AQM Problem Set 3

Ismael Martinez

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Problem 1 - Regularization and Condition Numbers

Part a.

Let $\Gamma = \gamma I$.

$$RSS = \frac{1}{2}(y - X\beta)^{T}(y - X\beta) - \frac{1}{2}\Gamma\beta^{T}\beta$$

$$\frac{\partial RSS}{\partial \beta} = -X^{T}(y - X\hat{\beta}) + \Gamma\hat{\beta} = 0$$

$$\implies -X^{T}y + X^{T}X\hat{\beta} + \Gamma\hat{\beta} = 0$$

$$\implies X^{T}X + \Gamma\hat{\beta} = X^{T}y$$

$$\implies (X^{T}X + \Gamma)\hat{\beta} = X^{T}y$$

$$\implies \hat{\beta} = (X^{T}X + \Gamma)^{-1}X^{T}y = (X^{T}X + \gamma I)^{-1}X^{T}y$$

Part b.

$$cond(A) = \frac{\frac{A^{-1}\Delta||b||}{A^{-1}||b||}}{\frac{\Delta||b||}{||b||}}$$

$$= \frac{||A^{-1}\Delta b||}{||A^{-1}b||} \cdot \frac{||b||}{||\Delta b||}$$

$$= \frac{||A^{-1}\Delta b||}{||\Delta b||} \cdot \frac{||b||}{||A^{-1}b||}$$

$$= ||A^{-1}|| \cdot \frac{1}{||A^{-1}||}$$

$$= ||A^{-1}|| \cdot ||A||$$

By the Cauchy-Schwarz inequality, we get

$$cond(A) = ||A^{-1}|| \cdot ||A|| \ge ||A^{-1}A|| = 1$$

Part c.

We assume that $\Delta A = 0$ and $\Delta b \leq b$.

$$\begin{aligned} \frac{\|\Delta\beta\|}{\|\beta\|} &= \frac{\|\Delta(A^{-1}b)\|}{\|A^{-1}b\|} \\ &= \frac{\|\Delta A^{-1}\Delta b\|}{\|A^{-1}b\|} \\ &< 1 < cond(A) \end{aligned}$$

Part d.

A is a positive definite matrix. $B = A + \alpha I$, $\alpha \in [0, 1]$. Looking at the eigenvalues of A and B, we see:

$$Ax = \lambda x$$

$$\implies Ax + \alpha Ix = \lambda x + \alpha x$$

$$\implies Bx = (\lambda + \alpha)x$$

Since the eigenvalues of B are larger than those of A, and A has all positive eigenvalues from being positive definite, then B has all positive eigenvalues

and B has an inverse.

$$cond(A) = ||A^{-1}|| ||A|| \ge ||A^{-1} + \alpha I|| ||A + \alpha I|| = cond(B)$$

Problem 2 - Elastic Net and Variable Selection

Part a.

We first define our functions and augmented data.

$$J_1(\beta) = \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda_2 \|\beta\|_2^2 + \lambda \|\beta\|$$

and

$$J_2(\widetilde{\beta}) = \|\widetilde{\mathbf{y}} - \widetilde{\mathbf{X}}\widetilde{\beta}\|_2^2 + \lambda_1 \|\widetilde{\beta}\|$$
where $c = (1 + \lambda_2)^{-\frac{1}{2}}$, $\beta = c\widetilde{\beta}$, and

$$\widetilde{\mathbf{X}} = c \begin{bmatrix} X \\ \sqrt{\lambda_2} \mathbf{I}_d \end{bmatrix}, \widetilde{\mathbf{y}} = \begin{bmatrix} \mathbf{y} \\ \mathbf{0}_{dx1} \end{bmatrix}$$

1.

