A non-asymptotic approach for model selection via penalization in mixture of experts models

TrungTin Nguyen¹, Hien Duy Nguyen², Faicel Chamroukhi³, Florence Forbes¹ ¹Inria Grenoble Rhone-Alpes, France, ²University of Queensland, Australia, ³UNICAEN, LMNO UMR CNRS, France.

Learning nonlinear regression models from complex data using GLoME models

Random sample: $(\mathbf{X}_i, \mathbf{Y}_i)_{i \in [n]} \in (\mathcal{X} \times \mathcal{Y})^n \subset (\mathbb{R}^D \times \mathbb{R}^L)^n$ of the multivariate response $\mathbf{Y} = (\mathbf{Y}_j)_{j \in [L]}$ and the set of covariates $\mathbf{X} = (\mathbf{X}_j)_{j \in [D]}$ with the corresponding observed values $(\mathbf{x}_i, \mathbf{y}_i)_{i \in [n]}, [n] := \{1, \dots, n\}$ (potentially $D \gg L$), arising from an unknown conditional density s_0 .

Our proposal: approximating s_0 by a Gaussian-gated Localized Mixture of Experts (GLoME) model due to its flexibility and effectiveness [3, 4, 5]:

 $s_{\boldsymbol{\psi}_{K,d}}(\mathbf{x}|\mathbf{y}) = \sum_{k=1}^{n} \underbrace{\mathbf{g}_{k}\left(\mathbf{y};\boldsymbol{\omega}\right)}_{\text{Gaussian-gated network}} \times \underbrace{\mathcal{N}_{D}\left(\mathbf{x};\boldsymbol{v}_{k,d}(\mathbf{y}),\boldsymbol{\Sigma}_{k}\right)}_{\text{Gaussian expert}}, \quad \mathbf{g}_{k}\left(\mathbf{y};\boldsymbol{\omega}\right) = \frac{\boldsymbol{\pi}_{k}\mathcal{N}_{L}\left(\mathbf{y};\mathbf{c}_{k},\boldsymbol{\Gamma}_{k}\right)}{\sum_{l=1}^{K} \boldsymbol{\pi}_{l}\mathcal{N}_{L}\left(\mathbf{y};\mathbf{c}_{l},\boldsymbol{\Gamma}_{l}\right)}, \forall k \in [K], K \in \mathbb{N}^{\star}, \text{where:}$

 $\boldsymbol{\psi}_{K,d} = (\boldsymbol{\omega},\boldsymbol{v},\boldsymbol{\Sigma}) \in \boldsymbol{\Omega}_K \times \boldsymbol{\Upsilon}_{K,d} \times \boldsymbol{V}_K =: \boldsymbol{\Psi}_{K,d}, \ \boldsymbol{\omega} = (\boldsymbol{\pi},\boldsymbol{c},\boldsymbol{\Gamma}) \in (\boldsymbol{\Pi}_{K-1} \times \mathbf{C}_K \times \boldsymbol{V}_K') =: \boldsymbol{\Omega}_K, \boldsymbol{\Pi}_{K-1} = \left\{ (\boldsymbol{\pi}_k)_{k \in [K]} \in (\mathbb{R}^+)^K, \sum_{k=1}^K \boldsymbol{\pi}_k = 1 \right\}, \ \mathbf{C}_K / \boldsymbol{\Upsilon}_{K,d} := (\boldsymbol{\omega},\boldsymbol{v},\boldsymbol{\Sigma}) \in \boldsymbol{\Omega}_K \times \boldsymbol{\Upsilon}_{K,d} \times \boldsymbol{V}_K =: \boldsymbol{\Psi}_{K,d}, \ \boldsymbol{\omega} = (\boldsymbol{\pi},\boldsymbol{c},\boldsymbol{\Gamma}) \in (\boldsymbol{\Pi}_{K-1} \times \mathbf{C}_K \times \boldsymbol{V}_K') =: \boldsymbol{\Omega}_K, \boldsymbol{\Pi}_{K-1} = \left\{ (\boldsymbol{\pi}_k)_{k \in [K]} \in (\mathbb{R}^+)^K, \sum_{k=1}^K \boldsymbol{\pi}_k = 1 \right\}, \ \mathbf{C}_K / \boldsymbol{\Upsilon}_{K,d} := (\boldsymbol{\pi}_k)_{k \in [K]} \in \boldsymbol{\Omega}_K \times \boldsymbol{\Upsilon}_{K,d} \times \boldsymbol{V}_K =: \boldsymbol{\Psi}_{K,d}, \ \boldsymbol{\omega} = (\boldsymbol{\pi},\boldsymbol{c},\boldsymbol{\Gamma}) \in (\boldsymbol{\Pi}_{K-1} \times \mathbf{C}_K \times \boldsymbol{V}_K') =: \boldsymbol{\Omega}_K, \boldsymbol{\Pi}_{K-1} = \left\{ (\boldsymbol{\pi}_k)_{k \in [K]} \in (\mathbb{R}^+)^K, \boldsymbol{\Sigma}_{K-1} \times \boldsymbol{\Sigma}_K \times \boldsymbol{\Sigma}_K \right\}$ K-tuples of mean vectors/functions of size $L \times 1/D \times 1$, V'_K/V_K : K-tuples of elements in $\mathcal{S}_L^{++}/\mathcal{S}_D^{++}$ (space of symmetric positive-definite matrices).

