TMA4267 - Linear statistical models

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Contents

Introduction	2
Course progress Keywords to know	2
Part 1 -	3
Multiple hypothesis testing	7
Design of experiment	8

Introduction

This is a brief summary of the course TMA4267 about linear statistical models. It includes the main content from the lecture held by ... recorded in, where some examples etc... are excluded.

The purpose of the notes is to give a good overview of the syllabus. I intend to add summaries of the lectures as I review them. I hope to include insights from projects / exercises where it is appropriate.

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Course progress

• First reading	□ Lecture 3	□ Lecture 15
✓ Lecture 1-19	\Box Lecture 4	☐ Lecture 16
☐ Lecture 20	\Box Lecture 5	□ Lecture 17
☐ Lecture 21	☐ Lecture 6	☐ Lecture 18
☐ Lecture 22	☐ Lecture 7	☐ Lecture 19
☐ Lecture 23	☐ Lecture 8	☐ Lecture 20
□ Lecture 24	☐ Lecture 9	☐ Lecture 21
	\Box Lecture 10	
☐ Lecture 25	\Box Lecture 11	☐ Lecture 22
• Skikkelig TeXing	\Box Lecture 12	☐ Lecture 23
\square Lecture 1	☐ Lecture 13	☐ Lecture 24
\square Lecture 2	☐ Lecture 14	\Box Lecture 25

Keywords to know

Part 1 -

 $\widehat{\boldsymbol{\beta}}, \, \boldsymbol{\beta}, \, \boldsymbol{\sigma}, \, \widehat{\boldsymbol{\sigma}}, \, \boldsymbol{\varepsilon}, \, \widehat{\boldsymbol{\varepsilon}}$

$$\widehat{\beta}$$
, β , σ , $\widehat{\sigma}$, ε , $\widehat{\varepsilon}$

theorem - trace formula

Theorem 1. (Trace formula)

$$\varepsilon(Y^TCY) = \operatorname{tr}(C\Sigma) + \mu^TC\mu$$

Proof. TODO:

theorem - ...

Lecture 8

Theorem 2. $\mathbf{Z} \sim N(0, I)$ and \mathbf{R} symmetric and idempotent of rank r. Then

$$oldsymbol{Z}^T oldsymbol{R} oldsymbol{Z} \sim \chi_r^2.$$

Lecture 9

Assumptions

- 1. \boldsymbol{X} is of cull column rank
- 2. $E\boldsymbol{\varepsilon} = \mathbf{0}$
- 3. Homostochastic: $Var(\varepsilon_i) = 0 \quad \forall i$.
- 4. If X is random, then 2 and 3 are conditioned on X.
- 5. Normality of errors: $\varepsilon \sim N(0, \sigma^2 I_n)$.

... obtain least squares estimators $\widehat{\pmb{\beta}}, \widehat{\pmb{\sigma}}^2$ of $\pmb{\beta}, \pmb{\sigma}^2$

Residuals ...

Parameter estimation

Two approaches: LSE and MLE ...

$$\widehat{oldsymbol{eta}} = rg \min_{oldsymbol{eta} \in \mathbb{R}^{k+1}} \sum_{i=1}^n (Y_i - oldsymbol{x}_i^T oldsymbol{eta})^2$$

... deducing that LSE and MLE give the same result ...

...

Hat matrix

ghifiodeifgoerjfkdworw9u0gryhj

Du fulgte ikke med nei

Lecture 10

Lecture 11

Lecture 12

questions about independence. Detour into sigma algebras etc ...

Theorem 3. Suppose X, Y are independent random variables and that f, g are two measurable functions. Then f(X), g(Y) are also independent.

ANOVA - Analysis of variance

Theorem 4. (ANOVA decomposition) Assuming the necesarry assumptions,

$$\underbrace{\sum_{i=1}^{n} (Y_i - \bar{Y})}_{\text{SST}} = \underbrace{\sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})}_{\text{SSR}} + \underbrace{\sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2}_{\text{SSE}}.$$

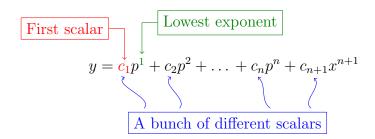
Proof. TODO: there aint space in the margin

R2 score ...

Lecture 13

Fictional model

"Fictional model" using x_{ij} as response for some fixed feature j.



