Verjetnostne metode v računalništvu - zapiski s predavanj prof. Marca

Tomaž Poljanšek

študijsko leto 2023/24

Kazalo

1 Introduction **1**
- 1.1 Probability . . . 1
- 1.2 Random variables . . . 2

2 Quicksort, min-cut **4**
- 2.1 Quicksort . . . 4
- 2.2 Min-cut . . . 6

3 Complexity classes **9**

4 Chernoff bounds **11**

5 Monte Carlo methods **16**
- 5.1 Example 1 . . . 16
- 5.2 Example 2 . . . 16
- 5.3 (ϵ,δ)-approximation . . . 17
- 5.4 DNF counting . . . 18

6 Polynomials **20**

7 Random graphs **24**
- 7.1 G(n,p) model . . . 24
- 7.2 Barbási-Albert Model . . . 27

8 Markov chains **29**

8.1 2-SAT 32
8.2 Generating a uniformly random element of a set 33
8.3 Metropolis algorithm 34
8.4 M.C. for 1-factor in bipartite graphs 37

9 Randomized incremented constructions (RIC) 39
9.1 Quicksort as RIC 40
9.2 Linear programming 41

10 Hashing 46
10.1 Chaining 49
10.2 2 level hashing 51
10.3 The power of 2 choices 52
10.4 Cockoo hashing 53

Poglavje 1

Introduction

1.1 Probability

(Ω, F, P_r):

- $\emptyset \in F$,
- $A \in F \implies A^c \in F$,
- $A_1, A_2 \dots \in F \implies \cup_{i=1}^{\infty} A_i \in F$.

$P_r(A) \geq 0$,
$P_r\left(\cup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P_r(A_i)$ if A_i disjoint,
$P_r\left(\cup_{i=1}^{\infty} A_i\right) \leq \sum_{i=1}^{\infty} P_r(A_i)$,
$\Omega = \{\omega_1, \omega_2 \dots\}$ - countable case.
$\begin{pmatrix} \omega_1 & \omega_2 & \dots \\ p_1 & p_2 & \dots \end{pmatrix}$

Primer.

```
Alg():
  while True:
    B = sample as random from {0,1}  # 1 with probability p
    if B = 1:
```

```
return
```

$\Omega = \{1, 01, 001, 0001 \dots\}$
$\begin{pmatrix} 1 & 01 & 001 & 0001 & \dots \\ p & (1-p)p & (1-p)^2p & (1-p)^3p & \dots \end{pmatrix}$.

1.2 Random variables

$X : \Omega \to \mathbb{Z}$.
$E[X] = \sum_{c \in \mathbb{Z}} c \cdot P_r(X = c)$ expected value of X.
Properties:

- $E[f(X)] = \sum_{c \in \mathbb{Z}} f(c) \cdot P_r(X = c)$,
- $E[aX + bY] = aE[X] + bE[Y]$,
- $E[X \cdot Y] = E[X] \cdot E[Y]$ if X, Y independent,
- $P_r(X \geq a) \leq \frac{E[X]}{a}$ $\forall a > 0$ $X \geq 0$ Markov inequality.

Primer. (Continuing from before).
$X =$ number of trials before return.
$X : \Omega \to \mathbb{Z}$.
$X : 1 \to 1, 01 \to 2, 003 \to 3 \dots$
$\begin{pmatrix} 1 & 2 & 3 & 4 & \dots \\ p & (1-p)p & (1-p)^2p & (1-p)^3p & \dots \end{pmatrix}$ - geometric distribution.

Trditev 1.2.1. $E[X] = \frac{1}{p}$.

Dokaz 1.2.2. $X = \sum_{i=1}^{\infty} X_i$.
$X_i = \begin{cases} 1 \text{ if trial } i \text{ is executed} \\ 0 \text{ else} \end{cases}$

$$\begin{aligned} E[X] &= E[\sum_{i=1}^{\infty} X_i] = \sum_{i=1}^{\infty} E[X_i] = \\ &= \sum_{i=1}^{\infty} (1-p)^{i-1} = \frac{i=0}{\infty}(1-p)^i = \frac{1}{1-(1-p)} = \frac{1}{p}. \end{aligned}$$

$E[X] = \frac{1}{p}$.
$P_r(X \geq 100 \cdot \frac{1}{p}) \leq \frac{E[X]}{\frac{1}{p}} = \frac{1}{100}$.

Definicija 1.2.3. $H_n = 1 + \frac{1}{2} + \frac{1}{3} + \cdots + \frac{1}{n} = \sum_{i=1}^{\infty} \frac{1}{i}$.

Izrek 1.2.4. $H_n \leq 1 + \ln(n)$.

Dokaz 1.2.5.

$$H_n = 1 + \sum_{i=2}^{n} \frac{1}{i} \overset{\text{integral}}{\leq} 1 + \int_1^n \frac{dx}{x} = 1 + \ln(x)|_1^n = 1 + \ln(n).$$

Poglavje 2

Quicksort, min-cut

2.1 Quicksort

```
Input: set (no equal element) (unordered list) S∈ ℝ
    (or whatever you can compare linearly)
Output: ordered list
Code:
  def Quicksort(S):
    if |S|= 0 or 1:
      return S
    else:
      a = uniformly at random from S
      S⁻ = {b ∈ S | b < a}
      S⁺ = {b ∈ S | a < b}
      return Quicksort(S⁻), a, Quicksort(S⁺)
```

$C(n)$ - random variable, the number of comparisons in evaluation of Quicksort with $|S| = n$.

Izrek 2.1.1. $E[C(n)] = O\left(N \log(n)\right)$.

Dokaz 2.1.2. $C(0) = C(1) = 0$.

$$E[C(n)] = n - 1 + \sum_{i=1}^{n} (E[C(i-1)] + E[C(n-i)]) \cdot P_r(a \text{ is } i\text{-it element}) \leq$$
$$\leq n + \frac{2}{n} \sum_{i=1}^{n-1} E[C(i)].$$

Induction:

$n = 1 : \checkmark$

$n - 1 \to n$:

$$\begin{aligned}
E[C(n)] &\leq n + \frac{2}{n} \sum_{i=1}^{n} E[C(i)] \leq \\
&\leq n + \frac{2}{n} \sum_{i=1}^{n} 5i \log i \leq \\
&\leq n + \frac{2}{n} \sum_{i=1}^{\lfloor \frac{n}{2} \rfloor} 5i \log i + \frac{2}{n} \sum_{i=1+\lfloor \frac{n}{2} \rfloor}^{n-1} 5i \log i \leq \\
&\leq n + \frac{2}{n} \sum_{i=1}^{\lfloor \frac{n}{2} \rfloor} 5i \log \frac{n}{2} + \frac{2}{n} \sum_{i=1+\lfloor \frac{n}{2} \rfloor}^{n-1} 5i \log n \leq \\
(\log \frac{n}{2} &= \log n - 1) \\
&\leq n + \frac{2}{n} \left(\sum_{i=1}^{n} 5i \log n - \sum_{i=1}^{\frac{n}{2}} 5i \right) = \\
&= n + \frac{10}{n} \left(\frac{n(n-1)}{2} \log n - \frac{\frac{n}{2}(\frac{n}{2}+1)}{2} \right) \leq \\
&\leq n + 5(n-1) \log n - n < \\
&< 5n \log n.
\end{aligned}$$

$P\left(C(n) \geq b \cdot 5n \log n\right) \overset{\text{Markov}}{\leq} \frac{1}{b}$.

Dokaz 2.1.3.

2:

Let $S_1, S_2 \dots S_n$ sorted elements of S.

Define random variable $X_{ij} = \begin{cases} 1 \text{ if } S_i \text{ and } S_j \text{ are compared} \\ 0 \text{ else} \end{cases}$

$C(n) = \sum_{1 \le i < j \le n} E[X_{ij}]$.
$E[X_{ij}] = P(S_i \text{ and } X_j \text{ compared})$.
S_{ij} - the last set including S_i and S_j.
$E[X_{ij}] = \frac{2}{|S_{ij}|} \le \frac{2}{j-i+1}$.
$|S_{ij}| \ge j - i + 1$.
S_{ij} has everything in between.

$$\begin{aligned}
\implies E[C(n)] \le \sum_{1 \le i < j \le n} \frac{2}{j-i+1} &= \\
\overset{k=j-i+1}{=} \sum_{i=1}^{n-1} \sum_{k=2}^{n-i+1} \frac{2}{k} &\le \\
\le 2 \cdot n \cdot H_n &\le \\
\le 2n(1 + \log n).
\end{aligned}$$

2.2 Min-cut

G multigraph.
Cut: $U \subset V(G),\ U \ne \emptyset, V(g)$.
$(U, V(G) \setminus U) = \{uv \in E(G) \mid u \in U, v \in V(G) \setminus U\}$.
Problem min-cut:
Input: G.
Output: $\min |(U, V(G) \setminus U)|$ - cut size.

```
Algorithm 1:
  x ∈ V(G)
  Call maxFlow(G, x, y) ∀ y ∈ V(G)
  Take min
```

maxFlow is Edmonds-Karp algorithm $O(|V||E|^2)$.

```
Algorithm 2 (Stoer Wagner)
```

Is $O(|E||V| + |V| log |V|)$.

```
Algorithm randMinCut:
  G_0 = G
  i = 0
  while |V(G_i)| > 2:
    e_i = uniformly at random from G_i
    G_{i+1} = G_i / e_i
    i = i + 1
  u, v = V(G_{n-2}) // n = |V(G)|
  U = {w ∈ V(G) | w is merged into u}
  return (U, V(G) \ U)
```

Izrek 2.2.1. Algorithm *randMinCut* gives you a minimal cut with probability greater or equal to $\frac{2}{n(n-1)}$.

Dokaz 2.2.2.
Fact 1: $minCut(G_i) \leq minCut(G_i)$;

$\ngtr$: $minCut$ remains.

