### Math 423 Linear Regression

Homework II

Frédéric Boileau

Prof. David A. Stephens

```
require(printr)

## Loading required package: printr

setwd(dirname(rstudioapi::getActiveDocumentContext()$path))

## Error: RStudio not running

salary<-read.csv("salary.csv",header=TRUE)

x1<-salary$SPENDING/1000
y<-salary$SALARY</pre>
```

we want to estimate the parameter  $\beta_1$  and  $\beta_0$ , namely the slope and the intercept. We use the least square estimators. This is a case of simple linear regression so we can use the following equations:

$$\hat{\beta}_1 = \frac{S_{xx}}{S_{xy}} \tag{1}$$

$$S_{xx} = \sum_{i=1}^{n} (x_i - \bar{x})^2 \tag{2}$$

$$S_{xy} = \sum_{i=1}^{n} y_i (x_i - \hat{x})$$
 (3)

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \hat{x} \tag{4}$$

```
xbar = mean(x1)
ybar = mean(y)
Sxx = sum((x1 - xbar)^2)
Sxy = sum(y*(x1 - xbar))
slope = Sxy/Sxx
intercept = ybar - slope*xbar
print(slope)
## [1] 3307.585
print(intercept)
## [1] 12129.37
```

#### b and c

The residual standard error is given by

$$\hat{\sigma}^2 = \frac{SS_{\text{Res}}}{n-2} \tag{5}$$

Moreover $SS_{Res}$  is the sum of squares of error:

$$SS_{Res} = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} y_i^2 - n\hat{y}^2 - \hat{\beta}_1 S_{xy}$$
 (6)

$$= \sum_{i=1}^{n} (y_i - \bar{y})^2 - \hat{\beta}_1 S_{xy}$$
 (7)

```
SSRes = sum((y - ybar)^2) - slope*Sxy
n = length(x1)
residualStdError = sqrt(SSRes/(n-2))
print(residualStdError)
## [1] 2324.779
```

#### $\mathbf{d}$

We wish to compute the standard error with the values in the table already given. This table gives us the degrees of freedom (49) and the t value.

$$t_0 = \frac{\hat{\beta}_0}{\text{se}(\hat{\beta}_0)} \Rightarrow \text{se}(\hat{\beta}_0) = \frac{\hat{\beta}_0}{t_0} = \frac{12129.4}{10.13}$$
 (8)

Now to do the computation directly from the data we use the actual formula for the standard error which is given by

$$\operatorname{se}(\hat{\beta}_1) = \sqrt{\frac{MS_{\text{Res}}}{S_{xx}}} \qquad MS_{\text{Res}} = \frac{SS_{\text{Res}}}{n-2} = \hat{\sigma}^2$$
 (9)

```
MSRes = residualStdError / sqrt(n-2)
print(MSRes)
## [1] 332.1113
```

 $\mathbf{e}$ 

We will derive simple expressions from known relationships

$$SS_{\mathrm{Res}} = SS_{\mathrm{T}} - \hat{\beta}_1 S_{xy}$$
  
 $SS_{\mathrm{T}} = SS_{\mathrm{R}} + SS_{\mathrm{Res}}$ 

It is easy to see then that

$$SS_{T} = SS_{Res} + \hat{\beta}_{1}S_{xy}$$
$$SS_{R} = \hat{\beta}_{1}S_{xy}$$

```
SST = SSRes + slope*Sxy
SSR = slope*Sxy
Rsqrd = SSR/SST
print(Rsqrd)
## [1] 0.6967813
```

 $\mathbf{f}$ 

```
p = 2
Fstat = (SSR/(p-1))/(SSRes/(n-p))
print(Fstat)
## [1] 112.5995
```

 $\mathbf{g}$ 

$$y^{\mathsf{T}}(I_n H_1)y = y^{\mathsf{T}}(I_n - H)y + y^{\mathsf{T}}(H - H_1)y$$

