

MODEL SELECTION: EXAMPLES

Given a collection of predictors, our goal is now to find

the simplest possible model that adequately explains the data

- the principle of *parsimony*.

We typically consider a ‘largest’ (most complex) model and look for simplifications of it. We may use partial F tests to do this.

Typical strategies are

- Fit all possible models using the available predictors.
- *Forward selection*: start from the intercept only model, and try to add predictors until the model cannot be improved further.
- *Backward selection*: start from the most complex model, and try to remove predictors until the model cannot be simplified further without compromising the quality of fit.
- *Stepwise selection*: use a combination of forward and backward steps.

In R, several functions are useful

- `drop1`: single term deletions;
- `add1`: adds terms to an existing model;
- `update`: adds terms to or deletes terms from an existing model;
- `step`: automated procedure for selection;
- in the `leaps` library: `regsubsets` for subset selection.

Stepwise model selection: Example

The following data set contains information on the rut depth wear on thirty one road surfaces (measured in mm per million wheel passes) prepared under various experimental conditions describing the construction method used for each “course”. These data are recorded in a data frame `p9.10` in the package `MPV` which contains the following columns:

- y : change in rut depth/million wheel passes (log scale)
- x_1 : viscosity (log scale);
- x_2 : percentage of asphalt in surface course;
- x_3 : percentage of asphalt in base course;
- x_4 : indicator recording which run (Run 1, Run2) each datum is recorded for;
- x_5 : percentage of fines in surface course;
- x_6 : percentage of voids in surface course.

Stepwise model selection: Example (cont.)

We aim to identify which of the predictors influence the response.

```
1 > library(MPV)
2 > Asphalt<-p9.10
3 > Asphalt$x4<-factor(Asphalt$x4,labels=c('Run1','Run2'))
4 > str(Asphalt)
5 'data.frame':   31 obs. of  7 variables:
6 $ y : num  0.829 1.114 1.169 1.1 0.916 ...
7 $ x1: num  0.447 0.146 0.146 0.519 0.23 ...
8 $ x2: num  4.68 5.19 4.82 4.85 4.86 5.16 4.82 4.86 4.78 5.16 ...
9 $ x3: num  4.87 4.5 4.73 4.76 4.95 4.45 5.05 4.7 4.84 4.76 ...
10 $ x4: Factor w/ 2 levels "Run1","Run2": 1 1 1 1 1 1 1 1 1 1 ...
11 $ x5: num  8.4 6.5 7.9 8.3 8.4 7.4 6.8 8.6 6.7 7.7 ...
12 $ x6: num  4.92 4.56 5.32 4.87 3.78 ...
```

We may assess the influence of each predictor in an informal way by inspecting their correlations with the response and with each other

Stepwise model selection: Example (cont.)

```
13 > round(cor(p9.10), 3)
14           y          x1          x2          x3          x4          x5          x6
15 y      1.000 -0.893  0.145  0.099 -0.912  0.257 -0.409
16 x1 -0.893  1.000 -0.126 -0.217  0.939 -0.303  0.469
17 x2  0.145 -0.126  1.000 -0.187 -0.116 -0.229 -0.368
18 x3  0.099 -0.217 -0.187  1.000 -0.120  0.437 -0.027
19 x4 -0.912  0.939 -0.116 -0.120  1.000 -0.234  0.406
20 x5  0.257 -0.303 -0.229  0.437 -0.234  1.000  0.116
21 x6 -0.409  0.469 -0.368 -0.027  0.406  0.116  1.000
```

From this we note that x_1 and x_4 are highly correlated with the response, but also with each other. There appear to be some other moderately large correlations.

Plots of the data reveal more of the structure: we plot

- Figure 1: scatterplots of y versus each predictor;
- Figure 2: boxplots of the continuous predictor versus x_4 .

We see that there is interesting structure amongst the x_j s.

