

1. Show properties one by one:

- $Y_0 = 0$ by definition ✓

- Independence: We use the fact that Multivariate-Gaussian $\Leftrightarrow \text{Cov} = 0 \Rightarrow \text{Independent}$
 for $r < s < t$, we have $(Y_t - Y_s, Y_r)$ is ~~not~~ Multivariate Gaussian because that is $(t \cdot B_{\frac{1}{t}} - s \cdot B_{\frac{1}{s}}, r \cdot B_{\frac{1}{r}})$ & by def

$$t \cdot B_{\frac{1}{t}} - s \cdot B_{\frac{1}{s}} = (t-s) B_{\frac{1}{t}} + s (B_{\frac{1}{t}} - B_{\frac{1}{s}}) = (t-s) B_{\frac{1}{t}} + s (B_{\frac{1}{s}} - B_{\frac{1}{t}})$$

is sum of 2d Gaussian \Rightarrow still Gaussian. Now we only have
 To show $\text{Cov} = 0$. to get $Y_t - Y_s \perp \sigma\{Y_r\}$

~~But this~~

↑ independent

Here we use $\text{Cov}(B_t, B_s) = \min\{t, s\}$ (alternative definition) discussed in class

to get $\text{cov}(Y_t - Y_s, Y_r) = \text{cov}\left(t B_{\frac{1}{t}} - s B_{\frac{1}{s}}, r B_{\frac{1}{r}}\right)$

$$= t \cancel{s} \text{cov}\left(B_{\frac{1}{t}}, B_{\frac{1}{r}}\right) - s r \underbrace{\text{cov}\left(B_{\frac{1}{s}}, B_{\frac{1}{r}}\right)}_{\text{cov}}$$

$$= t r \cdot \frac{1}{t} - s r \cdot \frac{1}{s} = r - r = 0$$

Hence $Y_t - Y_s \perp \sigma\{Y_r\}$. But this argument works for all $r \leq s$
 $\Rightarrow Y_t - Y_s \perp \sigma\{Y_r : r < s\}$

Hence Independence ✓

- Normal increments:

$$\forall t \geq s, \quad Y_t - Y_s = (t-s) B_{\frac{1}{t}} - s (B_{\frac{1}{s}} - B_{\frac{1}{t}}) =: X_1 + X_2$$

& $X_1 \perp X_2$ & $\{X_i \sim N(0, \frac{(t-s)^2}{t})\}$

by independence $| X_2 \sim N(0, \frac{(t-s) \cdot s}{t})$

$$\Rightarrow X_1 + X_2 \sim N\left(0, \frac{(t-s)^2}{t} + \frac{(t-s) \cdot s}{t}\right) = N(0, t-s) \quad \checkmark$$

Note: sign is ignored
 by ~~symmetric~~ distribution

• Continuity:

- for $t > 0$, w.p.1 it's cts since B_t^1 is cts on $(0, \infty)$.

- for $t=0$, $Y_t=0$ and we want to check w.p.1, $\lim_{t \rightarrow 0} Y_t = 0$.

But $\lim_{t \rightarrow 0} Y_t = \lim_{t \rightarrow 0} t \cdot B_t^1 = \lim_{t \rightarrow 0} \frac{1}{t} \cdot B_t = 0$

from last homework.

Thus, Y_t is cts at 0 with p.1.

Now combine the 2 prob 1 set \Rightarrow w.p.1. Y_t has cts path.

$$(\frac{1}{2}, \frac{1}{2})_{10} - (\frac{1}{2}, \frac{1}{2})_{10} \cdot 10 =$$

$$0 = 1 - \frac{1}{2} \cdot 10 - \frac{1}{2} \cdot 10 =$$

$$X_{10} = (\frac{1}{2}, \frac{1}{2})_0 - (\frac{1}{2}, \frac{1}{2})_{10} = X_0 - X_{10}$$

$$(\frac{1}{2}, \frac{1}{2})_{10} \cup X_0 \cap X_{10} = X_0 \cap X_{10}$$

$$(\frac{1}{2}, \frac{1}{2})_0 \cup X_0 \cap X_{10} = X_0 \cap X_{10}$$

$$(e^{10}, 0) = (\frac{1}{2}, \frac{1}{2})_0 + (\frac{1}{2}, \frac{1}{2})_{10} = X_0 + X_{10}$$

Tommenix Yu:

Q2: Just check: on $(\Omega, \bar{\mathcal{F}}_T, P)$.

