

A non-asymptotic model selection in mixture of experts models

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Learning nonlinear regression models for heterogeneous data using GLoME models

Random samples: $(\mathbf{X}_i, \mathbf{Y}_i)_{i \in [n]} \in (\mathcal{X} \times \mathcal{Y})^n \subset (\mathbb{R}^D \times \mathbb{R}^L)^n$ of the multivariate response $\mathbf{Y} = (\mathbf{Y}_j)_{j \in [L]}$ and the set of covariates $\mathbf{X} = (\mathbf{X}_j)_{j \in [D]}$ with the corresponding observed values $(\mathbf{x}_i, \mathbf{y}_i)_{i \in [n]}$, $[n] := \{1, \dots, n\}$, arising from an unknown conditional density s_0 .

Our proposal: approximating s_0 by a **Gaussian-gated localized mixture of experts (GLoME)** model due to its flexibility and effectiveness [3, 4]:

$$s_{\psi_K}(\mathbf{x}|\mathbf{y}) = \sum_{k=1}^K \underbrace{\mathbf{g}_k(\mathbf{y}; \boldsymbol{\omega})}_{\text{Gaussian gating function}} \underbrace{\Phi_D(\mathbf{x}; \mathbf{v}_k(\mathbf{y}), \boldsymbol{\Sigma}_k)}_{\text{Gaussian expert}}, \quad \mathbf{g}_k(\mathbf{y}; \boldsymbol{\omega}) = \frac{\pi_k \Phi_L(\mathbf{y}; \mathbf{c}_k, \boldsymbol{\Gamma}_k)}{\sum_{j=1}^K \pi_j \Phi_L(\mathbf{y}; \mathbf{c}_j, \boldsymbol{\Gamma}_j)}, \forall k \in [K], K \in \mathbb{N}^*, \text{ where:}$$

$\psi_K = (\boldsymbol{\omega}, \mathbf{v}, \boldsymbol{\Sigma}) \in \boldsymbol{\Omega}_K \times \boldsymbol{\Upsilon}_K \times \mathbf{V}_K =: \boldsymbol{\Psi}_K$, $\boldsymbol{\omega} = (\boldsymbol{\pi}, \mathbf{c}, \boldsymbol{\Gamma}) \in (\boldsymbol{\Pi}_{K-1} \times \mathbf{C}_K \times \mathbf{V}'_K) =: \boldsymbol{\Omega}_K$, $\boldsymbol{\Pi}_{K-1} = \left\{ (\boldsymbol{\pi}_k)_{k \in [K]} \in (\mathbb{R}^+)^K, \sum_{k=1}^K \pi_k = 1 \right\}$, $\mathbf{C}_K / \boldsymbol{\Upsilon}_K$: K -tuples of **mean vectors/functions** of size $L \times 1 / D \times 1$, $\mathbf{V}'_K / \mathbf{V}_K$: K -tuples of elements in $\mathcal{S}_L^{++} / \mathcal{S}_D^{++}$ (space of symmetric positive-definite matrices).

- **Model selection problem:** estimating the number of mixture components via penalized maximum likelihood estimators.
- **Non-asymptotic oracle inequality:** providing a lower bound on the penalty such that our estimator satisfies an oracle inequality.

Boundedness assumptions

$$S_m = \left\{ \mathcal{X} \times \mathcal{Y} \ni (\mathbf{x}, \mathbf{y}) \mapsto s_{\psi_K}(\mathbf{x}|\mathbf{y}) =: s_m(\mathbf{x}|\mathbf{y}) : \right.$$

$$\left. \psi_K = (\boldsymbol{\omega}, \mathbf{v}, \boldsymbol{\Sigma}) \in \tilde{\boldsymbol{\Omega}}_K \times \boldsymbol{\Upsilon}_K \times \mathbf{V}_K =: \tilde{\boldsymbol{\Psi}}_K \right\},$$

$$\tilde{\boldsymbol{\Omega}}_K = \left\{ \boldsymbol{\omega} \in \boldsymbol{\Omega}_K : \forall k \in [K], \|\mathbf{c}_k\|_\infty \leq A_c, \right.$$

$$\left. 0 < a_\Gamma \leq m(\boldsymbol{\Gamma}_k) \leq M(\boldsymbol{\Gamma}_k) \leq A_\Gamma, 0 < a_\pi \leq \pi_k \right\},$$

