frac{1}{\sqrt[3]{y^2}}$$

SUBSTITUTE

BECAUSE  $X \sim N(1,2)$

DISCRETE I KNOW:  $y = h(x)$   $P_X(1) = 0.2$   $P_X(2) = P_X(3) = 0.4 \Rightarrow \{1, 2, 3\} \rightarrow \{0, \sqrt{\pi}\}$

YOU HAVE EITHER 0 OR  $\sqrt{\pi}$  AS CODOMAIN/RANGE

$$\cdot P(0) = "2 CASES" = 0.2 + 0.4 = 0.6 \quad \cdot P(\sqrt{\pi}) = "1 CASE" = 0.4$$

## TRANSFORMATION $R^2 \rightarrow R^2$

2 VECTOR

$$\begin{cases} S = x+y \\ T = y \end{cases} \quad h(x,y) = (x+y, y) = \begin{pmatrix} x+y \\ y \end{pmatrix} \begin{pmatrix} x \\ y \end{pmatrix}$$

M MATRIX

$$1) \text{ CHECK } \det(M) \neq 0 \quad \det(M) = (1 \cdot 1) - (0 \cdot 1) = 1 \quad \cup$$

$\cdot$  IF  $\det(M) \neq 0$  M IS INV. BLE

$\cdot$  IF  $\det(M) = 0$  M IS NOT INV. BLE

$$2) \text{ FIND } M^{-1} \text{ (INVERSE of M)}$$

$$M^{-1}(x,y) = \begin{cases} U = x+v \\ V = y \end{cases}$$

EXPRESS X AND Y WRT U AND V

$$\begin{cases} x = 1U - 1V \\ y = 1U \end{cases}^*$$

$$\begin{matrix} 1 & -1 \\ 0 & 1 \end{matrix} = M^{-1}$$

1 BEFORE

$$M^{-1}(U,V) = \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} U \\ V \end{pmatrix}$$

SAME  
INVERT SIGN

$$3) \text{ FIND } \det(M^{-1}) = (1 \cdot 1) - (-1 \cdot 0) = 1$$

→ FROM BEFORE WE KNOW:  $f_{X,Y}(x,y) = \frac{1}{16}$  WITH  $0 \leq x \leq 2$  AND  $0 \leq y \leq 4$

4) SOLVE

$$f_{S,T}(U,V) = f_{X,Y} \left( \frac{U-V}{\sqrt{2}}, \frac{V}{\sqrt{2}} \right) \cdot \det(M^{-1}) \Rightarrow f_{X,Y}(U-V, V) = \frac{1}{16} \cdot 1 \Rightarrow \frac{1}{16}$$

OLD DOMAIN

$0 \leq x \leq 2 \rightarrow 0 \leq U-V \leq 2$

$0 \leq y \leq 4 \rightarrow 0 \leq V \leq 4$

NEW DOMAIN

$0 \leq U-V \leq 2$

$0 \leq V \leq 4$

### CHECK INDEPENDENCE

• INDEP.  $\Leftarrow$  DOMAIN IS RECTANGLE

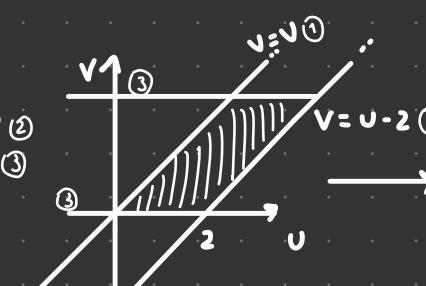
• NOT INDEP.  $\Leftarrow$  OTHER SHAPE

NEW DOMAIN

$0 \leq U-V \leq 2$

$0 \leq V \leq 4$

$\Rightarrow \begin{cases} V \leq U \quad ① \\ V \geq U-2 \quad ② \\ 0 \leq V \leq 4 \quad ③ \end{cases}$



