

## 1. Cálculo de momentos de distribución uniforme continua.

$$X \sim U(a, b), a < b \Leftrightarrow f_X(x) = \frac{1}{b-a} \quad \forall x \in (a, b)$$

$$F_X(x) = \frac{x-a}{b-a} \quad M_X(t) = E[e^{tx}] = \frac{e^{tb} - e^{ta}}{t(b-a)}, \quad t \in \mathbb{R}$$

• Momentos no centrados de orden  $k$ .

$$m_k = E[X^k] = \int_a^b \frac{x^k}{b-a} dx = \left[ \frac{x^{k+1}}{(k+1)(b-a)} \right]_{x=a}^{x=b} = \frac{b^{k+1} - a^{k+1}}{(k+1)(b-a)}$$

• Momentos centrados de orden  $k$ .

$$\begin{aligned} \mu_k = E[(X - m_1)^k] &= \int_a^b \frac{(x - m_1)^k}{b-a} dx = \frac{1}{(b-a)^{k+1}} \left( \left( \frac{b-a}{2} \right)^{k+1} - \left( \frac{a-b}{2} \right)^{k+1} \right) \\ &= \begin{cases} 0 & \text{si } k \in \mathbb{N} \text{ impar} \\ \frac{(b-a)^k}{2^k(k+1)} & \text{si } k \in \mathbb{N} \text{ par} \end{cases} \end{aligned}$$

## 2. Cálculo de las funciones generatrices de momentos de distribuciones: Unif. continua, Normal, exponencial.

$$X \sim U(a, b)$$

$$M_X(t) = E[e^{tx}] = \int_a^b \frac{e^{tx}}{b-a} dx = \frac{e^{tb} - e^{ta}}{t(b-a)}, \quad t \in \mathbb{R}$$

$$X \sim N(0, 1)$$

$$\begin{aligned} M_X(t) = E[e^{tx}] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tx - \frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t^2/2} \frac{e^{tx} e^{-x^2/2}}{e^{t^2/2}} dx \\ &= \frac{e^{t^2/2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tx - \frac{x^2}{2} - \frac{t^2}{2}} dx = \frac{e^{t^2/2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(x-t)^2} dx = e^{t^2/2} \end{aligned}$$

$$X \sim N(\mu, \sigma^2) \Leftrightarrow Z = \frac{X - \mu}{\sigma} \sim N(0, 1) \quad X = \sigma Z + \mu$$

$$\begin{aligned} M_X(t) &= E[e^{tX}] = E[e^{t(\sigma Z + \mu)}] = e^{t\mu} E[e^{t\sigma Z}] \\ &= e^{t\mu} M_Z(t\sigma) = e^{t\mu} e^{(t\sigma)^2/2} = e^{t\mu + \frac{t^2\sigma^2}{2}} \end{aligned}$$

$$X \sim \exp(\lambda), \quad \lambda > 0$$

$$M_X(t) = E[e^{tX}] = \int_0^{\infty} \lambda e^{x(t-\lambda)} dx = \left[ \frac{\lambda e^{x(t-\lambda)}}{t-\lambda} \right]_{x=0}^{x=\infty} = -\frac{\lambda}{t-\lambda}$$

3. Enunciar propiedades de la f. distribución de un vector aleatorio.

Sea  $X = (X_1, \dots, X_n): (\Omega, \mathcal{A}, P) \longrightarrow (\mathbb{R}^n, \mathcal{B}^n, P_X)$

se define la f. distribución

$$F_X: \mathbb{R}^n \longrightarrow [0, 1]$$

$$F_X(x) = P_X((-\infty, x]) = P[X \leq x]$$

$$\text{esto es, } F_X(x_1, \dots, x_n) = P[X_1 \leq x_1, \dots, X_n \leq x_n].$$

Propiedades:

•) Monótona no decreciente  $\forall i=1, \dots, n, \forall x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n$

$$\begin{aligned} x_i < x_i' &\Rightarrow F_X(x_1, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n) \\ &\leq F_X(x_1, \dots, x_{i-1}, x_i', x_{i+1}, \dots, x_n) \end{aligned}$$