$m(\mathbf{A})/M(\mathbf{A})$: smallest/largest eigenvalues of matrix \mathbf{A} ,

$$\boldsymbol{\Upsilon}_K = \boldsymbol{\Upsilon}_b^K, d_\Upsilon \in \mathbb{N}^*, T_\Upsilon \in \mathbb{R}^+,$$

$(\varphi_{\Upsilon,i})_{i \in [d_\Upsilon]}$: collection of bounded functions on \mathcal{Y} ,

$$\boldsymbol{\Upsilon}_b = \left\{ \mathbf{y} \mapsto \left(\sum_{i=1}^{d_\Upsilon} \alpha_i^{(j)} \varphi_{\Upsilon,i}(\mathbf{y}) \right)_{j \in [D]} : \|\boldsymbol{\alpha}\|_\infty \leq T_\Upsilon \right\},$$

$$\mathbf{V}_K = \left\{ \boldsymbol{\Sigma} = (\boldsymbol{\Sigma}_k)_{k \in [K]} = \left(B_k \mathbf{P}_k \mathbf{A}_k \mathbf{P}_k^\top \right)_{k \in [K]} : \right.$$

$$\left. B_- \leq B_k \leq B_+, \mathbf{P}_k \in SO(D), \mathbf{A}_k \in \mathcal{A}(\lambda_-, \lambda_+) \right\},$$

$$B_k = |\boldsymbol{\Sigma}_k|^{1/D}: \text{volume}, B_- \in \mathbb{R}^+, B_+ \in \mathbb{R}^+,$$

\mathbf{P}_k : eigenvectors of $\boldsymbol{\Sigma}_k \in$ special orthogonal $SO(D)$,

\mathbf{A}_k : diagonal matrix of normalized eigenvalues of $\boldsymbol{\Sigma}_k$,

such that $|\mathbf{A}_k| = 1$ and $0 < \forall i \in [D], \lambda_- \leq (\mathbf{A}_k)_{i,i} \leq \lambda_+$.

Model selection procedure

Goal: find the best model among $(S_m^*)_{m \in \mathcal{M}}$, $\mathcal{M} = [K_{\max}]$ based one $(\mathbf{x}_i, \mathbf{y}_i)_{i \in [n]}$ arising from an forward conditional density s_0^* .

1. Each $m \in \mathcal{M}$: estimate the forward MLE $(\hat{s}_m^*(\mathbf{y}_i|\mathbf{x}_i))_{i \in [N]}$ by inverse MLE \hat{s}_m via an **inverse regression trick** by GLLiM-EM algorithm (xLLiM package [2]).
2. Calculate η' -PMLE \hat{m} with $\text{pen}(m) = \kappa \dim(S_m^*)$.
3. **Large enough but not explicit value for κ !** Asymptotic criteria: AIC ($\kappa = 1$) and BIC ($\kappa = \frac{\ln n}{2}$). **Non-asymptotic criterion:** strong justification for **slope heuristic approach** (capushe package [1]) in a finite sample setting.

References

- [1] Jean-Patrick Baudry, Cathy Maugis, and Bertrand Michel. Slope heuristics: overview and implementation. *Statistics and Computing*, 22(2):455–470, 2012.
- [2] Antoine Deleforge, Florence Forbes, and Radu Horaud. High-dimensional regression with gaussian mixtures and partially-latent response variables. *Statistics and Computing*, 25(5):893–911, 2015.
- [3] Nhat Ho, Chiao-Yu Yang, and Michael I Jordan. Convergence Rates for Gaussian Mixtures of Experts. *arXiv preprint arXiv:1907.04377*, 2019.
- [4] Hien Duy Nguyen, TrungTin Nguyen, Faicel Chamroukhi, and Geoffrey McLachlan. Approximations of conditional probability density functions in Lebesgue spaces via mixture of experts models. *arXiv preprint arXiv:2012.02385*, 2020.
- [5] Trung Tin Nguyen, Hien Duy Nguyen, Faicel Chamroukhi, and Florence Forbes. A non-asymptotic penalization criterion for model selection in mixture of experts models. *arXiv preprint arXiv:2104.02640*, 2021.

