STA 304H1F-1003H Fall 2019

Week 11 - Two-Stage Cluster Sampling $_{Chapter\ 9}$

How to Draw a Cluster Sample

1. Simple one-stage cluster sample:

List all the clusters in the population, and

from the list, select the clusters – usually with simple random sampling (SRS) strategy.

All units (elements) in the sampled clusters are selected for the survey.

2. Simple two-stage cluster sample:

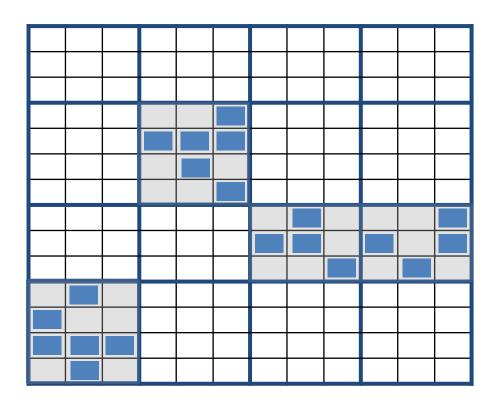
List all the clusters in the population, and

First, select the clusters, usually by simple random sampling (SRS).

The units (elements) in the selected clusters of the firststage are then sampled in the second-stage, usually by simple random sampling (or often by systematic sampling).

Remainder on basics (I)

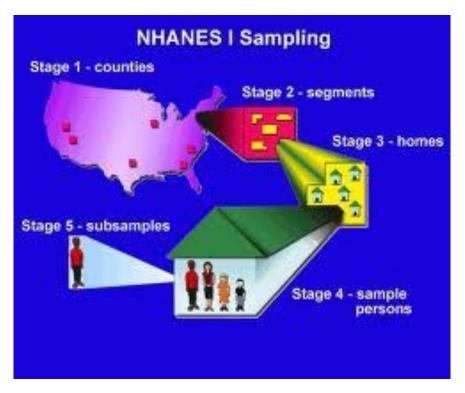
Basics: The population is divided into large number of (bigger) groups (clusters). **First stage**: sample of clusters selected. **Second stage**: sample of elements from each selected cluster. Sample: all selected elements.



Population size: 168.
8 clusters of size 9,
8 clusters of size 12.
Sample size: 20.
4 clusters selected
"at random" from
16 clusters.
6+6+4+4 elements
selected.

Remainder on basics (II)

The population is divided into clusters on several levels/stages - primary sampling units (PSUs), secondary sampling units, ternary sampling units, ... Sampling is performed at every stage. Sample: all sampling units selected at the last stage.



National Health and Nutrition Examination Survey (USA)

Five-stage cluster sampling

Stage 1: 1900 PSUs – cities, counties

Stage 2: Segments – each with 18

housing units/addresses

Stage 3: Households

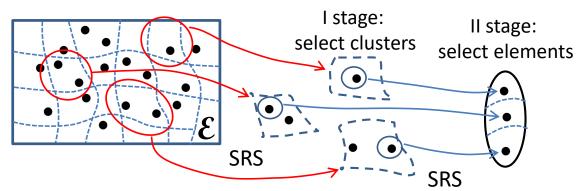
Stage 4: Family members for basic

examination (~ 28,000, 1-74 years old)

Stage 5: Subsample for detailed health

examination (25-74 years old)

General considerations, definition (I)



Simplest case: SRS of sampling units at every stage (PSU, SSU, TSU, ...)

Large number of (bigger or smaller) groups

- First stage: SRS of n groups primary sampling units (PSU)
- Second stage: SRS of elements secondary sampling units
 (SSU) from every PSU
- Sampling units: Clusters at every stage
- Design: Two-stage cluster sampling. All elements selected at the second stage are in the sample

General considerations, notation (II)

- N number of clusters in the population (# of sampling units)
- n number of clusters selected in the sample at stage 1
- $M_i i$ th clusters size

 m_i -samplesizeselected from ith clusters, $m_i \leq M_i$ | In one-stage $m_i = M_i$

In one-stage
$$m_i = M_i$$

$$M$$
 – population size, $M = M_1 + M_2 + \cdots + M_N = \sum_{i=1}^N M_i$

$$\overline{M}$$
 – average cluster size, $\overline{M} = \frac{M}{N} = \frac{1}{N} \sum_{i=1}^{N} M_i$