Scatterplot

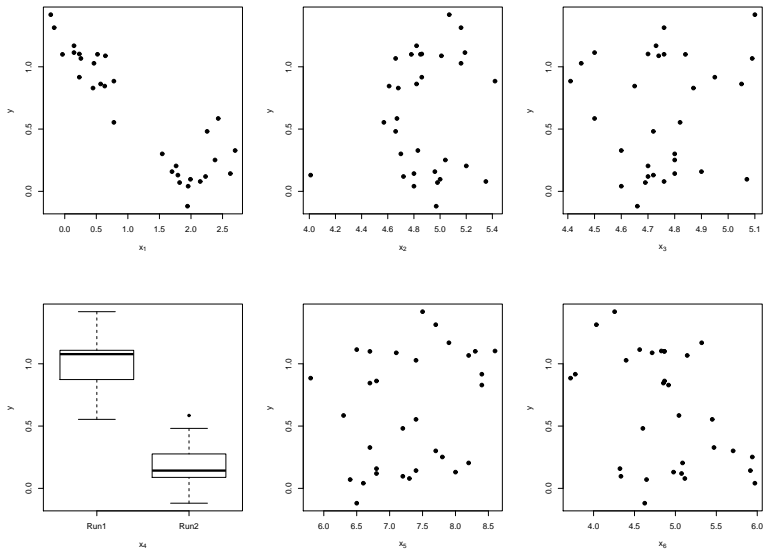


Figure: Scatterplots of y vs $x_j, j = 1, \dots, 6$

Box plot

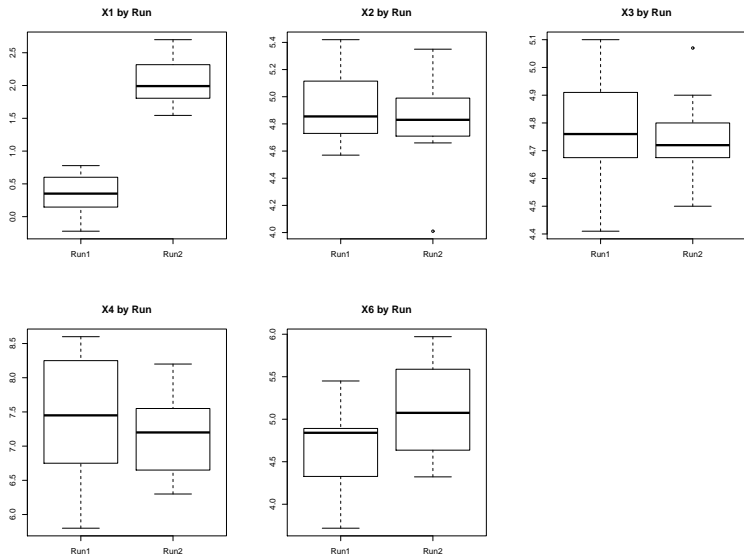


Figure: Boxplots of $x_j, j = 1, 2, 3, 5, 6$ vs x_4

We begin the best model identification by examining the additive model that fits all predictors as main effects:

```
22 > fit1<-lm(y~x1+x2+x3+x4+x5+x6,data=Asphalt)
23 > summary(fit1)
24 Coefficients:
25             Estimate Std. Error t value Pr(>|t|)
26 (Intercept)  1.31975    1.53757   0.858   0.3992
27 x1          -0.14658    0.12872  -1.139   0.2661
28 x2           0.05625    0.15180   0.371   0.7142
29 x3          -0.15483    0.24140  -0.641   0.5274
30 x4Run2      -0.56213    0.22176  -2.535   0.0182 *
31 x5           0.03681    0.06156   0.598   0.5555
32 x6          -0.01220    0.08053  -0.151   0.8809
33 ---
34 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
35
36 Residual standard error: 0.2021 on 24 degrees of freedom
37 Multiple R-squared:  0.8482,    Adjusted R-squared:  0.8102
38 F-statistic: 22.35 on 6 and 24 DF,  p-value: 1.009e-08
```

Model selection (cont.)

This model achieves a reasonably high R^2 , but only one predictor appears to be significantly influential; the factor predictor x_4 appears to yield a lower mean response for level Run2 than for the baseline Run1.

We now attempt stepwise elimination starting from the main effects model: we focus first on F -test comparisons:

```
39 > drop1(fit1,test='F')
40 Single term deletions
41 Model:
42 y ~ x1 + x2 + x3 + x4 + x5 + x6
43      Df Sum of Sq    RSS    AIC F value  Pr(>F)
44 <none>            0.98005 -93.078
45 x1      1  0.052949 1.03300 -93.447  1.2966 0.26607
46 x2      1  0.005607 0.98566 -94.901  0.1373 0.71422
47 x3      1  0.016798 0.99685 -94.551  0.4114 0.52736
48 x4      1  0.262391 1.24244 -87.724  6.4256 0.01818 *
49 x5      1  0.014602 0.99465 -94.620  0.3576 0.55546
50 x6      1  0.000937 0.98099 -95.049  0.0229 0.88086
```

Model selection (cont.)

