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THEORY OF BIVARIATE, CONTINUOUS RANDOM VECTORS

- 1) Definition [RANDOM VECTOR] RANDOM VARIABLES BELONGS TO SOME INTERVAL

$$(X, Y) : \Omega \rightarrow \mathbb{R}^2$$

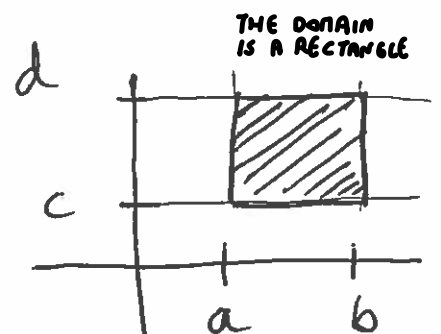
where both $X : \Omega \rightarrow \mathbb{R}$ and $Y : \Omega \rightarrow \mathbb{R}$ are random variables.

- 2) PROBABILITY DISTRIBUTION OF A CONTINUOUS
RANDOM VECTOR [$a < b$, $c < d$ fixed]
for x for y

$$P[(X, Y) \in (a, b) \times (c, d)] =$$

$$= P[X \in (a, b), Y \in (c, d)]$$

$$= \int_a^b \int_c^d f_{(X, Y)}(u, v) du dv$$



where $f_{(X, Y)} : \mathbb{R}^2 \rightarrow \mathbb{R}$

is the (JOINT) DENSITY FUNCTION
of the random vector (X, Y)

3) Definition [DENSITY FUNCTION]

A density function $f_{(X,Y)}: \mathbb{R}^2 \rightarrow \mathbb{R}$ is a function with the following properties:

3A) $f_{(X,Y)}(u,v) \geq 0 \quad \forall (u,v) \in \mathbb{R}^2$

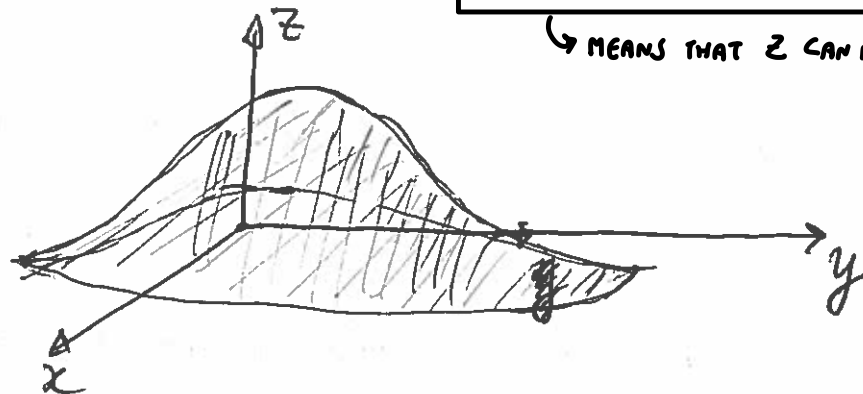
3B) $f_{(X,Y)}$ is integrable on every rectangle

3C) $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f_{(X,Y)}(u,v) du dv = 1$

Remarks

3A) Geometrically speaking, $f_{(X,Y)}(u,v) \geq 0$ means that the graph of $f_{(X,Y)}$ is a surface lying upon the xy -plane.

↳ MEANS THAT z CAN NOT BE NEGATIVE



3B) Recall that the CONTINUITY of $f_{(X,Y)}$ is a sufficient condition yielding integrability.

3c) $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f_{(X,Y)}(u,v) du dv = 1$ means that the volume included between the xy -plane and the surface defined by the graph of $f_{(X,Y)}$ is 1. It is related to the FIRST AXIOM of probability: $P(\Omega) = 1$.