• $\Omega \in \bar{\mathcal{F}}_T$ ✓

• complement:

if $A \cap \{T \leq t\} \in \bar{\mathcal{F}}_t$, then $A^c \cap \{T \leq t\} = \{T \leq t\} \setminus (A \cap \{T \leq t\}) \in \bar{\mathcal{F}}_t$

since $\bar{\mathcal{F}}_t$ is σ -alg. $\Rightarrow A^c \in \bar{\mathcal{F}}_T$. ✓

• countable union:

If $A_i, i \in \omega$ is a ctb set of events ~~not in~~ in $\bar{\mathcal{F}}_T$, then

$$(\bigcup_{i \in \omega} A_i) \cap \{T \leq t\} = \bigcup_{i \in \omega} (A_i \cap \{T \leq t\}) \in \bar{\mathcal{F}}_t$$

$$\Rightarrow \bigcup_{i \in \omega} A_i \in \bar{\mathcal{F}}_T. \quad \checkmark$$

So $\bar{\mathcal{F}}_T$ is indeed a σ -alg.

Q3: for $r = \sqrt{2(1+\varepsilon)}$, we have by reflection principle that, if we define

$$U_n := \left\{ \exists t \in [n-1, n] \text{ with } |B_t - B_{n-1}| \geq \sqrt{2(1+\varepsilon) \log n} \right\} = \left\{ M_n \geq \sqrt{2(1+\varepsilon) \log n} \right\}$$

then ~~$P(U_n)$~~ since $B_{n-1} - B_{n-2} \sim N(0, S)$, we can pass this by taking prob.

i.e.

$$\begin{aligned} P(U_n) &= P\left\{ \exists t \in [0, 1] \text{ with } |B_t| \geq \sqrt{2(1+\varepsilon) \log n} \right\} \leq 4 P\left\{ B_1 \geq \sqrt{2(1+\varepsilon) \log n} \right\} \\ &\leq e^{-(1+\varepsilon) \log n} \cdot \left(\frac{4}{\sqrt{2(1+\varepsilon) \log n}} \right) = c \cdot \frac{1}{n^{(1+\varepsilon)/2} \log n} \text{ is summable for all } \varepsilon > 0. \end{aligned}$$

Now, since ~~For~~ For n sufficiently large, $M_n \leq \sqrt{2(1+\varepsilon) \log n}$ for all ε

$\Rightarrow M_n \leq \sqrt{2 \log n}$ as it's closed.

$\Rightarrow r \geq \sqrt{2}$ is all fine.

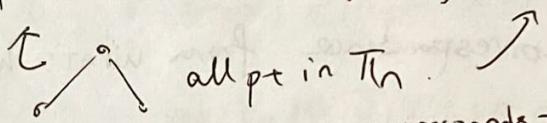
1) This is direct from definition: if $\Delta(z_{j-1, n+1}) + \Delta(z_j, n+1) \geq \Delta(z_{jn})$

i.e.



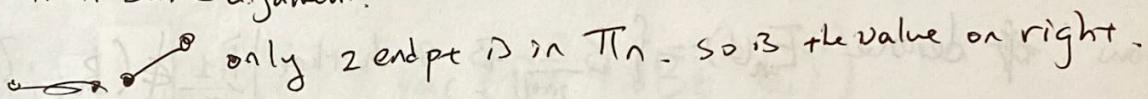
2nd variation is larger, we include the middle pt. in

the partition T_n . So $\sum_{j=1}^{2^n} [B(t_{j,n}) - B(t_{j-1,n})]^2 = \sum_{j=1}^{2^n} Y(j,n)$



corresponds to term here

if not, then same argument.