U AND V ARE NOT INDEPENDENT

## TRANSFORMATION $R^2 \rightarrow R^1$ (FROM SAMPLE EXAM 3E) FIND DENSITY OF $X+Y \Rightarrow Z = X+Y$

REWRITE THE DOMAIN:  $0 \leq y \leq 1-x \Rightarrow 0 \leq z-x \leq 1-x$

$$f_{(x,y)}(z) = \int_0^1 f_{(x,y)}(x, z-x) dx$$

$\Downarrow$   $\min(x) + \min(y) \leq z \leq \max(x) + \max(y)$  (SINCE  $Z = X+Y$ )

$$\begin{matrix} \downarrow & \downarrow & \downarrow & \downarrow \\ 0 & 0 & 1 & 1 \\ \downarrow & \downarrow & \downarrow & \downarrow \\ 0 \leq z \leq 2 \end{matrix}$$

I KNOW  $0 \leq y \leq 1-x$



$$\text{DOMAIN } \begin{cases} 0 \leq z-x \leq 1-x \\ 0 \leq z \leq 2 \end{cases}$$

ES:  $z = 1.5$   $\therefore$

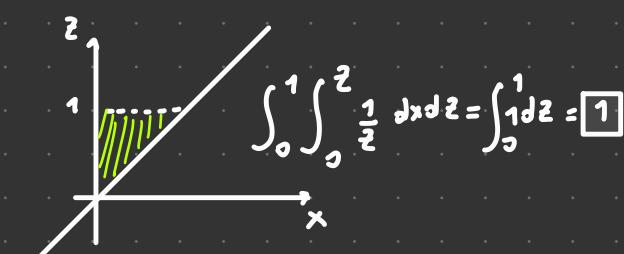
RESTRICT THE DOMAIN

$$z \in (0,1) \quad \text{MORE THAN 1 } \Rightarrow \text{ ① IS FALSE}$$

DON'T WRITE 1

$$\int_0^1 \frac{1}{x+y} dx = \int_0^1 \frac{1}{z} dx = \frac{1}{2} [z-1] = 1 \quad \cup$$

$$z \notin (0,1) \Rightarrow z \in (1,2) \quad \boxed{0} \text{ OUTSIDE THE DOMAIN}$$



$$f_{(x+y)}(z) = \begin{cases} 1 & \text{if } z \in (0,1) \\ 0 & \text{otherwise} \end{cases}$$

**Exercise 3.5.** Suppose that the discrete random variable  $X$  has cumulative distribution function given by

$$F(x) = \begin{cases} 0, & \text{if } x < 1 \\ 1/3, & \text{if } 1 \leq x < \frac{4}{3} \\ 1/2, & \text{if } \frac{4}{3} \leq x < \frac{3}{2} \\ 3/4, & \text{if } \frac{3}{2} \leq x < \frac{9}{5} \\ 1, & \text{if } x \geq \frac{9}{5}. \end{cases}$$

Find the possible values and the probability mass function  $X$ .



**3.5.** The possible values of a discrete random variable are exactly the values where the c.d.f. jumps. In this case these are the values  $1, 4/3, 3/2$  and  $9/5$ . The corresponding probabilities are equal to the size of corresponding jumps:

$$\begin{aligned} p_X(1) &= \frac{1}{3} - 0 = \frac{1}{3}, \\ p_X(4/3) &= \frac{1}{2} - \frac{1}{3} = \frac{1}{6}, \\ p_X(3/2) &= \frac{3}{4} - \frac{1}{2} = \frac{1}{4}, \\ p_X(9/5) &= 1 - \frac{3}{4} = \frac{1}{4}. \end{aligned}$$

**Exercise 3.7.** Suppose that the continuous random variable  $X$  has cumulative distribution function given by

$$F(x) = \begin{cases} 0, & \text{if } x < \sqrt{2} \\ x^2 - 2, & \text{if } \sqrt{2} \leq x < \sqrt{3} \\ 1, & \text{if } \sqrt{3} \leq x. \end{cases}$$



- (a) Find the smallest interval  $[a, b]$  such that of  $P(a \leq X \leq b) = 1$ .
- (b) Find  $P(X = 1.6)$ .
- (c) Find  $P(1 \leq X \leq \frac{3}{2})$ .
- (d) Find the probability density function of  $X$ .