Fact 2: $minCut(G) \leq \delta(G)$.
$k := minCut(G)$.
Let (A,B) be an optimal cut.
ϵ_i not in (A,B).

$$\begin{aligned}
&P_r(\text{Algorithm not returning } (A,B)) \\
&= P_r(\epsilon_0 \cap \cdots \cap \epsilon_{n-3}) \\
&= P_r(\epsilon_0 \cap \cdots \cap \epsilon_{n-4}) \cdot P_r(\epsilon_{n-3} \mid \epsilon_0 \cap \cdots \cap \epsilon_{n-4}) \\
&= P_r(\epsilon_{n-3} \mid \cap_{i=0}^{n-4} \epsilon_i) \cdot P_r(\epsilon_{n-3} \mid \cap_{i=0}^{n-4} \epsilon_i) \\
&\ldots P_r(\epsilon_1 \mid \epsilon_0) \cdot P_r(\epsilon_0).(*)
\end{aligned} \tag{2.1}$$

$$P_r(\overline{\epsilon_i} \mid \epsilon_{i-1} \cap \cdots \cap \epsilon_0) = \frac{k}{|E(G_i)|} \overset{(**)}{\leq} \frac{k}{\frac{(n-i)k}{2}} = \frac{2}{n-i}$$

$$|E(G_i)| \geq \frac{(n-i)\delta(G)}{2} \geq \frac{(n-i)k}{2}.(**) \tag{2.2}$$

$$P_r(\epsilon_i \mid \epsilon_{i-1} \cap \cdots \cap \epsilon_0) \geq 1 - \frac{2}{n-i} = \frac{n-2-i}{n-i}.$$

$$(*) \geq \frac{n-2}{n} \cdot \frac{n-3}{n-1} \cdots \frac{1}{3} = \frac{2}{n(n-1)}.$$

Izrek 2.2.3. Running *randMinCut* $n(n-1)$ times and taking best output gives correct solution with probability ≥ 0.86.

Dokaz 2.2.4. A_i - event that i-th run gives sub-optimal solution.

$$\begin{aligned} P_r(\text{solution not correct}) &= P_r(A_1 \cap \cdots \cap A_{n(n-1)}) \\ &= \prod_{i=1}^{n(n-1)} P_r(A_i) \leq (1 - \frac{2}{n(n-1)})^{n(n-1)} \\ &\leq e^{-\frac{2}{n(n-1)} \cdot n(n-1)} = e^{-2} \leq 0.14. \end{aligned}$$

$1 - x \leq e^x \; \forall x \in \mathbb{R}$.

If we run $n(n-1)log(n)$ times $\rightarrow O\left(\frac{1}{n}\right)$.
$O\left(n^2 \log n \cdot n\right)$.
Improved: $O\left(n^2 \log^3 n\right)$.

Poglavje 3

Complexity classes

Decision problem - yes/no question on a set of inputs = asking $w \in \Pi$.
Randomized algorithms:

- Las Vegas algorithms: always gives correct solution, example: *Quicksort*.
- Monte Carlo algorithms: it can give wrong answers. Monte Carlo algorithms subtypes:
  - type(1): $\begin{cases} \text{if } \omega \in \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & \geq \frac{1}{2} \\ \text{if } \omega \notin \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & = 0 \end{cases}$
  - type(2): $\begin{cases} \text{if } \omega \in \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & = 1 \\ \text{if } \omega \notin \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & \leq \frac{1}{2} \end{cases}$
  - type(3): $\begin{cases} \text{if } \omega \in \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & \geq \frac{3}{4} \\ \text{if } \omega \notin \Pi \implies \text{algorithm returns „}\omega \in \Pi\text{“ with probability} & \leq \frac{1}{2} \end{cases}$

  type(1) and type(2): one-sided error, type(3): 2-sided error.
  $\frac{1}{2}$, $\frac{3}{4}$ and $\frac{1}{4}$ arbitrary numbers, can be something different (for type(3) better than coin flip).

Primer. Decisional problem: does a graph G have $minCut \leq k$?

Run $randMinCut(G)$ $n(n-1)$ times.

```
Algorithm randMinCut:
  if one of runs gives |(A,B)| ≤ k:
    return true
  else:
    return false
```

Complexity classes:

- RP (randomized polynomial time): decisional problems for which there exists Monte Carlo algorithm of type(1) with polynomial time complexity (worst case).
- co-RP: decisional problems for which there exists Monte Carlo algorithm of type(2) with polynomial time complexity (worst case).
- BRP (bounded-error probabilistic polynomial time): decisional problems for which there exists Monte Carlo algorithm of type(3) with polynomial time complexity (worst case).
- ZPP (zero-error probabilistic polynomial time): decisional problems for which there exists Las Vegas algorithm with expected polynomial time complexity (worst case).

ZPP = RP $\cap$ co-RP.

Poglavje 4

Chernoff bounds

Izrek 4.0.1. Let $X_1, X_2 \dots X_n$ independent random variables with image $\{0, 1\}$.
Let $p_i = P_r(X_i = x_i), X = \sum_{i=1}^{n} X_i$ and $\mu = E(X) = p_1 + \cdots + p_n$.
For every $\delta \in$ (0,1):

$$
\begin{aligned}
P_r(X - \mu \geq \delta\mu) &\leq e^{-\frac{\delta^2\mu}{3}} \\
P_r(\mu - X \leq \delta\mu) &\leq e^{-\frac{\delta^2\mu}{2}} \\
\Longrightarrow P_r(|X - \mu| \geq \delta\mu) &\leq e^{-\frac{\delta^2\mu}{3}}.
\end{aligned}
$$

Probability falls extremely quickly after $E(X)$.

Dokaz 4.0.2.

$$\begin{aligned}
P_r(X-\mu \geq \delta\mu) &= P_r(X \geq \mu(1+\delta)) \\
&\overset{t\geq 0}{=} P_r(tX \geq t\mu(1+\delta)) \\
&\overset{e^y\geq 0}{=} P_r(e^{tX} \geq e^{t\mu(1+\delta)}) \\
&\overset{\text{Markov}}{\leq} \frac{E\left(e^{tX}\right)}{e^{t\mu(1+\delta)}} \\
&\overset{4.1}{\leq} \frac{e^{(e^t-1)\mu}}{e^{t\mu(1+\delta)}} \\
&\overset{4.3}{\leq} e^{-\mu\frac{\delta^2}{3}}.
\end{aligned}$$

$$\begin{aligned}
E(e^{tX}) &= E(e^{tX_1+\cdots+tX_n}) \\
&= E(e^{tX_1}\ldots e^{tX_n}) \\
&\overset{\text{independent}}{=} \prod_{i=1}^{n} E(e^{tX_i}) \\
&\overset{4.2}{\leq} \prod_{i=1}^{n} e^{p_i(e^t-1)} \\
&= e^{(e^t-1)\sum_{i=1}^{n} p_i} \\
&= e^{(e^t-1)\mu}.
\end{aligned} \tag{4.1}$$

$$E(e^{tX_i}) = p_i \cdot e^t + (1-p_i)\cdot e^0 = 1 + p_i(e^t-1) \overset{1+x\leq e^x}{\leq} e^{p_i(e^t-1)}. \tag{4.2}$$

Want:

$$e^t - 1 - t(1+\delta) \leq -\frac{\delta^2}{3} \ \forall\delta \in (0,1) \tag{4.3}$$

$$\begin{aligned}
&t = \ln(1+\delta) \\
&f(\delta) = 1+\delta-1-(1+\delta)\ln(1+\delta) + \frac{\delta^2}{3} \overset{?}{\leq} 0 \\
&f(0) = 0 \\
&f^{'}(\delta) = 1-\ln(1+\delta) - 1 + \frac{2}{3}\delta = \frac{2}{3}\delta - \ln(1+\delta) \overset{?}{\leq} 0 \\
&\frac{2}{3}\delta \leq \ln(1+\delta) \\
&\delta = 1 : \ \frac{2}{3} \overset{?}{\leq} \ln(2) \approx 0.69\checkmark
\end{aligned}$$

$$\begin{aligned}P_r(\mu - X \leq \delta\mu) &= P_r(X \geq \mu(1-\delta)) \\ &\overset{t\geq 0}{=} P_r(tX \geq t\mu(1-\delta)) \\ &\overset{e^y\geq 0}{=} P_r(e^{tX} \geq e^{t\mu(1-\delta)}) \\ &\leq \cdots \leq \frac{e^{(e^t-1)\mu}}{e^{t\mu(1-\delta)}}.\end{aligned}$$

Want: $e^t - 1 - t(1-\delta) \leq -\frac{\delta^2}{2}$ $\forall \delta \in (0,1)$:

$$\begin{aligned}&t = \ln(1-\delta) \\ &f(\delta) = 1 - \delta - 1 - (1-\delta)\ln(1-\delta) + \frac{\delta^2}{2} \overset{?}{\leq} 0 \\ &f(0) = 0 \\ &f'(\delta) = -1 + 1 - \ln(1-\delta) + \delta \overset{?}{\leq} 0 \\ &\frac{2}{3}\delta \leq \ln(1+\delta) \\ &\ln(1-\delta) \overset{?}{\leq} -\delta \checkmark\end{aligned}$$

$\blacksquare$

$X_i \sim \begin{pmatrix} 0 & 1 \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix}$
$X = \sum_{i=1}^n X_i$
$\mu = \frac{n}{2}$

$$\begin{aligned}P_r(|X-\mu| \geq \sqrt{\frac{3}{2}n\ln(n)}) &= P_r(|X-\mu| \geq \frac{n}{2}\sqrt{\frac{6}{n}\ln(n)}) \\ &\mu = \frac{n}{2}, \delta = \sqrt{\frac{6}{n}\ln(n)}, \\ &\text{for „big“ } n\delta \in (0,1) \\ &\overset{\text{Chernoff}}{\leq} 2e^{-\frac{\frac{n}{2}\frac{6}{n}\ln(n)}{3}} = \frac{2}{n}.\end{aligned}$$

$d = \sqrt{\frac{3}{2}n\ln(n)}$

$$\implies P_r(X \in (\mu - \sqrt{\frac{3}{2}n\ln(n)}, \mu + \sqrt{\frac{3}{2}n\ln(n)})) \geq 1 - \frac{2}{n}.$$

Trditev 4.0.3.
Let $X_1, X_2 \dots$ independent random variables with image {0,1}.
$P_r(X_i = 1) = \frac{1}{2}\ \forall i$.
Let $X = \sum_{i=1}^{cm} X_i$ where $c \geq 4$.
Then $P_r(X \leq m) \leq e^{-\frac{cm}{16}}$.