4) DISTRIBUTION FUNCTION of (X,Y)

↑ JOINT PROBABILITY of $x \leq s$ AND $y \leq t$
 ↓ FIXED NUMBERS

↓ JOINT DISTRIBUTION FUNCTION

$$F_{(X,Y)}(s,t) = P[X \leq s, Y \leq t] \quad (s,t) \in \mathbb{R}^2$$

INTEGRATE

$$= \int_{-\infty}^s \int_{-\infty}^t f_{(X,Y)}(u,v) du dv$$

5) MARGINAL DISTRIBUTIONS

Given the random vector $(X,Y): \Omega \rightarrow \mathbb{R}^2$ with joint density function $f_{(X,Y)}: \mathbb{R}^2 \rightarrow \mathbb{R}$ we define:

↑ OBTAIN SOMETHING THAT DEPENDS ON X , so you DERIVATE WRT Y

FIRST MARGINAL DENSITY: $(x \in \mathbb{R} \text{ fixed})$

$$f_X(x) = \int_{-\infty}^{+\infty} f_{(X,Y)}(x,y) dy$$

↓ OBTAIN SOMETHING THAT DEPENDS ON Y , so you DERIVATE WRT TO X $(y \in \mathbb{R} \text{ fixed})$

SECOND MARGINAL DENSITY:

$$f_Y(y) = \int_{-\infty}^{+\infty} f_{(X,Y)}(x,y) dx$$

Probabilistic meaning of the marginals.

Start by considering

$$P[X \leq s] = P[X \leq s, Y < +\infty] =$$

$s \in \mathbb{R}, \text{ fixed}$

$$= \lim_{t \rightarrow +\infty} P[X \leq s, Y \leq t].$$

DOES NOT ADD ANY CONSTRAINTS
 $P(A) = P(A \cap \Omega)$
↳ IT IS AGAIN A
↳ JOINT DISTRIBUTION FUNCTION

By resorting to the notion of (joint) distribution function, this fact can be written as:

$$F_X(s) = P[X \leq s] = \lim_{t \rightarrow +\infty} F_{(X,Y)}(s, t)$$

$$= \lim_{t \rightarrow +\infty} \int_{-\infty}^s \int_{-\infty}^t f_{(X,Y)}(x, y) dx dy$$

$$= \int_{-\infty}^s \int_{-\infty}^{+\infty} f_{(X,Y)}(x, y) dx dy$$

(x) (y)

We deduce that $F_X(s) = \int_{-\infty}^s f_X(x) dx$

with $f_X(x) = \int_{-\infty}^{+\infty} f_{(X,Y)}(x, y) dy$

FIRST MARGINAL FORMULA OF BEFORE

A similar argument works for the second marginal. $f_Y(y) = \int_{-\infty}^{+\infty} f_{(X,Y)}(x, y) dx$

6) INDEPENDENCE OF RANDOM VARIABLES

→ SAME AS $P(A \cap B) = P(A)P(B)$
BUT DO NOT USE $X = \text{something}$
BECAUSE FOR CONTINUOUS $= 0$

Remark. Independence is not a property of the two random variables X, Y , thought of as two disjoint entities. Rather, Independence is a property of the random vector (X, Y) . Then, if you only know the marginal distributions/densities, you cannot decide ~~whether~~ whether X, Y are independent or not.

Definition [For general random vectors]

Given the random vector (X, Y) with joint distribution function $F_{(X, Y)}$, we say that its component X and Y are independent if

$$F_{(X, Y)}(s, t) = F_X(s) \cdot F_Y(t), \quad \forall (s, t) \in \mathbb{R}^2$$

Theorem. If (X, Y) is continuous random vector with density $f_{(X, Y)}$, then

X and Y are independent iff

$$f_{(X,Y)}(u,v) = f_X(u) \cdot f_Y(v) \quad \forall (u,v) \in \mathbb{R}^2$$

Proposition. Let (X,Y) be a random vector with independent components X and Y . If $\varphi: \mathbb{R} \rightarrow \mathbb{R}$ and $\psi: \mathbb{R} \rightarrow \mathbb{R}$ are two functions then the transformed random vector $(\varphi(X), \psi(Y))$ has independent components, that is $\varphi(X)$ and $\psi(Y)$ are independent.