*i*th cluster elements: $y_{i1}, y_{i1}, ..., y_{iM_i}$

$$au_i = \sum_{j=1}^{M_i} y_{ij}$$
 -cluster total, $\mu_i = \frac{ au_i}{M_i}$ - cluster mean, $au_i = M_i \mu_i$

General considerations, notation (III) Summary, population:

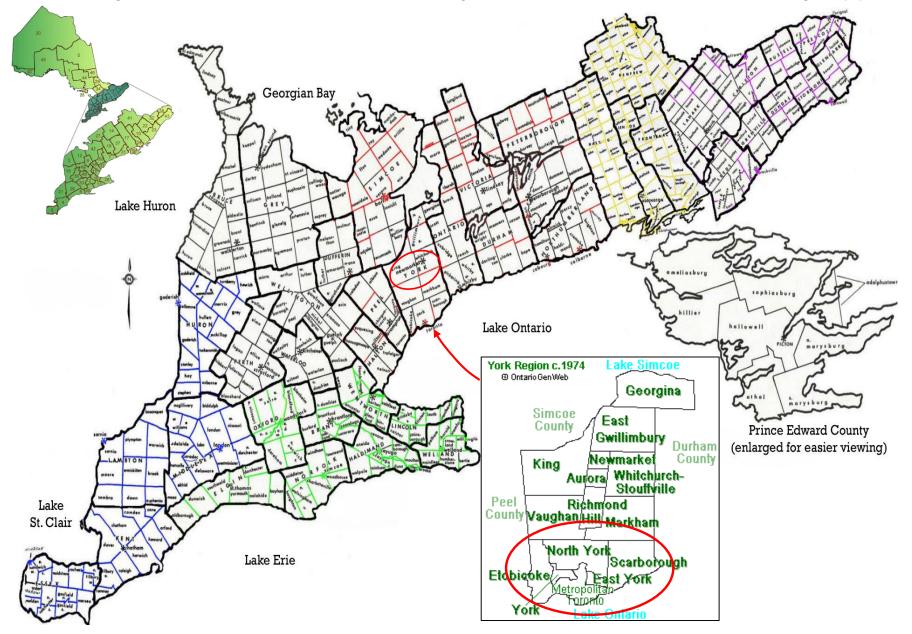
Cluster	1	2	• • •	N	Pop
Size	M_{1}	M_2	• • •	$oldsymbol{M}_N$	$\mid M \mid$
Total	$ au_1$	$ au_2$	• • •	${ au}_N$	$\mid au \mid$
Mean	μ_1	μ_2	• • •	$\mu_{\scriptscriptstyle N}$	μ

$$au = au_y = M \mu_y$$
 , $\mu = \mu_y = rac{ au_y}{M}$

$$\tau_{y} = \sum_{i=1}^{N} \tau_{i} = \sum_{i=1}^{N} M_{i} \mu_{i} = N \frac{1}{N} \sum_{i=1}^{N} \tau_{i} = N \mu_{t} = N \bar{\tau}$$

$$\mu_t = \overline{\tau} = \frac{1}{N} \sum_{i=1}^N \tau_i$$
 - average cluster total

Example: Southern Ontario, map of counties and townships (I)



Example: Multi-stage cluster sample from Southern Ontario (VI)

First stage sampling units: Counties Two strata: Metropolitan

Second stage sampling units: Townships Toronto, York region (why?)

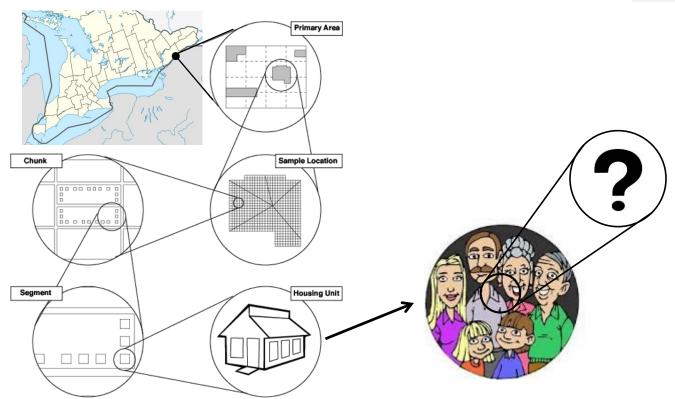
Third stage sampling units: Communities

Fourth stage sampling units: Neighbourhoods/wards

Fifth stage sampling units: ? Households/mailing addresses

Sixth stage sampling units: ? Persons





Example: Multi-stage cluster sample from Southern Ontario (VII)

Selecting Respondents within Households using telephone surveys:

- (a) The respondent (an adult) is asked to list all the eligible members of the household by gender and age; they are then ordered by age, first males and then females, and then one person is selected at random using preselected random number an unbiased but not very convenient procedure, creating higher refusal rate
- **(b)** A less invasive procedure: The respondent is asked two questions:
 - 1. How many persons 18 years or older live in your household, counting yourself?
 - 2. How many of them are men (women)?