**Problema 3 (10 punti).** Su uno spazio di probabilità  $(\Omega, \mathcal{F}, \mathbb{P})$  si consideri il vettore aleatorio  $(X, Y)$  avente come codominio l'insieme  $\{1, 2\} \times \{-1, 3\}$ . Per tale vettore sono note le leggi di probabilità marginali, come in tabella:

	$Y = -1$	$Y = 3$	
$X = 1$			0.3
$X = 2$			0.7
	0.6	0.4	1

Si affrontino i seguenti quesiti.

- 3a) Determinare  $E(X), E(Y), \text{Var}(X), \text{Var}(Y)$ .
- 3b) Determinare le distribuzioni congiunte di  $X$  e  $Y$  che, compatibilmente con le marginali assegnate, rendano massima, minima o nulla, rispettivamente,  $\text{Cov}(X, Y)$ .
- 3c) Dimostrare che la distribuzione congiunta che rende massima (minima, rispettivamente)  $\text{Cov}(X, Y)$  rende anche massima (minima, rispettivamente) l'espressione  $E(|X - Y|)$ .
- 3d) Stabilire se la distribuzione congiunta che rende nulla  $\text{Cov}(X, Y)$  rende anche  $X$  e  $Y$  stocasticamente indipendenti.

36)  $\text{cov}(x, y) = E[xy] - E[x]E[y]$

	$Y = -1$	$Y = 3$	
$X = 1$			0.3
$X = 2$			0.7
	0.6	0.4	

FILL WITH  $\alpha$

	$Y = -1$	$Y = 3$	
$X = 1$	$\alpha$	$0.3 - \alpha$	0.3
$X = 2$	$0.6 - \alpha$	$0.1 + \alpha$	0.7
	0.6	0.4	

3a)

$$\cdot E[X] = 0.3 \cdot 1 + 0.7 \cdot 2 = 1.7$$

$$\cdot E[X^2] = 0.3 \cdot 1^2 + 0.7 \cdot 2^2 = 3.1$$

$$\cdot E[Y] = 0.6 \cdot (-1) + 0.4 \cdot 3 = 0.6$$

$$\cdot E[Y^2] = 0.6 \cdot (-1)^2 + 0.4 \cdot 3^2 = 4.2$$

$$\cdot \text{VAR}(X) = E[X^2] - (E[X])^2 = 3.1 - (1.7)^2 = 0.21$$

$$\cdot \text{VAR}(Y) = E[Y^2] - (E[Y])^2 = 4.2 - (0.6)^2 = 3.84$$

$$E[xy] = 1 \cdot (-1) \cdot (\alpha) + 1 \cdot 3 \cdot (0.3 - \alpha) + 2 \cdot (-1) \cdot (0.6 - \alpha) + 2 \cdot 3 \cdot (0.1 + \alpha) =$$

$$= 4\alpha + 0.3$$

$$\text{cov}(x, y) = E[xy] - E[x]E[y] = 4\alpha + 0.3 - (1.7)(0.6) = 4\alpha - 0.72$$

$$\text{??}$$

WE KNOW:

$$\cdot \text{MAX} \quad \text{cov}(x, y) = 0.48$$

$$\cdot \text{MIN} \quad \text{cov}(x, y) = -0.72$$

	$Y = -1$	$Y = 3$	
$X = 1$	$\alpha$	$0.3 - \alpha$	0.3
$X = 2$	$0.6 - \alpha$	$0.1 + \alpha$	0.7
	0.6	0.4	

WE KNOW NULL: FOR  $0 = 4\alpha - 0.72 \Rightarrow \alpha = 0.18 \quad \text{cov}(x, y) = 0$

	$Y = -1$	$Y = 3$	
$X = 1$	$\alpha$	$0.3 - \alpha$	0.3
$X = 2$	$0.6 - \alpha$	$0.1 + \alpha$	0.7
	0.6	0.4	

	$Y = -1$	$Y = 3$	
$X = 1$	0.18	0.12	0.3
$X = 2$	0.92	0.28	0.7
	0.6	0.4	

- $X$  AND  $Y$  INDEP.  $\Rightarrow \text{cov}(x, y) = 0$
- $\text{cov}(x, y) \neq X$  AND  $Y$  INDEP.