* TWO RANDOM FUNCTIONS IF INDEPENDENT BEFORE \Rightarrow INDEPENDENT ALSO NOW

NOT INVOLVES Y NOT INVOLVES X

7) EXPECTATION

$$U = \varphi(X) \quad V = \psi(Y)$$

X INDEPENDENT OF $Y \Rightarrow V$ INDEPENDENT OF U

Let (X,Y) be a continuous random vector with joint density $f_{(X,Y)}$. Let $h: \mathbb{R}^2 \rightarrow \mathbb{R}$ be a function. Then, we put

$$\mathbb{E}[h(X,Y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} h(x,y) f_{(X,Y)}(x,y) dx dy$$

7) COVARIANCE

$$\text{Cov}(X, Y) = E[X \cdot Y] - E[X] \cdot E[Y]$$

where:

$$E[X \cdot Y] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_{(X,Y)}(x,y) dx dy$$

$$E[X] = \int_{-\infty}^{+\infty} x f_X(x) dx = m_X \quad \rightarrow \text{MARGINAL OF } X$$

$$E[Y] = \int_{-\infty}^{+\infty} y f_Y(y) dy = m_Y \quad \rightarrow \text{MARGINAL OF } Y$$

Proposition [Properties of the COVARIANCE]

A) $\text{Cov}(X, Y) = E[(X - m_X) \cdot (Y - m_Y)]$

B) $\text{Cov}(X, X) = \text{Var}(X)$

C) Covariance is BILINEAR, i.e.,

$$\begin{aligned} \text{Cov}(a_1 X_1 + a_2 X_2, Y) &= a_1 \text{Cov}(X_1, Y) + \\ a_1, a_2 \in \mathbb{R} \quad &+ a_2 \text{Cov}(X_2, Y) \end{aligned}$$

$$\begin{aligned} \text{Cov}(X, b_1 Y_1 + b_2 Y_2) &= b_1 \text{Cov}(X, Y_1) + \\ b_1, b_2 \in \mathbb{R} \quad &+ b_2 \text{Cov}(X, Y_2) \end{aligned}$$

A) $\text{Cov}(X, Y) = \text{Cov}(Y, X)$

E) Cauchy-Schwarz inequality:

$$|\text{Cov}(X, Y)| \leq \sqrt{\text{Var}(X) \cdot \text{Var}(Y)}$$

F) Geometrically speaking, if we restrict the Covariance to the space (linear space)

$$L_0^2 = \left\{ \underset{\text{R.V.}}{X: \Omega \rightarrow \mathbb{R}} \mid \mathbb{E}[X^2] < +\infty, \mathbb{E}[X] = 0 \right\}$$

then,

$$\boxed{\text{Cov}: L_0^2 \times L_0^2 \rightarrow \mathbb{R}} \text{ is a scalar product}$$

→ OPERATION TO DO ON RANDOM VECTOR

8) TRANSFORMATION OF RANDOM VECTORS.

Given the random vector X, Y
there are two kinds of transformation:

FIRST KIND

$$Z = g(X, Y), \quad g: \mathbb{R}^2 \rightarrow \mathbb{R}^{(1)}$$

Z is a random variable

E.g. $Z = X + Y$, or $Z = X \cdot Y$

SECOND KIND

$$(U, V) = h(X, Y), \quad h: \mathbb{R}^2 \rightarrow \mathbb{R}^{(2)}$$

(U, V) is a random vector.

E.g. $(U, V) = (X+Y, X-Y)$ or

$$(U, V) = (X+3Y, XY)$$

Problem. Given that (X, Y) is a continuous random vector, find f_Z and $f_{(U, V)}$.

Problem 1: f_Z

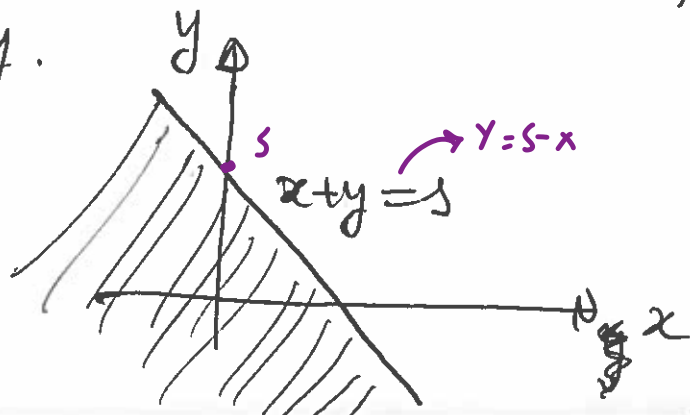
Split into 3 steps, starting from
 $F_Z(s) = P[Z \leq s], \quad s \in \mathbb{R}.$

