# Bayesian Methods for System Identification

Optimisation and Machine Learning in Process Systems
Engineering

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29th September 2020



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Linear Regression

Ridge Regression

Extension to Gaussian processes

Stable Spline Kernel

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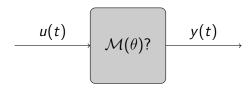
Stable Spline Kernel

Introduction

#### Problem Statement

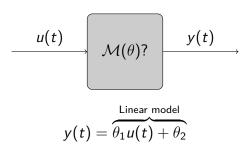
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#### Problem Statement



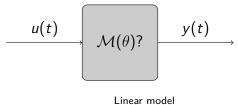
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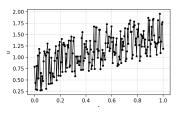
$$y(t) = \overbrace{\theta_1 u(t) + \theta_2}^{\text{Linear model}}$$

$$y(t) = \underbrace{\theta_1 u(t)^n + \theta_2 u(t)^{n-1} + ... + \theta_n}_{\text{n-th degree polynomial model}}$$

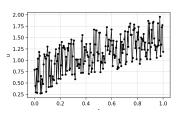
Finite Impulse Response (FIR) Model

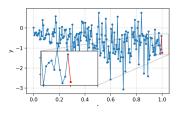
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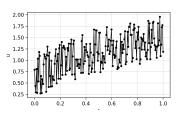


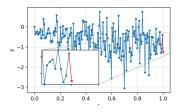
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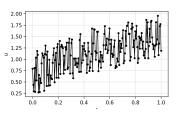


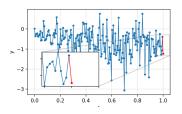


$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + ... + \theta_K u_{n-N} = \sum_{i=0}^{N} \theta_i \cdot u_{n-i}$$

Finite Impulse Response (FIR) Model

Given a time series of inputs  $u_0, u_1, u_2, ..., u_k$  and outputs  $y_0, y_1, y_2, ..., y_{k-1}$ , how to predict  $y_k$ ?





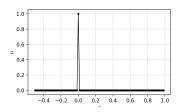
$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + ... + \theta_K u_{n-N} = \sum_{i=0}^N \theta_i \cdot u_{n-i}$$

Since the order K is finite, response is bounded and stable for a bounded input: bounded input, bounded output (BIBO).

$$\delta(t) = \begin{cases} 1, & \text{if } t = 0 \\ 0, & \text{otherwise} \end{cases}$$

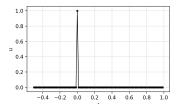
$$\delta(t) = \begin{cases} 1, & \text{if } t = 0 \\ 0, & \text{otherwise} \end{cases}$$
 i.e.  $\mathbf{u} = \begin{bmatrix} 1 & 0 & 0 & 0 & \dots & 0 \end{bmatrix}$ 

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Impulse Signal (Dirac delta function)

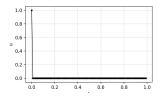
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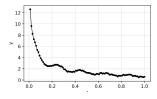


Using our previous FIR model, consider the moving horizon of K time-steps it is clear that eventually the output y will eventually become 0 as t > K, hence finite.

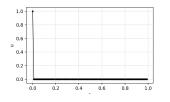
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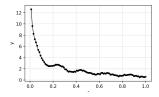
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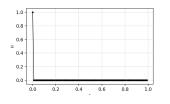
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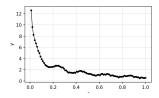




$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

Impulse Signal (Dirac delta function)

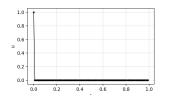




$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

$$y_K = \theta_0 \cdot 0 + \theta_1 \cdot 0 + \dots + \theta_{K-1} \cdot 0 + \theta_K \cdot 1$$

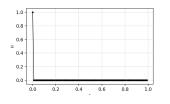
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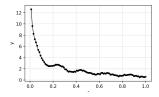


$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

$$y_{K-1} = \theta_0 \cdot 0 + \theta_1 \cdot 0 + ... + \theta_{K-1} \cdot 1 + \theta_K \cdot 0$$

Impulse Signal (Dirac delta function)

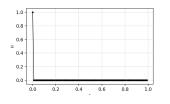




$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

$$y_1 = \theta_0 \cdot 0 + \theta_1 \cdot 1 + ... + \theta_{K-1} \cdot 0 + \theta_K \cdot 0$$

Impulse Signal (Dirac delta function)



$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

$$y_0 = \theta_0 \cdot 1 + \theta_1 \cdot 0 + ... + \theta_{K-1} \cdot 0 + \theta_K \cdot 0$$

$$y_k = \theta_0 u_k + \theta_1 u_{k-1} + \dots + \theta_K u_{n-N}$$

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▶ The parameters  $\theta$  of our FIR model can be seen as the outputs of the system if subjected to an impulse signal  $\rightarrow$  predicting the impulse response.

