

## MEASURE THEORETIC PROBABILITY III HW 4

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STAT 38300

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Exercise 1 - Subsequence of subsequences method:

Consider the sequence of random variables  $X_n$  defined as:

$$X_n = \begin{cases} (-1)^n \left(1 - \frac{1}{n}\right) & \text{with probability } \frac{1}{2} \\ n & \text{with probability } \frac{1}{2} \end{cases}$$

Show that the sequence converges in distribution but not in probability or almost surely.

Hint: Consider subsequences for even and odd  $n$ .

Exercise 2 - Kolmogorov's three-series theorem:

Let  $X_n$  be a sequence of independent random variables with  $E[X_n] = 0$  for all  $n$ . Suppose that the following conditions hold:

$\sum_{n=1}^{\infty} |E[X_n^2]| < \infty$   $\sum_{n=1}^{\infty} P(|X_n| > \varepsilon) < \infty$  for all  $\varepsilon > 0$   $\sum_{n=1}^{\infty} E[|X_n|^3] < \infty$  Use Kolmogorov's three-series theorem to show that the series  $\sum_{n=1}^{\infty} X_n$  converges almost surely.

Exercise 3 - Martingale convergence theorem:

Let  $X_n$  be a martingale with respect to a filtration  $F_n$ . Assume that there exists a constant  $K$  such that  $|X_n| \leq K$  for all  $n$ . Prove that  $X_n$  converges almost surely. Hint: Apply the Doob's martingale convergence theorem.

**Exercise 1.**

*Proof.*

□

**Exercise 2. Ex 13.1**

*Proof.* Prove that every bounded martingale is uniformly integrable. Hint: Use the definition of a bounded martingale and the definition of uniformly integrable martingales.

Let  $X_n$  be a martingale. Show that if  $X_n$  is uniformly integrable, then  $X_\infty = \lim_{n \rightarrow \infty} X_n$  exists almost surely, and  $E[|X_\infty|] < \infty$ . Hint: Use the martingale convergence theorem.

Let  $X_n$  be a martingale, and let  $T$  be a stopping time with respect to the filtration generated by  $X_n$ . Prove that if  $X_n$  is uniformly integrable, then  $X_T$  is integrable and  $E[X_T] = E[X_0]$ . Hint: Apply the optional stopping theorem to the stopped process  $X_{n \wedge T}$ .

Let  $Y_n$  be a sequence of i.i.d. random variables with  $E[|Y_1|] < \infty$ . Define the partial sum process  $S_n = \sum_{k=1}^n Y_k$ . Show that the martingale  $S_n$  is uniformly integrable if and only if  $E[Y_1^2] < \infty$ . Hint: Use the definition of uniformly integrable martingales and the properties of the partial sum process.

Let  $X_n$  be a martingale. Define a new martingale  $Y_n$  by  $Y_n = X_n^2 - \sum_{k=1}^{n-1} E[X_k^2]$ . Prove that if  $X_n$  is uniformly integrable, then  $Y_n$  is also uniformly integrable. Hint: Use the definition of uniformly integrable martingales and the properties of conditional expectation.

□

**Exercise 3. Ex 13.2.**

*Proof.*

Question 1: Let  $X_n$  be a martingale with respect to the filtration  $\mathcal{F}_n$ .

(a) Let  $X$  and  $Y$  be integrable random variables such that  $E[X|\mathcal{F}_n] = Y$ . Prove that  $E[\phi(X)|\mathcal{F}_n] = \phi(Y)$  for any bounded, continuous function  $\phi$ .

(b) Let  $X_n = \sum_{k=1}^n Y_k$ , where  $Y_k$  is a sequence of random variables with  $E[Y_k|\mathcal{F}_{k-1}] = 0$ . Prove that if  $\sup_n E[|X_n|^2] < \infty$ , then  $X_n$  is a uniformly integrable martingale.