One of four (or seven) selection matrices is randomly assigned in advance to each sampled telephone number and used depending on answers to Q. 1 and 2 - e.g., to speak with the "oldest" or "youngest" man or women

(c) The simplest: The respondent is asked to speak to the eligible member who had the "last" (or will have "the next") birthday





An interviewer

General considerations, sample (IV)

I stage: SRS of *n* clusters

II stage : SRS of m_i elements selected from cluster i selected in stage I

Sample from cluster $i: y_{i1}, y_{i1}, ..., y_{im_i}, i = 1, 2, ..., n$

$$y_i = \sum_{j=1}^{m_i} y_{ij} - i$$
th cluster sample total ($\neq \tau_i$)

$$\hat{\mu}_i = \overline{y}_i = \frac{y_i}{m_i}$$
 - *i*th cluster sample mean

Unbiased estimator

$$S_i^2 = \frac{1}{m_i - 1} \sum_{j=1}^{m_i} (y_{ij} - \overline{y}_i)^2$$
 - *i*th cluster sample variance

Inference: Unbiased estimation of mean and total (I)

First estimate τ_i , i = 1, 2, ..., n - totals of clusters in the sample

$$\hat{\mu}_i = \overline{y}_i \Longrightarrow \hat{\tau}_i = M_i \hat{\mu}_i = M_i \overline{y}_i \Longrightarrow \hat{\tau}_1, \hat{\tau}_2, ..., \hat{\tau}_n$$

$$\hat{\mu}_t = \hat{\overline{\tau}} = \frac{1}{n} \sum_{i=1}^n \hat{\tau}_i = \frac{1}{n} \sum_{i=1}^n M_i \overline{y}_i \implies \hat{\tau} = N \hat{\overline{\tau}} = \frac{N}{n} \sum_{i=1}^n M_i \overline{y}_i,$$

$$\hat{\mu} = \frac{\hat{\tau}}{M} = \frac{N}{M} \frac{1}{n} \sum_{i=1}^{n} \hat{\tau}_i = \frac{1}{\overline{M}} \frac{1}{n} \sum_{i=1}^{n} M_i \overline{y}_i,$$
 All unbiased

 $|\hat{V}ar(\hat{\mu}) = \frac{N-n}{N_{\bullet}} \frac{S_b^2}{n\overline{M}^2} + \frac{1}{nN\overline{M}^2} \sum_{i=1}^n M_i^2 \frac{M_i - m_i}{M_i} \frac{S_i^2}{m_i}|$

First stage component

Second stage component

$$S_b^2 = \frac{1}{n-1} \sum_{i=1}^n (M_i \bar{y}_i - \overline{M} \hat{\mu})^2 = \frac{1}{n-1} \sum_{i=1}^n (\hat{\tau}_i - \hat{\overline{\tau}})^2$$

Derivation of the formula for variance is not simple, see book.

Inference: Unbiased estimation of mean and total (II)

How to organize data: Calculation table

Cl		sam	ple		$\overline{\mathcal{Y}}_i$	S_i^2	$\hat{ au_i}$ =	$=M_i \overline{y}_i$
1	y_{11}	\mathcal{Y}_{12}	• • •	\mathcal{Y}_{1m_1}	$\overline{\mathcal{Y}}_1$	S_1^2	$\hat{\tau}_1$	$=M_1\overline{y}_1$
2	\mathcal{Y}_{21}	\mathcal{Y}_{22}	• • •	\mathcal{Y}_{2m_2}	$\overline{\mathcal{Y}}_2$	S_2^2	$\hat{ au}_2$	$=M_2\overline{y}_2$
•	•			•	•	•/	-	•
n	y_{n1}	\mathcal{Y}_{n2}	• • •	${\cal Y}_{nm_n}$	$\overline{\mathcal{Y}}_n$	S_n^2	$\hat{\tau}_n$	$=M_n \overline{y}_n$

$$\hat{\mu}_{t} = \hat{\overline{\tau}} = \frac{1}{n} \sum_{i=1}^{n} \hat{\tau}_{i} \implies \hat{\tau} = N \hat{\overline{\tau}}, \hat{\mu} = \frac{\hat{\tau}}{M}$$

$$S_{b}^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (\hat{\tau}_{i} - \hat{\overline{\tau}})^{2}$$

Inference: Ratio estimation of mean and total (I)

We can use ratio estimation in stead of unbiased.