Step 1. $\{g(x, y) \leq s\} = \{(x, y) \in \mathbb{R}^2 \mid g(x, y) \leq s\}$
It is a problem of ANALYTIC GEOMETRY,

E.g. $g(x, y) = x + y.$

$$\{(x, y) \in \mathbb{R}^2 \mid x + y \leq s\}$$

= HALF PLANE



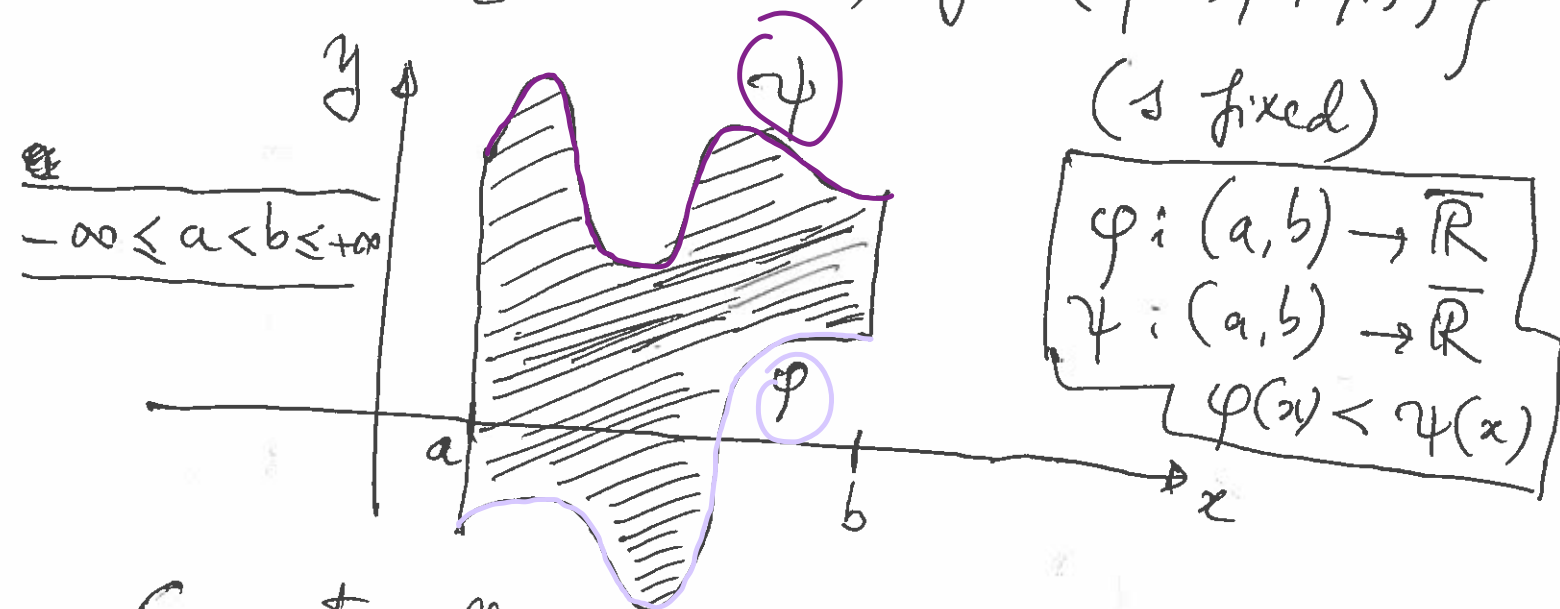
Step 2 Putting $B(s) = \{(x, y) \in \mathbb{R}^2 \mid f(x, y) \leq s\}$
 $\subseteq \mathbb{R}^2$

We consider the problem of evaluating
 $P[(X, Y) \in B(s)]$. [s fixed]

Assume that we can write

$$B(s) = \{x \in (a, b), y \in (\varphi(x), \psi(x))\}$$

(s fixed)



Geometrically, φ, ψ are two curves
 with φ lying below of ψ .

$$P[(X, Y) \in B(s)] = \int_a^b \left(\int_{\varphi(x)}^{\psi(x)} f(x, y) dy \right) dx$$

