

Introduction to Machine Learning

CMU-10701

2. MLE, MAP, Bayes classification

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Administration

http://www.cs.cmu.edu/~aarti/Class/10701_Spring14/index.html

- Blackboard manager & Peer grading: Dani
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Outline

Theory:

☐ Probabilities:

- Dependence, Independence, Conditional Independence

☐ Parameter estimation:

- Maximum Likelihood Estimation (MLE)
- Maximum a posteriori (MAP)

☐ Bayes rule

- Naïve Bayes Classifier

Application:

Naive Bayes Classifier for

- Spam filtering
- “Mind reading” = fMRI data processing

Independence

Independence

Independent random variables:

$$P(X, Y) = P(X)P(Y)$$

$$P(X|Y) = P(X)$$

Y and X don't contain information about each other.

Observing Y doesn't help predicting X.

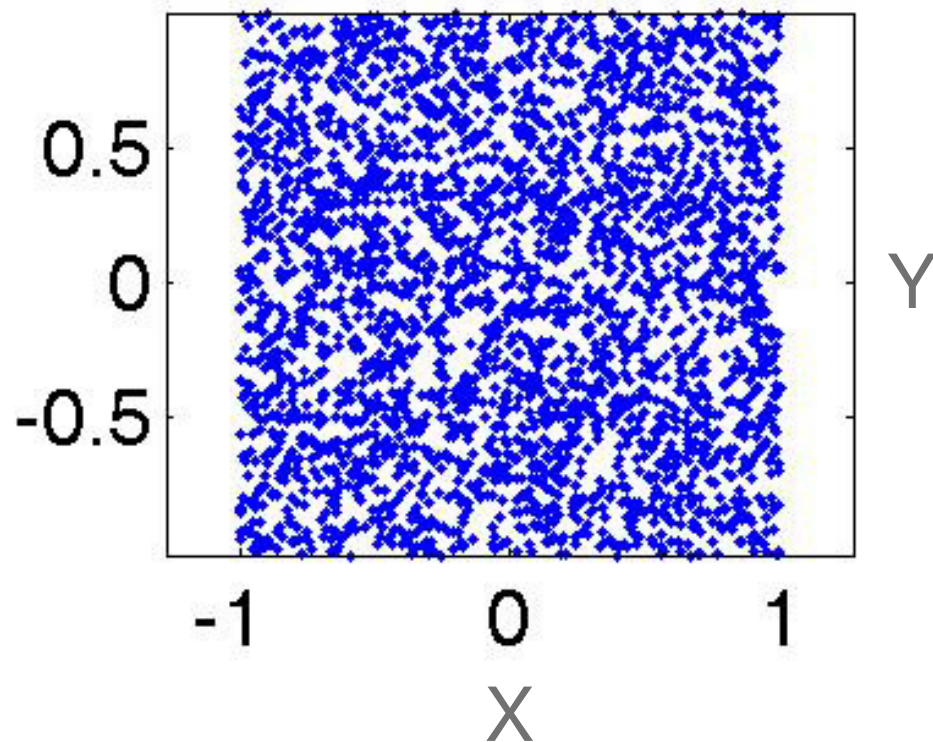
Observing X doesn't help predicting Y.

Examples:

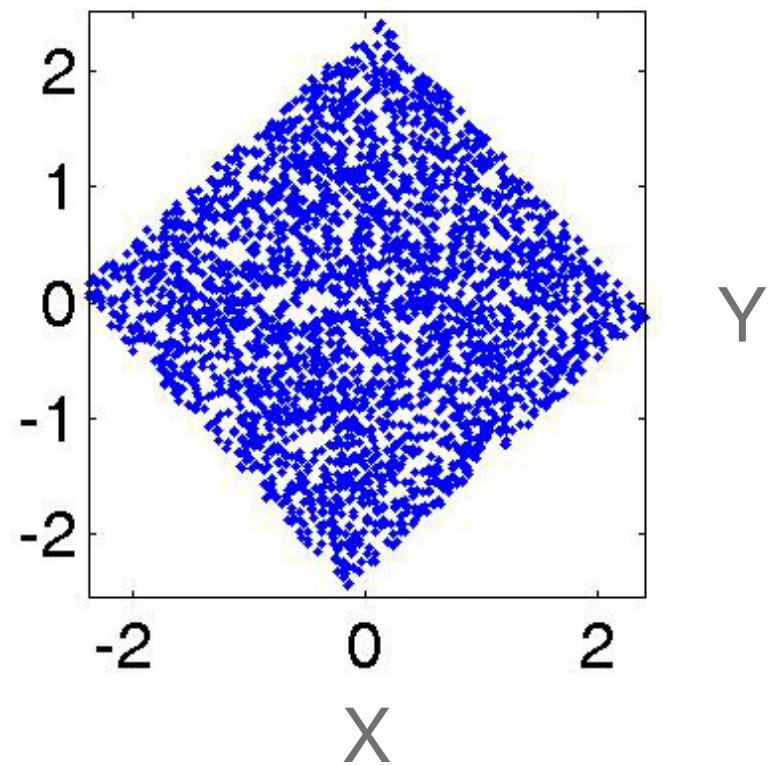
Independent: Winning on roulette this week and next week.