This suggests that we could potentially drop any predictor apart from x_4 . Standard elimination methods would eliminate one predictor at a time, starting with the apparently least significant.

Here we proceed in a more ambitious fashion, and attempt to drop x_2 , x_3 , x_5 and x_6 simultaneously: the `update` function is used to define the new model for fitting.

```
51 > fit2<-update(fit1, ~.-x2-x3-x5-x6)
52 > anova(fit2,fit1,test='F')
53 Analysis of Variance Table
54
55 Model 1: y ~ x1 + x4
56 Model 2: y ~ x1 + x2 + x3 + x4 + x5 + x6
57   Res.Df    RSS Df Sum of Sq    F Pr(>F)
58 1      28 1.01123
59 2      24 0.98005  4  0.031181 0.1909 0.9408
```

Model selection (cont.)

Line 59 reveals that the model $X_1 + X_2$ is an adequate simplification of the more complex model $X_1 + X_2 + X_3 + X_4 + X_5 + X_6$ (the p -value is 0.9408, indicating no significant loss in fit). The model summary reveals that the R^2 quantity is still quite high (line 70).

```
60 > summary(fit2)
61 Coefficients:
62             Estimate Std. Error t value Pr(>|t|)
63 (Intercept)  1.07490    0.05968  18.011  < 2e-16 ***
64 x1          -0.14885    0.10677  -1.394  0.17424
65 x4Run2      -0.57317    0.19860  -2.886  0.00743 **
66 ---
67 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
68
69 Residual standard error: 0.19 on 28 degrees of freedom
70 Multiple R-squared:  0.8434,    Adjusted R-squared:  0.8322
71 F-statistic: 75.37 on 2 and 28 DF,  p-value: 5.356e-12
```

Model selection (cont.)

If we want to simplify further we can check using the `drop1` function:

```
72 > drop1(fit2,test='F')
73 Single term deletions
74
75 Model:
76 y ~ x1 + x4
77      Df Sum of Sq    RSS      AIC F value    Pr(>F)
78 <none>          1.0112 -100.107
79 x1      1  0.070194  1.0814 -100.027   1.9436 0.174245
80 x4      1  0.300800  1.3120  -94.035   8.3288 0.007433 **
```

Model selection (cont.)

The indication is that x_1 can also be dropped.

```
81 > fit3<-update(fit2, ~.-x1); summary(fit3)
82           Estimate Std. Error t value Pr(>|t|)
83 (Intercept)  1.02455    0.04828   21.22 < 2e-16 ***
84 x4Run2       -0.83316    0.06940  -12.01 8.97e-13 ***
85 ---
86 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
87 Residual standard error: 0.1931 on 29 degrees of freedom
88 Multiple R-squared:  0.8325,    Adjusted R-squared:  0.8267
89 F-statistic: 144.1 on 1 and 29 DF,  p-value: 8.974e-13
```

However, we now try to add interaction terms: a natural place to start is the interaction $X_1 : X_4$, so we attempt to fit the model

$$X_1 + X_4 + X_1 : X_4$$

recalling the convention on interactions ('no interactions without main effects').

Model selection (cont.)

```
90 > fit4<-update(fit3,~.+x1+x1:x4)
91 > anova(fit3,fit4,test='F')
92 Analysis of Variance Table
93
94 Model 1: y ~ x4
95 Model 2: y ~ x4 + x1 + x4:x1
96      Res.Df      RSS Df Sum of Sq      F      Pr(>F)
97 1         29 1.08143
98 2         27 0.60764  2    0.47378 10.526 0.0004172 ***
99 ---
100 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
101
102 > summary(fit4)
103 Coefficients:
104             Estimate Std. Error t value Pr(>|t|)
105 (Intercept)  1.20274    0.05595  21.495 < 2e-16 ***
106 x4Run2       -1.40384    0.25111  -5.591 6.26e-06 ***
107 x1           -0.52676    0.12275  -4.291 0.000204 ***
108 x4Run2:x1     0.71501    0.16884   4.235 0.000237 ***
109 ---
110 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
111
112 Residual standard error: 0.15 on 27 degrees of freedom
```

Model selection (cont.)

```
113 Multiple R-squared:  0.9059,    Adjusted R-squared:  0.8954  
114 F-statistic: 86.62 on 3 and 27 DF,  p-value: 5.684e-14
```

Line 98 notes that the model X_4 is not an adequate simplification of the more complex model $X_1 + X_4 + X_1 : X_4$ as the null hypothesis that removes the X_1 main effect and interaction is rejected (p -value equal to 0.0004172).