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- Useful as if you wanted to find out model parameters you just have to 'kick' the system and observe the response.

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- Useful as if you wanted to find out model parameters you just have to 'kick' the system and observe the response.
- ▶ What if we can't 'kick' the system?

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N observations, K inputs: Given input matrix  $\mathbf{U} \in \mathbb{R}^{N \times K}$  and outputs  $\mathbf{y} \in \mathbb{R}^{N \times 1}$ 

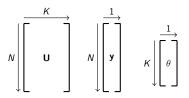
Given input matrix  $\mathbf{U} \in \mathbb{R}^{N \times K}$  and outputs  $\mathbf{y} \in \mathbb{R}^{N \times 1}$  along with a model  $\mathcal{M}$  in the form  $y = \mathbf{u}\theta^T$  where  $\theta \in \mathbb{R}^{K \times 1}$ :

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$$\begin{array}{c|c}
K & \downarrow \\
N & \downarrow \\
V &$$

Formulating the problem as least squares...

$$\hat{\theta} = \arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2$$

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$$\begin{array}{c|c}
K & \downarrow \\
N & \downarrow \\
\end{array}$$

$$\begin{array}{c|c}
K & \downarrow \\
V & \downarrow \\
\end{array}$$

$$\begin{array}{c|c}
K & \downarrow \\
0
\end{array}$$

Formulating the problem as least squares...

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Analytical solution of:

$$\hat{\theta} = (\mathbf{U}^T \mathbf{U})^{-1} \mathbf{U}^T \mathbf{y}$$

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$$\begin{bmatrix}
K \\
U
\end{bmatrix}$$

$$\begin{bmatrix}
K \\
V
\end{bmatrix}$$

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V \\
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What about bias vs variance?  $\rightarrow$  Needs to be considered in linear regression.



# Linear Regression

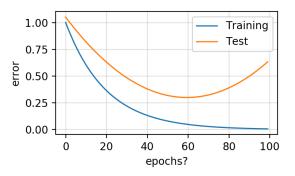
Trading off bias / variance.

For neural networks...

#### Linear Regression

Trading off bias / variance.

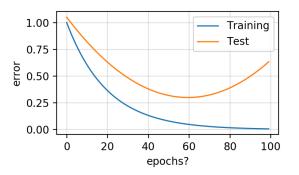
For neural networks...



### Linear Regression

Trading off bias / variance.

For neural networks...



What is the equivalent of epochs for a model with analytical parameters?

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Regularisation technique

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$$\hat{\boldsymbol{\theta}} = \arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2$$

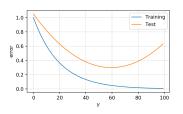
Regularisation technique

$$\hat{\theta} = \arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2 + \underbrace{\gamma ||\boldsymbol{\theta}||^2}_{\text{Ridge term}}$$

Regularisation technique

$$\hat{\theta} = \arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2 + \underbrace{\gamma ||\boldsymbol{\theta}||^2}_{\text{Ridge term}}$$

By varying  $\gamma$ , bias and variance can be traded off.



For any  $\gamma$ , analytical solution of  $\hat{\theta} = (\mathbf{U}^T \mathbf{U} + \gamma \mathbf{I})^{-1} \mathbf{U}^T \mathbf{y}$ 

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$$\underbrace{\gamma ||\boldsymbol{\theta}||^2 = \gamma \boldsymbol{\theta}^T \mathbf{I} \boldsymbol{\theta}}_{\text{Ridge Term}}$$

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What if instead of the identity matrix we used some other matrix?

$$\underbrace{\gamma||\theta||^2 = \gamma\theta^T \mathbf{I}\theta}_{\text{Ridge Term}}$$

What if instead of the identity matrix we used some other matrix?

$$\arg\min_{\boldsymbol{\theta}}||\mathbf{y}-\mathbf{U}\boldsymbol{\theta}||^2+\gamma\boldsymbol{\theta}^T\mathbf{I}\boldsymbol{\theta}$$

$$\underbrace{\gamma||\theta||^2 = \gamma\theta^T \mathbf{I}\theta}_{\text{Ridge Term}}$$

What if instead of the identity matrix we used some other matrix?