Dependent: Russian roulette

Dependent / Independent



Independent X,Y



Dependent X,Y

Conditionally Independent

Conditionally independent:

$$P(X, Y|Z) = P(X|Z)P(Y|Z)$$

Knowing Z makes X and Y independent

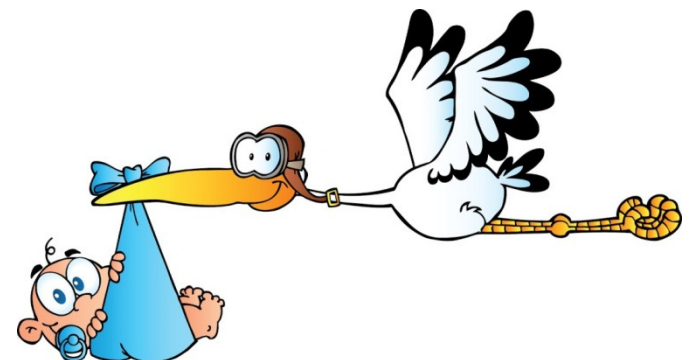
Examples:

Dependent: show size and reading skills

Conditionally independent: show size and reading skills given age

Storks deliver babies:

Highly statistically significant correlation exists between stork populations and human birth rates across Europe.

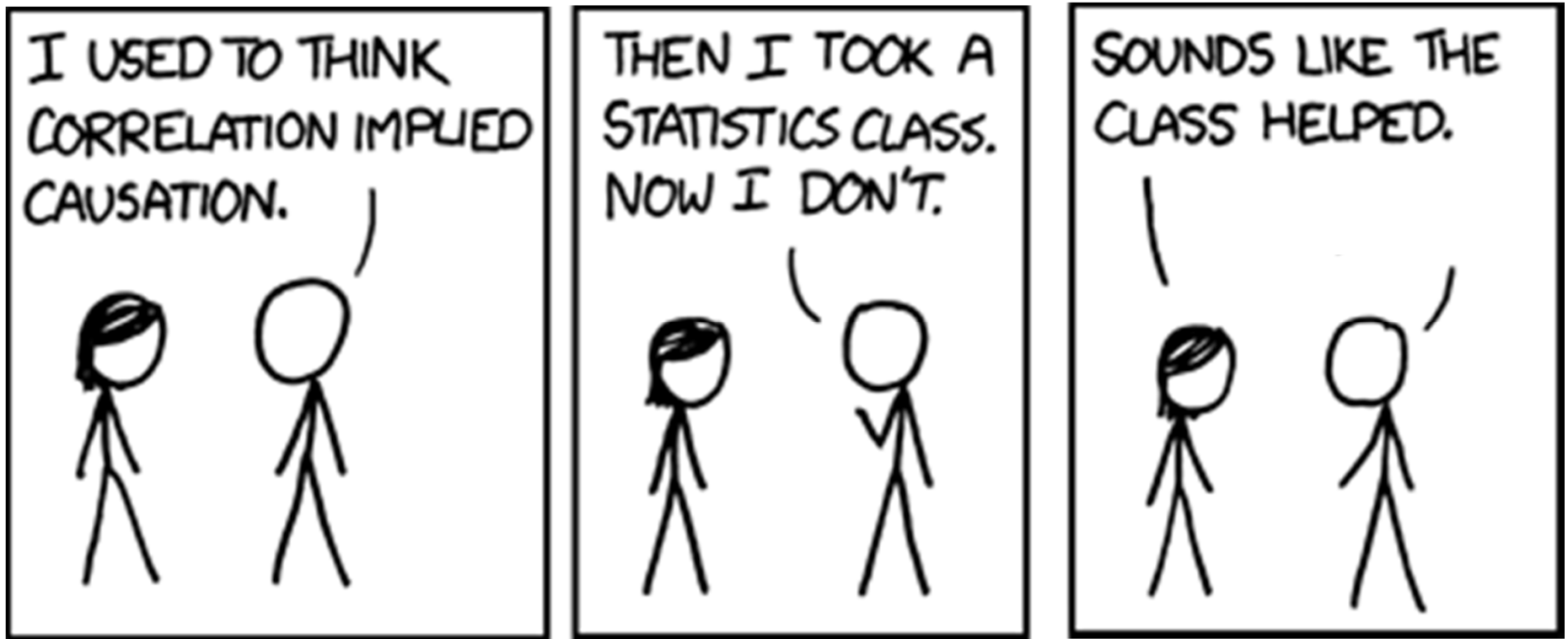


Conditionally Independent

London taxi drivers: A survey has pointed out a positive and significant correlation between the number of accidents and wearing coats. They concluded that coats could hinder movements of drivers and be the cause of accidents. A new law was prepared to prohibit drivers from wearing coats when driving.