We have to use it to estimate μ if M is not known.

$$\overline{M} = \frac{1}{N} \sum_{i=1}^{N} M_i \Rightarrow \boxed{\hat{\overline{M}} = \frac{1}{n} \sum_{i=1}^{n} M_i, \hat{M} = N \widehat{\overline{M}} = \frac{N}{n} \sum_{i=1}^{n} M_i} \quad \text{Unbiased}$$

$$\hat{\mu}_r = \frac{\hat{\tau}}{\hat{M}} = \frac{\frac{N}{n} \sum_{i=1}^n M_i \overline{y}_i}{\frac{N}{n} \sum_{i=1}^n M_i} = \frac{\sum_{i=1}^n M_i \overline{y}_i}{\sum_{i=1}^n M_i}, \hat{\tau}_r = M \hat{\mu}_r = M \frac{\sum_{i=1}^n M_i \overline{y}_i}{\sum_{i=1}^n M_i}$$
Biased

We can use \overline{M} if we know it

$$\hat{Var}(\hat{\mu}_r) = \frac{N-n}{N} \frac{S_r^2}{n\bar{M}^2} + \frac{1}{nN\bar{M}^2} \sum_{i=1}^n M_i^2 \frac{M_i - m_i}{M_i} \frac{S_i^2}{m_i}$$
 Two components of the variance

$$S_r^2 = \frac{1}{n-1} \sum_{i=1}^n (M_i \bar{y}_i - M_i \hat{\mu}_r)^2 = \frac{1}{n-1} \sum_{i=1}^n M_i^2 (\bar{y}_i - \hat{\mu}_r)^2$$

Inference: Comparison of ratio and unbiased estimation

We want to compare variances of $\hat{\mu} = \hat{\mu}_{UNB}$ and $\hat{\mu} = \mu_r$.

Assuming $\hat{\overline{M}} \approx \overline{M}$ (or just using \overline{M}),

$$\hat{V}ar(\hat{\mu}_r) < \hat{V}ar(\hat{\mu}_{UNB}) \Leftrightarrow S_r^2 < S_b^2$$

$$\sum_{i=1}^{n} (M_{i} \bar{y}_{i} - M_{i} \hat{\mu}_{r})^{2} < \sum_{i=1}^{n} (M_{i} \bar{y}_{i} - \overline{M} \hat{\mu}_{UNB})^{2}$$

or
$$\sum_{i=1}^{n} (\hat{\tau}_i - M_i \hat{\mu}_r)^2 < \sum_{i=1}^{n} (\hat{\tau}_i - \hat{\tau}_{UNB})^2$$

Here we have usual discussion about deviations. Ratio estimator usually performs better than unbiased.



Science and Medicine Library: Collection of statistical books (I)

Sampling population: 161 shelves

Sampling design: Two-stage cluster sampling

PSU – clusters: N = 25 bookcases (vertical collections of shelves)

Sample size: n = 10 bookcases (clusters), and m = 2 shelves per bookcase,

total 20 shelves

Bookcase



1_	2	3	4	5
1	1	1	1	1
2	2	2	2	2
3	3	3	3	3
4	4	4	4	4
5	! 5	5	5	5
6	6	6	6	6

_	6	7	8	9	10	11
	1	1	1	1	1	1
	2	2	2	2	2	2
	3	3	3	3	3	3
. [4	4	4	4	4	4
	5	5	5	5	5	5
, [6	6	6	6	6	6
· [7	7	7	7	7	7

