

First we derive cdf of  $U=Y^2$  when  $Y \sim f_Y(y)$

$$F_U(u) = P(U \leq u) = P(Y^2 \leq u)$$

if  $u < 0$ ; then  $F_U(u) = 0$

$$\text{if } u \geq 0; F_U(u) = P(Y^2 \leq u) = P(-\sqrt{u} \leq Y \leq \sqrt{u})$$

$$= \int_{-\sqrt{u}}^{\sqrt{u}} f_Y(y) dy = F_Y(\sqrt{u}) - F_Y(-\sqrt{u})$$

$$\text{So } F_U(u) = F_Y(\sqrt{u}) - F_Y(-\sqrt{u})$$

$$f_U(u) = \frac{dF_Y(\sqrt{u})}{du} \cdot \frac{1}{2\sqrt{u}} + \frac{dF_Y(-\sqrt{u})}{du} \cdot \frac{1}{2\sqrt{u}}$$

$$= f_Y(\sqrt{u}) \cdot \frac{1}{2\sqrt{u}} + f_Y(-\sqrt{u}) \cdot \frac{1}{2\sqrt{u}}$$

~~$$= \frac{1}{2\sqrt{u}} (f_Y(\sqrt{u}) + f_Y(-\sqrt{u}))$$~~

$$\text{So } f_U(u) = \begin{cases} f_Y(\sqrt{u}) \left( \frac{1}{2\sqrt{u}} \right) + f_Y(-\sqrt{u}) \left( \frac{1}{2\sqrt{u}} \right) \\ 0 \end{cases}$$

$u > 0$

$0.0$

Since  $Y \sim U(-1, 1) \Rightarrow f_Y(y) = \frac{1}{2} \quad -1 < y < 1$

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$$U = Y^2 \quad Y: -1 \rightarrow 1 \\ u: 0 \rightarrow 1$$

$$f_U(u) = \begin{cases} \frac{1}{2} \left( \frac{1}{2\sqrt{u}} \right) + \frac{1}{2} \left( \frac{1}{2\sqrt{u}} \right) & u > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\Rightarrow f_U(u) = \begin{cases} \frac{1}{2\sqrt{u}} \\ 0 \end{cases}$$

$$0 < u < 1$$

o.w

$$f_Y(y) = \begin{cases} 2y & 0 \leq y \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

$$U = -4Y + 3 ;$$

$$h(y) = -4y + 3 \Rightarrow u = -4y + 3 \Rightarrow u - 3 = -4y \Rightarrow 4y = 3 - u$$

$$\Rightarrow y = \frac{3-u}{4} = h^{-1}(u) \quad \Rightarrow \quad \frac{dh^{-1}(u)}{du} = -\frac{1}{4}$$

$$f_U(u) = f_Y(h^{-1}(u)) \left| \frac{dh^{-1}}{du} \right|$$

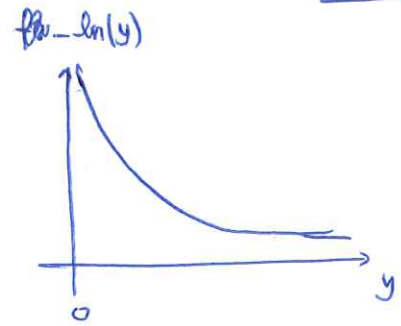
$$\Rightarrow f_U(u) = 2 \left( \frac{3-u}{4} \right) \cdot \left| -\frac{1}{4} \right| = 2 \left( \frac{3-u}{4} \right) \cdot \frac{1}{4} = \frac{3-u}{8} \quad 0 \leq \frac{3-u}{4} \leq 1$$

$$0 \leq \frac{3-u}{4} \leq 1 \Rightarrow 0 \leq 3-u \leq 4 \Rightarrow -4 \leq u-3 \leq 0 \Rightarrow -1 \leq u \leq 3$$