... TODO:

General F-test

TODO: important to have on yellow paper

We set up a much more general problem. Let $A \in \mathbb{R}^{r \times p}$, r < p, rank(A) = r, $\mathbf{d} \in \mathbb{R}^d$. We test the hypothesis:

$$H_0: A\boldsymbol{\beta} = \boldsymbol{d}, \qquad \quad H_1: A\boldsymbol{\beta} \neq \boldsymbol{d}.$$

Some special cases of this general setup are.

1. r = 1, d = 0, A = (0, ..., 1, ..., 0) with 1 at index i, gives the test

$$\mathbf{H}_0: \beta_i = 0, \qquad \quad \mathbf{H}_1: \beta_i \neq 0.$$

2. $r = 1, d = 0, A = (0, \dots, 1, \dots, -1, \dots, 0)$ with 1 at index i and -1 at index j, gives the test $H_0: \beta_i = \beta_j, \qquad H_1: \beta_i \neq \beta_j.$

3. $r = k, d = \mathbf{0} \in \mathbb{R}^k, A = (\mathbf{0}, \operatorname{diag}(1)) \in \mathbb{R}^{k \times p}$, gives the test $H_0: \beta_i = 0 \quad \forall i \in \{1, \dots, k\}, \qquad H_1: \beta_i \neq 0 \text{ for some } i \in \{1, \dots, k\}.$

Lecture 14

Let \mathcal{B} be the space of $\boldsymbol{\beta}$ satisfying H_0 . The restricted problem is:

$$\widehat{\boldsymbol{\beta}}^R = \underset{\boldsymbol{\beta} \in \mathcal{B}}{\operatorname{arg min}} (\boldsymbol{Y} - \boldsymbol{X}\boldsymbol{\beta})^T (\boldsymbol{Y} - \boldsymbol{X}\boldsymbol{\beta}).$$

Using lagrange multipliers and a bag of tricks, we obtain:

$$\widehat{\boldsymbol{\beta}}^R = \widehat{\boldsymbol{\beta}} - (\boldsymbol{X}^T \boldsymbol{X})^{-1} \boldsymbol{A}^T (\boldsymbol{A} (\boldsymbol{X}^T \boldsymbol{X})^{-1} \boldsymbol{A}^T)^{-1} (\boldsymbol{A} \widehat{\boldsymbol{\beta}} - \boldsymbol{d}).$$

Denoting $\Delta = \widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}^R$, we find:

$$SSE^R = SSE + \Delta^T \boldsymbol{X}^T \boldsymbol{X} \Delta$$

... IMPORTANT: the concrete expressions for the F statistic...

We claim that the under H_0 , we have

$$F = \frac{SSE^R - SSE/r}{SSE/(n-p)} \sim F_{r,n-p}.$$

Proof. what the

Lecture 15

... example ...

Transformations of data

Motivation: ...

box cox transformation

variance stabilising transformation

Suppose $\mu = \varepsilon(Y_i)$ and that $Var(Y_i)$ depends on μ

Lecture 16

...

Lecture 17

Suppose k covariates. Then 2^k possible models from maximal:

$$Y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_k x_{ik}.$$

to minimal:

$$Y_i = \beta_0$$
.

We want to arrive at a compromise between simplisity and goodness of fit.

1. Adjusted coefficient of determination:

$$R_{\text{adj}}^2 = 1 - \frac{\text{SSE}/(n-k-1)}{\text{SST}/(n-1)}$$

- 2.
- 3.
- 4.

example...

Multiple hypothesis testing

motivation ...

Lecture 18

. . .

FWER = probability of at least one false positive finding

... two representations

The Bonferrony method

The Šidák method

...

example 2019

. . .

example 2020

Lecture 19

Example with three groups and their means \dots rewrite to regression problem \dots

Analysis of varance (ANOVA)

p treatments, samples \dots

Lecture 20

... cont ... + brief on two way ANOVA

Design of experiment

two level factorial design \dots

Vi tester en sitering [1].

References

 $[1] \;\;$ test. $test\ bok.$ Ed. by Trond. UiO, 2030.

\mathbf{Index}

(ANOVA decomposition), 4 (Trace formula), 3

Bonferrony method, 7

 $\check{\mathrm{S}}\mathrm{id}\acute{\mathrm{a}}\mathrm{k}$ method, 7