Finally another study pointed out that people wear coats when it rains...

Correlation \neq Causation



Conditional Independence

Formally: X is **conditionally independent** of Y given Z:

$$P(X, Y|Z) = P(X|Z)P(Y|Z)$$

$$P(\text{Accidents, Coats} | \text{Rain}) = P(\text{Accidents} | \text{Rain})P(\text{Coats} | \text{Rain})$$

Equivalent to:

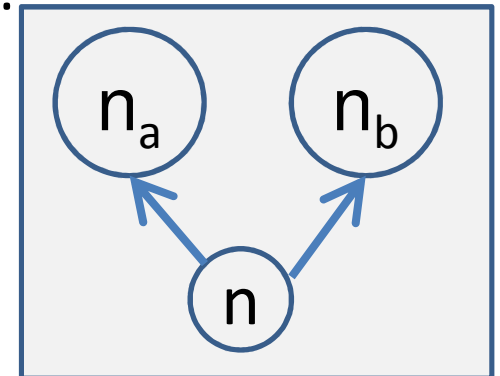
$$(\forall x, y, z) P(X = x | Y = y, Z = z) = P(X = x | Z = z)$$

$$P(\text{Thunder} | \text{Rain, Lightning}) = P(\text{Thunder} | \text{Lightning})$$

Note: does NOT mean Thunder is independent of Rain
But given Lightning knowing Rain doesn't give more info about Thunder

Conditional vs. Marginal Independence

- C calls A and B separately and tells them a number $n \in \{1, \dots, 10\}$
- Due to noise in the phone, A and B each imperfectly (and independently) draw a conclusion about what the number was.
- A thinks the number was n_a and B thinks it was n_b .
- Are n_a and n_b marginally independent?
 - No, we expect e.g. $P(n_a = 1 \mid n_b = 1) > P(n_a = 1)$
- Are n_a and n_b conditionally independent given n ?



- Yes, because if we know the true number, the outcomes n_a and n_b are purely determined by the noise in each phone.

$$P(n_a = 1 \mid n_b = 1, n = 2) = P(n_a = 1 \mid n = 2)$$

Our first machine learning problem:

Parameter estimation: MLE, MAP

Estimating Probabilities



Flipping a Coin

I have a coin, if I flip it, what's the probability that it will fall with the head up?

Let us flip it a few times to estimate the probability:



The estimated probability is: $\frac{3}{5}$ "Frequency of heads"

Flipping a Coin



The estimated probability is: $3/5$ "Frequency of heads"

Questions:

- (1) Why frequency of heads???
- (2) How good is this estimation???
- (3) Why is this a machine learning problem???

We are going to answer these questions

Question (1)

Why frequency of heads???

- Frequency of heads is exactly the *maximum likelihood estimator* for this problem
- MLE has nice properties
(interpretation, statistical guarantees, simple)

Maximum Likelihood Estimation

MLE for Bernoulli distribution

Data, $D =$



$$D = \{X_i\}_{i=1}^n, X_i \in \{H, T\}$$

$$P(\text{Heads}) = \theta, P(\text{Tails}) = 1 - \theta$$

Flips are **i.i.d.**:

- **Independent** events
- **Identically distributed** according to Bernoulli distribution

MLE: Choose θ that maximizes the probability of observed data

Maximum Likelihood Estimation

MLE: Choose θ that maximizes the probability of observed data

$$\begin{aligned}\hat{\theta}_{MLE} &= \arg \max_{\theta} P(D | \theta) \leftarrow P(x_1, x_2, \dots, x_n | \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n P(X_i | \theta) \quad \text{Independent draws} \\ &= \arg \max_{\theta} \prod_{i: X_i = H} \theta \prod_{i: X_i = T} (1 - \theta) \quad \text{Identically distributed} \\ &= \arg \max_{\theta} \underbrace{\theta^{\alpha_H} (1 - \theta)^{\alpha_T}}_{J(\theta)}\end{aligned}$$

n
 α_H
 α_T

Maximum Likelihood Estimation

MLE: Choose θ that maximizes the probability of observed data

$$\begin{aligned}\hat{\theta}_{MLE} &= \arg \max_{\theta} P(D | \theta) \\ &= \arg \max_{\theta} \underbrace{\theta^{\alpha_H} (1 - \theta)^{\alpha_T}}_{J(\theta)}\end{aligned}$$