Model selection (cont.)

Using the `drop1` function we verify that the interaction term should not be dropped.

```
115 > drop1(fit4,test='F')
116 Single term deletions
117
118 Model:
119 y ~ x4 + x1 + x4:x1
120      Df Sum of Sq      RSS      AIC F value    Pr(>F)
121 <none>            0.60764 -113.90
122 x4:x1      1      0.40359 1.01123 -100.11   17.933 0.0002372 ***
123 ---
124 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The success of the interaction model encourages us to try other interactions: to fit all second order interactions involving `x4` and the continuous predictors, we may type

```
125 > fit5 <- lm(y ~ x4 * (x1 + x2 + x3 + x5 + x6), data = Asphalt)
```

Model selection (cont.)

This yields the following model fit:

```
126 > summary(fit5)
127 Coefficients:
128             Estimate Std. Error t value Pr(>|t|)
129 (Intercept) -3.55427     3.12640  -1.137  0.26973
130 x4Run2       5.42570     3.59232   1.510  0.14740
131 x1          -0.43526     0.15342  -2.837  0.01054 *
132 x2           0.64022     0.33846   1.892  0.07389 .
133 x3           0.13316     0.26854   0.496  0.62567
134 x5           0.04104     0.05870   0.699  0.49299
135 x6           0.13774     0.12476   1.104  0.28335
136 x4Run2:x1    0.65813     0.20397   3.227  0.00444 **
137 x4Run2:x2   -0.74479     0.36646  -2.032  0.05633 .
138 x4Run2:x3   -0.53269     0.46713  -1.140  0.26832
139 x4Run2:x5    0.04697     0.10422   0.451  0.65730
140 x4Run2:x6   -0.21067     0.15399  -1.368  0.18725
141 ---
142 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
143
144 Residual standard error: 0.1521 on 19 degrees of freedom
145 Multiple R-squared:  0.9319,    Adjusted R-squared:  0.8925
146 F-statistic: 23.65 on 11 and 19 DF,  p-value: 9.046e-09
```

Model selection (cont.)

We now try to drop the apparently unimportant terms:

```
147 > fit6<-update(fit5, ~.-x3-x5-x6-x3:x4-x5:x4-x6:x4)
148 > anova(fit6,fit5,test='F')
149 Analysis of Variance Table
150
151 Model 1: y ~ x4 + x1 + x2 + x4:x1 + x4:x2
152 Model 2: y ~ x4 * (x1 + x2 + x3 + x5 + x6)
153      Res.Df      RSS Df Sum of Sq      F Pr(>F)
154 1         25 0.51104
155 2         19 0.43946  6    0.07158 0.5158 0.7891
```

From this analysis, we retain the model

$$X_1 + X_2 + X_4 + X_1 : X_4 + X_2 : X_4$$

Using drop1 we may check for further simplification

```
156 > drop1(fit6,test='F')
157 Single term deletions
158
159 Model:
160 y ~ x4 + x1 + x2 + x4:x1 + x4:x2
161      Df Sum of Sq      RSS      AIC F value    Pr(>F)
162 <none>            0.51104 -115.26
163 x4:x1      1      0.36505  0.87609 -100.56  17.8584 0.0002768 ***
164 x4:x2      1      0.09572  0.60675 -111.94   4.6825 0.0402258 *
165 ---
166 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Neither of the interactions $X_1 : X_4$ or $X_2 : X_4$ can be dropped on the basis of the two F -test results.

Model selection (cont.)

For a further comparison, we compare the simpler

$$X_1 + X_4 + X_1 : X_4$$

with the more complex model

$$X_1 + X_2 + X_4 + X_1 : X_4 + X_2 : X_4$$

```
167 > anova(fit4,fit6,test='F')
168 Analysis of Variance Table
169
170 Model 1: y ~ x4 + x1 + x4:x1
171 Model 2: y ~ x4 + x1 + x2 + x4:x1 + x4:x2
172    Res.Df    RSS Df Sum of Sq    F Pr(>F)
173 1       27 0.60764
174 2       25 0.51104  2   0.096608 2.363 0.1148
```

The model seems to perform adequately, giving a high R^2 value (line 189).