$$\cdot E[|x-y|]$$

$$Z = |x-y|$$

$$\rightarrow E[Z] = 2 \cdot \alpha + 2 \cdot (0.3 - \alpha) + 3 \cdot (0.6 - \alpha) + 1 \cdot (0.1 + \alpha) = -2\alpha + 2.5$$

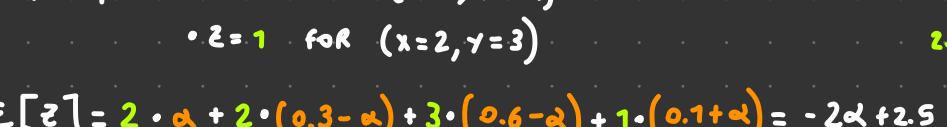
CHECK INDEPENDENCE

- $0.6 \cdot 0.3 = 0.18 \quad \checkmark$
- $0.6 \cdot 0.7 = 0.42 \quad \checkmark$
- $0.4 \cdot 0.3 = 0.12 \quad \checkmark$
- $0.4 \cdot 0.7 = 0.28 \quad \checkmark$

MIN: FOR  $\alpha = 0 \Rightarrow \text{cov}(x, y) = -0.72$

NULL: FOR  $0 = 4\alpha - 0.72 \Rightarrow \alpha = 0.18 \quad \text{cov}(x, y) = 0$

MAX: FOR  $\alpha = 0.3 \Rightarrow \text{cov}(x, y) = 0.48$



**Exercise 10.5.** Suppose that the joint density function of  $X$  and  $Y$  is

$$f(x, y) = \frac{12}{5}x(2 - x - y), \quad \text{for } 0 < x < 1 \text{ and } 0 < y < 1,$$

and zero otherwise.

- (a) Find  $f_{X|Y}(x|y)$ .
- (b) Compute  $P(X > 1/2 | Y = 3/4)$  and  $E[X | Y = 3/4]$ .



10.5. (a) The conditional probability density function is given by the formula:

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)},$$



if  $f_Y(y) > 0$ . Since the joint density is only nonzero for  $0 < y < 1$ , the  $Y$  variable will have a density which is only nonzero in  $0 < y < 1$ . In that case we have

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f_{X,Y}(w,y) dw = \int_0^1 \frac{12}{5}w(2-w-y) dw \\ &= \frac{12}{5}(w^2 - \frac{1}{3}w^3 - \frac{1}{2}yw^2)|_0^1 = \frac{12}{5}(1 - \frac{1}{3} - \frac{1}{2}y) = \frac{8}{5} - \frac{6}{5}y \end{aligned}$$

Thus, for  $0 < y < 1$  we have

$$f_{X|Y}(x|y) = \frac{\frac{12}{5}x(2-x-y)}{\frac{8}{5} - \frac{6}{5}y} = \frac{6x(2-x-y)}{4-3y}.$$

(b) We have

$$\begin{aligned} P(X > \frac{1}{2} | Y = \frac{3}{4}) &= \int_{\frac{1}{2}}^1 f_{X|Y}(x|y = \frac{3}{4}) dx = \int_{\frac{1}{2}}^1 \frac{6x(2-x-\frac{3}{4})}{4-\frac{9}{4}} dx \\ &= \frac{24}{7} \int_{\frac{1}{2}}^1 x(\frac{5}{4}-x) dx = \frac{24}{7}(\frac{5}{8}x^2 - \frac{1}{3}x^3)|_{\frac{1}{2}}^1 = \frac{24}{7}(\frac{5}{8} - \frac{1}{3} - \frac{5}{32} + \frac{1}{24}) \\ &= \frac{24}{7}(\frac{7}{24} - \frac{11}{96}) = \frac{24}{7}\frac{17}{96} = \frac{17}{28}. \end{aligned}$$

$$\begin{aligned} E[X | Y = \frac{3}{4}] &= \int_0^1 x \frac{6x(\frac{5}{4}-x)}{\frac{7}{4}} dx = \frac{24}{7} \int_0^1 x^2(\frac{5}{4}-x) dx = \frac{24}{7}(\frac{5}{12}x^3 - \frac{1}{4}x^4)|_0^1 \\ &= \frac{24}{7}\frac{1}{6} = \frac{4}{7}. \end{aligned}$$