$$\frac{\partial J(\theta)}{\partial \theta} = \alpha_H \theta^{\alpha_H - 1} (1 - \theta)^{\alpha_T} - \alpha_T \theta^{\alpha_H} (1 - \theta)^{\alpha_T - 1} \Big|_{\theta = \hat{\theta}_{MLE}} = 0$$

$$\left[\alpha_H (1 - \theta) - \alpha_T \theta \Big|_{\theta = \hat{\theta}_{MLE}} = 0 \right]$$

$$\begin{aligned} &\nearrow (1 - \theta)^{\alpha_T - 1} \alpha_H^{-1} \\ &\quad \cdot \theta \end{aligned}$$

$$\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$

That's exactly the “Frequency of heads”

Question (2)

How good is this MLE estimation???

$$\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$

$$\mathbb{E} \hat{\theta} = \theta$$

$$Bias = |\theta - \mathbb{E} \hat{\theta}|$$

How many flips do I need?

I flipped the coins 5 times: 3 heads, 2 tails

$$\hat{\theta}_{MLE} = \frac{3}{5}$$

What if I flipped 30 heads and 20 tails?

$$\hat{\theta}_{MLE} = \frac{30}{50}$$

- Which estimator should we trust more?
- The more the merrier???

Simple bound

Let θ^* be the true parameter.

For $n = \alpha_H + \alpha_T$, and $\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$

For any $\epsilon > 0$:

Hoeffding's inequality:

$$P(|\hat{\theta} - \theta^*| \geq \epsilon) \leq 2e^{-2n\epsilon^2}$$



Probably Approximate Correct (PAC) Learning

I want to know the coin parameter θ , within $\epsilon = 0.1$ error with probability at least $1 - \delta = 0.95$.

How many flips do I need?

$$P(|\hat{\theta} - \theta^*| \geq \epsilon) \leq \underbrace{2e^{-2n\epsilon^2}}_{\substack{e^{-2n\epsilon^2} \leq \frac{\delta}{2} \\ -2n\epsilon^2 \leq \ln(\frac{\delta}{2}) \\ \ln(\frac{2}{\delta}) \leq 2n\epsilon^2}} \leq \underbrace{\delta}_{\text{0.05}}$$
$$\Rightarrow n \geq \frac{1}{2\epsilon^2} \ln\left(\frac{2}{\delta}\right)$$

Sample complexity:

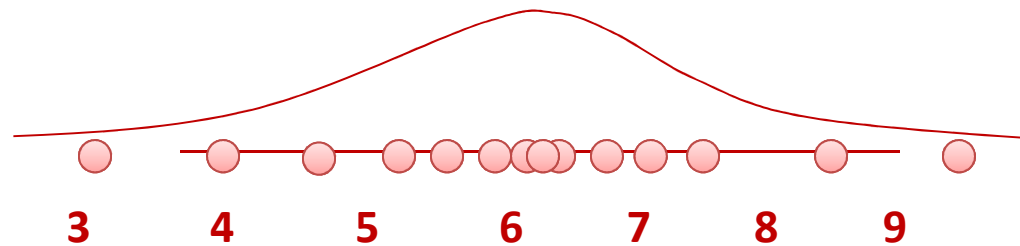
$$n \geq \frac{\ln(2/\delta)}{2\epsilon^2}$$

Question (3)

Why is this a machine learning problem???

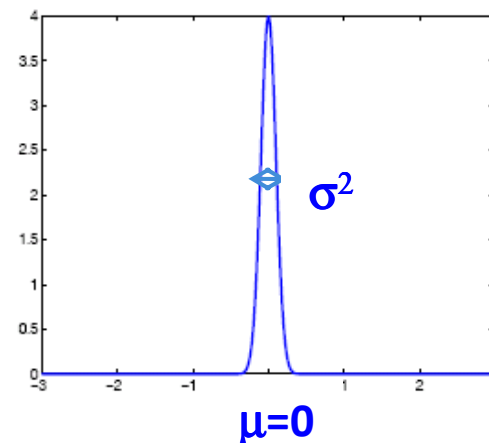
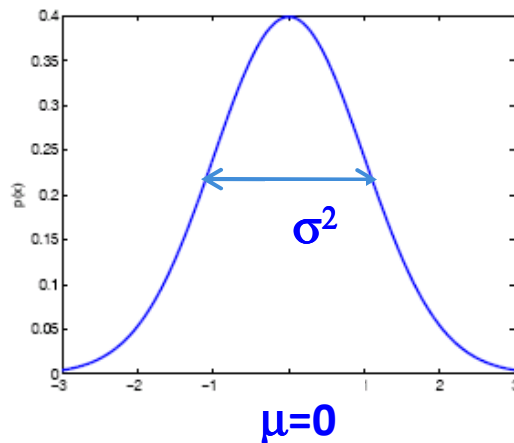
- improve their **performance** (accuracy of the predicted prob.)
- at some **task** (predicting the probability of heads)
- with **experience** (the more coins we flip the better we are)

What about continuous features?