•) Continua a la derecha  $\forall i=1, \dots, n \quad \forall x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n$

$$\lim_{x_i' \rightarrow x_i^+} F_X(x_1, \dots, x_{i-1}, x_i', x_{i+1}, \dots, x_n) = F_X(x_1, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n)$$

$$\bullet) \forall i=1, \dots, n \quad \forall x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n$$

$$\lim_{x_i \rightarrow -\infty} F_X(x_1, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n) = F_X(x_1, \dots, x_{i-1}, -\infty, x_{i+1}, \dots, x_n) = 0$$

$$\lim_{x_1, \dots, x_n \rightarrow +\infty} F_X(x_1, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n) = F_X(+\infty, \dots, +\infty) = 1$$

$$\bullet) \forall (x_1, \dots, x_n) \in \mathbb{R}^n \quad \forall (\varepsilon_1, \dots, \varepsilon_n) \in \mathbb{R}_+^n$$

$$\begin{aligned} F_X(x_1 + \varepsilon_1, \dots, x_n + \varepsilon_n) &= \sum_{i=1}^n F_X(x_1 + \varepsilon_1, \dots, x_{i-1} + \varepsilon_{i-1}, x_i, x_{i+1} + \varepsilon_{i+1}, \dots, x_n + \varepsilon_n) \\ &+ \sum_{i=1}^n \sum_{j=1}^n F_X(x_1 + \varepsilon_1, \dots, x_{i-1} + \varepsilon_{i-1}, x_i, x_{i+1} + \varepsilon_{i+1}, \dots, x_{j-1} + \varepsilon_{j-1}, x_j, x_{j+1} + \varepsilon_{j+1}, \dots, x_n + \varepsilon_n) \\ &+ \dots + (-1)^n F_X(x_1, \dots, x_n) \geq 0 \end{aligned}$$

4. Dem caracterización de independencia por conjuntos Borel.

Sea  $X = (X_1, \dots, X_n)$  vector aleatorio.  $X_1, \dots, X_n$  independientes

$$\Leftrightarrow P[X_1 \in B_1, \dots, X_n \in B_n] = P[X_1 \in B_1] \cdots P[X_n \in B_n] \quad \forall B_1, \dots, B_n \in \beta$$

Dem

$\Leftarrow$  si la relación cierta  $\forall B_1, \dots, B_n \in \beta$ , tomando  $B_i = (-\infty, x_i]$   $\forall i=1, \dots, n$ , se tiene la def de independencia.

$\Rightarrow$  Caso discreto  $X = (X_1, \dots, X_n)$  discreto.

$$\begin{aligned} P[X_1 \in B_1, \dots, X_n \in B_n] &= \sum_{x_i \in B_i} P[X_1 = x_1, \dots, X_n = x_n] \stackrel{\text{indep}}{=} \\ &= \sum_{x_i \in B_i} P[X_1 = x_1] P[X_n = x_n] = \sum_{x_i \in B_i} P[X_1 = x_1] \cdots \sum_{x_n \in B_n} P[X_n = x_n] \\ &= P[X_1 \in B_1] \cdots P[X_n \in B_n] \end{aligned}$$

caso continuo:  $X = (X_1, \dots, X_n)$  continuo.

$$\begin{aligned} P[X_1 \in B_1, \dots, X_n \in B_n] &= \int_{B_1} \dots \int_{B_n} f_X(x_1, \dots, x_n) dx_n \dots dx_1 \\ &\stackrel{\text{indep}}{=} \int_{B_1} \dots \int_{B_n} f_1(x_1) \dots f_n(x_n) dx_n \dots dx_1 = \int_{B_1} f_1(x_1) dx_1 \dots \int_{B_n} f_n(x_n) dx_n \\ &= P[X_1 \in B_1] \dots P[X_n \in B_n] \end{aligned}$$

5. Dem de la reproductividad de la Binomial.

$$X_i \sim B(n_i, p) \Rightarrow \sum_{i=1}^n X_i \sim B\left(\sum_{i=1}^n n_i, p\right)$$