Non-asymptotic oracle inequality [5]

Given a collection $(S_m)_{m \in \mathcal{M}}$ of GLoME models, $\rho \in (0, 1)$, $C_1 > 1$, assume that $\Xi = \sum_{m \in \mathcal{M}} e^{-z_m} < \infty, z_m \in \mathbb{R}^+, \forall m \in \mathcal{M}$, and there exist constants C and $\kappa(\rho, C_1) > 0$ s.t. $\forall m \in \mathcal{M}$, $\text{pen}(m) \geq \kappa(\rho, C_1) [(C + \ln n) \dim(S_m) + z_m]$. Then, the η' -PMLE $\hat{s}_{\hat{m}}$, defined by $\hat{m} = \text{argmin}_{m \in \mathcal{M}} (\sum_{i=1}^n -\ln(\hat{s}_m(\mathbf{x}_i|\mathbf{y}_i)) + \text{pen}(m)) + \eta'$, $\hat{s}_m = \text{argmin}_{s_m \in S_m} \sum_{i=1}^n -\ln(s_m(\mathbf{x}_i|\mathbf{y}_i))$, satisfies, for any $\rho \in (0, 1)$, $\text{JKL}_\rho^{\otimes n}(s, t) = \mathbb{E}_\mathbf{Y} \left[\frac{1}{n} \sum_{i=1}^n \frac{1}{\rho} \text{KL}(s(\cdot|\mathbf{Y}_i), (1-\rho)s(\cdot|\mathbf{Y}_i) + \rho t(\cdot|\mathbf{Y}_i)) \right]$,

$$\mathbb{E} [\text{JKL}_\rho^{\otimes n}(s_0, \hat{s}_{\hat{m}})] \leq C_1 \inf_{m \in \mathcal{M}} \left(\inf_{s_m \in S_m} \text{KL}^{\otimes n}(s_0, s_m) + \frac{\text{pen}(m)}{n} \right) + \frac{\kappa(\rho, C_1) C_1 \Xi}{n} + \frac{\eta + \eta'}{n}.$$

Numerical experiment

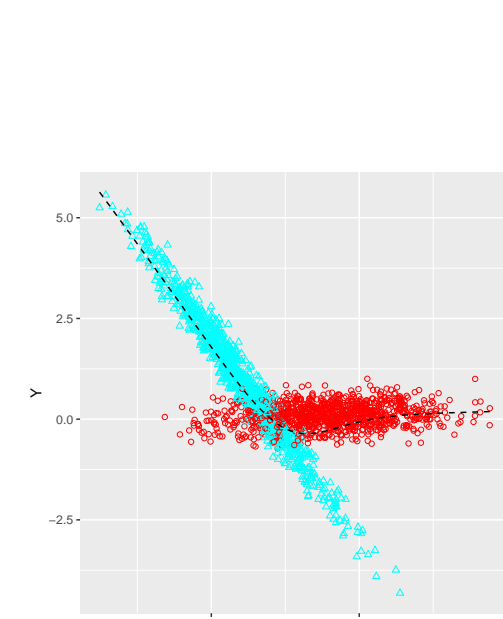
▮ **Well-Specified (WS)** : $s_0^* \in S_m^*$,

$$s_0^*(y|x) s_0^*(y|x) = \frac{\Phi(x; 0.2, 0.1) \Phi(y; -5x + 2, 0.09) + \Phi(x; 0.8, 0.15) \Phi(y; 0.1x, 0.09)}{\Phi(x; 0.2, 0.1) + \Phi(x; 0.8, 0.15)},$$

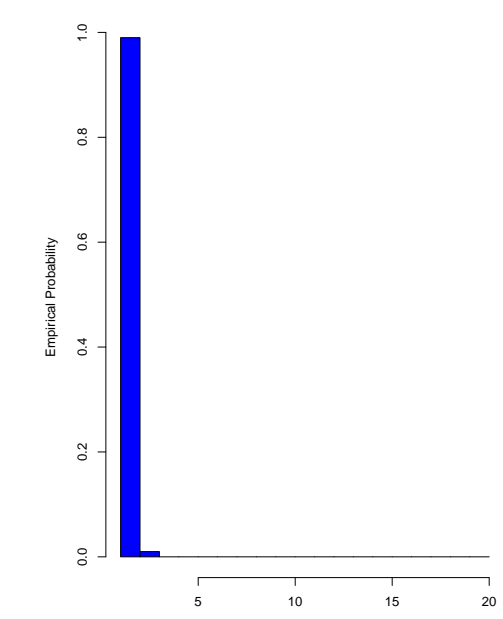
▮ **Misspecified (MS)** : $s_0^* \notin S_m^*$,

$$s_0^*(y|x) = \frac{\Phi(x; 0.2, 0.1) \Phi(y; x^2 - 6x + 1, 0.09) + \Phi(x; 0.8, 0.15) \Phi(y; -0.4x^2, 0.09)}{\Phi(x; 0.2, 0.1) + \Phi(x; 0.8, 0.15)}.$$