Let us try Gaussians...

$$p(x | \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) = \mathcal{N}_x(\mu, \sigma)$$



MLE for Gaussian mean and variance

Choose $\theta = (\mu, \sigma^2)$ that maximizes the probability of observed data

$$\begin{aligned}\hat{\theta}_{MLE} &= \arg \max_{\theta} P(D | \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n P(X_i | \theta) && \text{Independent draws} \\ &= \arg \max_{\theta} \prod_{i=1}^n \frac{1}{2\sigma^2} e^{-(X_i - \mu)^2 / 2\sigma^2} && \text{Identically distributed} \\ &= \arg \max_{\theta = (\mu, \sigma^2)} \underbrace{\frac{1}{2\sigma^2} e^{-\sum_{i=1}^n (X_i - \mu)^2 / 2\sigma^2}}_{J(\theta)}\end{aligned}$$

MLE for Gaussian mean and variance

$$\hat{\mu}_{MLE} = \frac{1}{n} \sum_{i=1}^n x_i$$
$$\hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2$$

Note: MLE for the variance of a Gaussian is **biased**

[Expected result of estimation is **not** the true parameter!]

Unbiased variance estimator: $\hat{\sigma}_{unbiased}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu})^2$

$$E[\hat{\sigma}_{MLE}^2] \neq \sigma^2 \quad E[\hat{\sigma}_{unbiased}^2] = \sigma^2$$

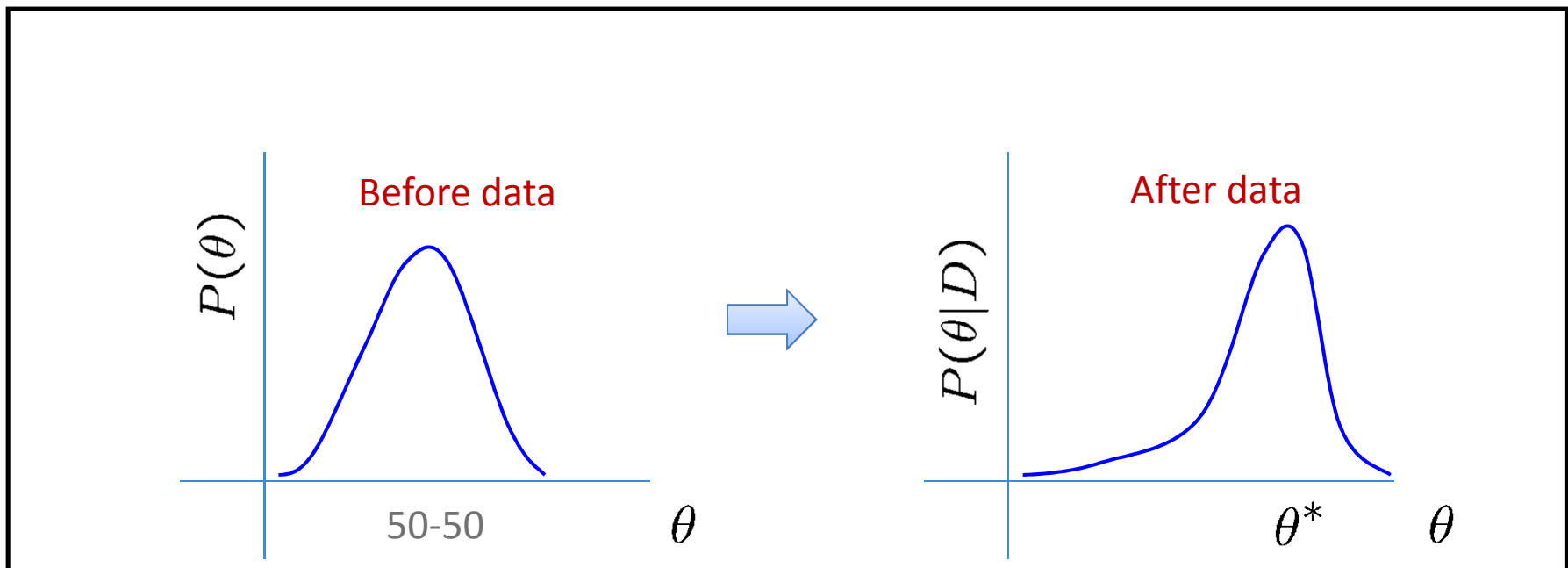
What about prior knowledge ?
(MAP Estimation)

What about prior knowledge?