Eg. For the half plane; $a = -\infty, b = +\infty$
 $\varphi(x) \equiv -\infty, \psi(x) = s - x$

Step 3 Once we have found

INTEGRATE
DENSITY → DISTRIBUTION
← DIFFERENTIATE

$$F_Z(s) = P[Z \leq s] = P[(X, Y) \in B(s)]$$

We conclude by resorting to

$$\boxed{f_Z(s) = F'_Z(s)} \quad \forall s \in \mathbb{R}$$

E.g. $g(x, y) = x + y \Rightarrow Z = X + Y$

$$f_Z(s) = \int_{-\infty}^{+\infty} f_{(X, Y)}(x, s-x) dx$$

In fact, from step 2, we get

$$F_Z(s) = \int_{-\infty}^{+\infty} \left(\int_{-\infty}^{s-x} f_{(X, Y)}(x, y) dy \right) dx$$

Now consider only the inner integral

$$\int_{-\infty}^{s-x} f_{(X, Y)}(x, y) dy = \int_{-\infty}^s f_{(X, Y)}(x, z-x) dz$$

by the transformation (change of variable) $y = z - x$.

Thus, by interchanging the order of integration (Fubini)

$$F_Z(s) = \int_{-\infty}^s \left(\int_{-\infty}^{+\infty} f_{(X, Y)}(x, z-x) dx \right) dz$$

we have
Conclude
~~and we are done~~

Second KIND of transformation.

This is more difficult, and we confine ourselves to considering $h: \mathbb{R}^2 \rightarrow \mathbb{R}^2$ such that

- i) h is one-to-one (injective and surjective)
- ii) h is $C^1(\mathbb{R}^2; \mathbb{R}^2)$ [differentiable with continuous derivatives]
- iii) h^{-1} is $C^1(\mathbb{R}^2; \mathbb{R}^2)$

$$f(u,v)(u,v) = f(x,y)(h^{-1}(u,v)) \cdot \left| \det(\text{Jac}[h^{-1}]_{(u,v)}) \right|$$

① $h^{-1}(u,v)$ is obtained by solving $h(x,y) = (u,v)$ and finding x & y in terms of u,v

② $\text{Jac}[h^{-1}](u,v) = \begin{pmatrix} \frac{\partial H_1}{\partial u} & \frac{\partial H_1}{\partial v} \\ \frac{\partial H_2}{\partial u} & \frac{\partial H_2}{\partial v} \end{pmatrix}$

$$h^{-1}(u,v) = H(u,v)$$

$$= (H_1(u,v), H_2(u,v))$$

[JACOBIAN MATRIX]

③

$$\left| \det \begin{pmatrix} a & b \\ c & d \end{pmatrix} \right| = |ad - bc|$$

9) CONDITIONAL DENSITIES

$$f_{Y|X}(y|x) = \frac{f(x,y)(x,y)}{f_X(x)}$$

\nearrow JOINT DENSITY
 \searrow MARGINAL

[if $f_X(x) > 0$]

$$f_{X|Y}(x|y) = \frac{f(x,y)(x,y)}{f_Y(y)}$$

[if $f_Y(y) > 0$]

10) CONDITIONAL EXPECTATIONS

\nearrow SAME AS UNIVARIATE CASE

Expectation of Y given $X=x$

\nearrow SEE BEFORE $\frac{f(x,y)(x,y)}{f_X(x)}$

$$E[Y|X=x] = \int_{-\infty}^{+\infty} y f_{Y|X}(y|x) dy$$

x fixed

More generally

$$E[g(Y)|X=x] = \int_{-\infty}^{+\infty} g(y) f_{Y|X}(y|x) dy$$

\nearrow SEE BEFORE

$$E[X|Y=y] = \int_{-\infty}^{+\infty} x f_{X|Y}(x|y) dx$$

\nearrow SEE BEFORE

y fixed

More generally,

$$E[g(X)|Y=y] = \int_{-\infty}^{+\infty} g(x) f_{X|Y}(x|y) dx$$

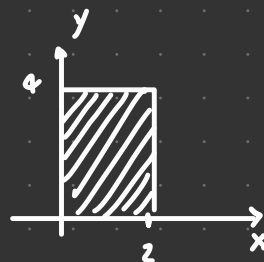
\nearrow SEE BEFORE

EX 2 SESSION JULY 15, 2022

$$f_{(X,Y)}(x,y) = \begin{cases} \frac{y}{16} & \text{if } (x,y) \in (0,2) \times (0,4) \\ 0 & \text{otherwise} \end{cases}$$