Dokaz 4.0.4.

$$\begin{aligned}
P_r(X \leq m) &= P_r(\frac{cm}{2} - X \geq \frac{cm}{2} - m) \\
&= P_r(\frac{cm}{2} - X \geq \frac{cm}{2}(1 - \frac{2}{c})) \\
&\overset{\text{Chernoff}}{\leq} e^{-\frac{\frac{cm}{2}(1-\frac{2}{c})^2}{2}} \\
&\quad 1 - \frac{2}{c} \geq \frac{1}{2} \text{ if } c \geq 4 \\
&\leq e^{-\frac{\frac{cm}{2}\frac{1}{4}}{2}} = e^{-\frac{cm}{16}}.
\end{aligned}$$

■

Back to Quicksort.

Izrek 4.0.5.
With probability $\geq 1 - \frac{1}{n}$ Quicksort uses at most $48n\ln(n)$ comparisons.

Dokaz 4.0.6.
For $s \in S$ define $S_1^S \dots S_{t_s}^S \neq \emptyset$ sets that include s, t_s - number of comparisons with s where s is not a pivot +1.
Define: iteration i is successful if $|S_{i+1}| \leq \frac{3}{4}|S_i|$ ($\frac{1}{2}$ is too strict).

$$X_i = \begin{cases} 1 \text{ if iteration } i \text{ is successful} \\ 0 \text{ else} \end{cases}$$

$P_r(X_i = 1) \geq \frac{1}{2}$
$S_i : n \to \frac{3}{4}n \to (\frac{3}{4})^2 n \to \dots \to 1$.
Notice: max number of iteration is $\log_{\frac{4}{3}}(n) = \frac{\ln(n)}{\ln(4)-\ln(3)}$.

Probability that we haven't succeeded in $\log_{\frac{4}{3}}(n)$ steps:

$$P_r(\sum_{i=1}^{c\log_{\frac{4}{3}}(n)} X_i < \log_{\frac{4}{3}}(n)) \leq P_r(\sum_{i=1}^{c\log_{\frac{4}{3}}(n)} Y_i < \log_{\frac{4}{3}}(n)) \tag{4.4}$$

$$\overset{\text{Chernoff}}{<} e^{-\frac{c\log_{\frac{4}{3}}(n)}{24}} \tag{4.5}$$

$$= e^{-\frac{c\ln(n)\log_{\frac{4}{3}}(e)}{24}} \tag{4.6}$$

$$= \frac{1}{n}\frac{c\log_{\frac{4}{3}}(e)}{24} \tag{4.7}$$

$$\log_{\frac{4}{3}}(e) \approx 3.4,\ c = 14 \tag{4.8}$$

$$\leq \left(\frac{1}{n}\right)^2 \tag{4.9}$$

4.4 because X_i not independent, $Y_i \sim \begin{pmatrix} 0 & 1 \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix}$ independent.

$P_r(t_s \geq c\log_{\frac{4}{3}}(n)) \geq \left(\frac{1}{n}\right)^2$ for one s.

$c = 14 \implies$ at least $48\ln(n)$ iterations with probability $\leq \left(\frac{1}{n}\right)^2$.

With probability as least $1 - \frac{1}{n}$ for all $s \in S$ it holds that s has $\leq 48\ln(n)$ comparisons with a pivot.

$\implies$ total number of comparisons $n \cdot 48\ln(n)$ with probability as least $1 - \frac{1}{n}$. ■

Poglavje 5

Monte Carlo methods

5.1 Example 1

Area of circle $= \frac{\pi}{4}$.

$X_i = \begin{cases} 1 \text{ if you hit the area of circle} \\ 0 \text{ else} \end{cases}$

$P_r(X_i = 1) = \frac{\frac{\pi}{2}}{1} = \frac{\pi}{4}$.

$E(X_i) = \frac{\pi}{4}$.

$X = \frac{\sum_{i=1}^{n} X_i}{n}$.

$E(X) = \frac{n \cdot E(X_i)}{n} = E(X_i)$.

5.2 Example 2

$I = \int_\Omega f(x)dx$ - volume.

$X_i = \begin{cases} 1 \ F(x_i, y_i) \leq z_i \\ 0 \text{ otherwise} \end{cases}$

$v \cdot E\left(\frac{\sum_{i=1}^{n} X_i}{n}\right) = I$.

5.3 (ϵ,δ)-approximation

Definicija 5.3.1 ((ϵ,δ)-approximation)**.** A random algorithm gives a (ϵ,δ)-approximation for value v if the output X satisfies:

$$P_r(|X - v| \leq \epsilon v) \geq 1 - \delta.$$

Izrek 5.3.2. Let $X_1 \dots X_n$ be independent and identically distributed indicator variables. Let $\mu = E(X_i)$, $Y = \frac{\sum_{i=1}^{m} X_i}{m}$. If $m \geq \frac{3\ln\left(\frac{2}{\delta}\right)}{\epsilon^2\mu}$, then $P_r(|Y - \mu| \geq \epsilon\mu) \leq \delta \implies Y$ is (ϵ,δ)-approximation for μ.

Dokaz 5.3.3.
$X = \sum_{i=1}^{n} X_i$
$E(X) = mE(x_i) = m\mu$
$m \geq \frac{3\ln\left(\frac{2}{\delta}\right)}{\epsilon^2\mu}$

$$\begin{aligned}
P_r(|Y - \mu| \geq \epsilon\mu) &= P_r(\left|\frac{X}{m} - \mu\right| \geq \epsilon\mu) \\
&= P_r(\frac{1}{m}\left|X - E(X)\right| \geq \frac{1}{m}\epsilon E(x)) \\
&\overset{\text{Chernoff}}{\leq} 2e^{-\frac{\epsilon^2 E(x)}{3}} \\
&= 2e^{-\frac{\epsilon^2\mu m}{3}} \\
&\leq 2e^{-\frac{\epsilon^2\mu}{3}\cdot\frac{3\ln\left(\frac{2}{\delta}\right)}{\epsilon^2\mu}} = \delta.
\end{aligned}$$

Back to example 1:
$E(Y) = \frac{\pi}{4}, \delta = \frac{1}{1000}$ (99.9% sure), $\epsilon = \frac{1}{10000}$
$\implies M = \frac{3\ln\left(\frac{2}{\frac{1}{1000}}\right)4}{\pi\left(\frac{1}{10000}\right)^2} \approx 29106.$
Problems for MC (Monte-Carlo):

- rare events, e.g. $X \sim \begin{pmatrix} 0 & 10^{100} \\ 1 - 10^{-20} & 10^{-20} \end{pmatrix}$, $E(X) = 10^{80}$

5.4 DNF counting

CNF: $(X_{i_1} \vee \overline{X_{i_2}} \vee X_{i_4}) \wedge (X_{i_1} \vee \overline{X_{i_3}}) \wedge \dots$
DNF: $(\overline{X_{i_1}} \wedge X_{i_2} \vee \overline{X_{i_4}}) \vee \dots$ - easy to determine if solution exists.
Question: number of solutions to a given DNF?
Observation: CNF F has a solution $\iff$ DNF $\neg F$ has less than 2^n solutions, n is number of samples.

```
ALG_1(F):
  x = 0
  for i in range(1,m+1):
    x_1 ... x_n uniformly random from {0,1}^n
    if F(x_1 ... x_n) = 1:
      x += 1
  return x/m · 2^n
```

$Y = \frac{\sum_{i=1}^{m} X_i}{m}$
(ϵ,δ)-approximation for Y
$E(Y) = \frac{\text{number of solutions of } F}{2^n} = \frac{c(F)}{2^n}$
$m \geq \frac{3\ln\left(\frac{2}{\delta}\right)}{\epsilon^2 E(X)} = \frac{3\ln\left(\frac{2}{\delta}\right)}{\epsilon^2} \cdot \frac{2^n}{x(F)}$
$c(F)$ very small $\rightarrow$ m exponentially big $\rightarrow$ not good (we need a lot of samples).

Definicija 5.4.1.
$SC_i = \{(a_1 \dots a_n) \in \{0,1\}^n$ such that $F = F_1 \vee \dots \vee F_t,\ F_i(a_1 \dots a_n) = 1\}$.

$|SC_i| = 2^{n-l_i}$, l_i: number of values in F_i
$U = \{(i,a) \mid i \in \{1,2 \dots t\}, a \in SC_i\}$
$U = \sum_{i=1}^{t} |SC_i|$ - $O(tn)$ (space smaller than $\{0,1\}^n$)
$S = \{(i,a) \in U \mid a \in SC_i,\ a \notin SC_j\ 1 \leq j < i\}$
$|S| = |SC_1| + \dots + |SC_t| = c(F)$.

```
ALG_2(F):
  x = 0
  for i in range(1,m+1):
```

```
    (i, a) uniformly random from U (**)
    if (i, a) ∈ S: (*)
      x += 1
  return x/m · |U|
```

$(*)$ $a \in SC_i \to O(n)$, $a \notin SC_j$ $j = 1 \dots i-1 \to O(tn) \implies O(tn), m$ times.

$(**)$: watch for details on how to, e.g. $x_2, x_2 \wedge x_3$: x_2 is more probable than $x_2 \wedge x_3 \to O(1)$.

Izrek 5.4.2. For $m = \lceil \frac{3t \ln\left(\left(\frac{2}{\delta}\right)\right)}{\epsilon^2} \rceil$ algorithm returns (ϵ,δ)-approximation in $O\left(\frac{t^n n \ln\left(\frac{2}{\delta}\right)}{\epsilon^2}\right)$ time.

Dokaz 5.4.3. $O(t \cdot n \cdot m)$.
Insert $m = ...$

Prove

$$P_r(Y|U| - c(F) > \epsilon c(F)) < \delta :$$

$c(F) = |S|, E(Y) = \frac{|S|}{|U|}$

$$P_r(Y|U| - c(F) > \epsilon c(F)) = P_r(|U|(Y - E(Y)) > \epsilon |U| E(Y)) \leq \delta$$

if

$$m \geq \frac{3 \ln\left(\frac{2}{\delta}\right)}{\epsilon^2 E(Y)} \geq \frac{3 \ln\left(\frac{2}{\delta}\right) t}{\epsilon^2}$$

where

$$E(Y) = \frac{|S|}{|U|} \geq \frac{1}{t}$$

(= if disjoint).
In new space $E(Y)$ much larger $\implies$ m smaller.