We know the coin is “close” to 50-50. What can we do now?

The Bayesian way...

Rather than estimating a single θ , we obtain a distribution over possible values of θ



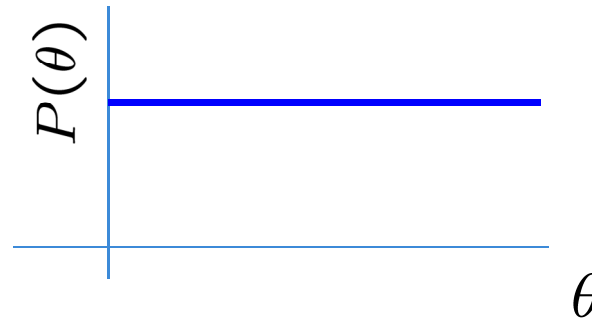
Prior distribution

What prior? What distribution do we want for a prior?

- Represents expert knowledge (philosophical approach)
- Simple posterior form (engineer's approach)

Uninformative priors:

- Uniform distribution



Conjugate priors:

- Closed-form representation of posterior
- $P(\theta)$ and $P(\theta|D)$ have the same form

In order to proceed we will need:

Bayes Rule



Bayes, Thomas (1763) An essay towards solving a problem in the doctrine of chances. *Philosophical Transactions of the Royal Society of London*, 53:370-418

Chain Rule & Bayes Rule

Chain rule:

$$P(X, Y) = P(X|Y)P(Y) = P(Y|X)P(X)$$

Bayes rule:

$$P(X|Y) = \frac{P(Y|X)P(X)}{P(Y)}$$

Bayes rule is important for reverse conditioning.

Bayesian Learning

- Use Bayes rule:

$$P(\theta \mid \mathcal{D}) = \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

- Or equivalently:

$$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta)P(\theta)$$

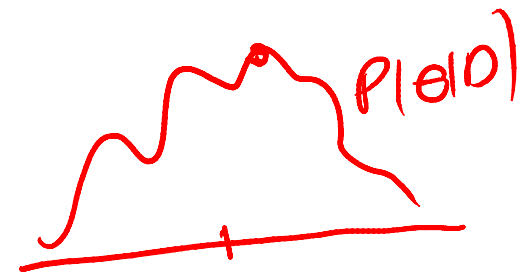
posterior likelihood prior

MLE vs. MAP

- Maximum Likelihood estimation (MLE)

Choose value that maximizes the probability of observed data

$$\hat{\theta}_{MLE} = \arg \max_{\theta} P(D|\theta)$$



- Maximum *a posteriori* (MAP) estimation

Choose value that is most probable given observed data and prior belief

$$\begin{aligned}\hat{\theta}_{MAP} &= \arg \max_{\theta} P(\theta|D) \\ &= \arg \max_{\theta} P(D|\theta)P(\theta)\end{aligned}$$

When is MAP same as MLE?

MAP estimation for Binomial distribution

Coin flip problem: Likelihood is Binomial

$$P(\mathcal{D} \mid \theta) = \binom{n}{\alpha_H} \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

If the prior is Beta distribution,

$$P(\theta) = \frac{\theta^{\beta_H-1} (1 - \theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$$

\Rightarrow posterior is Beta distribution

Beta function: $B(x, y) = \int_0^1 t^{x-1} (1 - t)^{y-1} dt$

MAP estimation for Binomial distribution

Likelihood is Binomial: $P(\mathcal{D} | \theta) = \binom{n}{\alpha_H} \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$

Prior is Beta distribution: $P(\theta) = \frac{\theta^{\beta_H-1} (1 - \theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$