$i=1 \dots n$  indep

Dem

$$X \sim B(n, p) \quad M_X(t) = E[e^{tX}] = (pe^t + (1-p))^n, \quad t \in \mathbb{R}, p \in [0, 1], n \in \mathbb{N}.$$

$$\begin{aligned} M_{\sum_{i=1}^n X_i}(t) &= E[e^{t \sum_{i=1}^n X_i}] = E\left[\prod_{i=1}^n e^{tX_i}\right] = \prod_{i=1}^n E[e^{tX_i}] \\ &= \prod_{i=1}^n (pe^t + (1-p))^{n_i} = (pe^t + (1-p))^{\sum_{i=1}^n n_i} \sim B\left(\sum_{i=1}^n n_i, p\right) \end{aligned}$$

~~6. Enunciar Tma multiplicación esperanzas~~

~~$X_1, \dots, X_n$  independientes y  $\exists E[X_i], \forall i=1 \dots n$  entonces:~~

~~a)  $\exists E[X_1, \dots, X_n]$~~

~~b)  $E[X_1, \dots, X_n] = E[X_1] \dots E[X_n]$~~

7. Dem del Tma descomposición de la varianza

Sean  $X, Y$  var. aleatorias definidas en el mismo esp. prob.  
tal que  $\exists E[X^2]$ , entonces.

$\exists E[\text{Var}[X|Y]]$  y  $\text{Var}[E[X|Y]]$  y además,

$$\text{Var}[X] = E[\text{Var}[X|Y]] + \text{Var}[E[X|Y]]$$

Dem

$$\text{Var}[X] = E[(X - E[X])^2]$$

$$= E[E[(X - E[X])^2/Y]]$$

$$= E[E[(X^2 + E[X]^2 - 2XE[X])/Y]]$$

$$= E[E[X^2/Y] + E[X]^2 - 2E[Y]E[X/Y]]$$

$$= E[E[X^2/Y] - E[X/Y]^2 + E[X/Y]^2 + E[X]^2 - 2E[X]E[X/Y]]$$

$$= E[E[X^2/Y] - E[X/Y]^2] + E[(E[X/Y] - E[X])^2]$$

$$= E[\text{Var}[X/Y]] + \text{Var}[E[X/Y]]$$

8. Propiedades de la esperanza condicionada.

1)  $E[C/Y] = C$

2) linealidad. Si  $\exists E[X], a, b \in \mathbb{R}$

$$\Rightarrow \exists E[aY + b/Y] = aE[X/Y] + b$$

3)  $X_1, \dots, X_n$  va,  $\exists E[X_i] \forall i=1, \dots, n$

$$\Rightarrow \forall a_1, \dots, a_n \in \mathbb{R}, \exists E[a_1X_1 + \dots + a_nX_n/Y] = a_1E[X_1/Y] + \dots + a_nE[X_n/Y]$$

4) Si  $X \geq 0$  tal que  $\exists E[X] \Rightarrow E[X/Y] \geq 0$

$$\text{y } E[X/Y] = 0 \Leftrightarrow P[X=0] = 1$$

5) Si  $\exists E[X], E[Y], X_1 \leq X_2 \Rightarrow E[X_1/Y] \leq E[X_2/Y]$

6)  $X, Y$ , indep,  $\exists E[g(X)] \Rightarrow E[g(X)/Y] = E[g(X)]$

en particular,  $E[X/Y] = E[X]$

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6. Enunciat el T<sup>ma</sup> multiplicaci3n de esperanzas.

1) Si  $x_1, \dots, x_n$  indep.  $\Rightarrow \exists E[x_1] \dots E[x_n] \Rightarrow \exists E[x_1 \dots x_n] = E[x_1] \dots E[x_n]$

2) Si  $x_1, \dots, x_n$  indep.  
 $g: \mathbb{R} \rightarrow \mathbb{R}$  med.  
 $\exists E[g(x)] \Rightarrow \exists E[g_1(x_1), \dots, g_n(x_n)] = E[g_1(x_1)] \dots E[g_n(x_n)]$   
 $g(x_1), \dots, g(x_n)$  indep.

3)  $x, y$  indep.  $\Rightarrow \text{cov}(x, y) = 0$   
 ~~$\Leftrightarrow$~~

4) Si  $x_1, \dots, x_n$  indep. amb  $\text{MNC} < \infty$

$$\text{var} \left( \sum_{i=1}^n a_i x_i \right) = \sum_i a_i^2 \text{var}(x_i)$$