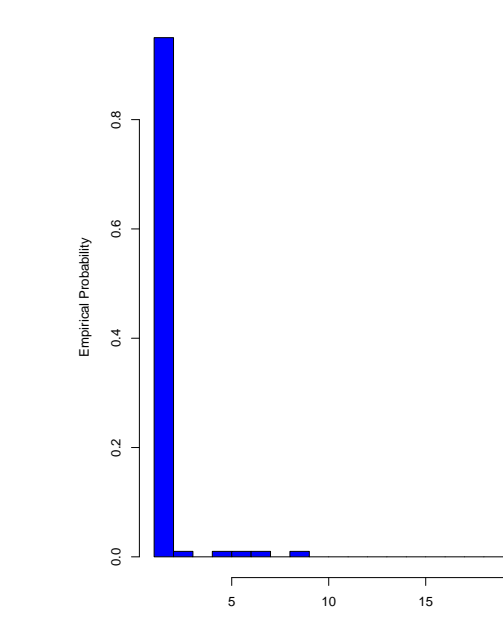
1. Clustering deduced from the estimated conditional density of GLoME by a MAP principle with 2000 data points. The dash and solid black curves present the true and estimated mean functions.
2. Histogram of selected K using slope heuristic over 100 trials.
3. Box-plot of the Kullback-Leibler divergence over 100 trials.
4. Rate of error decay in a log-log scale, using 30 trials.