ALREADY DONE

$$f_X(x) = \begin{cases} 1/2 & \text{if } x \in (0,2) \\ 0 & \text{otherwise} \end{cases}$$



2a) CHECK THAT $f_{(X,Y)}$ IS DENSITY AND FIND f_X, f_Y

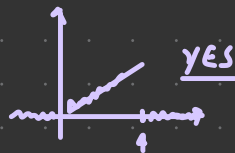
ALREADY DONE

Y FIXED

$$f_Y(y) = \int_0^2 \frac{y}{16} dx = \frac{y}{16} [x]_0^2 = \frac{y}{8}$$

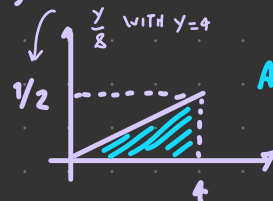
$$f_Y(y) = \begin{cases} y/8 & \text{if } y \in (0,4) \\ 0 & \text{otherwise} \end{cases}$$

EXTRA: CHECK THE COMPUTATION IS IT TRUE THAT $f_Y(y) \geq 0$ (AKA NON NEGATIVE)



$$\int_0^4 \frac{y}{8} dy = \frac{1}{8} \left[\frac{y^2}{2} \right]_0^4 = 1$$

6) CALCULATE AREA OF RECTANGLE



$$\text{AREA} = 2 \cdot \frac{1}{2} = 1$$

2b) FIND $E(Y)$ AND $\text{VAR}(Y)$

$$E[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy = \int_0^4 y \cdot \frac{y}{8} dy = \frac{1}{8} \int_0^4 y^2 dy = \frac{1}{8} \left[\frac{y^3}{3} \right]_0^4 = \frac{1}{8} \cdot \frac{64}{3} = \frac{8}{3}$$

$$\text{VAR}(Y) = E[Y^2] - (E[Y])^2 = 8 - \left(\frac{8}{3} \right)^2 = \frac{72 - 64}{9} = \frac{8}{9}$$

$$\int_{-\infty}^{\infty} y^2 f_Y(y) dy = \int_0^4 y^2 \cdot \frac{y}{8} dy = \frac{1}{8} \left[\frac{y^4}{4} \right]_0^4 = 8$$

$$\text{VAR}(Y) = 8 - \left(\frac{8}{3} \right)^2$$

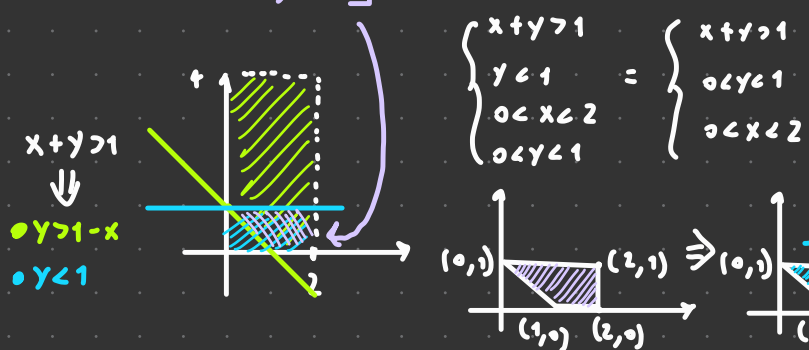
$$3a) P(X+Y > 1 | Y < 1) = \frac{P(X+Y > 1, Y < 1)}{P(Y < 1)}$$

$$P(Y < 1) = \int_0^1 f_Y(y) dy = \int_0^1 \frac{y}{8} dy = \frac{1}{8} \left[\frac{y^2}{2} \right]_0^1 = \frac{1}{16}$$

$$P(X+Y > 1, Y < 1) = \int_0^1 \int_{1-y}^2 \frac{y}{16} dx dy = \int_0^1 \frac{y}{16} (2 - (1-y)) dy = \int_0^1 \frac{y}{16} (1+y) dy = \frac{1}{16} \left[\frac{y^2}{2} + \frac{y^3}{3} \right]_0^1 = \frac{5}{96}$$