Model selection (cont.)

```
175 > summary(fit6)
176 Coefficients:
177             Estimate Std. Error t value Pr(>|t|)
178 (Intercept)  -0.1990     0.7739  -0.257 0.799154
179 x4Run2        0.7180     1.0080   0.712 0.482847
180 x1          -0.4924     0.1185  -4.155 0.000332 ***
181 x2           0.2833     0.1560   1.816 0.081450 .
182 x4Run2:x1     0.6849     0.1621   4.226 0.000277 ***
183 x4Run2:x2    -0.4337     0.2004  -2.164 0.040226 *
184 ---
185 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
186
187 Residual standard error: 0.143 on 25 degrees of freedom
188 Multiple R-squared:  0.9208,    Adjusted R-squared:  0.905
189 F-statistic: 58.16 on 5 and 25 DF,  p-value: 5.828e-13
```


Model selection (cont.)

For a final analysis, we attempt to add in a three-way interaction.

```
190 > fit7<-update(fit6, ~.+x1:x2:x4)
191 > drop1(fit7,test='F')
192 Single term deletions
193
194 Model:
195 y ~ x4 + x1 + x2 + x4:x1 + x4:x2 + x4:x1:x2
196           Df Sum of Sq      RSS      AIC F value Pr(>F)
197 <none>                 0.46666 -114.08
198 x4:x1:x2    2   0.044379 0.51104 -115.26   1.0936 0.3518
```

However, there is no evidence that the three-way interaction is useful in the model. Hence a reasonable model seems to be

$$X_1 + X_2 + X_4 + X_1 : X_4 + X_2 : X_4 = (X_1 + X_2) * X_4$$

which uses a six parameter conditional mean.

The model fits a different two-dimensional plane (in variables (x_1 , x_2)) through the data for each of the levels of the Run factor predictor (x_4).

Model selection (cont.)

The model can be written

$$\mathbb{E}_{Y_i|X_i}[Y_i|\mathbf{x}_i] = \begin{cases} \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} & x_4 = \text{Run1} \\ (\beta_0 + \beta_{41}^C) + (\beta_1 + \beta_{141}^C)x_{i1} + (\beta_2 + \beta_{241}^C)x_{i2} & x_4 = \text{Run2} \end{cases}$$

where

$\hat{\beta}_0$	Line 178	(Intercept)	-0.1990
$\hat{\beta}_{41}^C$	Line 179	x4Run2	0.7180
$\hat{\beta}_1$	Line 180	x1	-0.4924
$\hat{\beta}_2$	Line 181	x2	0.2833
$\hat{\beta}_{141}^C$	Line 182	x4Run2 : x1	0.6849
$\hat{\beta}_{241}^C$	Line 183	x4Run2 : x2	-0.4337

Residual plots

Residual plots seem reasonable, although there may be some mild evidence that the variance is not constant.

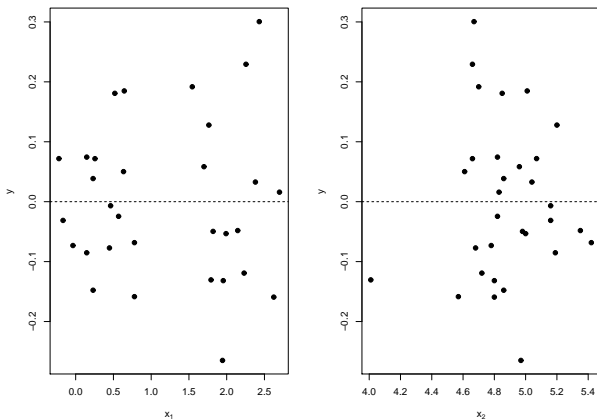


Figure: Residuals vs X_1 (left) and X_2 (right)

We have the model for three continuous predictors X_1, X_2, X_3

$$Y_i = 2 + 2x_{i1} + 2x_{i2} - 2x_{i1}x_{i2} + \epsilon_i$$

with $\sigma^2 = 1$. We have $n = 200$.