Main contributions: • Model selection criterion: choosing number of mixture components and mean functions' degree via a penalized maximum likelihood estimator.

• Finite-sample oracle inequality: establishing non-asymptotic risk bounds provided a lower bound on the penalty holds.

Boundedness assumptions

$$\widetilde{\mathbf{\Omega}}_{K} = \left\{ \boldsymbol{\omega} \in \mathbf{\Omega}_{K} : \forall k \in [K], \| \mathbf{c}_{k} \|_{\infty} \leq A_{\mathbf{c}}, \\
0 < a_{\mathbf{\Gamma}} \leq m \left(\mathbf{\Gamma}_{k} \right) \leq M \left(\mathbf{\Gamma}_{k} \right) \leq A_{\mathbf{\Gamma}}, 0 < a_{\mathbf{\pi}} \leq \boldsymbol{\pi}_{k} \right\}, \\
m(\mathbf{\Gamma}_{k})/M(\mathbf{\Gamma}_{k}): \text{ the smallest/largest eigenvalues of } \mathbf{\Gamma}_{k}, \\
\mathbf{\Upsilon}_{b,d} = \left\{ \mathbf{y} \mapsto \left(\sum_{i=1}^{d} \boldsymbol{\alpha}_{i}^{(j)} \varphi_{\mathbf{\Upsilon},i}(\mathbf{y}) \right)_{j \in [D]} : \| \boldsymbol{\alpha} \|_{\infty} \leq T_{\mathbf{\Upsilon}} \right\}, \\
\mathbf{\Upsilon}_{K,d} = \bigotimes_{k \in [K]} \mathbf{\Upsilon}_{k,d} = \mathbf{\Upsilon}_{b,d}^{K}, T_{\mathbf{\Upsilon}} \in \mathbb{R}^{+}, \\
(\varphi_{\mathbf{\Upsilon},i})_{i \in [d]}: \text{ collection of bounded functions on } \mathcal{Y}, \\
\mathbf{V}_{\mathbf{T},i} = \left\{ (\mathbf{\Sigma}_{i}) \right\}_{i \in [d]} = \left\{ \mathbf{P}_{i} \mathbf{P}_{i} \mathbf{\Lambda}_{i} \mathbf{P}^{\top} \right\}_{i \in [d]}.$$