Proof.

$$c(\underset{\widetilde{\beta}}{\operatorname{argmin}} J_{2}(\widetilde{\beta})) = c[\underset{\widetilde{\beta}}{\operatorname{argmin}} \|\widetilde{\mathbf{y}} - \widetilde{\mathbf{X}}\widetilde{\beta}\|_{2}^{2} + c\lambda_{1}\|\widetilde{\beta}\|]$$

$$= c[\underset{\beta}{\operatorname{argmin}} \frac{1}{c} [\sum_{j=1}^{N} (\mathbf{y} - c\mathbf{X}\frac{\beta}{c})^{2} + \sum_{j=1}^{D} (0 - c\sqrt{\lambda_{2}}\frac{\beta}{c})^{2} + c\lambda_{1} \sum_{j=1}^{D} |\frac{\beta}{c}|]]$$

$$= \frac{c}{c} [\underset{\beta}{\operatorname{argmin}} \sum_{j=1}^{N} (\mathbf{y} - \mathbf{X}\beta)^{2} + \lambda_{2} \sum_{j=1}^{D} \beta^{2} + \lambda_{1} \|\beta\|]$$

$$= \underset{\beta}{\operatorname{argmin}} J_{1}(\beta)$$

2.

$$J_{W_1}(\beta) = (\mathbf{y} - \mathbf{X}\beta)^T \mathbf{W} (\mathbf{y} - \mathbf{X}\beta) + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\beta\|$$
where $\mathbf{W} = \begin{bmatrix} \frac{1}{\sigma_1^2} & 0 & \dots & 0\\ 0 & \frac{1}{\sigma_2^2} & \dots & 0\\ \vdots & \vdots & \ddots & \vdots\\ 0 & 0 & \dots & \frac{1}{\sigma_n^2} \end{bmatrix}$

$$\hat{\beta} = \operatorname*{argmin}_{\beta} J_{W_1}(\beta)$$

3.

$$J_{W_2}(\widetilde{\beta}) = (\widetilde{\mathbf{y}} - \widetilde{\mathbf{X}}\widetilde{\beta})^T \widetilde{\mathbf{W}} (\widetilde{\mathbf{y}} - \widetilde{\mathbf{X}}\widetilde{\beta}) + \lambda_1 \|\widetilde{\beta}\|$$
where $\widetilde{\mathbf{W}} = \begin{bmatrix} \mathbf{W} \\ \mathbf{I}_d \end{bmatrix}$

$$\hat{\beta} = \underset{\widetilde{\beta}}{\operatorname{argmin}} J_W(\widetilde{\beta})$$

4.

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} \beta^T \left(\frac{\mathbf{X}^T \mathbf{W} \mathbf{X} + \lambda_2 \mathbf{I}}{1 + \lambda_2} \right) \beta - 2 \mathbf{y}^T \mathbf{W} \mathbf{X} \beta + \lambda_1 \|\beta\|_1$$

Part b.

$$J_W(\beta) = \beta^T \left(\frac{X^T W X}{1 + \lambda_2} \right) \beta - 2y^T W X \beta + \lambda_1 \|\beta\|_1$$

$$\frac{\partial J_W}{\partial \beta} = \beta^T \left(\left(\frac{X^T W X}{1 + \lambda_2} \right)^T + \left(\frac{X^T W X}{1 + \lambda_2} \right) \right) - 2y^T W X + \lambda_1 \partial \|\beta\|_1$$
where $\partial \|\beta\|_1 = \begin{cases} -\lambda_1 & \text{if } \beta_j < 0 \\ [-\lambda_1 . \lambda_1] & \text{if } \beta_j = 0 \\ \lambda_1 & \text{if } \beta_j > 0 \end{cases}$

Part c.

Run the file GradientDescentMethods.py which calls the proximalGradientDescent and subgradientDescent.py algorithms. By testing a range of λ_1 and λ_2 values, we see that the Proximal Gradient Descent method is much faster.

Part d.

- 1. Find the parallelized algorithm in parallelizedGradientDescent.py.
- 2. Find the Bootstrap CI Algorithm in bootstrapBoston.py

Part f.

2. We are clustering a set of normally distributed arrays of size 30, where each point is drawn from a $\mathcal{N}(0, \Sigma)$. It can be important to distinguish between points overlapped in multiple clusters as they show a level of uncertainty to which cluster a point belongs and allows us to view those points as having a probability of belonging to one or the other. When a point must be a part of a single cluster, having this non-overlapping approach is more useful.