$$= \begin{cases} \frac{3-u}{8} & -1 \leq u \leq 3 \\ 0 & \text{o.w.} \end{cases}$$

$$Y \sim U(0,1) ; U = -2\ln(Y)$$

$$\begin{cases} Y=0 \rightarrow 1 \\ -\ln(Y): 0 \rightarrow \infty \end{cases}$$



$$f_Y(y) = 1 \quad 0 < y < 1$$

$$u = -2\ln(y) \Rightarrow -\frac{u}{2} = \ln(y) \Rightarrow y = e^{-u/2}; \quad \frac{\partial y}{\partial u} = -e^{-u/2} \cdot \frac{1}{2}$$

$$|j| = \left| \frac{\partial y}{\partial u} \right| = e^{-u/2} \cdot \frac{1}{2}$$

$$f_U(u) = f_Y(y) \cdot |j| = f_Y(e^{-u/2}) \cdot e^{-u/2} \cdot \frac{1}{2}$$

$$= \frac{1}{2} \cdot e^{-u/2} = \frac{1}{2} e^{-u/2} \quad u \geq 0 \Rightarrow U \sim \text{Exp}(2)$$

$$Y \sim \text{Bin}(n, p) \Rightarrow \text{mgf of } Y : m_Y(t) = (1 - p + pe^t)^n$$

$$X = n - Y \Rightarrow$$

$$m_X(t) = E(e^{tx}) = E(e^{t(n-Y)}) = e^{tn} E(e^{-ty})$$

$$= e^{nt} m_Y(-t) = e^{nt} (1 - p + pe^{-t})^n = (e^t - pe^t + p)^n$$

$$= (p + (1-p)e^t)^n \Rightarrow X \sim \text{Bin}(n, 1-p)$$

So  $X$  is # of failures observed in the experiment.

From Slide 7 and page 305 of textbook, we have

$$Y \sim f_Y(y) \quad \text{Then } U = Y^2$$

$$f_u(u) = \begin{cases} \frac{1}{2\sqrt{u}} \left( f_Y(\sqrt{u}) + f_Y(-\sqrt{u}) \right) & u > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\text{Here } Z = \frac{X - \mu}{\sigma} \sim N(0, 1) \Rightarrow f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

$$\text{Now } U = Z^2$$

$$\begin{aligned} f_U(u) &= \frac{1}{2\sqrt{u}} \left( f_Z(\sqrt{u}) + f_Z(-\sqrt{u}) \right) \\ &= \frac{1}{2\sqrt{u}} \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{u}{2}} + \frac{1}{\sqrt{2\pi}} e^{-\frac{u}{2}} \right) \\ &= \begin{cases} \frac{1}{\sqrt{2\pi u}} e^{-\frac{u}{2}} & u > 0 \\ 0 & \text{o.w.} \end{cases} = \frac{u^{-1/2} e^{-u/2}}{2^{1/2} \sqrt{\pi}} \end{aligned}$$

$$X \sim \text{Gamma}(\alpha, \beta) \quad f(x) = \frac{x^{\alpha-1} e^{-x/\beta}}{\beta^\alpha \Gamma(\alpha)}$$

$$\Rightarrow U \sim \text{Gamma}\left(\frac{1}{2}, 2\right) \quad ; \quad \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

"  $\chi^2_1$

$$Z \sim N(0,1) \Rightarrow m_{Z^2}(t) = E(e^{tZ^2}) = \int_{-\infty}^{\infty} e^{tZ^2} f(z) dz = \int_{-\infty}^{\infty} e^{tZ^2} \cdot \frac{e^{-\frac{z^2}{2}}}{\sqrt{2\pi}} dz$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{z^2}{2}\right)(1-2t)} dz \quad ; \quad \text{s.t.} \quad 1-2t > 0 \Rightarrow t < \frac{1}{2}$$

$$= \frac{\exp\left(-\left(\frac{z^2}{2}\right)(1-2t)\right)}{\sqrt{2\pi}} = \frac{\exp\left(-\left(\frac{z^2}{2}\right)/(1-2t)^{-1}\right)}{\sqrt{2\pi}}$$

$$\Rightarrow m_{Z^2}(t) = \frac{1}{(1-2t)^{1/2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi} (1-2t)^{-1/2}} \exp\left(-\left(\frac{z^2}{2}\right)/(1-2t)^{-1}\right) dz$$