Other Criteria: Example (cont.)

```
1  > fit0<-lm(Y~1)
2  > fit1<-lm(Y~x1)
3  > fit2<-lm(Y~x2)
4  > fit3<-lm(Y~x3)
5  > fit12<-lm(Y~x1+x2)
6  > fit13<-lm(Y~x1+x3)
7  > fit23<-lm(Y~x2+x3)
8  > fit123<-lm(Y~x1+x2+x3)
9  > fit12i<-lm(Y~x1*x2)
10 > fit13i<-lm(Y~x1*x3)
11 > fit23i<-lm(Y~x2*x3)
12 > fit123i<-lm(Y~x1*x2*x3)
13
14 > bigs.hat<-summary(fit123i)$sigma
```

Other Criteria: Example (cont.)

```
15  criteria.eval<-function(fit.obj,nv,bigsig.hat){
16      cvec<-rep(0,5)
17      SSRes<-sum(residuals(fit.obj)^2)
18      p<-length(coef(fit.obj))
19
20      #R squared
21      cvec[1]<-summary(fit.obj)$r.squared
22      #Adjusted R squared
23      cvec[2]<-summary(fit.obj)$adj.r.squared
24      #Cp
25      cvec[3]<-SSRes/bigsig.hat^2-nv+2*p
26      #AIC in R computes
27      #n*log(sum(residuals(fit.obj)^2)/n)+
28      #2*(length(coef(fit.obj))+1)+n*log(2*pi)+n
29      cvec[4]<-AIC(fit.obj)
30      #BIC in R computes
31      #n*log(sum(residuals(fit.obj)^2)/n)+
32      #log(n)*(length(coef(fit.obj))+1)+n*log(2*pi)+n
33      cvec[5]<-BIC(fit.obj)
34
35      return(cvec)
36  }
```

Other Criteria: Example (cont.)

```
37 > cvals<-matrix(0,nrow=12,ncol=5)
38 > cvals[1,]<-criteria.eval(fit0,n,bigs.hat)
39 > cvals[2,]<-criteria.eval(fit1,n,bigs.hat)
40 > cvals[3,]<-criteria.eval(fit2,n,bigs.hat)
41 > cvals[4,]<-criteria.eval(fit3,n,bigs.hat)
42 > cvals[5,]<-criteria.eval(fit12,n,bigs.hat)
43 > cvals[6,]<-criteria.eval(fit13,n,bigs.hat)
44 > cvals[7,]<-criteria.eval(fit23,n,bigs.hat)
45 > cvals[8,]<-criteria.eval(fit123,n,bigs.hat)
46 > cvals[9,]<-criteria.eval(fit12i,n,bigs.hat)
47 > cvals[10,]<-criteria.eval(fit13i,n,bigs.hat)
48 > cvals[11,]<-criteria.eval(fit23i,n,bigs.hat)
49 > cvals[12,]<-criteria.eval(fit123i,n,bigs.hat)
50 >
51 > Criteria<-data.frame(cvals)
52 > names(Criteria)<-c('Rsqr','Adj.Rsqr','Cp','AIC','BIC')
53 >
54 > rownames(Criteria)<-c('1','x1','x2','x3','x1+x2','x1+x3','x2+x3',
55 + 'x1+x2+x3','x1*x2','x1*x3','x2*x3','x1*x2*x3')
```


Other Criteria: Example (cont.)

```
56 > round(Criteria,4)
57           Rsq Adj.Rsq           Cp           AIC           BIC
58 1           0.0000  0.0000 1808.7895 1016.2116 1022.8082
59 x1          0.4426  0.4398  922.5695  901.3145  911.2094
60 x2          0.0207  0.0158 1769.2514 1014.0284 1023.9233
61 x3          0.0056  0.0006 1799.5465 1017.0879 1026.9829
62 x1+x2       0.5266  0.5218  755.9216  870.6292  883.8224
63 x1+x3       0.4827  0.4775  844.0514  888.3734  901.5667
64 x2+x3       0.0237  0.0138 1765.1450 1015.4060 1028.5992
65 x1+x2+x3    0.5268  0.5195  757.6328  872.5684  889.0599
66 x1*x2       0.9041  0.9026   0.4429  553.3132  569.8048
67 x1*x3       0.7600  0.7564  289.5338  736.7485  753.2401
68 x2*x3       0.1734  0.1608 1466.7742  984.1201 1000.6117
69 x1*x2*x3    0.9043  0.9008   8.0000  560.8523  590.5372
```

Line 67 reveals the model $X_1 * X_2 = X_1 + X_2 + X_1 : X_2$ as most appropriate model.