12	13	14	15	16
1	1	1	1	1
2	2	2	2	2
3	3	3	3	3
4	4	4	4	4
5	5	5	5	5
6	6	6	6	6

	17	18	19	20
1	1	1	1	1
2	2	2	2	2
3	3	3	3	3
4	4	4	4	4
5	5	5	5	5
6	6	6	6	6

	21	22	23	24	25
1	1	1	1	1	1
2	2	2	2	2	2
3	3	3	3	3	3
4	4	4	4	4	4
5	5	5	5	5	5
6	6	6	6	6	6
7	7	7	7	7	7



Science and Medicine Library: Collection of statistical books (II)

Sample, first stage: SRS of 10 bookcases (clusters) from 25 bookcases

Table of random numbers (two digits, modulo 50):

46 32 74 60 48 22 35 47 47 75 63 13 11 96 20 98 16 56 87 48 38 30 35 78 21

24 10 22

25 13 13 11 46 20 48 16 6 37

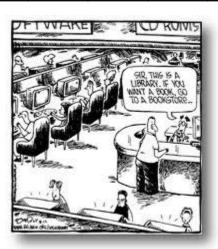
28 **21**

Bookcases selected in the first stage: 6, 10, 11, 13, 16, 20, 21, 22, 24, 25

Second stage: From each selected bookcase two shelves are selected (using one digit)

Bookcase	6	10	11	13	16	20	21	22	24	25
# of shelves	7	7	7	6	6	6	7	7	7	7
Shelves selected	1, 6	6, 4	2, 1	1, 6	4, 5	3, 4	7, 3	4, 3	7, 1	2, 1









Science and Medicine Library: Collection of statistical books (III)

Sample values: (first value y, second value x)

1	2	3	4	5

	6	7	8	9	10	11
1	24,16					21,19
2						17,17
3						
4					30,17	
5						
6	21,17				20,19	
7						

12	13	14	15	16
	21,16			
	+ +			
				32,22
				28,17
	19,13			

	17	18	19	20
1				
2				
3				26,15
4				26,15 26,18
5				
6				

	21	22	23	24	25
1				21,8	17,10
2					17,7
3	18,5	15,7			
4		21,11			
5					
6					
7	18,16			14,0	

Notice: No bookcase was selected from first group of 5 bookcases, just by chance. And, what is that chance?





Science and Medicine Library: Collection of statistical books (IV)

Number of clusters (bookcases) N = 25, number of elements (shelves) M = 161 Sample and calculation organized:

Cluster	1	2	3	4	5	6	7	8	9	10
M_i	7	7	7	6	6	6	7	7	7	7
m_i	2	2	2	2	2	2	2	2	2	2
y_{il}	24	30	21	21	32	26	18	15	21	17
y_{i2}	21	20	17	19	28	26	18	21	14	17
$\overline{\mathcal{Y}}_i$	22.5	25	19	20	30	26	18	18	17.5	17
S_i^2	4.5	50	8	2	8	0	0	18	24.5	0
$\hat{ au}_{i}$	157.5	175	133	120	180	156	126	126	122.5	119

Unbiased estimation:

$$\overline{M} = \frac{161}{25} = 6.44$$
, $\sum_{i=1}^{10} M_i = 7 \times 7 + 3 \times 6 = 67$, $\hat{\overline{M}} = \frac{67}{10} = 6.7$

Average # of books per bookcase

$$\hat{\bar{\tau}} = \hat{\mu}_t = \frac{1}{n} \sum M_i \bar{y}_i = \frac{1}{10} (157.5 + 175 + 133 + 120 + \dots + 122.5 + 119) = 141.5$$

$$\hat{\mu} = \frac{N}{M} \frac{1}{n} \sum M_i \bar{y}_i = \frac{N}{M} \hat{\mu}_t = \frac{25}{161} \times 141.5 = 21.97$$

Average # of books per shelve





Science and Medicine Library: Collection of statistical books (V)

Accuracy, unbiased:

$$S_b^2 = \frac{1}{n-1} \sum (\hat{\tau}_i - \hat{\tau})^2 = \frac{1}{10-1} \sum (\hat{\tau}_i - 141.5)^2 = 550.33$$

$$\sum M_i^2 \frac{M_i - m_i}{M_i} \frac{S_i^2}{m_i} = \frac{1}{2} \sum M_i (M_i - 2) S_i^2 = 1957.5$$

$$\hat{V}ar(\hat{\mu}) = \frac{N-n}{N} \frac{1}{n\overline{M}^2} S_b^2 + \frac{1}{nN\overline{M}^2} \sum M_i^2 \frac{M_i - m_i}{M_i} \frac{S_i^2}{m_i} \qquad \overline{M} = 6.44$$

$$= \frac{25-10}{25} \times \frac{1}{10 \times (6.44)^2} \times 550.33 + \frac{1}{10 \times 25 \times (6.44)^2} \times 1957.5 = 0.985$$

$$\hat{S}d(\hat{\mu}) = 0.992$$

Comparison with SRS, samples ize 20: $\hat{\mu} = 22.9$, $\hat{S}d(\hat{\mu}) = 1.507$