But this integral equals to 1; if  $t < \frac{1}{2}$

$$m_{Z^2}(t) = \frac{1}{(1-2t)^{1/2}} = (1-2t)^{-1/2}$$

$$X \sim \text{Gamma}(d, \beta) \quad ; \quad m_X(t) = (1-\beta t)^{-d}$$

$$\text{So } Z^2 = U \sim \text{Gamma}(d = \frac{1}{2}, \beta = 2) \equiv \chi^2_{(1)}$$

"   
 Gamma( $\nu/2, 2$ )

$$Y_1 \sim \text{Poi}(\mu_1) \Rightarrow m_{Y_1}(t) = \exp(\mu_1(e^t - 1))$$

$$Y_2 \sim \text{Poi}(\mu_2) \Rightarrow m_{Y_2}(t) = \exp(\mu_2(e^t - 1))$$

$$m_U(t) = E(e^{tU}) = E(e^{t(Y_1 + Y_2)}) = E(e^{tY_1 + tY_2}) = E(e^{tY_1}) \cdot E(e^{tY_2})$$

$$= m_{Y_1}(t) \cdot m_{Y_2}(t) = \exp(\mu_1(e^t - 1)) \exp(\mu_2(e^t - 1))$$

$$= \exp((e^t - 1)(\mu_1 + \mu_2)) \Rightarrow U \sim \text{Poi}(\mu_1 + \mu_2)$$



$$y_i \sim N(\mu_i, \sigma_i^2) \quad m_{y_i}(t) = E(e^{ty_i}) = e^{\mu_i t + \frac{t^2 \sigma_i^2}{2}}$$

$$m_{a_i y_i} = E(e^{ta_i y_i}) = m_{y_i}(ta_i)$$

$$m_U(t) \stackrel{\text{independent}}{=} m_{a_1 y_1}(t) * m_{a_2 y_2}(t) * \dots * m_{a_n y_n}(t)$$

$$= m_{y_1}(ta_1) * m_{y_2}(ta_2) * \dots * m_{y_n}(ta_n)$$

$$= e^{\sum_{i=1}^n \left( \mu_i a_i t + \frac{t^2 a_i^2 \sigma_i^2}{2} \right)}$$

$$= e^{t \sum_{i=1}^n a_i \mu_i + \frac{t^2}{2} \sum_{i=1}^n a_i^2 \sigma_i^2}$$

$$\text{So } U \sim N\left(\sum_{i=1}^n a_i \mu_i, \sum_{i=1}^n a_i^2 \sigma_i^2\right)$$

$$Y_i \sim \chi^2_{(v_i)} \equiv \text{Gamma}\left(\frac{v_i}{2}, 2\right) \Rightarrow m_{Y_i}(t) = (1-2t)^{-v_i/2}$$

$$U = \sum_{i=1}^n Y_i \Rightarrow m_U(t) = E\left(e^{t \sum_{i=1}^n Y_i}\right) = \prod_{i=1}^n m_{Y_i}(t) = m_{Y_1}(t) * m_{Y_2}(t) * \dots * m_{Y_n}(t)$$

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$$= \prod_{i=1}^n m_{Y_i}(t) = \prod_{i=1}^n (1-2t)^{-v_i/2} = (1-2t)^{-\sum_{i=1}^n v_i/2}$$

$$\Rightarrow U \sim \text{Gamma}\left(\frac{\sum_{i=1}^n v_i}{2}, 2\right) \equiv \chi^2_{\left(\sum_{i=1}^n v_i\right)}$$