MATH 423/533 – BIVARIATE NORMAL MODELLING

We may decide to treat the predictor x as a random variable, and make a bivariate Normal distribution assumption for the data pairs (x_i, y_i) , i = 1, ..., n. We specify that

$$\left[\begin{array}{c} X \\ Y \end{array}\right] \sim \operatorname{Normal}\left(\left[\begin{array}{c} \mu_X \\ \mu_Y \end{array}\right], \left[\begin{array}{cc} \sigma_X^2 & \sigma_{XY} \\ \sigma_{XY} & \sigma_Y^2 \end{array}\right]\right) \equiv \operatorname{Normal}\left(\mu, \Sigma\right).$$

Writing

$$\rho = \frac{\sigma_{XY}}{\sigma_X \sigma_Y}$$

the joint density is

$$f_{X,Y}(x,y) = \frac{1}{2\pi} \frac{1}{\sigma_X \sigma_Y \sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_X}{\sigma_X}\right)^2 - \frac{2\rho(x-\mu_X)(y-\mu_Y)}{\sigma_X \sigma_Y} + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2 \right] \right\}$$

- σ_{XY} is the covariance parameter
- ρ is the correlation parameter.

We may factorize the joint density

$$f_{X,Y}(x,y) = f_{Y|X}(y|x)f_X(x)$$

where $X \sim \text{Normal}(\mu_X, \sigma_X^2)$ and

$$Y|X = x \sim \text{Normal}\left(\mu_Y + \frac{\sigma_{XY}}{\sigma_X^2}(x - \mu_X), \sigma_Y^2 - \frac{\sigma_{XY}^2}{\sigma_X^2}\right) \equiv \text{Normal}\left(\mu_Y + \rho \frac{\sigma_Y}{\sigma_X}(x - \mu_X), \sigma_Y^2(1 - \rho^2)\right)$$

Equating the conditional expectation of Y given X = x

$$\mathbb{E}_{Y|X}[Y|x] = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X}(x - \mu_X)$$

with a simple linear regression

$$\mathbb{E}_{Y|X}[Y|x] = \beta_0 + \beta_1 x$$

we identify

$$\beta_0 = \mu_Y - \mu_X \rho \frac{\sigma_Y}{\sigma_X}$$
 $\beta_1 = \rho \frac{\sigma_Y}{\sigma_X}.$

The sample correlation is

$$r = \frac{\sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum_{i=1}^{n} (x_i - \overline{x})^2 \sum_{i=1}^{n} (y_i - \overline{y})^2}} = \frac{S_{xy}}{\sqrt{S_{xx}SS_T}}$$

so that

$$\hat{\beta}_1 = \left(\frac{\mathrm{SS}_{\mathrm{T}}}{S_{xx}}\right)^{1/2} r \qquad \Longrightarrow \qquad r^2 = \frac{\mathrm{SS}_{\mathrm{R}}}{\mathrm{SS}_{\mathrm{T}}} = R^2.$$

We may test the hypothesis $H_0: \rho = 0$ using the statistic

$$t = \frac{r\sqrt{n-2}}{\sqrt{1-r^2}}$$

which is compared against the relevant quantile of the Student(n-2) distribution.

Bivariate Modelling: Simulation

```
set.seed(234)
 2
   library (MASS)
 3
  n<-10000
 4 muX<-50; sigmaX<-20; muY<-100; sigmaY<-50; rho<-0.5
 5 Sig<-matrix(c(sigmaX^2,rho*sigmaX*sigmaY,rho*sigmaX*sigmaY,sigmaY^2),2,2)
  XY<-mvrnorm(n, mu=c(muX, muY), Sigma=Sig)</pre>
 7 be0<-muY-rho*(sigmaY/sigmaX)*muX;be1<-rho*(sigmaY/sigmaX)</pre>
   Fitting a linear model to the simulated bivariate (x_i, y_i) pairs yields estimates that are close to
   the true values.
   > c(be0,be1)
   [1] 37.50 1.25
10 > coef(lm(XY[,2] \sim XY[,1]))
  (Intercept)
                       XY[, 1]
12
      37.269659
                      1.253241
                       400
                       300
                       200
                       100
                       -100
                                                        100
                                                                  150
                                          Boxplots of residuals
                         150
                         9
                         20
                         -20
```

Figure 1: The residual plots indicate that the fit of the model is adequate.

(10,20]