$$\arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2 + \gamma \boldsymbol{\theta}^T \mathbf{K}^{-1} \boldsymbol{\theta}$$

$$\underbrace{\gamma||\theta||^2 = \gamma \theta^T \mathbf{I}\theta}_{\text{Ridge Term}}$$

What if instead of the identity matrix we used some other matrix?

$$\arg\min_{\boldsymbol{\theta}} ||\mathbf{y} - \mathbf{U}\boldsymbol{\theta}||^2 + \gamma \boldsymbol{\theta}^T \mathbf{K}^{-1}\boldsymbol{\theta}$$

This has an analytical solution of:

$$\hat{\theta} = (K\mathbf{U}^T\mathbf{U} + \gamma I_K)^{-1}K\mathbf{U}^T\mathbf{y}$$

$$\hat{\theta} = \underbrace{(K\mathbf{U}^T\mathbf{U} + \gamma I_K)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{regularised least squares estimate}}$$

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For jointly Gaussian variables  $(\theta \sim \mathcal{N}(0, K), e_{\mathbf{y}} \sim \mathcal{N}(0, \sigma))$ 

$$\begin{bmatrix} \boldsymbol{\theta} \\ \mathbf{y} \end{bmatrix} \sim \mathcal{N} \left( \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \boldsymbol{K} & \boldsymbol{K} \mathbf{U}^T \\ \mathbf{U} \boldsymbol{K} & \mathbf{U} \boldsymbol{K} \mathbf{U}^T + \sigma^2 \boldsymbol{I}_N \end{bmatrix} \right)$$

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The posterior distribution of  $\theta$  given  $\mathbf{y}$  is given as:

$$heta | \mathbf{y} \sim \mathcal{N}(\hat{ heta}, K^*)$$

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The posterior distribution of  $\theta$  given **y** is given as:

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$$\hat{\theta} = \underbrace{(K\mathbf{U}^T\mathbf{U} + \sigma^2 I_N)^{-1} K\mathbf{U}^T \mathbf{y}}_{\text{Posterior distribution of } \theta}$$

$$\underbrace{\left(\boldsymbol{K}\boldsymbol{\mathsf{U}}^T\boldsymbol{\mathsf{U}} + \gamma\boldsymbol{I}_{\boldsymbol{K}}\right)^{-1}\boldsymbol{K}\boldsymbol{\mathsf{U}}^T\boldsymbol{\mathsf{y}}}_{\text{kernel-regularised least squares estimate}}\underbrace{\left(\boldsymbol{K}\boldsymbol{\mathsf{U}}^T\boldsymbol{\mathsf{U}} + \sigma^2\boldsymbol{I}_{\boldsymbol{N}}\right)^{-1}\boldsymbol{K}\boldsymbol{\mathsf{U}}^T\boldsymbol{\mathsf{y}}}_{\text{Posterior distribution of }\boldsymbol{\theta}}$$

$$\underbrace{(K\mathbf{U}^T\mathbf{U} + \gamma I_K)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{kernel-regularised least squares estimate}}\underbrace{(K\mathbf{U}^T\mathbf{U} + \sigma^2 I_N)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{Posterior distribution of }\theta}$$

They're equivalent! What does this mean?

$$\underbrace{(K\mathbf{U}^T\mathbf{U} + \gamma I_K)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{kernel-regularised least squares estimate}}\underbrace{(K\mathbf{U}^T\mathbf{U} + \sigma^2 I_N)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{Posterior distribution of }\theta}$$

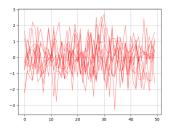
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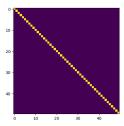
▶ Ridge regression and Gaussian processes are equivalent when the covariance matrix = *I*. e.g. when prior samples look like this...

$$\underbrace{(K\mathbf{U}^T\mathbf{U} + \gamma I_K)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{kernel-regularised least squares estimate}} \underbrace{(K\mathbf{U}^T\mathbf{U} + \sigma^2 I_N)^{-1}K\mathbf{U}^T\mathbf{y}}_{\text{Posterior distribution of }\theta}$$

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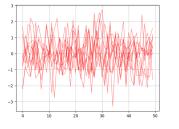
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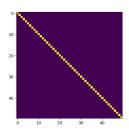




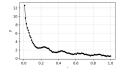
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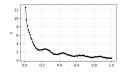




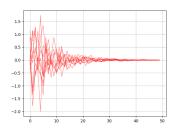
Now we can look at this system identification problem from a Bayesian point of view; can we fully embrace the framework and incorperate prior knowledge into our prediction of  $\hat{\theta}$ ?