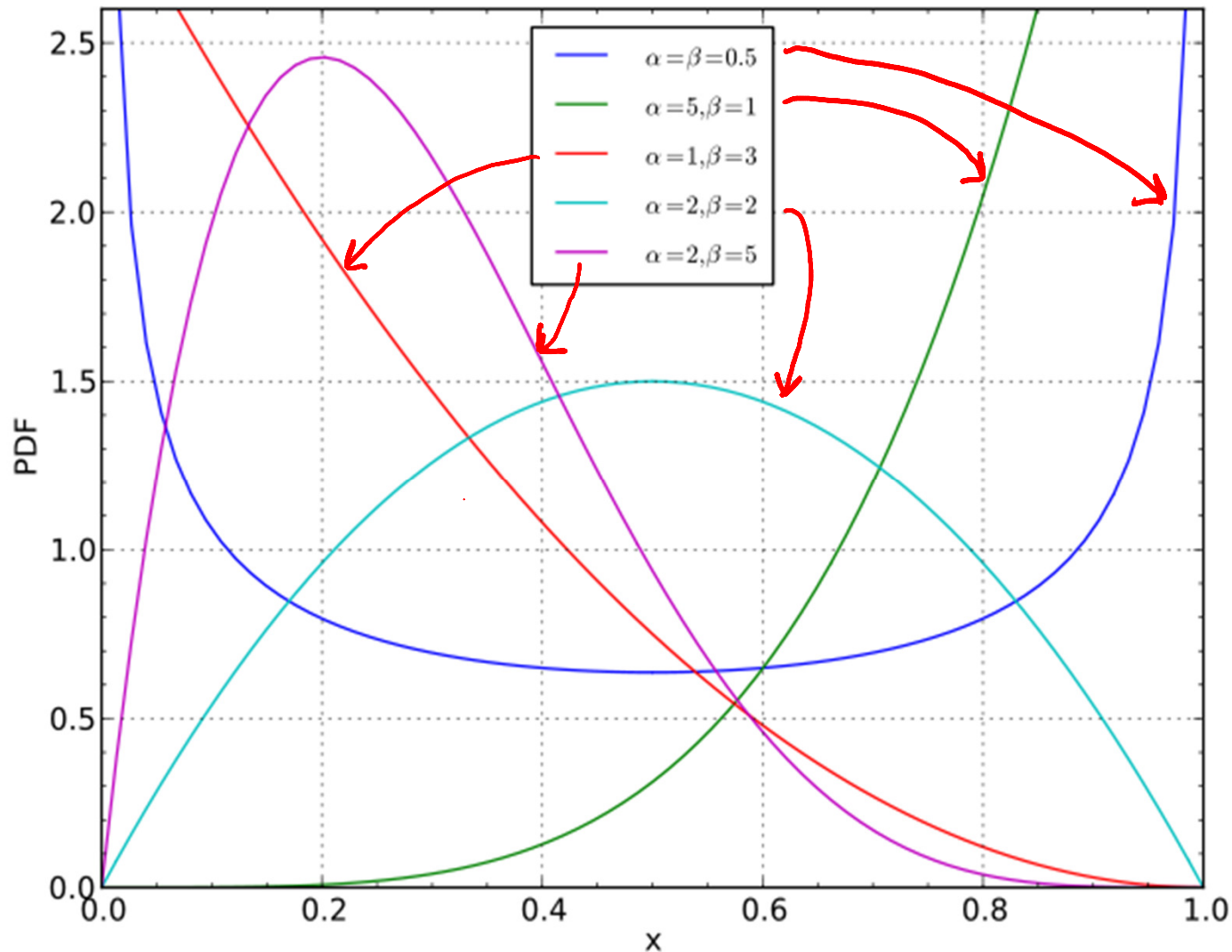
⇒ posterior is Beta distribution

$$P(\theta | D) \sim \text{Beta}(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

$P(\theta)$ and $P(\theta | D)$ have the same form! [Conjugate prior]

$$\begin{aligned} \hat{\theta}_{MAP} &= \arg \max_{\theta} P(\theta | D) = \arg \max_{\theta} P(D | \theta) P(\theta) \\ &= \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2} \end{aligned}$$

Beta distribution

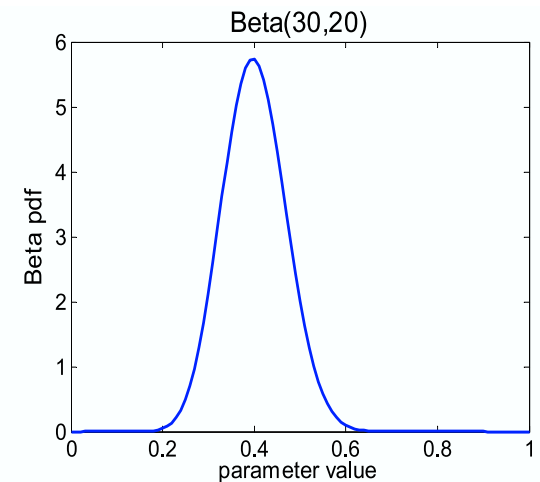
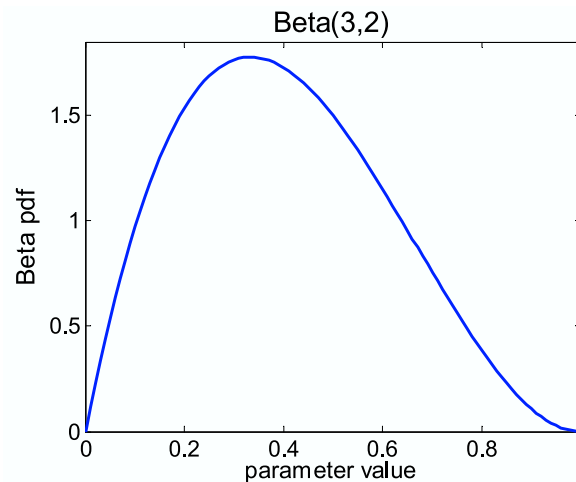
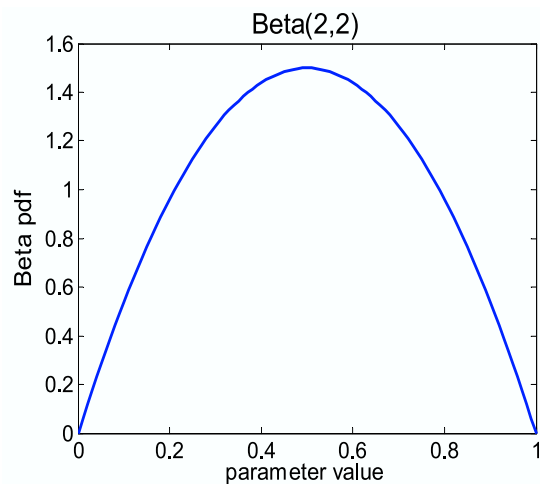