Problem 1 (8 Points) In a certain community, 36% of the families own a dog, and 22% of the families that own a dog also own a cat. In addition, 30% of the families own a cat. Find:

$$P(CAT) = 0.30$$

$$P(DOG) = 0.36$$

$$P(CAT|DOG) = 0.22$$

LAW OF TOTAL PROBABILITY  
 $P(A \cup B) = P(A) + P(B) - P(AB)$

$$P(A \cup B) = P(A) + P(B) - P(AB)$$

1a) the probability that a randomly selected family owns both a dog and a cat;

$$P(CAT \cap DOG) = P(CAT|DOG)P(DOG) = 0.22 \cdot 0.36 = 0.0792$$

1b) the conditional probability that a randomly selected family own a dog, given that it owns a cat.

$$P(DOG|CAT) = \frac{P(DOG \text{ AND } CAT)}{P(CAT)} = \frac{P(DOG \cap CAT)}{P(CAT)} = \frac{0.0792}{0.30} = 0.264$$

Then, select at random, independently, two families out of this community. Find:

1c) the conditional probability that both the families own a cat, given that at least one of the two families owns a dog.

$$P(CAT_1 \cap CAT_2 | DOG_1 \cup DOG_2) = \frac{P((CAT_1 \cap CAT_2) \cap (DOG_1 \cup DOG_2))}{P(DOG_1 \cup DOG_2)} = \frac{0.0412}{0.5904} = 0.0698$$

$$\bullet P((CAT_1 \cap CAT_2) \cap (DOG_1 \cup DOG_2)) = P((CAT_1 \cap CAT_2 \cap DOG_1) \cup (CAT_1 \cap CAT_2 \cap DOG_2)) - (CAT_1 \cap CAT_2 \cap DOG_1 \cap DOG_2)$$

$$= P(CAT_2) P(CAT_1 \cap DOG_1) + P(CAT_1) P(CAT_2 \cap DOG_1) - P(CAT_1 \cap DOG_1) P(CAT_2 \cap DOG_1)$$

$$= 0.30 \cdot 0.0792 + 0.30 \cdot 0.0792 - 0.0792 \cdot 0.0792 = 0.0412$$

$$\bullet P(DOG_1 \cup DOG_2) = P(DOG_1) + P(DOG_2) - P(DOG_1 \cap DOG_2) = 0.36 + 0.36 - (0.36)^2 = 0.5904$$

Problem 2 (10 Points). The joint probability mass function  $P(X,Y)$  of the random vector  $(X,Y): \Omega \rightarrow \{1,2\} \times \{1,2\}$  is given by

$$\begin{aligned} P(X,Y)(1,1) &= \frac{1}{8} & P(X,Y)(1,2) &= \frac{1}{4} \\ P(X,Y)(2,1) &= \frac{1}{8} & P(X,Y)(2,2) &= \frac{1}{2} \end{aligned}$$

2a) Compute the conditional mass function of  $X$  given  $Y$ .

### CONDITIONAL MASS FUNCTION OF $X$ GIVEN $Y$

VISUALIZE ON A MATRIX AND COMPUTE THE MARGINALS

		X		$\frac{1}{4}$
		1	2	
1	1	$\frac{1}{8}$	$\frac{1}{8}$	$\frac{3}{4}$
	2	$\frac{1}{8}$	$\frac{1}{2}$	
		$\frac{3}{8}$	$\frac{5}{8}$	