Textbook Example: step

```
70 > fit5<-lm(y~x4*(x1+x2+x3+x5+x6),data=Asphalt) #AIC version
71 > fit.by.step.aic<-step(fit5,k=2)
72 Start: AIC=-107.94
73 y ~ x4 * (x1 + x2 + x3 + x5 + x6)
74
75           Df Sum of Sq      RSS      AIC
76 - x4:x5    1  0.004698  0.44415 -109.613
77 <none>                                0.43946 -107.942
78 - x4:x3    1  0.030077  0.46953 -107.890
79 - x4:x6    1  0.043290  0.48275 -107.030
80 - x4:x2    1  0.095536  0.53499 -103.844
81 - x4:x1    1  0.240794  0.68025  -96.398
82
83 Step: AIC=-109.61
84 y ~ x4 + x1 + x2 + x3 + x5 + x6 + x4:x1 + x4:x2 + x4:x3 + x4:x6
85
86           Df Sum of Sq      RSS      AIC
87 - x4:x3    1  0.025434  0.46959 -109.886
88 <none>                                0.44415 -109.613
89 - x5       1  0.030762  0.47492 -109.537
90 - x4:x6    1  0.039507  0.48366 -108.971
91 - x4:x2    1  0.112030  0.55618 -104.640
92 - x4:x1    1  0.241310  0.68546  -98.161
93
94 Step: AIC=-109.89
95 y ~ x4 + x1 + x2 + x3 + x5 + x6 + x4:x1 + x4:x2 + x4:x6
96
97           Df Sum of Sq      RSS      AIC
98 - x3       1  0.00083  0.47041 -111.83
99 - x5       1  0.02093  0.49052 -110.53
100 - x4:x6    1  0.02586  0.49545 -110.22
```

Textbook Example: step (cont.)

```
101 <none> 0.46959 -109.89
102 - x4:x2 1 0.08960 0.55919 -106.47
103 - x4:x1 1 0.34502 0.81461 -94.81
104
105 Step: AIC=-111.83
106 y ~ x4 + x1 + x2 + x5 + x6 + x4:x1 + x4:x2 + x4:x6
107
108      Df Sum of Sq      RSS      AIC
109 - x5    1  0.02012 0.49054 -112.533
110 - x4:x6 1  0.02656 0.49698 -112.129
111 <none>      0.47041 -111.832
112 - x4:x2 1  0.12486 0.59528 -106.534
113 - x4:x1 1  0.35250 0.82291 -96.496
114
115 Step: AIC=-112.53
116 y ~ x4 + x1 + x2 + x6 + x4:x1 + x4:x2 + x4:x6
117
118      Df Sum of Sq      RSS      AIC
119 - x4:x6 1  0.01826 0.50880 -113.400
120 <none>      0.49054 -112.533
121 - x4:x2 1  0.10585 0.59639 -108.476
122 - x4:x1 1  0.36400 0.85454 -97.326
123
124 Step: AIC=-113.4
125 y ~ x4 + x1 + x2 + x6 + x4:x1 + x4:x2
126
127      Df Sum of Sq      RSS      AIC
128 - x6    1  0.00224 0.51104 -115.264
129 <none>      0.50880 -113.400
130 - x4:x2 1  0.09138 0.60018 -110.280
131 - x4:x1 1  0.34584 0.85464 -99.323
```

Textbook Example: step (cont.)

```
132
133 Step: AIC=-115.26
134 y ~ x4 + x1 + x2 + x4:x1 + x4:x2
135
136           Df Sum of Sq      RSS      AIC
137 <none>                0.51104 -115.26
138 - x4:x2  1      0.09572 0.60675 -111.94
139 - x4:x1  1      0.36505 0.87609 -100.56
140
141 > summary(fit.by.step.aic)
142 Coefficients:
143             Estimate Std. Error t value Pr(>|t|)
144 (Intercept)  -0.1990     0.7739  -0.257  0.799154
145 x4Run2        0.7180     1.0080   0.712  0.482847
146 x1           -0.4924     0.1185  -4.155  0.000332 ***
147 x2            0.2833     0.1560   1.816  0.081450 .
148 x4Run2:x1     0.6849     0.1621   4.226  0.000277 ***
149 x4Run2:x2    -0.4337     0.2004  -2.164  0.040226 *
150 ---
151 Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
152
153 Residual standard error: 0.143 on 25 degrees of freedom
154 Multiple R-squared:  0.9208,    Adjusted R-squared:  0.905
155 F-statistic: 58.16 on 5 and 25 DF,  p-value: 5.828e-13
```

