

Linear 6

Tonghao Zhang

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6.25

The analyst can modify the Y_i by deducing the X_{i2} component: $Y_i^* = Y_i - \beta_2 X_{i2}$. Then fit Y_i^* with covariates X_1 , X_3 will yield desiring result.

6.26

Simple regression \hat{Y} on Y will yield the exact set of coefficient H , for there exist only one unique least square solution. To be specific, we are looking for c_0 , c_1 such that $\sum (c_0 + c_1 \hat{Y}_i - Y_i)^2$ is minimized. As $\hat{Y}_i = b_0 + b_1 X_{1i} + \dots + b_p X_{pi}$, the criterion is to minimize $\sum (c_0 + c_1 b_0 + c_1 b_1 X_{1i} + \dots + c_1 b_p X_{pi})^2$. This is still a least square solution of X_i on Y , the uniqueness of least square ensures $c_0 = 0$, $c_1 = 1$, so the regression fit is identical to that of X_i and Y . Now that the fitted values are the same, SSE is the same, so will be R^2 .

6.27

- b: $[b_0, b_1, b_2] = [33.932, 2.784, -0.264]$
- e: $[-2.700, -1.230, -1.637, -1.330, -0.090, 6.987]$
- H:

##		[,1]	[,2]	[,3]	[,4]	[,5]	[,6]
##	[1,]	0.23143293	0.25167585	0.21178735	0.1488684	-0.05475543	0.21099091
##	[2,]	0.25167585	0.31240459	0.09437844	0.2662773	-0.14787283	0.22313666
##	[3,]	0.21178735	0.09437844	0.70442026	-0.3191744	0.10446672	0.20412159
##	[4,]	0.14886839	0.26627729	-0.31917435	0.6142563	0.14143492	0.14833743
##	[5,]	-0.05475543	-0.14787283	0.10446672	0.1414349	0.94039955	0.01632707
##	[6,]	0.21099091	0.22313666	0.20412159	0.1483374	0.01632707	0.19708635

- SSR: 3009.92
- $s^2\{b\}$:

##		X1	X2
##		715.43035	-34.1569695
##	X1	-34.15697	1.6615716
##	X2	-13.59416	0.6440307

- $\hat{Y}_h = 53.847$
- $s^2\{\hat{Y}_h\} = 5.424$

7.1

- $SSR(X_1|X_2)$ df = 1
- $SSR(X_2|X_1, X_3)$ df = 1
- $SSR(X_1, X_2|X_3, X_4)$ df = 2
- $SSR(X_1, X_2, X_3|X_4, X_5)$ df = 3

7.3

```
data7.3<-read.table("CH06PR05.txt",header=FALSE)
colnames(data7.3)<-c("Y","X1","X2")
lm7.3<-lm(Y~X1+X2,data = data7.3)
anova(lm7.3)

## Analysis of Variance Table
##
## Response: Y
##           Df Sum Sq Mean Sq F value    Pr(>F)
## X1          1 1566.45  1566.45  215.947 1.778e-09 ***
## X2          1  306.25   306.25   42.219 2.011e-05 ***
## Residuals 13    94.30     7.25
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

a)

$SSR(X_1) = 1566.45$, $SSR(X_2|X_1) = 306.25$.

b)

$H_0 : \beta_2 = 0$, $H_A : \beta_2 \neq 0$

decision rule: if $F^* > F(0.99; 1, 13)$, reject H_0 .

$F^* = \frac{306.25}{7.25}$, while threshold = 9.0738057, reject H_0 . The p-value is 2.0056483×10^{-5} .

7.6

```
## Analysis of Variance Table
##
## Response: Y
##           Df Sum Sq Mean Sq F value    Pr(>F)
## X1          1 8275.4   8275.4  81.8026 2.059e-11 ***
## X2          1  480.9    480.9   4.7539 0.03489 *
## X3          1  364.2    364.2   3.5997 0.06468 .
## Residuals 42 4248.8    101.2
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

a)

$SSR(X_2, X_3|X_1) = 480.9 + 364.2$

$H_0 : \beta_2 = \beta_3 = 0$, H_A : at least one $\beta_i \neq 0$.

decision rule: if $F^* > F(0.975; 2, 42)$, reject H_0 .

$F^* = \frac{480.9+364.2}{101.2} = 4.175$, while the threshold $F(0.975; 2, 42) = 4.0327099$, reject H_0 . p-value is 0.0221872.

7.11

a)

- $R_{Y1}^2 = \frac{SSR(X_1)}{SSTO} = 0.550$, it measures the usefulness of fitting X_1 to Y in linear regression model. Here 0.550 is far from 1, the fit is not very useful.
- $R_{Y2}^2 = \frac{SSR(X_2)}{SSTO} = 0.409$, it measures the usefulness of fitting X_2 to Y in linear regression model. Here 0.409 means this model is not very useful.
- $R_{12}^2 = \frac{SSR(X_1)}{SSE(X_2)} = 0$, it measures the correlation of X_1 and X_2 . 0 basically says there is no correlation between X_1 and X_2 .
- $R_{Y1|2}^2 = \frac{SSR(X_1|X_2)}{SSE(X_2)} = 0.930$, it measures the effectiveness of adding X_1 into linear regression model given X_2 is already in the model. 0.930 indicates it's very effective.
- $R_{Y2|1}^2 = \frac{SSR(X_2|X_1)}{SSE(X_1)} = 0.907$, it measures the effectiveness of adding X_2 into linear regression model given X_1 is already in the model. 0.907 indicates it's very effective.

b)

Because the two variable are uncorrelated. $R_{12}^2 = 0$. Furthermore the extra sum of variance explain by adding new variable is the same regardless of X_1 , X_2 sequence.

Problem 9

part A

Just plug $X = X^*C^t$ into $X^tY = X^tX\hat{\beta}$, $CX^{*t}Y = CX^{*t}X^*C^t\hat{\beta}$, cancel C on both side yield the desire result $X^{*t}Y = X^{*t}X^*\hat{\beta}^*$.

part B

```
Xt%*%X
```

```
##      x1 x2  x3
##      6  0  0   0
## x1  0 70  0   0
## x2  0  0 84   0
## x3  0  0  0 180
```

```
round(solve(Xt%*%X),3)
```

```
##      x1      x2      x3
## 0.167 0.000 0.000 0.000
## x1 0.000 0.014 0.000 0.000
## x2 0.000 0.000 0.012 0.000
## x3 0.000 0.000 0.000 0.006
```

```
round(solve(Xt%*%X)%*%Xt%*%Y,3)
```

```
##      [,1]
##      4.400
## x1 0.583
## x2 0.239
## x3 0.074
```

```
round(t(C)%*(solve(Xt*X)%*Xt*Y),3)
```

```
##      [,1]  
## [1,] 3.353  
## [2,] 1.360  
## [3,] 2.243  
## [4,] 1.924
```

part C