$\bullet P(X=1, Y=1) = \frac{P(X=1, Y=1)}{P(Y=1)} = \frac{\frac{1}{8}}{\frac{1}{4}} = \frac{1}{2}$   
 $\bullet P(X=1, Y=2) = \frac{P(X=1, Y=2)}{P(Y=2)} = \frac{\frac{1}{8}}{\frac{3}{4}} = \frac{1}{6}$   
 $\bullet P(X=2, Y=1) = \frac{P(X=2, Y=1)}{P(Y=1)} = \frac{\frac{1}{8}}{\frac{1}{4}} = \frac{1}{2}$   
 $\bullet P(X=2, Y=2) = \frac{P(X=2, Y=2)}{P(Y=2)} = \frac{\frac{1}{2}}{\frac{3}{4}} = \frac{2}{3}$

2b) Compute  $E[X]$ ,  $E[Y]$ ,  $\text{Var}(X)$ ,  $\text{Var}(Y)$

$$E[X] = \frac{1 \cdot 3}{8} + 2 \cdot \frac{5}{8} = \frac{13}{8}$$

$$E[Y] = 1 \cdot \frac{1}{4} + 2 \cdot \frac{3}{4} = \frac{7}{4}$$

$$E[X^2] = 1^2 \cdot \frac{3}{8} + 2^2 \cdot \frac{5}{8} = \frac{23}{8}$$

$$E[Y^2] = 1^2 \cdot \frac{1}{4} + 2^2 \cdot \frac{3}{4} = \frac{13}{4}$$

$$\text{Var}(X) = E[X^2] - (E[X])^2 =$$

$$\text{Var}(Y) = E[Y^2] - (E[Y])^2 =$$

$$= \frac{23}{8} - \left(\frac{13}{8}\right)^2 = 0.2343$$

$$= \frac{13}{4} - \left(\frac{7}{4}\right)^2 = 0.1875$$

2d) Are  $X, Y$  independent?

$$\text{TEST } (1,1): \frac{3}{8} \cdot \frac{1}{4} \stackrel{?}{=} \frac{1}{8} \text{ NO} \Rightarrow \text{NOT INDEPENDENT}$$

2c) Compute  $\text{Cov}(X, Y)$

$$\text{cov}(x,y) = E[xy] - E[x]E[y] = \frac{23}{8} - \underbrace{\frac{13}{8} \cdot \frac{7}{4}}_{\frac{1}{2}} = \frac{1}{32}$$

$$\downarrow 1 \cdot 1 \cdot \frac{1}{8} + 2 \cdot 1 \cdot \frac{1}{8} + 1 \cdot 2 \cdot \frac{1}{4} + 2 \cdot 2 \cdot \frac{1}{2} = \frac{23}{8}$$

2e) Compute  $P[XY \leq 3]$ ,  $P[X+Y > 2]$  and  $P[X/Y > 1]$

$P(XY \leq 3)$  IS SATISFIED IN  $(1,1), (2,1), (1,2)$

$$\hookrightarrow = \frac{1}{8} + \frac{1}{8} + \frac{1}{4} = \frac{1}{2}$$

$P(X+Y > 2)$  IS SATISFIED IN  $(2,1), (1,2), (2,2)$

$$\hookrightarrow \frac{1}{8} + \frac{1}{4} + \frac{1}{2} = \frac{7}{8}$$

$P(X/Y > 1)$  IS SATISFIED IN  $(2,1)$

$$\hookrightarrow \frac{1}{8}$$

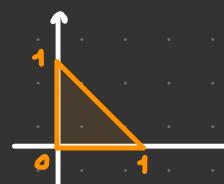
Problem 3 (12 Points). Let  $D \subseteq \mathbb{R}^2$  be the triangle with vertices  $(0,0)$ ,  $(0,1)$  and  $(1,0)$ . Consider the continuous random vector  $(X,Y) : \Omega \rightarrow \mathbb{R}^2$  with joint density function

$$f_{(X,Y)}(x,y) = \begin{cases} \frac{1}{x+y} & \text{if } (x,y) \in D \\ 0 & \text{otherwise} \end{cases}$$

Find:

3a) the marginal densities  $f_X$  and  $f_Y$

3b)  $E[X]$ ,  $\text{Var}(X)$ ,  $E[Y]$ ,  $\text{Var}(Y)$



### MARGINALS

#### FIRST MARGINAL ( $X$ )

BECUSE  $X$  IS "FIXED" AND  $Y$  CAN VARY ONLY IN  $1-X$  → BECAUSE  $X$  AND  $Y$  INFLUENCE EACH OTHER

$$f_X = \int_0^{1-x} \frac{1}{x+y} dy = [\ln(x+y)]_0^{1-x} = \ln(x+1-x) - \ln(x+0) = -\ln(x)$$

#### SECOND MARGINAL ( $Y$ )

$$f_Y = \int_0^{1-y} \frac{1}{x+y} dx = -\ln(y)$$

$$\begin{aligned} E[X] &= \int_0^1 x f_X dx = - \int_0^1 x \ln(x) dx = -[\ln(x) \cdot \frac{x^2}{2}]_0^1 + \int_0^1 \frac{1}{2} \cdot \frac{x^2}{2} dx = \left[ \frac{x^2}{4} \right]_0^1 = \frac{1}{4} = E[Y] \\ &\quad \text{by symmetry} \\ &\quad \cancel{\int \ln(x) \cancel{\int^1_0} x^1 dx} \\ &\quad \cancel{\cancel{\int^1_0} x^2/2} \end{aligned}$$

$$\text{VAR}(X) = E[X^2] - (E[X])^2 = \frac{1}{3} - \left(\frac{1}{4}\right)^2 = \frac{1}{3} - \frac{1}{16} = \frac{7}{48} = \text{VAR}(Y)$$

$$\begin{aligned} E[X^2] &= \int_0^1 x^2 f_X dx = \int_0^1 x^2 \ln(x) dx = -\left[\frac{x^3}{3} \ln(x)\right]_0^1 + \int_0^1 \frac{2}{3} \cdot \frac{x^3}{3} \cdot \frac{1}{x} dx = \left[\frac{x^3}{9}\right]_0^1 = \frac{1}{9} = E[Y^2] \\ &\quad \text{by symmetry} \\ &\quad \cancel{\int \ln(x) \cancel{\int^1_0} x^1 dx} \\ &\quad \cancel{\cancel{\int^1_0} x^3/3} \end{aligned}$$

3c) the conditional density  $f_{Y|X}(y|x)$

$$f_{Y|X}(y|x) = \frac{f_{(x,y)}(x,y)}{f_X(x)} = \frac{\frac{1}{x+y}}{-\ln(y)} = \frac{-1}{(x+y)\ln(y)} \Rightarrow \begin{cases} \frac{-1}{(x+y)\ln(y)} & \text{if } y \in (0, 1-x) \\ 0 & \text{otherwise} \end{cases}$$

3d)  $\text{Cov}(X,Y)$

$$\begin{aligned} \text{Cov}(X,Y) &= E[XY] - E[X]E[Y] = \frac{1}{18} - \frac{1}{4} \cdot \frac{1}{4} = \boxed{\frac{1}{18} - \frac{1}{16}} \\ &\downarrow \\ &\iint xy \frac{1}{x+y} dxdy = \int_0^1 y \int_0^{1-y} \frac{x}{x+y} dxdy = \int_0^1 y \int_0^{1-y} \frac{x+y-y}{x+y} dxdy = \int_0^1 y \int_0^{1-y} 1 - \frac{y}{x+y} dxdy = \int_0^1 y \left[ x - y \ln(x+y) \right]_0^{1-y} dy = \int_0^1 y \left[ 1 - y - 0 + y \ln(y) \right] dy \\ &= \int_0^1 y - y^2 + y^2 \ln(y) dy = \left[ \frac{y^2}{2} - \frac{y^3}{3} \right]_0^1 + \left[ \frac{y^3}{3} \ln(y) \right]_0^1 - \int_0^1 \frac{y^2}{3} dy = \frac{1}{6} - \left[ \frac{y^3}{9} \right]_0^1 = \frac{1}{6} - \frac{1}{9} = \frac{1}{18} \\ &\quad \cancel{\int \ln(y) \cancel{\int^1_0} y^1 dy} \\ &\quad \cancel{\cancel{\int^1_0} y^3/3} \end{aligned}$$