$$\begin{aligned} & \left(\varphi_{\Upsilon,i} \right)_{i \in [d]} \text{: collection of bounded functions on } \mathcal{Y}, \\ & \mathbf{V}_{K} = \left\{ \left(\mathbf{\Sigma}_{k} \right)_{k \in [K]} = \left(B_{k} \mathbf{P}_{k} \mathbf{A}_{k} \mathbf{P}_{k}^{\top} \right)_{k \in [K]} : \\ & 0 < B_{-} \leq B_{k} \leq B_{+}, \ \mathbf{P}_{k} \in SO(D), \ \mathbf{A}_{k} \in \mathcal{A} \left(\lambda_{-}, \lambda_{+} \right) \right\}, \\ & B_{k} = |\mathbf{\Sigma}_{k}|^{1/D} \text{: volume, } SO(D) \text{: eigenvectors of } \mathbf{\Sigma}_{k}, \\ & \mathcal{A} \left(\lambda_{-}, \lambda_{+} \right) \text{: set of diagonal matrices of normalized} \\ & \text{eigenvalues of } \mathbf{\Sigma}_{k} \text{ s.t. } \forall i \in [D], 0 < \lambda_{-} \leq (\mathbf{A}_{k})_{i,i} \leq \lambda_{+}, \\ & \mathbf{m} \in \mathcal{M} = \left\{ (K, d) : K \in [K_{\text{max}}], d \in [d_{\text{max}}] \right\}, \\ & S_{\mathbf{m}} = \left\{ (\mathbf{x}, \mathbf{y}) \mapsto s_{\psi_{K,d}}(\mathbf{x}|\mathbf{y}) =: s_{\mathbf{m}}(\mathbf{x}|\mathbf{y}) : \\ & \psi_{K,d} \in \widetilde{\Omega}_{K} \times \Upsilon_{K,d} \times \mathbf{V}_{K} =: \widetilde{\Psi}_{K} \right\}. \end{aligned}$$

Model selection procedure

GLLiM model: finding the best data-driven model among $(S_{\mathbf{m}}^*)_{m \in \mathcal{M}}$, $\mathcal{M} = [K_{\max}] \times \{1\}$, based on $(\mathbf{x}_i, \mathbf{y}_i)_{i \in [n]}$ arising from a forward conditional density s_0^* .

- 1. For each $\mathbf{m} \in \mathcal{M}$: estimate the forward MLE $(\widehat{s}_{\mathbf{m}}^*(\mathbf{y}_i|\mathbf{x}_i))_{i\in[N]}$ by inverse MLE $\widehat{s}_{\mathbf{m}}$ via an inverse regression trick by GLLiM-EM algorithm.
- 2. Calculate PMLE $\widehat{\mathbf{m}}$ with pen(\mathbf{m}) = $\kappa \dim(S_{\mathbf{m}}^*)$. Large enough but not explicit value for $\kappa!$ Asymptotic: AIC: $\kappa = 1$; BIC: $\kappa = \frac{\ln n}{2}$. Non-asymptotic: partially justification for slope heuristic criterion in a finite-sample setting.

References

- Jean-Patrick Baudry, Cathy Maugis, and Bertrand Michel. Slope heuristics: overview and implementation. Statistics and Computing, 22(2):455–470, 2012.
- Antoine Deleforge, Florence Forbes, and Radu Ho-High-dimensional regression with gaussian mixtures and partially-latent response variables. Statistics and Computing, 25(5):893–911, 2015.
- Nhat Ho, Chiao-Yu Yang, and Michael I Jordan. Convergence Rates for Gaussian Mixtures of Experts. arXiv preprint arXiv:1907.04377, 2019.
- Hien Duy Nguyen, TrungTin Nguyen, Faicel Chamroukhi, and Geoffrey John McLachlan. Approximations of conditional probability density functions in Lebesgue spaces via mixture of experts models. Journal of Statistical Distributions and Applications, 8(1):13, 2021.
- Trung Tin Nguyen, Hien Duy Nguyen, Faicel Chamroukhi, and Florence Forbes. A non-asymptotic penalization criterion for model selection in mixture of experts models. arXiv preprint arXiv:2104.02640, 2021.