Science and Medicine Library: Collection of statistical books (VI)

Ratio estimation

$$\hat{\mu}_r = \frac{\sum M_i \bar{y}_i}{\sum M_i} = \frac{\sum \hat{\tau}_i}{\sum M_i} = \frac{1415}{67} = 21.12 \qquad S_r^2 = \frac{1}{n-1} \sum (\hat{\tau}_i - M_i \hat{\mu})^2 = 802.243$$

$$\hat{V}ar(\hat{\mu}_r) = \frac{N-n}{N} \frac{1}{nM^2} S_r^2 + \frac{1}{nNM^2} \sum M_i^2 \frac{M_i - m_i}{M_i} \frac{S_i^2}{m_i} \qquad \overline{M} = 6.44$$

$$= \frac{25-10}{25} \times \frac{1}{10 \times (6.44)^2} \times 802.243 + \frac{1}{10 \times 25 \times (6.44)^2} \times 1957.5 = 1.349$$

$$\hat{S}d(\hat{\mu}) = 1.162$$

Comparison

Unbiased cluster : $\hat{\mu} = 21.97, \hat{S}d(\hat{\mu}) = 0.992$

Ratio cluster: $\hat{\mu}_r = 21.12, \hat{S}d(\hat{\mu}_r) = 1.162$

Simple random: $\hat{\mu}_{SRS} = 22.90, \hat{S}d(\hat{\mu}_{SRS}) = 1.507$

Two-stage cluster sampling behaves better than SRS in this problem, but in general need not, if clusters are homogeneous.

Population proportion, summary (Ch. 9.5)

$$y_{ij} = \begin{cases} 0 & \text{Number of elements} \\ 1 & \text{,} \quad \tau_i = \sum_j y_{ij} & \text{in cluster } i \text{ with given} \\ \text{property} \end{cases}$$

Number of elements

 $y_i = a_i$ in sample Same, but

 $\hat{\mu}_i = \hat{p}_i = \overline{y}_i = \frac{a_i}{m_i}$ - sample proportion in *i*th cluster $\Rightarrow \hat{\tau}_i = M_i \hat{p}_i$

$$\Rightarrow \hat{\overline{\tau}}_{UNB} = \frac{1}{n} \sum_{i=1}^{n} M_i \hat{p}_i, \ \hat{p}_{UNB} = \frac{\hat{\overline{\tau}}_{UNB}}{\overline{M}} = \frac{N}{M} \frac{1}{n} \sum_{i=1}^{n} M_i \hat{p}_i, \ \hat{p}_r = \frac{\sum_{i=1}^{n} M_i \hat{p}_i}{\sum_{i=1}^{n} M_i}, \ \hat{\overline{\tau}}_r = \overline{M} \hat{p}_r$$

$$\hat{Var}(\hat{p}) = \frac{N-n}{N} \frac{S_p^2}{n\overline{M}^2} + \frac{1}{nN\overline{M}^2} \sum_{i=1}^n M_i^2 \frac{M_i - m_i}{M_i} \frac{\hat{p}_i \hat{q}_i}{m_i - 1}$$

$$S_p^2 = \begin{cases} S_b^2, \text{ unbiased} \\ S_r^2, \text{ ratio} \end{cases}$$

$$S_p^2 = \begin{cases} S_b^2, \text{ unbiased} \\ S_r^2, \text{ ratio} \end{cases}$$

$$S_r^2 = \frac{1}{n-1} \sum_{i=1}^n (M_i \hat{p}_i - M_i \hat{p}_r)^2 = \frac{1}{n-1} \sum_{i=1}^n M_i^2 (\hat{p}_i - \hat{p}_r)^2$$

$$S_b^2 = \frac{1}{n-1} \sum_{i=1}^n (M_i \hat{p}_i - \overline{M} \hat{p}_{UNB})^2 = \frac{1}{n-1} \sum_{i=1}^n (\hat{\tau}_i - \hat{\overline{\tau}}_{UNB})^2$$

 $ig| \overline{M}$ with unbiased $\hat{\overline{M}}$ withratio (?)