②. I realized that for a_1, a_2 if $a = |a_1 + a_2|$, then always

$\Delta(j,n)$ is larger. \Rightarrow since $(a+b)^2 \geq a^2 + b^2$ for $a, b \geq 0$.

so if $B_{\frac{2j+1}{2^{n+1}}} \notin (B_{\frac{j-1}{2^n}}, B_{\frac{j}{2^n}})$, then there's no choice of pt in middle.

+ the \Leftarrow direction of the statement
is obvious.

WLOG assume it is ~~well~~ order
an interval,
if not switch

in other words, we're choosing the "finest" choice of path with maximal variation, given the specification of information (only know information about $t = \frac{k}{2^{n+1}}$)
actually used in later parts!

2) First, $\{B_t : t \in [\frac{i}{2^n}, \frac{i+1}{2^n}]\}_{i=0,1..2^n}$ is a collection of independent σ -alg.

But really, given observation above (and it's shown by $(a+b)^2 \geq a^2 + b^2$)

$Y(k,n)$ only depends on the path of B_t for $t \in [\frac{k-1}{2^n}, \frac{k}{2^n}]$, where pass to $t \in [\frac{k-1}{2^n}, \frac{k}{2^n}]$ by continuity.

So $Y_{(k,n)}$ is msrb w.r.t $\sigma\{B_t : t \in [\frac{k-1}{2^n}, \frac{k}{2^n}]\}$

So $Y_{(k,n)}$ are independent of each other.

They are iid because $B_{\frac{1}{2^n}} - B_0 \sim B_{\frac{1}{2^n}} \sim \dots \sim B_{\frac{2^n}{2^n}} - B_{\frac{1}{2^n}}$.

& $B_{\frac{1}{2^n}} - B_0 \sim 2^{-n/2} B_1$.

So there's a 1-1 correspondence from what's happening in $\underbrace{}_0 \frac{1}{2^n}$ and $\underbrace{}_0 1$
by $f_n : x \mapsto 2^{-n/2} x$ where $f_n(B_t) = B_{(\frac{t}{2^n})}$, f_n acts on appropriate function space

Now ~~the~~ denote $E_n(1) := \left\{ B_{\frac{1}{2^{n+1}}} \text{ is such that } \left| B_{\frac{1}{2^{n+1}}} \right| \geq \left| B_{\frac{1}{2^n}} \right| \right\}$

Then $\exists \tilde{E}(1) := \left\{ \left| B_{\frac{1}{2}} \right| \geq \left| B_1 \right| \right\}$ from which $\tilde{f}_n(\tilde{E}(1)) = E_n(1)$
is 1-1 onto.

Now, if $E_n(1) \ni \omega$, $\Rightarrow Y_{(1,n)} \Delta(z_j^{-1}, n+1) + \Delta(z_j, n+1) = B_{\frac{1}{2^{n+1}}}^2 + (B_{\frac{1}{2}} - B_{\frac{1}{2^{n+1}}})^2$

& $\tilde{f}_n(\omega)$ is s.t. $\tilde{E}(1)$ holds true

$$\Rightarrow Y = B_{\frac{1}{2}}^2 + (B_1^2 - B_{\frac{1}{2}}^2) = 2^{-n} \left(B_{\frac{1}{2^{n+1}}}^2 + \left(B_{\frac{1}{2}} - B_{\frac{1}{2^{n+1}}} \right)^2 \right)$$

if $\omega \notin E_n(1)$, $\Rightarrow Y_{(1,n)} = B_{\frac{1}{2^{n+1}}}^2$

& for $\tilde{f}_n^{-1}(\omega)$, $Y = 2^{-n} B_{(\frac{1}{2^{n+1}})}$

So for each $\omega \in \mathcal{R}$, $\exists \tilde{f}_n^{-1}(\omega) \in \mathcal{R}$ s.t. $Y = 2^{-n} Y_{(1,n)}$.

Thus, $Y_{(1,n)} \stackrel{d}{\sim} Y \cdot 2^{-n}$

after discussion with classmates:

a much easier way is just to find a copy of $\stackrel{iid}{\sim} B_t$, then

show both are $\stackrel{d}{\sim}$ to new B_t 's Y . ~~But Quantile method above is~~

~~correct so~~

3) After fact we use $\text{Var}(\max(X, Y)) \leq \text{Var}(X) + \text{Var}(Y)$
 for $X, Y \geq 0$.