Non-asymptotic oracle inequality [5]

Theorem. Given a collection $(S_{\mathbf{m}})_{\mathbf{m}\in\mathcal{M}}$ of GLoME models, $\rho\in(0,1),\ C_1>1$, assume that $\Xi = \sum_{\mathbf{m} \in \mathcal{M}} e^{-z_{\mathbf{m}}} < \infty, z_{\mathbf{m}} \in \mathbb{R}^+, \forall \mathbf{m} \in \mathcal{M}, \text{ and there exist constants } C \text{ and } \kappa(\rho, C_1) > 0 \text{ s.t.}$ $\forall \mathbf{m} \in \mathcal{M}, \ \text{pen}(\mathbf{m}) \geq \kappa(\rho, C_1) [(C + \ln n) \dim(S_{\mathbf{m}}) + z_{\mathbf{m}}]. \ \text{Then, a PMLE-} \widehat{s}_{\widehat{\mathbf{m}}}, \ \text{defined by } \widehat{\mathbf{m}} = 0$ $\arg\min_{\mathbf{m}\in\mathcal{M}} \left(\sum_{i=1}^{n} -\ln\left(\widehat{s}_{\mathbf{m}}\left(\mathbf{x}_{i}|\mathbf{y}_{i}\right)\right) + \operatorname{pen}(\mathbf{m})\right), \ \widehat{s}_{\mathbf{m}} = \arg\min_{s_{\mathbf{m}}\in S_{\mathbf{m}}} \sum_{i=1}^{n} -\ln\left(s_{\mathbf{m}}\left(\mathbf{x}_{i}|\mathbf{y}_{i}\right)\right), \ \text{with}$ the loss $\operatorname{JKL}_{\rho}^{\otimes n}(s,t) = \mathbb{E}_{\mathbf{Y}_{[n]}} \left| \frac{1}{n} \sum_{i=1}^{n} \frac{1}{\rho} \operatorname{KL}\left(s\left(\cdot | \mathbf{Y}_{i}\right), \left(1-\rho\right) s\left(\cdot | \mathbf{Y}_{i}\right) + \rho t\left(\cdot | \mathbf{Y}_{i}\right)\right) \right|$, satisfies

$$\mathbb{E}_{\mathbf{Y}_{[n]}}\left[\mathrm{JKL}_{\rho}^{\otimes \mathrm{n}}\left(s_{0},\widehat{s}_{\widehat{\mathbf{m}}}\right)\right] \leq C_{1}\inf_{\mathbf{m}\in\mathcal{M}}\left(\inf_{s_{\mathbf{m}}\in S_{\mathbf{m}}}\mathrm{KL}^{\otimes \mathrm{n}}\left(s_{0},s_{\mathbf{m}}\right) + \frac{\mathrm{pen}(\mathbf{m})}{n}\right) + \frac{\kappa\left(\rho,C_{1}\right)C_{1}\Xi}{n}.$$

Numerical experiments

Well-Specified (WS): $s_0^* \in S_{\mathbf{m}}^*$,

$$s_0^*(y|x) = \frac{\mathcal{N}(x; 0.2, 0.1)\mathcal{N}(y; -\mathbf{5}x + \mathbf{2}, 0.09) + \mathcal{N}(x; 0.8, 0.15)\mathcal{N}(y; \mathbf{0.1}x, 0.09)}{\mathcal{N}(x; 0.2, 0.1) + \mathcal{N}(x; 0.8, 0.15)},$$

 $\mathbf{Misspecified} \ (\mathbf{MS}) : s_0^* \notin S_{\mathbf{m}}^*,$

$$s_0^*(y|x) = \frac{\mathcal{N}(x; 0.2, 0.1)\mathcal{N}(y; \boldsymbol{x^2} - \boldsymbol{6x} + \boldsymbol{1}, 0.09) + \mathcal{N}(x; 0.8, 0.15)\mathcal{N}(y; -\boldsymbol{0.4x^2}, 0.09)}{\mathcal{N}(x; 0.2, 0.1) + \mathcal{N}(x; 0.8, 0.15)}.$$