$\min(x) + \min(y)$        $\max(x) + \max(y)$

DEFINE  $Z = X+Y$  AND FIND ITS DOMAIN:  $Z \in (0, 2)$

I TRANSFORM:  $f_{(X,Y)} \Rightarrow f_{(X,Z-X)}$  FROM  $1 \rightarrow 2$

$$f_{X+Y}(z) = \int_0^1 \frac{1}{x+(z-x)} dx = \frac{1}{z} \int_0^1 \mathbf{1}_{\{z-x \in (0, 1-x)\}} dx = \frac{1}{z} \left[ x \right]_0^1 = 1 ??$$

(+IF  $z-x \in (0, 1-x)$   $\mathbf{1}_{\{ \cdot \}} = 1$   
ELSE  $\mathbf{1}_{\{ \cdot \}} = 0$ )

Problem 4 (10 Points) Let  $(X, Y)$  be a continuous random vector having bivariate GAUSSIAN distribution with mean vector  $(m_x, m_y) = (2, 0)$  and covariance matrix  $\begin{pmatrix} \sigma_x^2 & \rho \sigma_x \sigma_y \\ \rho \sigma_x \sigma_y & \sigma_y^2 \end{pmatrix} = \begin{pmatrix} 1 & 1 \\ 1 & 9 \end{pmatrix}$

I KNOW:

$$m_x = 2 \quad \sigma_x^2 = 1$$

$$m_y = 0 \quad \sigma_y^2 = 9$$

$$\rho \sigma_x \sigma_y = 1$$

"RHO" RELATIONAL COEFFICIENT

FIRST STEP: FIND  $\rho$

$$I \text{ KNOW: } \sigma_x^2 = 1 \Rightarrow \sigma_x = 1 \quad \text{so, } \rho \cdot 1 \cdot 3 = 1; \quad \boxed{\rho = \frac{1}{3}}$$

$$\sigma_y^2 = 9 \Rightarrow \sigma_y = 3$$

4a)  $E[3X - 4Y]$

$$E[3x - 4y] = 3E[x] - 4E[y] = 3 \cdot 2 - 4 \cdot 0 = 6$$

4b)  $\text{Cov}(2X+Y, -X+3Y)$

FIND ALL THE COMBINATIONS POSSIBLE (LIKE  $(2x+y)(-x+3y)$ )

$$\begin{aligned} \text{Cov}(2X+Y, -X+3Y) &= -2\text{Cov}(x, x) + 6\text{Cov}(x, y) - \text{Cov}(y, x) + 3\text{Cov}(y, y) = \\ &= -2\text{VAR}(x) + 6\text{Cov}(x, y) - \text{Cov}(x, y) + 3\text{VAR}(y) = \\ &= -2\text{VAR}(x) + 5\text{Cov}(x, y) + 3\text{VAR}(y) = \\ &= -2 \cdot 1 + 5 \cdot 1 + 3 \cdot 9 = \boxed{30} \\ \text{!} \quad \text{cov}(x, y) &= \rho \sigma_x \sigma_y = \frac{1}{3} \cdot 1 \cdot 3 = \boxed{1} \end{aligned}$$

REMEMBER

$$\text{VAR}(3X) = 3 \text{VAR}(X)$$

$$\text{Cov}(X, Y) = \text{Cov}(Y, X)$$

$$\text{Cov}(X, X) = \text{VAR}(X)$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$\text{IF INDEPENDENT: } E[XY] = E[X]E[Y]$$

$$\rho = \frac{\text{Cov}(X, Y)}{\sqrt{\sigma_x^2 \sigma_y^2}} = \frac{\text{Cov}(X, Y)}{\sigma_x \sigma_y}$$

• VARIANCE IS ALWAYS

• COVARIANCE CAN BE POSITIVE OR NEGATIVE

• (+) OPPOSITE RELATION

• (+) DIRECTLY PROPORTIONAL

• (0) NO CORRELATION

3c) the distribution of the random vector  $(U, V) = (2X+Y, -X+3Y)$

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3d) the conditional distribution of  $X$  given  $Y$ .

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