Textbook Example: regsubsets

```
156 > library(leaps)
157 > fit.by.leaps<-regsubsets(y~x4*(x1+x2+x3+x5+x6),data=Asphalt,method='exhaustive')
158 > summary(fit.by.leaps)
159 Subset selection object
160 Call: regsubsets.formula(y ~ x4 * (x1 + x2 + x3 + x5 + x6), data = Asphalt,
161     method = "exhaustive")
162 11 Variables (and intercept)
163      Forced in Forced out
164 x4Run2      FALSE      FALSE
165 x1          FALSE      FALSE
166 x2          FALSE      FALSE
167 x3          FALSE      FALSE
168 x5          FALSE      FALSE
169 x6          FALSE      FALSE
170 x4Run2:x1   FALSE      FALSE
171 x4Run2:x2   FALSE      FALSE
172 x4Run2:x3   FALSE      FALSE
173 x4Run2:x5   FALSE      FALSE
174 x4Run2:x6   FALSE      FALSE
175 1 subsets of each size up to 8
176 Selection Algorithm: exhaustive
177      x4Run2 x1  x2  x3  x5  x6  x4Run2:x1 x4Run2:x2 x4Run2:x3 x4Run2:x5 x4Run2:x6
178 1  ( 1 ) " "   " " " " " " " " " " " " " " " " " " " " " " " " " " " "
179 2  ( 1 ) " "   "*" " " " " " " " " " " " " " " " " " " " " " " " "
180 3  ( 1 ) " "   "*" " " " " " " " " " " " " " " " " " " " " " " " "
181 4  ( 1 ) " "   "*" "*" " " " " " " " " " " " " " " " " " " " " " "
182 5  ( 1 ) " "   "*" "*" " " " " " " " " " " " " " " " " " " " " " "
183 6  ( 1 ) "*"   "*" "*" " " " " " " " " " " " " " " " " " " " " " "
184 7  ( 1 ) "*"   "*" "*" " " " " " " " " " " " " " " " " " " " " " "
185 8  ( 1 ) "*"   "*" "*" " " " " " " " " " " " " " " " " " " " " " "
```

Textbook Example: regsubsets

```
186 > summary(fit.by.leaps)$which
187 (Intercept) x4Run2 x1 x2 x3 x5 x6 x4Run2:x1 x4Run2:x2 x4Run2:x3
188 1 TRUE FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE TRUE
189 2 TRUE FALSE TRUE FALSE FALSE FALSE FALSE FALSE FALSE TRUE
190 3 TRUE FALSE TRUE FALSE FALSE FALSE FALSE TRUE FALSE TRUE
191 4 TRUE FALSE TRUE TRUE FALSE FALSE FALSE TRUE TRUE FALSE
192 5 TRUE FALSE TRUE TRUE FALSE FALSE FALSE TRUE TRUE FALSE
193 6 TRUE TRUE TRUE TRUE FALSE TRUE FALSE TRUE TRUE FALSE
194 7 TRUE TRUE TRUE TRUE FALSE TRUE FALSE TRUE TRUE TRUE
195 8 TRUE TRUE TRUE TRUE FALSE TRUE TRUE TRUE TRUE TRUE FALSE
196 x4Run2:x5 x4Run2:x6
197 1 FALSE FALSE
198 2 FALSE FALSE
199 3 FALSE FALSE
200 4 FALSE FALSE
201 5 TRUE FALSE
202 6 FALSE FALSE
203 7 FALSE FALSE
204 8 FALSE TRUE
205 > summary(fit.by.leaps)$rsq
206 [1] 0.8364571 0.8474287 0.9081912 0.9192309 0.9211305 0.9228623 0.9249253 0.9271302
207 > summary(fit.by.leaps)$rss
208 [1] 1.0557586 0.9849311 0.5926761 0.5214089 0.5091459 0.4979660 0.4846483 0.4704142
209 > summary(fit.by.leaps)$bic
210 [1] -49.26311 -47.98186 -60.29351 -60.83104 -58.13485 -55.38915 -52.79552 -50.28564
```