Poglavje 6

Polynomials

Let $\mathbb{F}$ be a field.
$\mathbb{F}$ can be $\mathbb{R}, \mathbb{C}, \mathbb{Z}_p, \mathbb{F}_{p^n}$.
$\mathbb{F}[x_1 \dots x_n]$ algebra of polynomials with values $x_1 \dots x_n$.
$f \in \mathbb{F}[x_1 \dots x_n]$
$deg(f[x_1 \dots x_n]) := deg(f[x \dots x])$.

Izrek 6.0.1. Let $p(x_1 \dots x_n) \in \mathbb{F}[x_1 \dots x_n]$ have the degree $d \geq 0$ and $p \neq 0$. Let $s \subset \mathbb{F}$ be finite. If $(r_1 \dots r_n)$ is uniformly at random element from S^n. Then $P_r(p(r_1 \dots r_n) = 0) \leq \frac{d}{|S|}$.

Dokaz 6.0.2. Induction on n.
$n = 1$:

$$\begin{aligned} & p(x) = (x - z_1)(x - z_2) \dots (x - z_j) q(z) \\ & \text{number of zeros } \leq \text{ degree - fact} \\ & P_r(p(r_1) = 0) = \frac{\text{number of zeros}}{|S|} \leq \frac{d}{|S|}. \end{aligned}$$

$n - 1 \to n$:

$$\begin{aligned}
&\text{rewrite } p: \\
&p(x_1 \dots x_n) = \sum_{i=0}^{j} x^i p_i(x_2 \dots x_n) \\
&j \le d
\end{aligned}$$

$$\begin{aligned}
P_r(p(r_1 \dots r_n) = 0) &= P_r(p(r_1 \dots r_n = 0) \mid p_j(r_2 \dots r_n) = 0) \cdot P_r(p_j(r_2 \dots r_n) = 0) \\
&+ P_r(p(r_1 \dots r_n = 0) \mid p_j(r_2 \dots r_n) \neq 0) \cdot P_r(p_j(r_2 \dots r_n) \neq 0) \\
&\le 1 \cdot \frac{d-j}{|S|} + \frac{j}{|S|} \cdot 1,
\end{aligned}$$

because

$$\begin{aligned}
&P_r(p(r_1 \dots r_n = 0) \mid p_j(r_2 \dots r_n) \neq 0) \le \frac{d-j}{|S|} \\
&P_r(p_j(r_2 \dots r_n) \neq 0) \le \frac{j}{|S|}.
\end{aligned}$$

Problem:

Let A,B,$C \in \mathbb{F}^{n \times n}$, is $A \cdot B = C$?

Computing $A \cdot B$:

- school-book algorithm: $O(n^3)$,
- Strassen algorithm: $O(n^{2,807\dots})$,
- galactic algorithm: $O(n^{2.372\dots})$ - has enormous constants.

```
RAND_ACB(A,B,C):
  for i in range(1,k+1):
    x uniformly at random from {0,1}^n
    if A · (B · x) ≠ x:
      return false
  return true
```

$O\left(kn^2\right)$.

If $A \cdot B = C$, algorithm returns true.
If $A \cdot B \neq C$:

$$\begin{aligned} P_r(ABx = Cx) &= P_r((AB - C)x = 0) \\ &= P_r(||(AB - C)x||^2 = 0) \overset{\text{Poly}}{\leq} \frac{2}{3}. \end{aligned}$$

$||(AB_C)x||^2$ - polynomial in $x_1 \ldots x_n$ of degree 2.
If $A \cdot B \neq C$, then algorithm return false with probability at least $1 - \left(\frac{2}{3}\right)^k$.
<u>Problem</u>:
1-factor in bipartite graphs.
$|V(g)| = 2n$.
Represent G with $n \times n$ matrix $Z = (Z_{ij})_{i,j=1}^n$

$$Z_{ij} = \begin{cases} X_{ij} \text{ if } a_i b_j \in E(x) & \text{(X: variable)} \\ 0 \text{ else} \end{cases}$$

$$\begin{aligned} detZ(x_{11} \ldots x_{nn}) &= \sum_{\pi \in S_n} sign(\pi) z_{1,\pi(1)} \ldots z_{n,\pi(n)} \\ &= \sum_{\pi \in S_n, \pi \text{ defines 1-factor}} sign(\pi) x_{1,\pi(1)} \ldots x_{n,\pi(n)}. \end{aligned}$$

$detZ \neq 0 \iff G$ has 1-factor.

```
Rand_1factor(G):
  construct Z with variables x11 ... xnn
  for i in range(1,k+1):
    u <- uniformly at random from  1,2..2n-1^(n^2) (r11 ... rnn)
    compuze d = det Z(r11 ... rnn)
    if d != 0:
      return true
  return false
```

Complexity: $k\cdot$ computing determinant: $O\left(n^3\right)$ (Gaussian elimination).
or apply approximation algorithm:

- if G has no 1-factor it always returns false,
- if G has 1-factor, it returns true with probability at least $1-\left(\frac{n}{2n}\right)^k = 1-\left(\frac{1}{2}\right)^k$ (k konstant, larger set $\implies$ smaller k needed).

Poglavje 7

Random graphs

7.1 G(n,p) model

G is a random Erdös-Rény graph if it has n vertices and each pair of vertices is connected with probability p.

Primer. $G\left(5, \frac{1}{2}\right)$.

$E($ edges in G fron $G(n,p)) = \sum_{1 \leq i < j \leq n} E(X_{ij}) = \binom{n}{2} p$.

$X_{ij} = \begin{cases} 1 \text{ if } i \text{ and } j \text{ have edge} \\ 0 \text{ otherwise} \end{cases}$

p can be function of n.

Y_v : degree of v.

$E(Y_v) = (n-1)p$.

Definicija 7.1.1.

We say that a random graph has some property almost surely (A.S.) if $P_r(G \in G(n,p)$ has property$) \overset{n \to \infty}{\to} 1$.

Trditev 7.1.2.

Let p be constant. Then $G \in G(n,p)$ has diameter 2 A.S.

Dokaz 7.1.3.

Let $u,v \in V(G)$
$X_w = \begin{cases} 1 \text{ if } uw \in E(G) \text{ in } vw \in E(G) \end{cases}$
$P_r(X_w = 1) = p^2$
$P_r(X_w = 0 \text{ for all } w \neq u,v) = (1-p^2)^{n-2}$.
$P_r(G \text{ has diameter } > 2)$
$= P_r(X_w = 0 \text{ for all } w \notin u,v \text{ for some } u,v)$
$\leq \binom{n}{2}(1-p^2)^{n-2} \overset{n\to\infty}{\to} 0;$
$\binom{n}{2}$ - polynomial, $e^{\cdots}$ - exponent.

$p = f(n)$
$\frac{1}{n}, \frac{1}{n^3}, \frac{\log n}{n}$

Izrek 7.1.4. (without proof)
Let p be a function of n: let $G \in G(n,p)$:

- $np < 1$ - G A.S. disconnected with connected components of size $O(\log n)$
- $np = 1$ - G A.S. has 1 large component of size $O\left(n^{\frac{2}{3}}\right)$
- $np = c > 1$ - G A.S. has giant component of size dn, $d \in (0,1)$
- $np \leq (1-\epsilon)\ln n$ - G A.S. disconnected with isolated vertices
- $np > (1-\epsilon)\ln n$ - G A.S. connected.

Izrek 7.1.5.
Let $np = \omega(n)\ln(n)$ for $\omega(n) \to \infty$ „very slowly“ think of $\omega(n) = \log(\log n)$, then $diam(G)$ in $\Theta\left(\frac{\ln n}{\ln(np)}\right)$ for G in $G(n,p)$.

Lema 7.1.6.
Let $S \subset V(G), |S| = cn$ for $c \in (0,1]$ and $v \notin S$.
then $cnp(1-\omega^{-\frac{1}{3}}) \leq N_S(v) \leq cnp(1+\omega^{-\frac{1}{3}})$ A.S. ($\omega^{-\frac{1}{3}} \to 0$ very slowly).

Dokaz 7.1.7. (Lemma):
$E(N_s(v)) = c \cdot n \cdot p, \delta = \omega^{-\frac{1}{3}}$

$$P_r(|N_s(v) - cnp| \geq \delta cnp) \overset{\text{Chernoff}}{\leq} 2e^{-\frac{\omega^{-\frac{2}{3}} cnp}{3}}$$
$$= 2e^{-\frac{cnp}{3\omega(n)^{\frac{2}{3}}}} \overset{n\to\infty}{\to} 0.$$

For all v: $n \cdot 2e^{-\frac{cnp}{3\omega(n)^{\frac{2}{3}}}} \overset{n\to\infty}{\to} 0$.

Dokaz 7.1.8. (Theorem):
k be such that $\sum_{i=0}^{k-1} |N_i| \leq \frac{n}{2}, \sum_{i=0}^{k} |N_i| > \frac{n}{2}$.
$|N_0| = 1$
$|N_i| \leq |N_{i-1}| \cdot n \cdot p \cdot (1 + \omega^{-\frac{1}{3}})$:
$|S| \leq n,\ np(1 + \omega^{-\frac{1}{3}})$-each element.
$k = \frac{\log\left(\frac{n}{3}\right)}{\log\left(n \cdot p \cdot \left(1+\omega^{-\frac{1}{3}}\right)\right)}$
$= \log_{np(1+\omega^{-\frac{1}{3}})} \frac{n}{3} = \Theta\left(\frac{\ln(n)}{\ln(np)}\right)$.
$|N_{\leq k}| = |N_1 \cup \cdots \cup N_k|$.

$$\begin{aligned}
|N_{\leq k}| &\leq \sum_{i=0}^{k} (np(1 + \omega^{-\frac{1}{3}}))^i \\
&= \frac{(np(1 + \omega^{-\frac{1}{3}}))^{k+1} - 1}{np(1 + \omega^{-\frac{1}{3}}) - 1} \\
&< \frac{np(1 + \omega^{-\frac{1}{3}})^{k+1}}{\frac{1}{2} np(1 + \omega^{-\frac{1}{3}})} \\
&= 2np(1 + \omega^{-\frac{1}{3}})^k \\
&\overset{k}{=} 2 \cdot \frac{n}{3} \text{ haven't covered all} \\
&\implies diam(G) > k \text{ bound from below.}
\end{aligned}$$

$N_i \subseteq S$
$\frac{1}{2} np \left(1 - \omega^{-\frac{1}{3}}\right) \cdot |N_{i-1}| \leq |N_i|$

$$\begin{aligned}
n &\geq \sum_{i=0}^{k} |N_i| \\
&\geq \sum_{i=0}^{k} \left(\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)\right)^i \\
&= \frac{\left(\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)\right)^{k+1}-1}{\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)-1} \\
&\geq \left(\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)\right)^k \qquad /\ln
\end{aligned}$$

$\frac{\ln n}{\ln(np)} \approx \frac{\ln n}{\ln\left(\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)\right)} \geq k.$
$\implies w \in S'$.
Number of neighbors in N_k A.S. ≥ 1,
$|N_k| \geq \left(\frac{1}{2}np\left(1-\omega^{-\frac{1}{3}}\right)\right)^k \approx c \cdot n$
$\implies diam(G) = k+1$ A.S.