This is proven by indicator function + MCT

indicator is obvious : $\underbrace{(a+b+c)}_A \quad \text{i.e. } P(A) = a+b+c$

$$\text{then } \text{Var}(\max\{\mathbb{1}_A, \mathbb{1}_B\}) = (a+b+c)^2 \leq (a+b)^2 + (b+c)^2 = \text{Var}(\mathbb{1}_A) + \text{Var}(\mathbb{1}_B).$$

Now MCT gives the result as $\text{Var}(Y) = \mathbb{E}[Y^2] - \mathbb{E}[Y]^2$ and limit passes into both.

$$\begin{aligned} \text{So } \text{Var}[Y] &\leq \text{Var}(B_1^2) + \text{Var}(B_{\frac{1}{2}}^2 + (B_{\frac{1}{2}} - B_1)^2) \\ &= \mathbb{E}[B_1^4] + \text{Var}[B_{\frac{1}{2}}^2] + \text{Var}[(B_{\frac{1}{2}} - B_1)^2] \\ &= (3)!! + \left(\frac{1}{2}\right)^4 [3!!] \times 2 = 3 + \frac{1}{8}3 = \frac{C}{8} < \infty \end{aligned}$$

for $\mathbb{E}[Y]$, we note $\mathbb{E}[B_1^2] = 1$.

and by def of Y , $\mathbb{E}[Y] \geq \mathbb{E}[B_1^2]$

and we only need to check $P(B_{\frac{1}{2}} > B_1) > 0$

$P(A) > 0$ where $A := \{|B_{\frac{1}{2}}| > |B_1| \text{ and } \text{sign}(B_{\frac{1}{2}}) = \text{sign}(B_1)\}$

But that's obviously true since

$$P(A) = \overbrace{\mathbb{P}(\dots \wedge B_1 \wedge B_{\frac{1}{2}}, \text{sign}(B_1) = \text{sign}(B_{\frac{1}{2}}))}^{\text{sign issue}} > 0$$

$$= \frac{1}{2} \sum_{x_1} \mathbb{P}(B_{\frac{1}{2}} = x_1) \cdot \mathbb{1}_{\{|B_1| < x_1\}} \stackrel{P > 0}{\rightarrow} P > 0.$$

sign
issue

$$\Rightarrow \mathbb{E}[Y] = P(A) \cdot (1+C) + P(A^c) \cdot 1 > 1$$

4) in 2) we've shown $Y_{(i,n)}$ iid & $Y_{(i,n)} \sim 2^{-n} Y$

By strong LLN in triangular form:

$$\begin{matrix} Y_{(1,0)} \\ Y_{(1,1)} & Y_{(1,2)} \\ Y_{(1,2)} & Y_{(2,2)} & Y_{(3,2)} & Y_{(4,2)} \\ \vdots & & & & \end{matrix}$$

we have

$$\lim_{n \rightarrow \infty} \frac{\sum_{i=1}^{2^n} Y_{(i,n)}}{2^n} - \mathbb{E}[Y_{(1,n)}] \xrightarrow{\text{a.s.}} 0$$

$$\text{and } \mathbb{E}[Y_{(1,n)}] = 2^{-n} \cdot \mathbb{E}[Y]$$

So w.p. 1,

$$\lim_{n \rightarrow \infty} \frac{\sum_{i=1}^{2^n} Y_{(i,n)}}{2^n} - 2^{-n} \cdot \mathbb{E}[Y] = 0$$

and $\sum_{i=1}^{2^n} Y_{(i,n)} = \sum_{j=1}^{k_n} [B(t_j, n) - B(t_{j-1}, n)]$ by definition of T_n & $Y_{(i,n)}$

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{S_n - \mathbb{E}[Y]}{2^n} = 0$$

$$\Rightarrow \lim S_n = \mathbb{E}[Y] \text{ w.p. 1}$$

Q5:

$$1) P\{K_n=0\} = 1 - P\{B_s=0 \text{ for some } 1 \leq s \leq 2\} = : 1 - P(A).$$