The first statement we want to show is

$$trace(I_n - H_1) = n - 1$$

Well the matrix  $I_n$  has  $a_{ii}=1\,\forall\,i\in[1,n]$  and  $h_{ii}=1/n\,\forall\,i\in[1,n]$  By definitions:

$$\operatorname{trace}(I_n - H_1) = \sum_{i=1}^n (a_{ii} - h_{ii})$$
 (10)

$$=\sum_{i=1}^{n} (1 - 1/n) \tag{11}$$

$$= n(1 - 1/n) = n - 1 \tag{12}$$

The second statement we need to prove is that:

$$trace(H - H_1) = p - 1 \tag{13}$$

We use the properties of the trace operator:

$$trace(H - H_1) = trace(H) - trace(H_1)$$
(14)

$$\operatorname{trace}(H) = \operatorname{trace}(X(X^{\mathsf{T}}X)^{-1}X^{\mathsf{T}}) \tag{15}$$

$$= \operatorname{trace}(X^{\mathsf{T}} X (X^{\mathsf{T}} X)^{-1}) \tag{16}$$

$$= \operatorname{trace}(I_p) \quad \text{since} \quad X^{\mathsf{T}} X \in \mathbb{R}^{p \times p} \tag{17}$$

$$= p \tag{18}$$

As shown before the trace of  $H_1$  is 1 and this with the previous derivation proves (13).

#### Numerical part

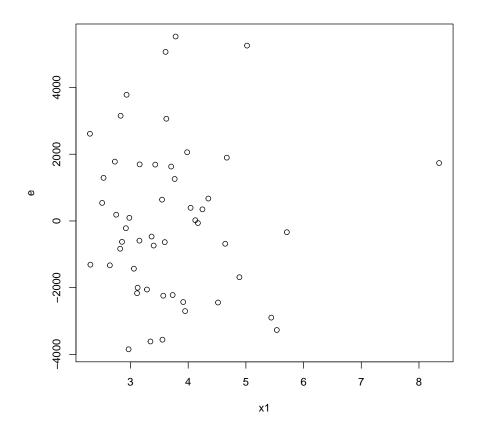
```
require(MASS)
## Loading required package: MASS
## Warning: package 'MASS' was built under R version 3.3.1
bigx = cbind(matrix(1,length(x1)),x1)

n1 =length(x1)
H1 = matrix(1/n1,n1,n1)
sum(diag((diag(n1) - H1)))
## [1] 50

H = bigx %*% ginv(t(bigx) %*% bigx) %*% t(bigx)
sum(diag(H - H1))
## [1] 1
```

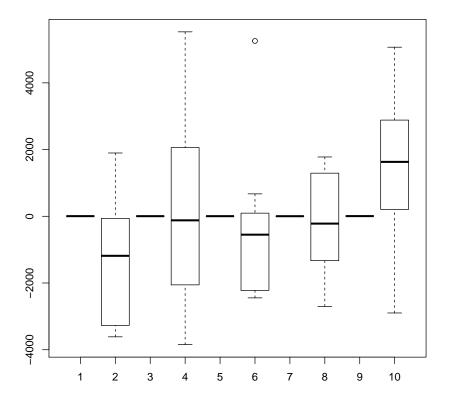
## $\mathbf{h}$

```
yhat = intercept + slope*x1
e = y - yhat
plot(x1,e)
```



```
mean(x1)
## [1] 3.696608
```

The residuals have zero mean.



However a simple box plot shows quite clearly that they do not have constant variance.

sum(e) #The sum of the residuals is zero, i.e. they are orthogonal to each other
## [1] -8.731149e-11
bigx = cbind(matrix(1,length(x1)),x1)
t(bigx)%\*%e#the residuals are orthogonal to the regressors

	0
x1	0

t(yhat)%\*%e#the residuals are orthogonal to the fitted values

-2.4e-06

i

```
prediction = intercept + slope*4.8
print(prediction)
## [1] 28005.78
```

# j

Let  $x_0 := x_1^{new}$ 

$$\widehat{\mathbf{E}(y|x_0)} = \hat{\mu}_{u|x_0} = \hat{\beta}_0 + \hat{\beta}_1 x_0 \tag{19}$$

(20)

The variance on a prediction is the variance of its estimator which is 19, so we compute as follows:

$$\operatorname{Var}(\hat{\mu}_{y|x_0}) = \operatorname{Var}(\hat{\beta}_0 + \hat{\beta}_1 x_0) \tag{21}$$

$$= \operatorname{Var}[\bar{y} + \hat{\beta}_1(x_0 - \bar{x})] \tag{22}$$

$$= \frac{\sigma^2}{n} + \frac{\sigma^2 (x_0 - \bar{x})^2}{S_{xx}}$$
 (23)

$$= \sigma^2 \left[ \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}} \right]$$
 (24)

Now to estimate this value we use:

$$\hat{\sigma}^2 = MS_{Res}$$

```
x1new = 4800/1000
MSRes*(1/n1 + (x1new - xbar)^2/Sxx)
## [1] 13.78083
```