7.1.1 Scale free property

$G \in G(n,p)$.

In real world: $p(k)$ = proportion of degree k vertices.

$\log(p(k)) = -\gamma \cdot \log k$

$p(k) = k^{-\gamma}$.

Internet: $\gamma \approx 3.42$,

protein reactions: $\gamma \approx 2.89$.

7.2 Barbási-Albert Model

B.A. model.

Start with m modes.

Grow:

- add node v,

- add m edges from v (to u),
- for each new edge: $P(v \sim u) = \frac{degu}{\sum_x degx}$.

Izrek 7.2.1.

B.A. model has scale free property, in particular
$p_k = \frac{2m(m+1)}{k(k+1)(k+2)}$.

Definicija 7.2.2.

$p_n(k)$: expected proportion of degree k vertices in graph with k vertices,
$p_k := \lim_{n\to\infty} p_n(k)$.

Dokaz 7.2.3.

$p_n(k) \cdot n$: expected number of degree k vertices,
$p_n(k)n \cdot \frac{k}{\sum_u degu} m = p_n(k) \cdot \frac{k}{2}$: expected number of degree k vertices changing into degree $k+1$ vertices.
$\sum_u degu = 2|E|$
$p_{n+1}(k) \cdot (n+1) = p_n(k) \cdot n - p_n(k) \cdot \frac{k}{2} + p_n(k-1) \cdot \frac{k-1}{2}$, where
$p_n(k) \cdot n$: degree $k \to k$,
$p_n(k) \cdot \frac{k}{2} : k \to k+1$,
$p_n(k-1) \cdot \frac{k-1}{2} : k-1 \to k$.
For n very big (very close to limit):
$p_n \cdot (n+1) = p_k \cdot n - p_{k-1} \cdot \frac{k}{2} + p_{k-1} \cdot \frac{k-1}{2}$
$\implies p_k = \frac{k-1}{k+2} p_{k-1}$.
For degree m:
$(n+1) \cdot p_{n+1}(m) = p_n(m) \cdot n - p_n(m) \cdot \frac{m}{2} + 1$

$$\begin{aligned}
p_m &= \frac{2}{m+2} \\
\implies p_{m+1} &= \frac{2}{m+2} \cdot \frac{m}{m+3} \\
\implies p_{m+2} &= \frac{2m(m+1)}{(m+2)(m+3)} \\
\implies p_k &= \frac{2m(m+1)}{k(k+1)(k+2)}.
\end{aligned}$$

Poglavje 8

Markov chains

Ω: finite set (of states).

Definicija 8.0.1 (Markov chain)**.**
(Discrete time) Markov chain is a sequence of random variables $X = X_0, X_1, X_2 \ldots$ with image Ω and properties:

- $P(X_{i+1} = x \mid X_i = x_i, X_{i-1} = x_{i-1} \ldots X_0 = x_0) =$
  $P(X_{i+1} = x \mid X_i = x_i)$,

- $PX_{i+1} = x \mid X_i = y = P(X_1 = x \mid X_0 = y)$ - time is homogenous.

Primer.
$\Omega = \mathbb{Z}_5$
$P(X_{i+1} = x + 1 \mid X_i = x) = \frac{1}{2}$
$P(X_{i+1} = x - 1 \mid X_i = x) = \frac{1}{2}$.

Definicija 8.0.2 (Transition matrix)**.**
$\Omega = \{x_1 \ldots x_n\}$
$p_{ij} = P(X_{t+1} = j \mid X_t = i)$

$$\begin{bmatrix} p_{11} & \dots & \\ p_{1n} & & \\ \vdots & & \vdots \\ p_{n1} & \dots & p_{nn} \end{bmatrix}.$$

Definicija 8.0.3 (Transition graph)**.**
Edge between states i and j exists if $p_{ij} > 0$.

P is stochastic matrix:
$p_{ij} \in [0,1]$
$\sum_j p_{ij} = 1$.
We choose beginning state randomly.
$q(0) = (q_1(0) \dots q_n(0))$
$P(X_0 = i) = q_i(0)$.
Let $q(t) = (q_1(t) \dots q_n(t))$
$P(X_t = i) = q_i(t)$.
It holds: $q(t) = q(t-1) \cdot P = q(0) \cdot P^t$.

$$\begin{bmatrix} 0 & \frac{1}{2} & 0 & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & 0 & 0 & \frac{1}{2} \end{bmatrix}$$

$q(0) = (1, 0, 0, 0, 0)$
$q(1) = (1, \frac{1}{2}, 0, 0, \frac{1}{2})$
$q(2) = (\frac{1}{2}, 0, \frac{1}{4}, \frac{1}{4}, 0)$
$\vdots$

Definicija 8.0.4.

- Distribution π is stationary if $\pi = \pi \cdot P$,
- f_{ij}: probability that $X_t = x_j$ for some t assuming $X_0 = x_i$,

- h_{ij}: expected number of steps needed to get to state X_j strting in X_i (hitting time),

- $N(i, t, q(0))$: expected number of times we visit x_i after t steps starting with distribution $q(0)$,

- $\forall f_{ij} > 0 \iff$ transition graph is strongly connected $\iff$ we say the chain is irreducible,

- M.C. is aperiodic if there is no $c \in \{2, 3, 4 \dots\}$ such that all lengths of cycled are divisible by c.

Izrek 8.0.5.
Let X be finite irreducible M.C. Then:

a) there exists unique stationary distribution $\pi = (\pi_1 \dots \pi_n)$,

b) $f_{ii} = 1, h_{ii} = \frac{1}{\pi_i}$,

c) $\lim_{t\to\infty} \frac{N(i,t,q(0))}{t} = \pi_i$ - approaches π regardless of $q(0)$,

d) if X is aperiodic: $\lim_{t\to\infty} q(0) \cdot P^t = \pi$.

Primer.

$$P = \begin{bmatrix} 0 & \frac{1}{2} & \dots & \frac{1}{2} \\ \frac{1}{2} & 0 & \dots & 0 \\ \vdots & & & \vdots \\ \dots & & \frac{1}{2} & 0 \end{bmatrix}$$

$\pi = (\frac{1}{n} \dots \frac{1}{n})$

$h_{i,i} = n$

$n = h_{i,i} = 1 + \frac{1}{2}h_{i-1,i} + \frac{1}{2}h_{i+1,i}, \quad h_{i-1,i} = h_{i+1,i}$

$n - 1 = h_{i-1,i}$

$E(\text{steps around}) \leq h_{0,1} + h_{1,2} + \dots + h_{n-1,n} \leq n(n-1)$.

8.1 2-SAT

Recall: k-SAT:

$$F = C_1 \wedge \cdots \wedge C_m$$
$$C_i = X_{i1} \vee \cdots \vee X_{ik}.$$

3-SAT: NP complete.

```
Algorithm:
  def rand2SAT(F):
    b^0 = (b_0^0 ... b_n^0)
    for i in range(t):
      if F(b^i) = 1:
        return True
      Cl <- clause trat is False
      xj <- uniformly at random from xl1 and xl2
      b^{i+1} = (b_0^i ... not b_j^i ... b_n^i)
    if F(X^t) = 1:
      return True
    return False
```

Izrek 8.1.1.
If $k = 8n^2$, then $P(\text{rand2SAT} = \text{True} \mid \text{correct answer is True}) \geq \frac{3}{4}$.

Dokaz 8.1.2. Let $a = (a_1 \dots a_n)$ be a correct solution.
Let X_i = Hamming distance from b^i to a.
Goal: bount $h_{n,0}$.
$P(\text{distance of } b^{i+1} \text{ to } a \text{ is } j-1 \mid \text{distance of } b^i \text{ to } a \text{ is } j) \geq \frac{1}{2}$.
$$P = \begin{bmatrix} 0 & 1 & \dots & 0 \\ \frac{1}{2} & 0 & \dots & 0 \\ \vdots & & & \vdots \\ \dots & & 1 & 0 \end{bmatrix}$$

$\pi \stackrel{?}{=} \pi P$
$\pi = (\frac{1}{2n}, \frac{1}{n} \dots \frac{1}{n}, \frac{1}{2n})$
By theorem
$h_{i,i} = \frac{1}{\pi_i} = n$ for $i = 1, 2 \dots n-1$
$h_{0,0} = h_{n,n} = 2n$
$n = h_{i,i} = 1 + \frac{1}{2}h_{i+1,i} + \frac{1}{2}h_{i-1,i}$
$h_{i+1,i} \leq 2n$
$i = 0: \ 2n = h_{0,0} = 1 + h_{1,0} \implies h_{1,0} < 2n$
$h_{n,0} \leq h_{n,n-1} + \dots + h_{1,0} \leq 2n^2$
$E(\text{steps in algorithm to reach corrce solution}) = E(Z) \leq 2n^2$
$P(\text{algorithm hasn't reached correct solution after } 8n^2 \text{ steps})$
$= P(Z > 8n^2) \stackrel{\text{Markov}}{\leq} \frac{E(Z)}{8n^2} \leq \frac{1}{4}$.