Computing $P(A)$ we have:

$$\begin{aligned} P(A) &= \int_{-\infty}^{\infty} P(A | B_1=x) \underbrace{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx}_{d\mu_{B_1}} = \int_{-\infty}^{\infty} P(A | B_1=x) \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= 2 \int_0^{\infty} P(A | B_1=x) \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

and

$$\begin{aligned} P(A | B_1=x) &= P\left\{\max_{0 \leq s \leq 1} B_s > x\right\} \stackrel{\text{same distribution}}{=} 2 \int_x^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \stackrel{\text{Reflection}}{=} 2 \int_x^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \\ \Rightarrow P(A) &= 4 \int_0^{\infty} \left[\int_x^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \right] \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= 4 \cdot \frac{1}{2\pi} \cdot \int_{\frac{\pi}{4}}^{\frac{\pi}{2}} \int_0^{\infty} e^{-\frac{r^2}{2}} r dr d\theta = 1 - \frac{2}{\pi} \arctan 1. \end{aligned}$$

$$\Rightarrow P\{K_n=0\} = 1 - P(A) = \frac{2}{\pi} \arctan 1 = \frac{1}{2}$$

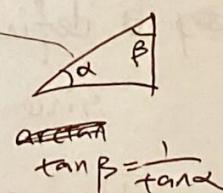
2) Ok we've computed it above so I'll not really compute again for

$\frac{1}{j}, \frac{1}{j-1}, \dots, \frac{1}{2}, \frac{1}{1}$, but as is shown in class, we have by scaling that

$$P(E_{j,n}) = P\left\{B_t=0 \text{ for } \exists 1 \leq t \leq \frac{j}{j-1} = 1 + \frac{1}{j-1}\right\}$$

$$= 1 - \frac{2}{\pi} \arctan \left(\sqrt{\frac{1}{j-1}} \right) = 1 - \frac{2}{\pi} \arctan \left(\sqrt{j-1} \right)$$

$$= \frac{2}{\pi} \arctan \left(\frac{1}{\sqrt{j-1}} \right)$$



Now $\frac{1}{\Gamma_{j-1}}$ is small we can just add things up using Taylor.

Say $j \geq 5$, then $\frac{1}{\Gamma_{j-1}} < \frac{1}{2}$ and taylor:

$\arctan(x) = x - \frac{x^3}{3} + \frac{x^5}{5} - \frac{x^7}{7} + \dots$, But we actually need $j \geq 2$
which is convenient

we get

$$\cancel{\text{arctan}(x) = x - \frac{x^3}{3} + \frac{x^5}{5} - \frac{x^7}{7} + \dots} \quad \frac{2}{\pi} \arctan\left(\frac{1}{\Gamma_{j-1}}\right)$$

$$= \frac{1}{2^{n/2}} \cdot \sum_{j=2^n+1}^{2^{n+1}} \cdot \frac{2}{\pi} \left(\frac{1}{\Gamma_{j-1}} + O\left(\frac{1}{2^{3n/2}}\right) \right)$$

$$= \frac{2}{\pi} \cdot 2^{-n/2} \cdot \sum_{j=2^n+1}^{2^{n+1}} \left(\frac{1}{\Gamma_{j-1}} + O\left(\frac{1}{2^{3n/2}}\right) \right)$$

$$= \left(\frac{2}{\pi} \cdot 2^{-n/2} \cdot \sum_{j=2^n+1}^{2^{n+1}} \frac{1}{\Gamma_{j-1}} \right) + \frac{2}{\pi} \cdot 2^{-n/2} O(2^{-3n/2}) \cdot 2^n \xrightarrow{\text{summation}}$$

$$= \frac{2}{\pi} \cdot 2^{-n} \cdot \sum_{j=2^n+1}^{2^{n+1}} \frac{1}{\Gamma_{j-1}} + O\left(\frac{1}{n}\right)$$

$$\text{So } \lim_{n \rightarrow \infty} 2^{-n/2} \mathbb{E}[K_n] = \lim_{n \rightarrow \infty} \frac{2}{\pi} 2^{-n} \sum_{j=2^n+1}^{2^{n+1}} \frac{1}{\Gamma_{j-1}} \frac{1}{2^n}$$