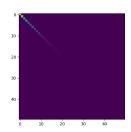


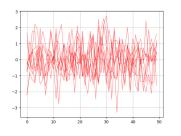
Instead of doing Ridge regression (i.e. using the identity prior) or some other covariance function (i.e. squared exponential) what about if we let it decay exponentially?

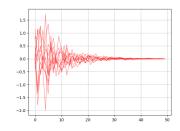


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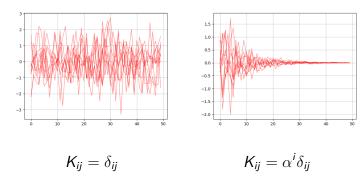






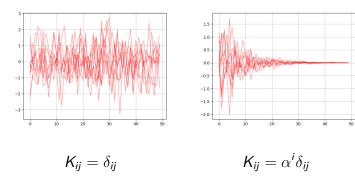
$$K_{ij} = \delta_{ij}$$

$$K_{ij} = \alpha^i \delta_{ij}$$



#### Remaining issue:

Each neighbouring value of  $\theta$ , i.e.  $\theta_1, \theta_2...$  is uncorrelated (off diagonal terms in the covariance matrix are all 0)  $\rightarrow$  non-smooth.



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Each neighbouring value of  $\theta$ , i.e.  $\theta_1, \theta_2...$  is uncorrelated (off diagonal terms in the covariance matrix are all 0)  $\rightarrow$  non-smooth.

This is a false assumption!



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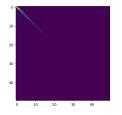
System Identification

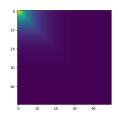
Linear Regression

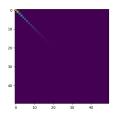
Ridge Regression

Extension to Gaussian processes

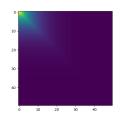
Stable Spline Kernel



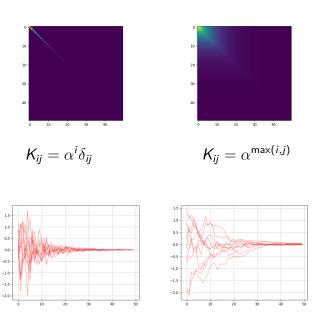




$$K_{ij} = \alpha^i \delta_{ij}$$



$$K_{ij} = \alpha^{\max(i,j)}$$



The stable spline kernel allows for prior information of the FIR parameters  $\theta$  to be included within a Bayesian estimation framework.

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► Stable (parameters decay to 0)

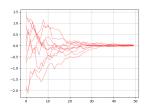
The stable spline kernel allows for prior information of the FIR parameters  $\theta$  to be included within a Bayesian estimation framework.

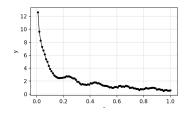
- ► Stable (parameters decay to 0)
- Smooth (neighbouring parameters are correlated)

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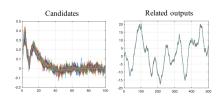
Can also extend the estimation of the parameter vector to continuous time. Continuous sytem identification  $\rightarrow$  function approximation using an informative prior.

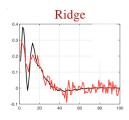




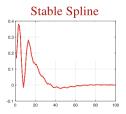
#### Results

100 different impulse responses generated, with a single 'true' response. Note how the predicted outputs with each impulse response varies by very little.









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- Regularised linear regression and Gaussian processes inherently linked
- Incorperating correct assumptions into prior distributions is extremely effective
- Continuous system identification enabled using GPs as a function approximator.

# Thanks!

#### References

- Pillonetto, G., Dinuzzo, F., Chen, T., De Nicolao, G., & Ljung, L. (2014). Kernel methods in system identification, machine learning and function estimation: A survey. Automatica, 50(3), 657–682. https://doi.org/10.1016/j.automatica.2014.01.001
- ▶ Pillonetto, G. (n.d.). REGULARIZED KERNEL-BASED APPROACHES TO SYSTEM IDENTIFICATION. Retrieved from https://www.google.com/url?sa=t& rct=j&g=&esrc=s&source=web&cd=&cad=rja&uact= 8&ved=2ahUKEwjWse28 Y3sAhUKSxUIHXVNCWOQFjABegQIARAB&url=https%3A% 2F%2Fwww.kth.se%2Fsocial%2Ffiles% 2F5b716a5b56be5bd565e07e4f%2Fgianluigi pillonetto sysid2018.pdf&usg= AOvVaw2uRQRyV2Tb9jo3104T06zj