Examples: book

Equal cluster sizes, Ch. 9.6 (I)

 $M_i = M, i = 1,2,...,N$ - equal cluster sizes, no ratio estimation Use $m_i = m, i = 1, 2, ..., N$ - equal sample sizes from each cluster

$$\hat{\mu} = \hat{\mu}_{UNB} = \frac{1}{\overline{M}} \frac{1}{n} \sum_{i=1}^{n} M_i \overline{y}_i = \frac{1}{n} \sum_{i=1}^{n} \overline{y}_i \Longrightarrow$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^{n} \overline{y}_i = \frac{1}{n \times m} \sum_{i,j} y_{ij} = \overline{y}$$
 Just sample mean

$$S_b^2 = \frac{1}{n-1} \sum_{i=1}^n (M_i \bar{y}_i - \overline{M} \hat{\mu})^2 = \frac{\overline{M}^2}{n-1} \sum_{i=1}^n (\bar{y}_i - \hat{\mu})^2 = \overline{M}^2 S_{\bar{y}}^2$$

$$MSB = \frac{m}{n-1} \sum_{i=1}^{n} (\bar{y}_i - \hat{\mu})^2 = mS_{\bar{y}}^2, \quad MSW = \frac{1}{n} \sum_{i=1}^{n} S_i^2$$

$$\hat{V}ar(\hat{\mu}) = \frac{N - n}{N} \frac{S_{\bar{y}}^2}{n} + \frac{\overline{M} - m}{\overline{M}} \frac{1}{mN} \frac{1}{n} \sum_{i=1}^n S_i^2 = (1 - \frac{n}{N}) \frac{MSB}{n \times m} + (1 - \frac{m}{\overline{M}}) \frac{1}{N} \frac{MSW}{m}$$

Equal cluster sizes (II)

Summary:

From
$$\hat{Var}(\hat{\mu}) = (1 - \frac{n}{N})\frac{MSB}{n \times m} + (1 - \frac{m}{\overline{M}})\frac{1}{N}\frac{MSW}{m}$$

1)
$$N \text{ large}: \hat{V}ar(\hat{\mu}) \approx (1 - \frac{n}{N}) \frac{MSB}{n \times m} \approx \frac{MSB}{n \times m}$$

2) If
$$m = \overline{M}$$
 - one - stage : $\widehat{Var}(\widehat{\mu}) = (1 - \frac{n}{N}) \frac{MSB}{n \times m}$

3) If
$$n = N$$
 - stratified, $L = N$, equal allocation, $n_i = m$:

$$\hat{V}ar(\hat{\mu}) = (1 - \frac{m}{\overline{M}}) \frac{MSW}{n \times m}$$

Equal cluster sizes: optimal sample size (I)

Two cases to select sample size n:

- Given cost of sampling, minimize variance (error bound)
- Given error bound (accuracy, standard error), minimize cost

Cost model – linear model:
$$C = C(n) = c_0 + c_1 \times n + c_2 \times n \times m$$

 c_0 – fixed cost,

 c_1 – cost of sampling one cluster,

Find optimal n and m

 c_2 – cost of sampling one unit

Theoretical variance: $Var(\hat{\mu}) = \frac{N-n}{N} \frac{\widetilde{\sigma}_{\mu}^2}{n} + \frac{1}{n} \frac{\overline{M} - m}{\overline{M}} \frac{\overline{\widetilde{\sigma}}^2}{n}$

$$\widetilde{\sigma}_{\mu}^{2} = \frac{1}{N-1} \sum_{i=1}^{N} (\mu_{i} - \mu)^{2}, \overline{\widetilde{\sigma}}^{2} = \frac{1}{N} \sum_{i=1}^{N} \widetilde{\sigma}_{i}^{2} = \frac{1}{N} \sum_{i=1}^{N} \frac{1}{\overline{M} - 1} \sum_{i=1}^{\overline{M}} (y_{ij} - \mu_{i})^{2}$$

PPS and two-stage cluster (I)