8.2 Generating a uniformly random element of a set

Ω: set.
Let G be a symmetric graph on Ω.
We form M.C:
$$P_{x,y} = \begin{cases} \frac{1}{M} \text{ if } x \neq y \wedge x \sim y \\ 0 \text{ if } x \neq y \wedge x \nsim y \\ 1 - \frac{|N(x)|}{M} \text{ if } x = y \end{cases}$$
$M \geq \max_{v \in \Omega} |N(v)|$.
If G is connected $\implies$ M.C. is irrecudible.
$\pi = (\frac{1}{|\Omega|} \dots \frac{1}{|\Omega|})$
$\pi \stackrel{?}{=} \pi P$

$$
\begin{aligned}
(\pi P)_x &= \sum_y \pi_y P_{y,x} \\
&= \sum_{y \in N(x)} \frac{1}{M} \cdot \frac{1}{|\Omega|} + \frac{1}{|\Omega|}\left(1 - \frac{|N(x)|}{M}\right) = \frac{1}{|\Omega|} = \pi_x.
\end{aligned}
$$

$\implies$ if we walk on the Markov chain long enough, we end up in state x with probability $\pi_x = \frac{1}{|\Omega|}$
$\implies$ we can sample uniformly.

Primer.
G graph, finding largest independent set $(\forall u,v : u \nsim v)$ is NP-complete.
Lets try sampling a uniformly random independent set
$\Omega = \{\text{independent sets}\}$
$u \sim v$ if $|u \triangle v| = 1$ $((u \cup \{el\}) = v)$
M.C.: X_0 = arbitrary independent set
X_{i+1}:

- pick uniformly at random $v \in V(G)$,
- if $v \in U$ then $X_{i+1} = U \backslash \{v\}$,
- if $U \cup \{v\}$ is independent then $X_{i+1} = U \cup \{v\}$,
- else $X_{i+1} = U$.

M is number of vertices
$\implies$ $\forall u \in \Omega : \lim_{t\to\infty} P(X_t = u) = \frac{1}{|\Omega|}$.
Note: irredudicle; $U \to \emptyset \to V$, aperiodic.

8.3 Metropolis algorithm

Ω: set,
π: chosen distribution on Ω.
Make G graph on Ω

$$P_{x,y} = \begin{cases} \frac{1}{M} \cdot \min\left(1, \frac{\pi_y}{\pi_x}\right) \text{ if } x \neq y \wedge x \sim y \\ 0 \text{ if } x \neq y \wedge x \nsim y \\ 1 - \sum_{y \in N(x)} \text{ if } x = y \end{cases}$$
$M \geq \max_{v \in \Omega} |N(v)|$
$\pi \stackrel{?}{=} \pi P$

$$\begin{aligned}(\pi P)_x &= \sum_y \pi_y P_{y,x} = \sum_{y \in N(x)} \pi_y \frac{1}{M} \min\left(\left(1, \frac{\pi_y}{\pi_x}\right)\right) + \pi_x \left(1 - \sum_{y \in N(x)} \frac{1}{M} \min\left(1, \frac{\pi_y}{\pi_x}\right)\right) \\ &= \sum_{y \in N(x), \pi_y \geq \pi_x} \pi_y \frac{1}{M} \cdot 1 + \sum_{y \in N(x), \pi_y < \pi_x} \pi_y \frac{1}{M} \frac{\pi_y}{\pi_x} + \pi_x \\ &- \sum_{y \in N(x), \pi_y \geq \pi_x} \pi_x \frac{1}{M} \frac{\pi_y}{\pi_x} - \sum_{y \in N(x), \pi_y < \pi_x} \frac{1}{M} \cdot 1 \\ &= \pi_x.\end{aligned}$$

Primer.
$\Omega = \mathbb{Z} \cap [-1000{,}1000]$
$\pi \sim e^{-\frac{(x-\mu)^2}{2\delta}}$

```
X_0 arbitrary
for i = in range(1,m):
  y <- uniformly from X_i+1,X_i-1
  M <- uniformly from [0,1]
  if M ≤ π(y)/π(x):
    X_{i+1} = y
  else:
    X_{i+1} = X_i
return X_m
```

Primer.
Find maximum of a positive function f.

Use metropolis algorithm to sample proportional to f.
Note: all I need to know is ratios $\frac{f(y)}{f(x)}$.

Back to independent sets.
$G = (V, E)$
$\Omega =$ independent sets.
$\lambda \in (1, \infty)$
$\pi(u) \sim \lambda^{|u|}$
$\pi(u) = \frac{\lambda^{|u|}}{\sum_{v \text{ independent set}} \lambda^{|v|}}$.
How to calculate the sum?
No problem: only need proportions.
X_0: arbitrary independent set.
$X_i \to X_{i+1}$:

- we pick $v \in V$ uniformly at random,
- if $v \in X_i \implies$
  - $X_{i+1} = X_i \setminus \{v\}$ qith probability $\frac{1}{\lambda} = \min\{1, \frac{\pi_y}{\pi_x}\}$,
  - $X_{i+1} = X_i$ with probability $1 - \frac{1}{\lambda}$,
- if $v \in X_j$ and $X_i \cup \{v\}$ is independent $\implies \quad X_{i+1} = X_i \cup \{v\}$,
- otherwise $X_{i+1} = X_i$.

Primer. Bayes: $P(A \mid B) = \frac{P(A \cap B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}$.
$B \leftarrow$ machine is giving values, e.g. $y_1 = 0.05, y_2 = -0.1, y_3 = 0.07, y_4 = 3$.
We believe $B \sim N(\mu, 0{,}05)$.
$\mu = laplacian(0, 0.01)$.
$P(\mu \mid B) = \frac{e^{\frac{|\mu|}{0.01}} e^{-\sum \frac{(x_i - \mu)^2}{0.05}}}{\int \cdots}$.
Integral is difficult to calculate.
Sample μ with Metropolis algorithm.

8.4 M.C. for 1-factor in bipartite graphs

G regular graph
$|A| = |B|$.
How to find 1-factor?
Augmenting paths.
Let M be (suboptimal) matching.
If we find $s - t$ path, we switch edges and get bigger matching.
Starting point.
G d-regular graph.
Graph $G = (A \cup B, E)$, M suboptimal matching.

- Add s and add directed edges to vertices in A that are not matched with weight d,
- add t and add directed edges to vertices in B that are not matched with weight d,
- orient edges in M from B to A that weight $d - 1$,
- orient edges in $E \setminus M$ from B to A that weight 1,
- we add edge from t to s that weight $(|A| - |M|)d$.

Observation:

- for each vertex $x :\ deg^-(x) = deg^+(x)$ (out weights = in weights),
- if $|A| > |M|$, then graph is eulerian $\implies$ there is an augmenting path.

How to find $s - t$ path?
Do a random walk.
Expected time to get from s to t is $h_{s,t}$
$\frac{1}{\pi(s)} = h_{s,s} = h_{s,t} + 1$.

Lema 8.4.1.
Let X be a M.C. defined as a random walk on directed (weighted) graph with $deg^-(x) = deg^+(x)$ for each x. Then the stationary distribution is

$\pi = \left[\frac{deg^+(x_i)}{|E|}\right]_{i=1}^n$.

w_{ij}: weight from i to j.

Dokaz 8.4.2.
$\pi P = \pi \left[\frac{w_{ij}}{deg^+(x_i)}\right]_{i,j=1}^n = \left[\frac{\sum_j w_j i}{|E|}\right]_{i=1}^n = \left[\frac{deg^-(x_i)}{|E|}\right]_{i=1}^n = \left[\frac{deg^+(x_i)}{|E|}\right]_{i=1}^n$.

$h_{s,s} = \frac{1}{\pi_s} \leq \frac{|E|}{deg^+(s)} \leq \frac{3(|A|-|M|)d+|M|(d-1)+(|A|-|M|)d+|M|(d-1)}{(|A|-|M|)d} \leq \frac{4|A|}{|A|-|M|}$.
Expected time to find augmenting path $\leq \frac{4|A|}{|A|-|M|}$.
$|A| = n$
Expected time to find 1-factor $\leq \frac{4n}{n-1} = 4n \sum_{i=1}^{n-1} \frac{1}{i} \leq 4n(1 + \ln n)$ - in $O(n \log n)$.

8.4.1 Network centrality

Degree as measure - natural idea.
Use M.C: walk randomly on the network, those that are visited more oftenly are more important.
Pagerank.
Let A be the adjacency matrix of G.
$P_{ij} = \alpha \frac{A_{ij}}{degi} + (1-\alpha)\frac{1}{n}$;
α: normal random walk,
$1-\alpha$: jump to any.
$\alpha = 0.85$.

Poglavje 9

Randomized incremented constructions (RIC)

Observation:
Let S be a set of n distinct elements.
Let $X_1 \dots X_n$ be a random permutation of the elements.
Let $S_i = \{X_1 \dots X_i\}$.
$P(X_i = \min(S_i)) = \frac{1}{i}$.
$Y = |\{j \in \{1 \dots n\} \mid j = \text{ minimal of } S_j\}|$
$Y = Y_1 + \dots + Y_n$
$Y_j = \begin{cases} 1 \text{ if } i = \min S_i \\ 0 \text{ otherwise} \end{cases}$
$E(Y) = \sum_{i=1}^{n} E(Y_i) = \sum_{i=1}^{n} \frac{1}{i}$ in $O(\log n)$.

```
Alg():
  X1 .. Xn = random permutation of S
  min = X1
  for i in range(1,n+1):
    if Xi < min:
      print("HA")
      min = Xi
```

We get $O(\log n)$ „HA“ printed.
Incremental construction (IC).
Input $S = \{s_1 \dots s_n\}$.
We will build structures $DS(S_i)$:
$DS(S_1 \to \dots \to DS(S_n))$.
$DS(S_n)$ will help us give answer.
Randomized: permute S at the beginning.

9.1 Quicksort as RIC

S: set of elements we want to order.
$X_1 \dots X_n$: random permutation of S.
$S_i = \{X_1 \dots X_i\}$.
S_i splits $\mathbb{R}$.
Define $DS(S_i)$:

- save intervals: each interval will be saved by endpoints,
- for each interval we will be saving its points,
- for each X_j, $j > i$ we will save in which interval it is,
- for each left point of the interval we will save the right point.

```
QuicksortRIC(S):
  # start of DS(Si)
  I=[(-∞,∞)]
  P[(-∞,∞)] = S
  for each Xi:
    Int(Xi) = (-∞,∞)
  Next(∞) = ∞
  # end of DS(Si)
  for i in range(1,n+1):
    Ii = Int(Xi) = (Xj,Xk)
```

```
    Ii1 = (Xj,Xi)
    Ii2 = (Xi,Xk)
    for Xl ≠ Xi, Xi ∈ P(I):
      add Xl to P(Ii1) or P(Ii2) depending on Xl < Xi or Xl > Xi
    Next(Xj) = Xi
    Next(Xi) = Xk
  return [Next(−∞), Next(Next(−∞)) ..]
```

Similarity to quicksort: spliting intervals.