$$= \frac{2}{\pi} \int_1^2 \frac{1}{\sqrt{x}} dx = \frac{2}{\pi} (2\sqrt{2} - 2) \text{ by Riemann integration.}$$

$$3). \quad P(E_{j,n} \cap E_{k,n}) = P(E_{k,n} | E_{j,n}) \cdot P(E_{j,n})$$

We have bound for $P(E_{j,n})$ so let's look at $P(E_{k,n} | E_{j,n})$

say to define $\sigma_j := \max \{ s < \frac{j}{2^n}; B_s = 0 \}$

Since condition on $E_{j,n}$, $\sigma_j \in [\frac{j-1}{2^n}, \frac{j}{2^n}]$

Then problem becomes $P\{B_t = 0 \text{ for } t \in \left(\frac{k-1}{2^n}, \frac{k}{2^n}\right)\}$ P_{10}

And now by same argument

$$P\{B_t=0 \text{ for some } t \in \left(\frac{k-1}{2^n} - \sigma_j, \frac{k}{2^n} - \sigma_j\right)\}; \text{ here scale: } \frac{k-1}{2^n} - \sigma_j \rightsquigarrow 1.$$

$$= -\frac{\pi}{2} \arctan\left(\frac{1}{\sqrt{\frac{1}{k-1} - \sigma_j^2}}\right) = \frac{\pi}{2} \arctan\left(\frac{1}{\sqrt{k-1 - 2^n \sigma_j^2}}\right).$$

Now just bound issue

$$P(E_{j,n} \cap E_{k,n}) = \frac{4}{\pi^2} \cdot \arctan\left(\frac{1}{\sqrt{k-1 - 2^n \sigma_j^2}}\right) \cdot \arctan\left(\frac{1}{\sqrt{\sigma_{j-1}}}\right)$$

$$\text{But } 2^n \sigma_j \in \left[2^{n-j-1}, 2^n \frac{j-1}{2^n}\right] = [j-1, j]$$

$$\Rightarrow k-1 - 2^n \sigma_j \in [k-j, k-j-1] \quad \text{ok we first forget when } k=j+1$$

$$\text{and } j-1 \in [2^n, 2^{n+1}]$$

$$\begin{aligned} \text{So Taylor gives } & \arctan\left(\frac{1}{\sqrt{k-1 - 2^n \sigma_j}}\right) \cdot \arctan\left(\frac{1}{\sqrt{\sigma_{j-1}}}\right) \\ & \leq \frac{1}{\sqrt{k-1 - 2^n \sigma_j}} \cdot \frac{1}{\sqrt{\sigma_{j-1}}} \leq \frac{1}{\sqrt{k-j} \cdot 2^{n/2}} \end{aligned}$$

for the other direction, we have to keep cubic term but roughly the same:

$$\begin{aligned} & \arctan\left(\frac{1}{\sqrt{k-1 - 2^n \sigma_j}}\right) \cdot \arctan\left(\frac{1}{\sqrt{\sigma_{j-1}}}\right) \\ & \geq \left(\frac{1}{\sqrt{k-1 - 2^n \sigma_j}} - \frac{1}{3(\sqrt{k-1 - 2^n \sigma_j})^3} \right) \left(\frac{1}{\sqrt{\sigma_{j-1}}} - \frac{1}{3(\sqrt{\sigma_{j-1}})^3} \right) \geq \left(\frac{1}{\sqrt{k-j-1}} - \frac{1}{3(k-j)^{3/2}} \right) \left(2^{\frac{(n+1)}{2}} - \frac{1}{3} \cdot 2^{-\frac{3n}{2}} \right) \\ & \geq \frac{1}{\sqrt{k-j-1} \cdot 2^{\frac{(n+1)}{2}}} - \frac{1}{3(k-j)^{3/2} \cdot 2^{\frac{(n+1)}{2}}} - \frac{1}{3\sqrt{k-j-1} \cdot 2^{n/2}} \\ & = \left(\frac{\frac{1}{\sqrt{k-j}}}{\sqrt{k-j-1}} - \frac{1}{3 \cdot (k-j) \cdot \sqrt{2}} - \frac{\frac{1}{\sqrt{k-j}}}{3\sqrt{k-j-1} \cdot 2^n} \right) \cdot \frac{1}{\sqrt{k-j} \cdot 2^{n/2}} \\ & \stackrel{?}{=} \left(\frac{1}{3\sqrt{2}} \left(\frac{\sqrt{k-j}}{\sqrt{k-j-1}} - \frac{1}{k-j} \right) + \frac{\sqrt{k-j}}{3\sqrt{k-j-1}} \left(\frac{1}{\sqrt{2}} - \frac{1}{2^n} \right) + \frac{\sqrt{k-j}}{3\sqrt{2} \cdot \sqrt{k-j-1}} \right) \cdot \frac{1}{(k-j) \cdot 2^{n/2}} \end{aligned}$$