Sample design: First stage – PPS of clusters,

Second stage – SRS of elements

Selest cluster i with probability proportional to its size: $\pi_i = \frac{M_i}{M}$

$$\left| \hat{\mu}_{PPS} = \frac{1}{n} \sum_{i=1}^{n} \overline{y}_{i} , \hat{\tau}_{PPS} = M \hat{\mu}_{PPS} , \overline{y}_{i} = \frac{1}{m_{i}} \sum_{j=1}^{m_{i}} y_{ij} \right|$$

Unbiased

$$|\hat{V}ar(\hat{\mu}_{PPS}) = \frac{1}{n} S_{\bar{y}}^2 = \frac{1}{n} \times \frac{1}{n-1} \sum_{i=1}^{n} (\bar{y}_i - \hat{\mu}_{PPS})^2$$

$$\left| \hat{V}ar(\hat{\tau}_{PPS}) = \hat{V}ar(M\hat{\mu}_{PPS}) = M^2 \frac{1}{n} S_{\bar{y}}^2 \right|$$



Example: PPS and cattle farms (I)

Population: 2072 farms divided into 53 clusters of unequal size (by area),

M = 2072/53 = 39.094, the average cluster size.

Goal: Estimate the average number of cattle per farm.

Variables: Number of cattle on farm (y).

Parameters to be estimated: Average number of cattle per farm (μ_y) and total

number of cattle on farms.

Sampling design: First stage: 14 clusters with probabilities of selection

proportional to cluster size (number of farms), with replacements.

Second stage: SRS of ¼ farms from the cluster.

Method of estimation: Unbiased PPS estimation.

Sample: PPS sample of 14 clusters selected from 53 clusters with probabilities

proportional to number of farms (clusters in the sample ordered by cluster

size). (see next slide)

Example: PPS and cattle farms (II)

Comple	Number	Number of	Number of	Avaraga aattla nar	Salastian
Sample	Number		Number of	C 1	Selection
cluster,	of farms,	farms in	cattle in	sampled farm,	Probability
i	M_{i}	sample, m_i	sample, y_i	$\overline{y}_i = y_i / m_i$	$\pi_i = M_i/M$
1	13	3	30	10.00	13/2072
2	15	3	58	19.33	15/2072
3	19	5	14	2.80	19/2072
4	28	7	73	10.43	
5	39	10	162	16.20	
6	41	11	88	8.00	
7	46	12	102	8.50	
8	46	12	102	8.50	
9	48	12	203	16.92	
10	51	13	134	10.31	
11	59	14	195	13.93	
12	74	19	272	14.32	
13	83	20	242	12.10	83/2072
14	83	20	242	12.10	83/2072
Total	645	161	1917	163.43	

$$\overline{y}_1 = \frac{30}{3} = 10,$$
 $\overline{y}_2 = \frac{58}{3} = 19.33,$

 $m_i \approx M_i/4$

Alternative first stage PPS sampling: Selection probabilities proportional to the cluster area.



Example: PPS and cattle farms (III)

Estimation:

$$\hat{\mu}_{pps} = \frac{1}{n} \sum_{i=1}^{n} \frac{y_i}{m_i} = \frac{1}{n} \sum_{i=1}^{n} \overline{y}_i = \frac{163.43}{14} = 11.674$$

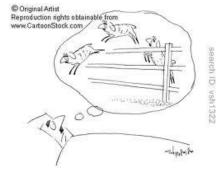
$$\left| \hat{\tau}_{pps} = M\hat{\mu} = 2072 \times 11.674 = 24,188.5 \right|$$

$$\hat{V}ar(\hat{\mu}_{pps}) = \frac{1}{n} \left[\frac{1}{(n-1)} \sum_{i=1}^{n} (\bar{y}_{i} - \hat{\mu}_{pps})^{2} \right] = \frac{1}{n} S_{\bar{y}}^{2} = \frac{1}{14} 18.283 = 1.306$$

$$\hat{\sigma}(\hat{\mu}) = \sqrt{1.306} = 1.143$$

$$\hat{V}ar(\hat{\tau}_{pps}) = M^2 \hat{V}ar(\hat{\mu}_{pps}) = 2072^2 \times 1.306$$

$$\hat{\sigma}(\hat{\tau}_{pps}) = 2072 \times 1.143 = 2,367.8$$



Compare with one stage PPS estimation: $\hat{\mu} = 11.713$, $\hat{\sigma}(\hat{\mu}) = 0.813$. Number of sampled farms: for one stage sample 662, for two stage sample 161. Two stage sample is four times smaller, but (slightly?) less efficient.