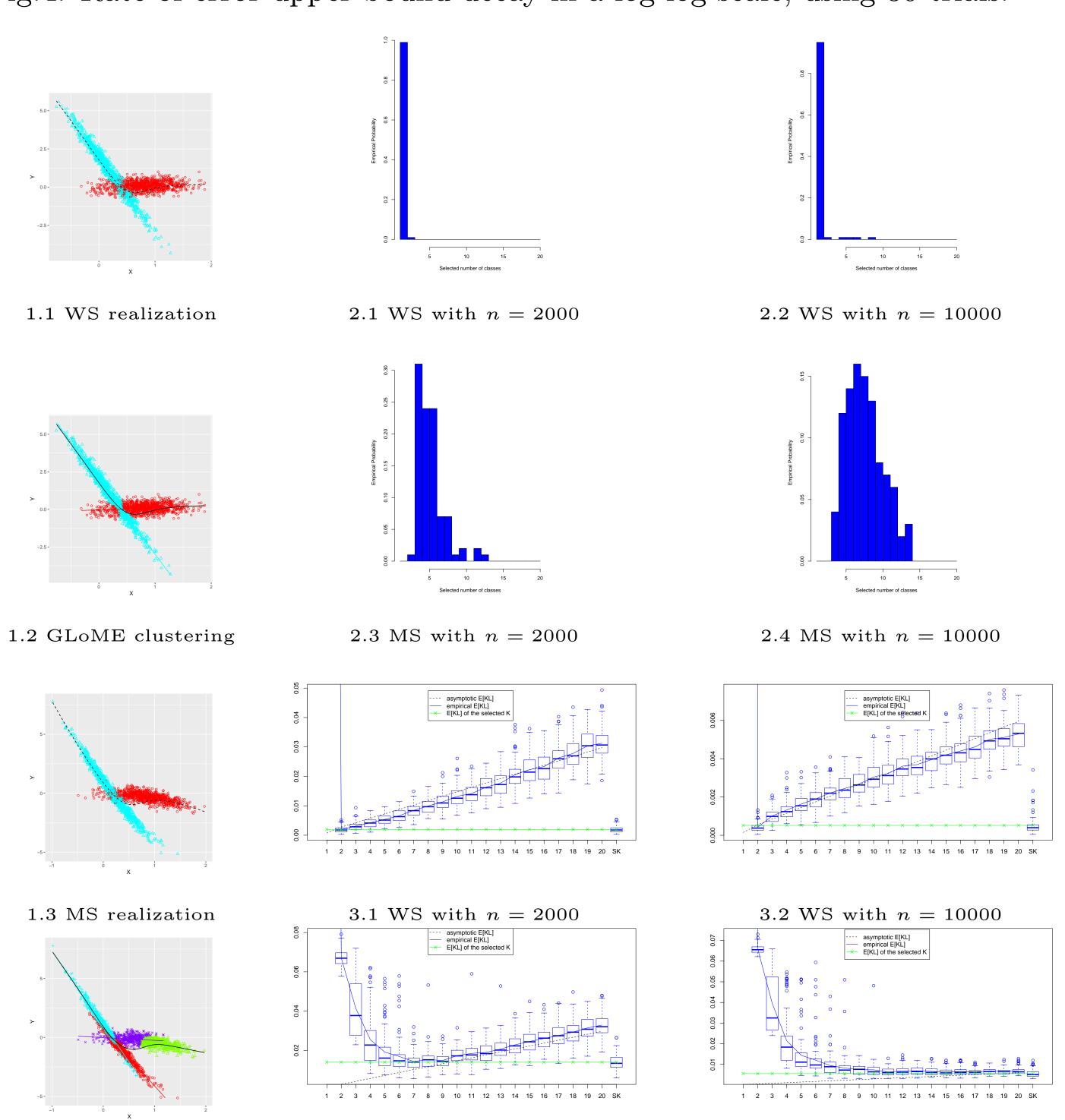
Estimation by EM (xLLiM package [2]) and model selection via the slope heuristic (capushe package [1]). Numerical results:

Fig.1: Clustering deduced from the estimated conditional density of GLoME via the Bayes' optimal allocation rule with 2000 data points. The dash and solid black curves present the true and estimated mean functions. Fig.2: Histogram of selected K using slope heuristic over 100 trials.

3.4 MS with n = 10000

Fig.3: Box-plot of the Kullback-Leibler divergence over 100 trials.

Fig.4: Rate of error upper bound decay in a log-log scale, using 30 trials.



3.3 MS with n = 2000

1.4 GLoME clustering