$$= \frac{\frac{5}{96}}{\frac{1}{16}} = \frac{5}{6}$$

$$P(X+Y > 1, Y < 1) =$$



$$P(X+Y > 1, Y < 1) = P[(X,Y) \in B] = \iint_B f_{(X,Y)}(x,y) dx dy = \iint_T f_{(X,Y)}(x,y) dx dy + \iint_R f_{(X,Y)}(x,y) dx dy = \int_0^1 \int_{1-y}^2 \frac{y}{16} dx dy + \int_1^2 \int_0^1 \frac{y}{16} dx dy =$$

$$= \frac{1}{16} \int_0^1 \left[\frac{x^2}{2} \right]_{1-y}^2 dy + \frac{1}{16} \int_1^2 \left[\frac{x^2}{2} \right]_0^1 dy = \frac{1}{32} \int_0^1 (4 - (1-y)^2) dy + \frac{1}{32} \int_1^2 1 dy = \frac{1}{32} \int_0^1 (3 + 2y - y^2) dy + \frac{1}{32} \int_1^2 1 dy = \frac{1}{32} \left[3y + y^2 - \frac{y^3}{3} \right]_0^1 + \frac{1}{32} \left[y \right]_1^2 = \frac{1}{32} \left(3 + 1 - \frac{1}{3} \right) + \frac{1}{32} (2 - 1) = \frac{1}{32} \left(\frac{10}{3} + 1 \right) = \frac{13}{96}$$

2b) $S = X+Y$; $T = Y$ (S, T NEW TRANSFORMED RANDOM VECTOR). FIND JOINT DENSITY $f_{(S,T)}$

$$h_{(X,Y)} = (X+Y, Y) \text{ LINEAR TRANSFORMATION}$$

$$\begin{pmatrix} X+Y \\ Y \end{pmatrix} = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} X \\ Y \end{pmatrix}$$

h IS INVERTIBLE (ONE-TO-ONE) IFF M IS NON-SINGULAR

$$\det(M) \neq 0 \quad \det(M) = (1 \cdot 1 - 1 \cdot 0) = 1 \Rightarrow M \text{ IS INVERTIBLE}$$

FIND THE INVERSE

$$h(x, y) = \begin{pmatrix} u \\ v \end{pmatrix} = M \begin{pmatrix} x \\ y \end{pmatrix} = \begin{pmatrix} x + y \\ y \end{pmatrix} \quad \rightarrow \quad h^{-1}(w, v) = \begin{cases} x = w - v \\ y = v \end{cases}$$

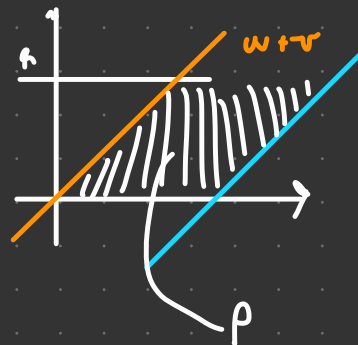
$$h^{-1}(w, v) = M^{-1} \begin{pmatrix} w \\ v \end{pmatrix} \quad \begin{cases} x + y = w; & x = w - v \\ y = v \end{cases}$$

COMPUTE THE JACOBIAN

$$h^{-1}(w, v) = \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} w \\ v \end{pmatrix}$$

$$\text{JAC}[h^{-1}] = M^{-1} = \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} \quad \left| \det(\text{JAC}(h^{-1})) \right| = \left| \det \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} \right| = 1$$

$$f_{(s, r)}(w, v) = f_{(x, y)}(w - v, v) \cdot 1 = \begin{cases} \frac{v}{16} & \text{if } 0 \leq w - v \leq 2 \\ & 0 \leq w \leq 4 \\ 0 & \text{OTHERWISE} \end{cases}$$



$$0 \leq w - v \leq 2 \Leftrightarrow v \leq w \\ v - v \leq 2 \Leftrightarrow w \leq 2 + v$$

$$\rho = \begin{cases} \frac{v}{16} & \text{if } w, v \in \rho \\ 0 & \text{OTHERWISE} \end{cases}$$