Question 2: Consider a sequence of i.i.d. random variables  $X_n$  with  $E[X_1] = 0$  and  $\text{Var}(X_1) = \sigma^2 < \infty$ . Define the partial sum process  $S_n = \sum_{k=1}^n X_k$ .

(a) Prove that the process  $M_n = S_n^2 - n\sigma^2$  is a martingale.

(b) Let  $T$  be the stopping time defined as  $T = \inf n \geq 1 : S_n = a \text{ or } S_n = -b$ , where  $a > 0$  and  $b > 0$ . Apply the optional stopping theorem to the martingale  $M_n$  and stopping time  $T$  to derive an expression for  $E[S_T^2]$ .

(c) Use the result from part (b) to compute the probability  $P(S_T = a)$ .

Question 3: Consider a continuous-time martingale  $X_t$  with respect to the filtration  $\mathcal{F}_t$ .

(a) State the Doob-Meyer decomposition theorem for continuous-time submartingales.

(b) Let  $X_t$  be a continuous-time submartingale. Prove that there exists an increasing, predictable process  $A_t$  such that  $M_t = X_t - A_t$  is a continuous-time martingale.

(c) Let  $N_t$  be a Poisson process with intensity  $\lambda > 0$ . Show that the compensated Poisson process  $M_t = N_t - \lambda t$  is a continuous-time martingale.

For each question, make sure to provide detailed proofs and arguments, using the definitions and properties of martingales, filtrations, stopping times, and any relevant theorems or lemmas.

Thus, we have shown that for any simple function  $g$ ,  $E[g(X)\mathbb{1}_A] = E[g(Y)\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ . Now, let  $\phi$  be a continuous bounded function. Since  $\phi$  is Borel-measurable, we can approximate  $\phi$  by a sequence of simple functions  $(g_n)$ , such that  $g_n \rightarrow \phi$  pointwise.

Question 1:

(a) Given that  $E[X\mathbb{1}_A] = E[Y\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ , we want to show that  $E[\phi(X)|\mathcal{F}_n] = \phi(Y)$ .

To show this, we will use the definition of conditional expectation. We need to show that for all  $A \in \mathcal{F}_n$ , we have  $E[\phi(X)\mathbb{1}_A] = E[\phi(Y)\mathbb{1}_A]$ .

First, let's establish that  $\phi(Y)$  is  $\mathcal{F}_n$ -measurable. Since  $Y$  is  $\mathcal{F}_n$ -measurable, and  $\phi$  is a continuous function, the composition  $\phi(Y)$  is also  $\mathcal{F}_n$ -measurable.

Now, we cannot replace  $\phi(X)$  with  $\phi(Y)$  directly, as you correctly pointed out. Instead, we will use an approximation argument. Since  $\phi$  is a continuous function on a compact set (because  $\phi$  is bounded), it is uniformly continuous. This means that for any  $\varepsilon > 0$ , there exists a  $\delta > 0$  such that if  $|x - y| < \delta$ , then  $|\phi(x) - \phi(y)| < \varepsilon$ .

Let's approximate  $\phi(X)$  by a simple function  $s(X) = \sum_{i=1}^m c_i \mathbb{1}_{B_i}(X)$ , where  $B_i$  are Borel sets in the domain of  $X$ , and  $c_i$  are constants, such that  $|s(X) - \phi(X)| < \varepsilon$  almost surely. Then, for each  $i$ , we have  $E[s(X)\mathbb{1}_{A \cap B_i}] = E[s(Y)\mathbb{1}_{A \cap B_i}]$ . This implies  $E[s(X)\mathbb{1}_A] = E[s(Y)\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ .

Now, let's consider the difference  $E[\phi(X)\mathbb{1}_A] - E[s(X)\mathbb{1}_A]$ . We have:

$$|E[\phi(X)\mathbb{1}_A] - E[s(X)\mathbb{1}_A]| \leq E[|\phi(X) - s(X)|\mathbb{1}_A] \leq E[\varepsilon\mathbb{1}_A] = \varepsilon P(A)$$

where the inequality follows from the approximation we made. Similarly, we have

$$|E[\phi(Y)\mathbb{1}_A] - E[s(Y)\mathbb{1}_A]| \leq \varepsilon P(A). \text{ Hence,}$$

$$|E[\phi(X)\mathbb{1}_A] - E[\phi(Y)\mathbb{1}_A]| \leq 2\varepsilon P(A).$$

Since this holds for any  $\varepsilon > 0$ , we have  $E[\phi(X)\mathbb{1}_A] = E[\phi(Y)\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ . Therefore,

Therefore, by the definition of conditional expectation, we have shown that  $E[\phi(X)|\mathcal{F}_n] = \phi(Y)$ . This completes the proof.

In summary, we approximated  $\phi(X)$  by a simple function  $s(X)$  and showed that the expectations of the approximations satisfy  $E[s(X)\mathbb{1}_A] = E[s(Y)\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ . Then, we used the uniform continuity of  $\phi$  to argue that the difference between the expectations of  $\phi(X)$  and  $s(X)$ , and  $\phi(Y)$  and  $s(Y)$ , is arbitrarily small. This allowed us to conclude that  $E[\phi(X)\mathbb{1}_A] = E[\phi(Y)\mathbb{1}_A]$  for all  $A \in \mathcal{F}_n$ , which implies that  $E[\phi(X)|\mathcal{F}_n] = \phi(Y)$ .

(b) From the given condition,  $E[Y_k|\mathcal{F}_{k-1}] = 0$ , so  $X_n$  is a martingale. Now we need to show that  $X_n$  is uniformly integrable. Note that:

$$\sup_n E[|X_n|^2] = \sup_n E\left[\left|\sum_{k=1}^n Y_k\right|^2\right] \leq \sup_n \sum_{k=1}^n E[|Y_k|^2]$$

By Cauchy-Schwarz inequality, we have:

$$E[|X_n|^2] \leq \left(E\left[\sum_{k=1}^n |Y_k|^2\right]\right)^2 \leq n \sum_{k=1}^n E[|Y_k|^2]$$

Since  $\sup_n E[|X_n|^2] < \infty$ , it follows that  $X_n$  is uniformly integrable.

Question 2:

(a) We want to show that  $E[M_{n+1}|\mathcal{F}_n] = M_n$ . Observe that:

$$E[M_{n+1}|\mathcal{F}_n] = E[(S_{n+1}^2 - (n+1)\sigma^2)|\mathcal{F}_n] = E[(S_n + X_{n+1})^2 - (n+1)\sigma^2|\mathcal{F}_n]$$

Expanding, we get:

$$E[M_{n+1}|\mathcal{F}_n] = E[(S_n^2 + 2S_nX_{n+1} + X_{n+1}^2 - (n+1)\sigma^2)|\mathcal{F}_n]$$

Using the linearity of conditional expectation:

$$E[M_{n+1}|\mathcal{F}_n] = S_n^2 - n\sigma^2 + 2E[S_nX_{n+1}|\mathcal{F}_n] + E[X_{n+1}^2|\mathcal{F}_n] - \sigma^2$$

As  $E[X_{n+1}|\mathcal{F}_n] = 0$  and  $E[X_{n+1}^2|\mathcal{F}_n] = \sigma^2$ , we have:

$$E[M_{n+1}|\mathcal{F}_n] = S_n^2 - n\sigma^2 = M_n$$

Thus,  $M_n$  is a martingale.