Analysis:
for set i, we need $O\left(|P(I_i)|\right)$,
$E\left(|P(I_i)|\right) = ?$
e.g.
if $x_4 = a_4$:
if $x_4 = a_2$:
$P(X_i = a_j) = \frac{1}{i}\ j \in \{1, 2 \ldots i\}$.
Expected value of steps in iteration i
$\sum_{j=1}^{i} \frac{1}{i}\left(P\left((a_{j-1}, a_j)\right) + P\left((a_j, a_{j+1})\right)\right) \leq \frac{1}{i} 2(n-i) \leq \frac{2n}{i}$

$$
\begin{aligned}
E\,(\text{number of steps in QuicksortRIC}) &\leq \sum_{i=1}^{n} \frac{2n}{i} \\
&\leq 2n(1 + \log n) \quad \rightarrow \text{ in } O(n \log n).
\end{aligned}
$$

9.2 Linear programming

Task: maximize $f(x_1 \ldots x_n) = c_1 x_1 + \cdots + c_d x_d$.
Constraints:

$$
\begin{aligned}
& a_{11} x_1 + \cdots + a_{1d} x_d \leq b_1 \\
& \vdots \\
& a_{n1} x_1 + \cdots + a_{nd} x_d \leq b_n.
\end{aligned}
$$

Geometric interpretation.
Cases:

- infeasible region
- unbounded
- .

Alg:

- symplex algorithm worst case $O(2^n)$,
- interior point method (polynomial algorithm).

Seidel's algorithm:
running in expected $O(n)$ time when d is constant.
One dimension.

$$\begin{aligned}\max\ & cx\\ & a_1x \leq b_1\\ & \vdots\\ & a_nx \leq b_n,\end{aligned}$$

where n is number of constraints.

- a_i positive: $(-\infty, \frac{b_i}{a_i}]$,
- a_i negative: $[\frac{b_i}{a_i}, \infty)$.

$a_i \neq 0$.
Alg:

$R = \min_i\{\frac{b_i}{a_i}; a_i > 0\}$,

$L = \max_i\{\frac{b_i}{a_i}; a_i < 0\}$,

if $L > R$: program infeasible,

else:

if $c > 0$: return R,

if $c < 0$: return L.

2-dim: assume general position.

$$\begin{aligned} \max c_1x + c_2y & \\ & a_{11}x + a_{12}y \leq b_1 \\ & \vdots \\ & a_{n1}x + a_{n2}y \leq b_n \\ & x \leq M \text{ or } x \geq -M \\ & y \leq M \text{ or } y \geq -M. \end{aligned}$$

$\leq, \geq$ depending on c_1, c_2.
Notation:

h_i: halfspace defined by $a_{i1}x + a_{i2}y \leq b_i$,

m_i: added halfspaces,

l_i: line that bounds.

Alg:

- first randomly permute h_i,
- $H_i = \{m_1, m_2, h_1 \ldots h_i\}$,
- $v_i \in \cap H_i$ optimal solution after i constraints,
- $v_0 = (\pm M, \pm M)$,
- inductively add h_i.

Cases:

if $v_{i-1} \in h_i \implies v_i = v_{i-1}$,

if $v_{i-1} \notin h_i \implies v_i \in h_i$:

$a_{i1}x + a_{i2}y = b_i$

a_{i1} or $a_{i2} \neq 0$, e.g. a_{i1};

$x = \frac{b_i - a_{i2}y}{a_{i1}}$.

Insert x in all constraints $\Longrightarrow$ linear program in 1-dim, i (i-1?) constraints $\Longrightarrow$ get v_i in $O(i)$.

Analysis:

- worst case: $\sum_{i=1}^{n} O(i) = O(n^2)$,
- expected: $E(X) = \sum_{i=1}^{n} E(X_i)$,
- X_i = running time of i-th iteration,
- $X_i = \begin{cases} O(1); \text{ case } 1 \\ O(i); \text{ case } 2 \end{cases}$
- $P(\text{case } 2) \leq \frac{2}{i}$ - optimal point on at most 2 lines,
- $E(X) \leq \sum_{i=1}^{n} O(1) \cdot 1 + O(i) \cdot \frac{2}{i} = O(n)$.

d-dim

- constraints define half-spaces,
- boundary is hyperplane ($d-1$ dimensional),
- general position: intersection of $d-i$ hyperplanes is i dimensional, intersection of $d+1$ hyperplanes is $\emptyset$.

Alg:

first add $X_i \leq M$ or $X_i \geq -M$ depemding on c_i,

random permutation $(h_1 \ldots h_n)$,

$H_i = \{m_1 \ldots m_d, h_1 \ldots h_i\}$,

$v_0 \in \cap \partial m_i$,

inductively add h_i:

$v_{i-1} \in h_i \implies v_i = v_{i-1}$,

$v_{i-1} \notin h_i \implies$ we need to solve LP in $d-1$ dimensions with i constraints ($O(i)$ expected),

$P(v_{i-1} \notin h_i) \leq \frac{d}{i}$,

$E(X) \leq \sum_{i=1}^{n} O(1) + \frac{d}{i} O(i) = O(n)$.

X: running time.
Careful implementation runs in $O(d!\, n) \implies$ very useful for low dimensions.
Problem: let P be convex polygon given by ordered set of vertices
$y = a_i x + b_i$.
Find largest disc embeddable in P.
Input: $P_1 \dots P_n$,
output: $(s_1, s_2), r$.

$$\max r$$

$$d = \left| \frac{s_2 - a_i s_1 - b_i}{\sqrt{a_i^2 + 1}} \right|$$

$$\frac{s_2 - a_i s_1 - b_i}{\sqrt{a_i^2 + 1}} \geq r \text{ - line above } P$$

$$-\frac{s_2 - a_i s_1 - b_i}{\sqrt{a_i^2 + 1}} \leq -r \text{ - line below } P$$

$\implies$ LP in 3 dim.
Note: $\frac{s_2 - a_i s_1 - b_i}{\sqrt{a_i^2+1}}$ positive if (s_1, s_2) above the line, negative otherwise.

Poglavje 10

Hashing

A hash function is a randon function,
$h : U \rightarrow \{0, 1 \dots n-1\} = M$,
U - universe,
$u = |U|$,
$m = |M|$.
Ideally we would like for h to be as completely rondom: $P(h(x) = t) = \frac{1}{m}$.
Standard application.
Let $V \subset U,\ |V| << |U|$.
We would like to quickly answer if $x \in V$ for ever $x \in V$.
Solution:

- take $h : U \rightarrow M$,
- make a table $T = [0, 1 \dots n-1]$,
- for $v \in V$:

  $T[h(v)] = 1$,

  $T[y] = 0\ \forall y \in h(V)$.

- Let $x \in V$. Check

– if $T[h(x)] = 1: \ x \in V$,

– else: $x \notin V$.

Note: this is not OK: h not injective.

For $x \in U$, tell if $x \in V$ in $O(1)$.
$h = \text{SHA256} : U \to \{0, 1\}^{256}$.
Approach:

- design a family of hash functions,
- study collisions $P_h(h(x) = h(y))$,
- H meeds to be „simple“.

Bad example: H = all functions from U to M storing $h \in H$ would take $|U| \log_2 |M|$ bits.

Definicija 10.0.1. A family of hash functions to be universal if for
$\forall x, y \in U, x \neq y, h \in H: \ P(h(x) = h(y)) \leq \frac{1}{m}$ (probability of collision).

k-independent if $\forall x_1 \ldots x_k \in U$ pairwise different, $\forall t_1 \ldots t_k \in M \ P_r(h(x_i) = t_i \ \forall i) \leq \frac{1}{m^k}$.

Primer.
$U = \{0, 1, 2, 3\}$,
$M = \{0, 1\}$,
$H = \{h_0, h_1, h_2\}$,
$h_0 : \{0 \to 0, 1 \to 0, 2 \to 1, 3 \to 1\}$,
$h_1 : \{0 \to 0, 1 \to 1, 2 \to 0, 3 \to 1\}$,
$h_2 : \{0 \to 0, 1 \to 1, 2 \to 1, 3 \to 0\}$.
$P(h(0) = h(2)) = \frac{2}{3} < \frac{1}{2}$ - not universal.

Why universal?
U, V, M
H universal: $\forall x, y: \ P(h(x) = h(y)) \leq \frac{1}{m}$.
X: number of collisions of V.

$E(X) = E\left(\sum_{x,y\in V, x\neq y} X_{x,y}\right)$

$X_{x,y} = \begin{cases} 1 \text{ if } h(x) = h(y) \\ 0 \text{ else} \end{cases}$

$E(X) = \sum_{x,y\in V, x\neq y} E(X_{x,y}) \leq \binom{n}{2} \cdot \frac{1}{n}$.

U, V, M, H

$T[0 \ldots m-1]$

$\forall v \in V$

$T[h(v)] = v$.

For $x \in V$ we check $T[h(x)]$ if equals x,

for $y \in U \setminus V$, $T[h(y)] \neq y$.

For $z \in V$, $T[h(z)]$ can happen $\neq z$ if h has collisions in V.

Lema 10.0.2. Let $m \geq n^2$ and H universal. Then the probability that h has no collisions in $V \ \geq \frac{1}{2}$.

Dokaz 10.0.3.

X: number of collisions

$E(X) \leq \binom{n}{2} \cdot \frac{1}{m} < \frac{n^2}{2} \cdot \frac{1}{n^2} = \frac{1}{2}$

$P(X \geq 1) \overset{\text{Markov}}{\leq} \frac{E(X)}{1} = \frac{1}{2}$

$P(X = 0) \geq \frac{1}{2}$.

Primer (Universal hash family).

$U = \{0, 1 \ldots u-1\}$ (bits $\equiv$ numbers)

$M = \{0, 1 \ldots m-1\}$.

Define: let $p \geq u$, p prime number.

Define for $a, b \in \mathbb{Z}_p,\ a \neq 0$.

$h_{a,b} = (ax+b) \mod m$

$ax + b \in \mathbb{Z}_p$

$H = \{h_{a,b} \mid a, b \in \mathbb{Z}_p,\ a \neq 0\}$.