Where $\frac{\overline{f_{k-j}}}{\overline{f_{k-j-1}}} - \frac{1}{k-j} = \frac{(k-j)^{3/2} - \overline{f_{k-j-1}}}{(k-j)\overline{f_{k-j-1}}} > 0$ since nominator > 0.

and we see we can bound below by 0 or thus.

Same for the second term.

So for the third term $\frac{1}{3\sqrt{2}} \cdot \frac{\overline{f_{k-j}}}{\overline{f_{k-j-1}}} \geq \frac{1}{3\sqrt{2}}$ uniform in n .

Thus. for $C_2 = \frac{2}{\pi}$, $C_1 = \frac{1}{3\sqrt{2}} \cdot \frac{2}{\pi}$

we have (for $k=j+1$, just adjust constant C_1, C_2 to suit that case)
 \downarrow
 $\tilde{C}_1 \neq C_1$ (still uniform in n).

$$\frac{\tilde{C}_1}{2^{n/2} \overline{f_{k-j}}} \leq P(E_{j,n} \cap E_{k,n}) \leq \frac{\tilde{C}_2}{2^{n/2} \overline{f_{k-j}}}.$$

$$4). \mathbb{E}[k_n^2] = \mathbb{E} \left[\sum_{\substack{j, k \in \mathbb{Z} \\ j, k \in [2^n, 2^{n+1}]}} \mathbb{1}_{\{E_{j,n} \cap E_{k,n}\}} \right]$$

$$= \underbrace{\sum_{j=2^n+1}^{2^{n+1}} P(E_{j,n})}_{\text{part (2)}} + 2 \sum_{j < k} P(E_{j,n} \cap E_{k,n})$$

part (3)

$$\leq \frac{2}{\pi} (2\sqrt{2}-2) \cdot 2^{n/2} + 2 \cdot C_2 \cdot \sum_{\substack{j < k \\ j \in \text{range}}} \frac{1}{\overline{f_{k-j}}} \cdot 2^{n/2}$$

$$\leq C_3 \cdot 2^{n/2} + \tilde{C}_2 \cdot 2^{n/2} \cdot \sum_{i=1}^{2^n-1} \frac{2^n-i}{\overline{f_i}}$$

$$\leq C_3 \cdot 2^{n/2} + \tilde{C}_2 \cdot 2^{n/2} \cdot \sum_{i=1}^{2^n-1} \frac{1}{\overline{f_i}} \quad \leftarrow \text{use left Riemann sum}$$

$$\leq C_3 \cdot 2^{n/2} + \tilde{C}_2 \cdot 2^{n/2} \cdot \left(\int_1^{2^n} \frac{1}{\overline{f_x}} dx \right)$$

$$\leq C_3 \cdot 2^{n/2} + \tilde{C}_2 \cdot 2^{n/2} \left(2 \cdot 2^{\frac{n}{2}} - 2 \right) = (C_3 - 2\tilde{C}_2) 2^{n/2} + 2 \cdot \tilde{C}_2 \cdot 2^n$$

$$\leq (|C_3 - 2\tilde{C}_2| + 2\tilde{C}_2) \cdot 2^n =: C \cdot 2^n$$