(b) Since  $M_n$  is a martingale and  $T$  is a bounded stopping time, by the optional stopping theorem:

$$E[M_T] = E[M_0]$$

So, we have:

$$E[M_T] = E[M_0]$$

Substitute  $M_T = S_T^2 - T\sigma^2$  and  $M_0 = S_0^2 = 0$ , we get:

$$E[S_T^2] - E[T\sigma^2] = 0$$

Thus,

$$E[S_T^2] = E[T\sigma^2]$$

(c) Using the result from part (b), we want to compute the probability  $P(S_T = a)$ . Note that:

$$E[S_T^2] = a^2 P(S_T = a) + (-b)^2 P(S_T = -b)$$

We also know from part (b) that  $E[S_T^2] = E[T\sigma^2]$ . Using the law of total expectation, we can write  $E[T\sigma^2]$  as:

$$E[T\sigma^2] = a^2 P(S_T = a) + (-b)^2 P(S_T = -b)$$

Now,  $P(S_T = a) + P(S_T = -b) = 1$ , so we can solve for  $P(S_T = a)$ :

$$P(S_T = a) = \frac{E[T\sigma^2] + b^2}{a^2 + b^2}$$

Question 3:

(a) The Doob-Meyer decomposition theorem states that for any continuous-time submartingale  $X_t$ , there exists a unique increasing, predictable process  $A_t$  such that  $M_t = X_t - A_t$  is a continuous-time martingale.

(b) To prove the existence of such process  $A_t$ , we define the following process:

$$A_t = \int_0^t E[dX_s|\mathcal{F}_{s-}]$$

Since  $X_t$  is a submartingale,  $E[dX_s|\mathcal{F}_{s-}] \geq 0$ , and therefore,  $A_t$  is an increasing process. Furthermore,  $A_t$  is predictable as it is adapted to the filtration  $\mathcal{F}_t$ .

Now, we have:

$$M_t = X_t - A_t = X_t - \int_0^t E[dX_s|\mathcal{F}_{s-}]$$

We want to show that  $M_t$  is a martingale. Observe that:

$$E[M_{t+h}|\mathcal{F}_t] = E[X_{t+h} - \int_0^{t+h} E[dX_s|\mathcal{F}_{s-}]|\mathcal{F}_t]$$

Using the linearity of conditional expectation and the fact that  $X_t$  is a submartingale, we have:

$$E[M_{t+h}|\mathcal{F}_t] = X_t + E[\int_t^{t+h} E[dX_s|\mathcal{F}_{s-}]|\mathcal{F}_t] - \int_0^t E[dX_s|\mathcal{F}_{s-}]$$

Since the integrand is predictable, we can apply the tower property:

$$E[M_{t+h}|\mathcal{F}_t] = X_t - \int_0^t E[dX_s|\mathcal{F}_{s-}] = M_t$$

Thus,  $M_t$  is a martingale, and the process  $A_t$  exists.

(c) For a Poisson process  $N_t$  with intensity  $\lambda > 0$ , we want to show that the compensated Poisson process  $M_t = N_t - \lambda t$  is a continuous-time martingale.

First, we note that  $N_t$  has independent increments, so the process  $M_t$  also has independent increments. To show that  $M_t$  is a martingale, we need to verify that  $E[M_{t+h}|\mathcal{F}_t] = M_t$  for all  $t \geq 0$  and  $h > 0$ . We have:

$$E[M_{t+h}|\mathcal{F}_t] = E[N_{t+h} - \lambda(t+h)|\mathcal{F}_t] = E[(N_{t+h} - N_t) - \lambda h|\mathcal{F}_t]$$

Since  $N_{t+h} - N_t$  is the number of events in the interval  $(t, t+h]$  and has a Poisson distribution with parameter  $\lambda h$ , its expectation is  $\lambda h$ . Therefore:

$$E[M_{t+h}|\mathcal{F}_t] = E[(N_{t+h} - N_t) - \lambda h|\mathcal{F}_t] = \lambda h - \lambda h = M_t$$

Thus,  $M_t$  is a continuous-time martingale.

□

**Exercise 4.** 13.3*Proof.*

**Exercise 5. 14.1***Proof.*