Dokaz 10.0.4. $P(h_{a,b}(x) = h_{a,b}(y)) =?$

x, y fixed.

For any a, b denote

$ax + b = t_x$

$ay + b = t_y$:

$a \sqcup +b \in \mathbb{Z}_p$.

$\begin{bmatrix} x & 1 \\ y & 1 \end{bmatrix} \begin{bmatrix} a \\ b \end{bmatrix} = \begin{bmatrix} t_x \\ t_y \end{bmatrix}$

$\det \begin{bmatrix} x & 1 \\ y & 1 \end{bmatrix} \neq 0$, because $x \neq y$

$\begin{bmatrix} a \\ b \end{bmatrix} = \begin{bmatrix} x & 1 \\ y & 1 \end{bmatrix}^{-1} \begin{bmatrix} t_x \\ t_y \end{bmatrix}$.

For each t_x, t_y there exists 1 a, b mapping to t_x, t_y.

$h_{a,b}(x) = h_{a,b}(y) \iff t_x = t_y \mod m$.

This holds for $p\left(\lceil \frac{p}{m} \rceil + 1\right)$

p: choise of t_y

$t_x = t_y + km$

$P\left(h_{a,b}(x) = h_{a,b}(y)\right) \leq \frac{p\left(\lceil \frac{p}{m} \rceil - 1\right)}{p(p-1)} \leq \frac{\frac{p-1}{m}}{p-1} = \frac{1}{m}$.

Function random for 2 elements, fixed for ≥ 3.

Higher k-independent: better.

10.1 Chaining

$V, U, h : U \to V$.

Answer $x \in V$ in $O(1)$.

$T[0 \ldots m-1]$

$n = |V|$

$\forall v \in V$:

$h(v_1) = h(v_2) \ \to \ [v_1 \ v_2 \ldots]$ - linked list.

Now:

$x \in U$.

Check if x is in list at $T[h(x)]$.

Check takes $O(\text{length of a list at } h(x)) = 1+$ number of collisions with x.

X_x: number of collisions with x.

$E(X_x) = \sum_{y \in V} E(X_{x,y}) \leq n \cdot \frac{1}{m}$ if hash function is universal.
$\alpha = \frac{n}{m}$: load factory (how many elements in 1 place).
$E(X_x) = 1$
$E(\max_x X_x) \neq \max_x E(X_x) = 1$.

Izrek 10.1.1. Assume we throw n balls into n bins uniformly at random. Then with high probability the fullest contains $\theta\left(\frac{\log n}{\log(\log n)}\right)$ balls.

Dokaz 10.1.2.
$\overset{?}{\leq} \frac{3 \ln n}{\ln \ln n}$.
Let X_j be the number of balls in bin j.
$P\left(X_j \geq \frac{3 \ln n}{\ln \ln n}\right) = P($ there exists subset S of balls thrown to bin$j)$.
$|S| = k$

$$\begin{aligned}
&P\left(\cup_{S \text{ balls}, |S|=k} \text{balls from } S \text{ are thrown to bin } j\right) \\
&\leq \sum_{S \text{ balls}, |S|=k} P(\text{balls from } S \text{ are thrown to } j) \\
&= \binom{n}{k} \left(\frac{1}{n}\right)^k \\
&\leq \frac{n^k}{k!} \cdot \frac{1}{n^k} = \frac{1}{k!} = (*).
\end{aligned}$$

Note: $e^x = \sum_{i=1}^{\infty} \frac{k^i}{i!} \geq \frac{k^k}{k!}$.

$$\begin{aligned}
(*) &\leq \frac{e^k}{k^k} \\
&= \left(\frac{e \ln n}{3 \ln \ln n}\right)^{\frac{3 \ln n}{\ln \ln n}} \\
&\leq e^{\frac{3 \ln n}{\ln \ln n} \cdot (\ln \ln \ln n - \ln \ln n)} \\
&= e^{-3 \ln n + \frac{\ln \ln \ln n \cdot (\ln n \cdot 3)}{\ln \ln n}} = (**)
\end{aligned}$$

$\frac{\ln \ln \ln n}{\ln \ln n} \to 0$

$$(**) \leq e^{-3 \ln n + \ln n} = \frac{1}{n^2}.$$

$P(\text{at least for 1 bin } j \geq k) = n \cdot \frac{1}{n^2} = \frac{1}{n}$.

U, V, H hash family, $h : U \to M$
$v \in V$
$n = |V|$
max load $O\left(\frac{\log n}{\log(\log n)}\right)$.
Perfect hashing: we would like

- $O(1)$ lookup (worst case)
- $O(n)$ size of table.

10.2 2 level hashing

Input: V
$n = |V|$.
Take hash function from universal family with $m = |M| = n$.
Count total collisions X.
$E(X) \leq \binom{n}{2} \cdot \frac{1}{m} \leq \frac{n}{2}$
$P(x \geq n) \overset{\text{Markov}}{\leq} \frac{1}{2}$
$\implies$ by repeating sample h we can guarantee

- for each $i \in M$ we store at $T[i]$ another hash table of size C_i^2, where C_i = number of elements of V, hashed in i,
- we sample h_i from universal hash family with $M_i = C_i^2$.

$P(h_i \text{ has no collisions}) \geq \frac{1}{2}$ (by lemma).
We resample if h_i has collisions.
$E(\text{sampling } h_i) = 2$.
Construction time:

- step 1: $O(n)$
- step 2: $O(C_1 + \cdots + C_n) = O(n)$;

together $O(n)$.

Lookup time: $O(1)$ (evaluating $h(x)$ and $h_{h(x)}(x)$).
Space: $O(C_1^2 + \cdots + C_n^2)$ in $O(n)$.
By first step $n >$ number of collisions of $h = \sum_{i=1}^n \binom{C_i}{2} = \sum_{i=1}^n \frac{C_i^2 - C_i}{2}$
$\implies \sum_{i=1}^n C_i^2 < 2n + \sum_{i=1}^n c_i = 3n$.

10.3 The power of 2 choices

Variant: placing n balls in n bins but for each ball we choose d balls uniformly at random and put the ball in bin with minimal load.

Izrek 10.3.1. The above process with $d \geq 2$ results in at most maximum load of $O\left(\frac{\ln(\ln n)}{\ln d}\right)$.

Dokaz 10.3.2. (sketch).
b_i = upper bound of the number of bins with load at most i.
Height of a ball = the number of balls in the bin, where the ball is placed.
P(a ball has height at least $i+1$) $\leq \left(\frac{b_i}{n}\right)^d$ (choose d times independently).
X^{i+1}: number of balls with height $\geq i+1$.
$X^{i+1} = \sum_{j=1}^n X_j^{i+1}$
X_j^{i+1}: indicator variable of j-th ball having height $i+1$.
$E(X^{i+1}) \leq \sum_{j=1}^n \left(\frac{b_i}{n}\right)^d = d \cdot \left(\frac{b_i}{n}\right)^d$.
Chernoff bound: with high probability $X^{i+1} \leq 2n \left(\frac{b_i}{n}\right)^d$.
$X^{i+1} \geq$ number of bins with load at least $i+1$.
Define (set)
$b_{i+1} = \frac{\sum b_i^d}{n^{d-1}}$
$b_4 = \frac{n}{4}$
$b_{i+4} = \frac{n}{2^{2 \cdot d^i - \sum_{j=0}^{i-1} d^j}}$

$$i = 0\colon\ b_4 = \frac{n}{2^{2^1}} = \frac{n}{4}$$

$i \to i+1$:

$$
\begin{aligned}
b_{i+4} &= \frac{2 \cdot b_{i+3}}{n^{d-1}} \\
&\overset{IH}{=} \frac{2 \cdot \left(\frac{n}{2^{2 \cdot d^i - \sum_{j=0}^{i-1} d^j}} \right)^d}{n^{d-1}} \\
&= \frac{2^1 \cdot n^d}{n^{d-1} \cdot 2^{2 \cdot d^{i+1} - \sum_{j=1}^{i} d^j}} \\
&= \frac{n}{2^{2 \cdot d^{i+1} - \sum_{j=0}^{i} d^j}}.
\end{aligned}
$$

In particular: $b_{i+4} \leq \frac{n}{2^{d^i}} < 1$ when?

$$
\begin{aligned}
n &< 2^{d^i} \\
\log_2 n &< d^i \\
\log_d \log_2 n &< i
\end{aligned}
$$

$\implies$ for $i = \frac{\log(\log_2 n)}{\log d}$ is $b_i < 1 \implies$ no bins with load $> \frac{\log(\log_2 n)}{\log d}$.

Application:
We sample 2 hash functions $h_1, h_2 : U \to M$.
For element $v \in V$ we insert in $T[h_1(v)]$ or $T[h_2(v)]$ depending on which list is shorter.
Max load in $O(\log(\log n))$.

10.4 Cockoo hashing

Idea: use 2 hash functions but allow moving elements later.
We want to have at most 1 element at each entry in the table.
Inserting:

- if empty: insert,
- if not empty: push other element to its other choise, repeat recursively.

Questions:

- how many do I need to move,
- how many elements can I insert before problems?

We can think of positions in the table as vertices and elements of V as edges. $|V|$ edges are inserted uniformly at random (if ideal hash function) $\implies$ random graph.
Erdös-reny model: $G_{n,m} \approx G_{n,p}$ if $m = \binom{n}{2}p$ (A.S. properties).
If $np < 1 - \epsilon$: all connected components have size at most $O(\log n)$, components are trees or at most 1 cycle per component, expected size of a component is $O(1)$.
Fact: if graph has at msot 1 cycle per component, then inserting can be done and takes at most 2·(size of component) time (each edge changes direction at most 2 times).

Izrek 10.4.1.
Let $n = |U|$, $h_1, h_2 : U \to M$, $m = |M| = 2 \cdot (1 + \epsilon) \cdot n$, then with high probability cockoo hashing works correctly with

- inserting time:
    - $O(\log n)$ time worst case,
    - $O(1)$ expected case,
- space: $O(n)$,
- lookup time: $O(1)$.

Dynamically add element:
$m = 2 \cdot (1 + \epsilon) \cdot n$
$p = \frac{m'}{\binom{n'}{2}} = \frac{2m'}{n'(n'-1)}$
$pn' = \frac{2m'}{(n'-1)} = \frac{2n'}{2(1+\epsilon)n'} = \frac{1}{1+\epsilon} < 1 + \epsilon'$