More concentrated as values of α, β increase

Beta conjugate prior

$$P(\theta) \sim \text{Beta}(\beta_H, \beta_T)$$

$$P(\theta|D) \sim \text{Beta}(\beta_H + \alpha_H, \beta_T + \alpha_T)$$



As $n = \alpha_H + \alpha_T$
increases

As we get more samples, effect of prior is “washed out”

From Binomial to Multinomial

Example: Dice roll problem (6 outcomes instead of 2)

Likelihood is $\sim \text{Multinomial}(\theta = \{\theta_1, \theta_2, \dots, \theta_k\})$

$$P(\mathcal{D} \mid \theta) = \theta_1^{\alpha_1} \theta_2^{\alpha_2} \dots \theta_k^{\alpha_k}$$



If prior is Dirichlet distribution,

$$P(\theta) = \frac{\prod_{i=1}^k \theta_i^{\beta_i - 1}}{B(\beta_1, \dots, \beta_k)} \sim \text{Dirichlet}(\beta_1, \dots, \beta_k)$$

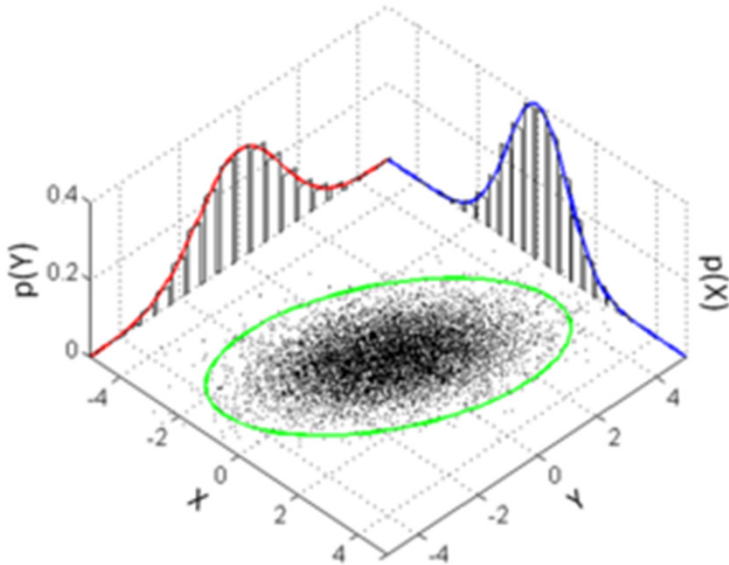
Then posterior is Dirichlet distribution

$$P(\theta \mid D) \sim \text{Dirichlet}(\beta_1 + \alpha_1, \dots, \beta_k + \alpha_k)$$

For Multinomial, conjugate prior is Dirichlet distribution.

http://en.wikipedia.org/wiki/Dirichlet_distribution

Conjugate prior for Gaussian?



$$(2\pi)^{-\frac{k}{2}} |\Sigma|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{x}-\mu)'\Sigma^{-1}(\mathbf{x}-\mu)},$$

Conjugate prior on mean: **Gaussian**

Conjugate prior on covariance matrix: **Inverse Wishart**

$$\frac{|\Psi|^{\frac{\nu}{2}}}{2^{\frac{\nu p}{2}} \Gamma_p\left(\frac{\nu}{2}\right)} |\mathbf{X}|^{-\frac{\nu+p+1}{2}} e^{-\frac{1}{2} \text{tr}(\Psi \mathbf{X}^{-1})}$$

Bayesians vs. Frequentists

You are no good when sample is small



You give a different answer for different priors