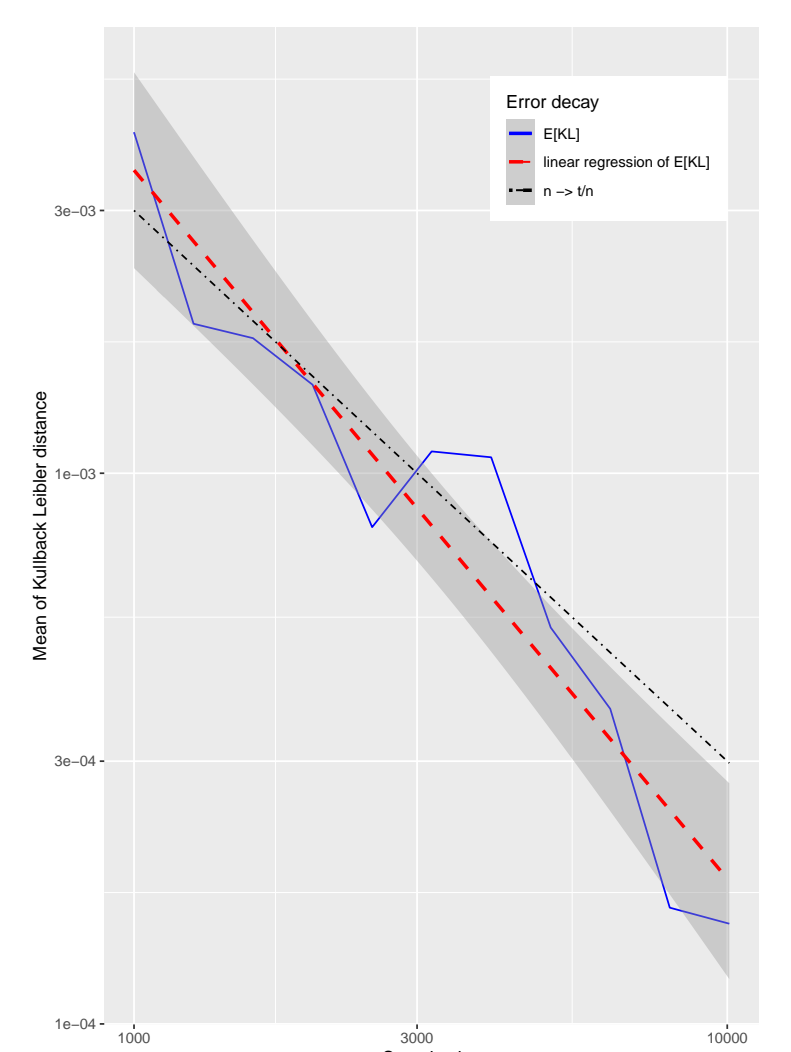
1.1 WS realization



2.1 WS with $n = 2000$



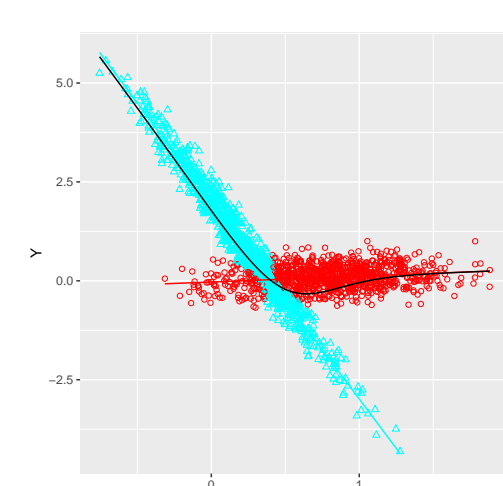
2.2 WS with $n = 10000$



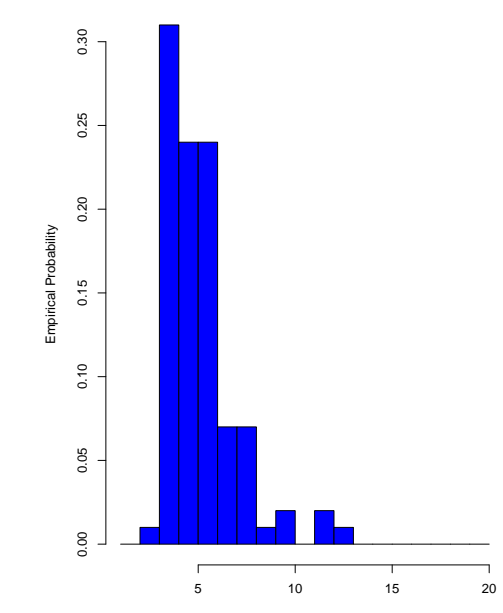
4.1 WS:

free regression's slope

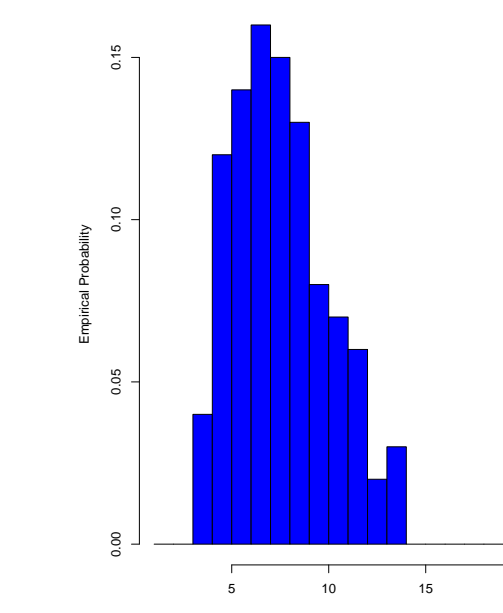
$$\approx -1.287 \text{ and } t = 3.$$



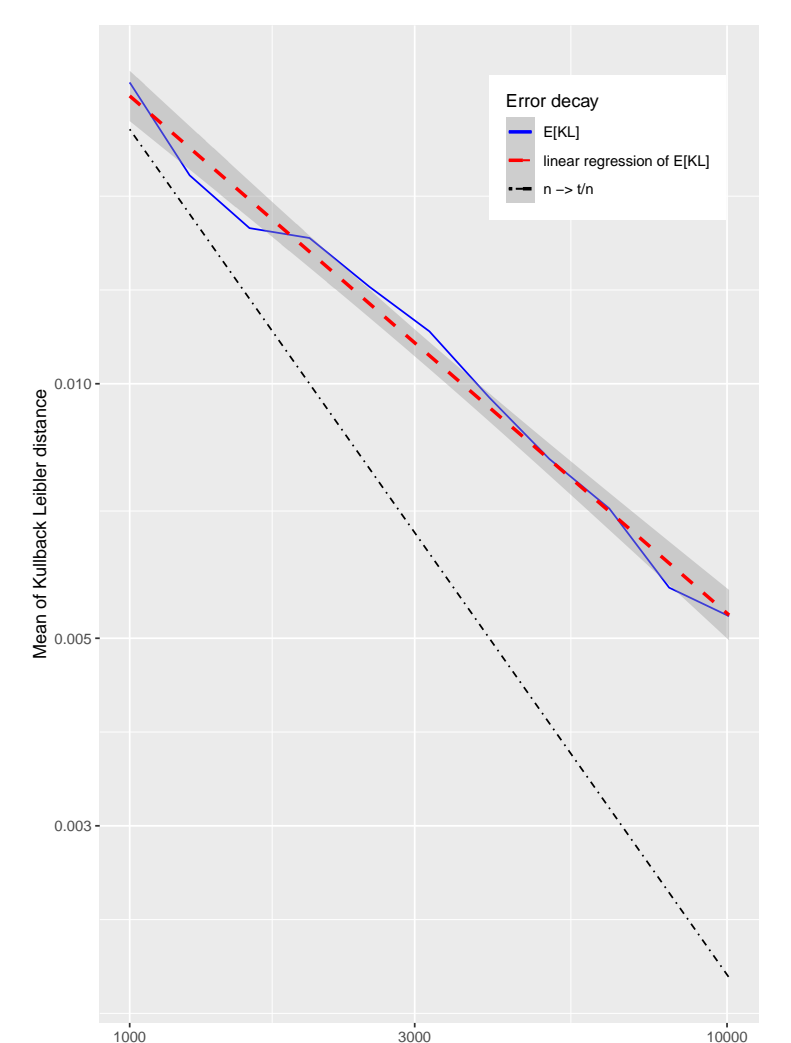
1.2 GLoME clustering



2.3 MS with $n = 2000$



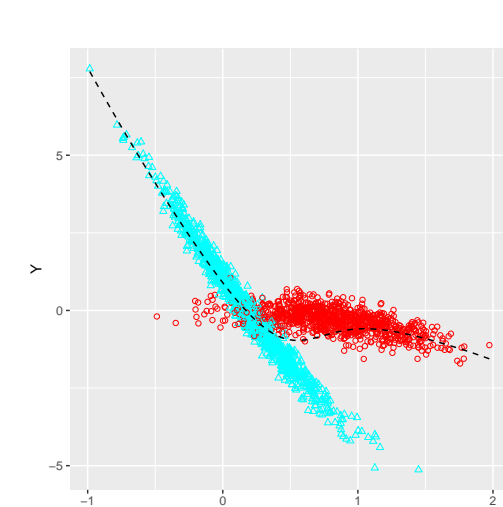
2.4 MS with $n = 10000$



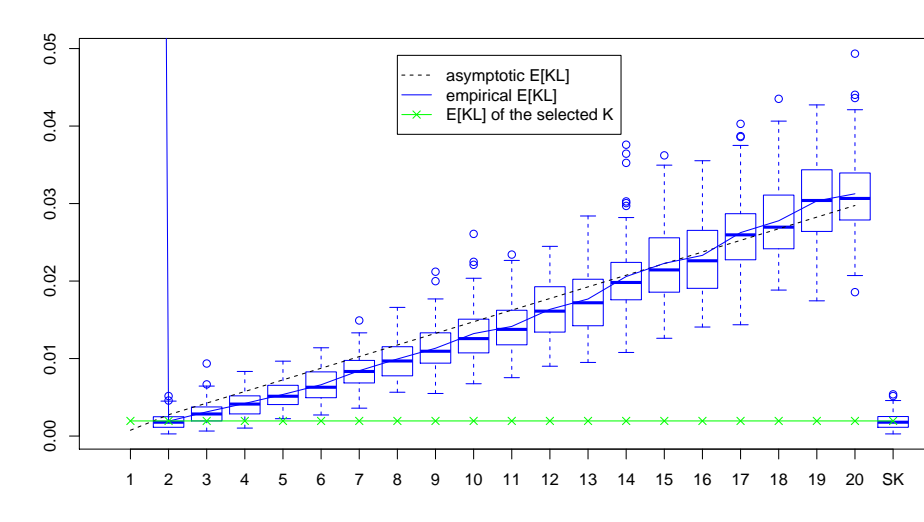
4.2 MS:

free regression's slope

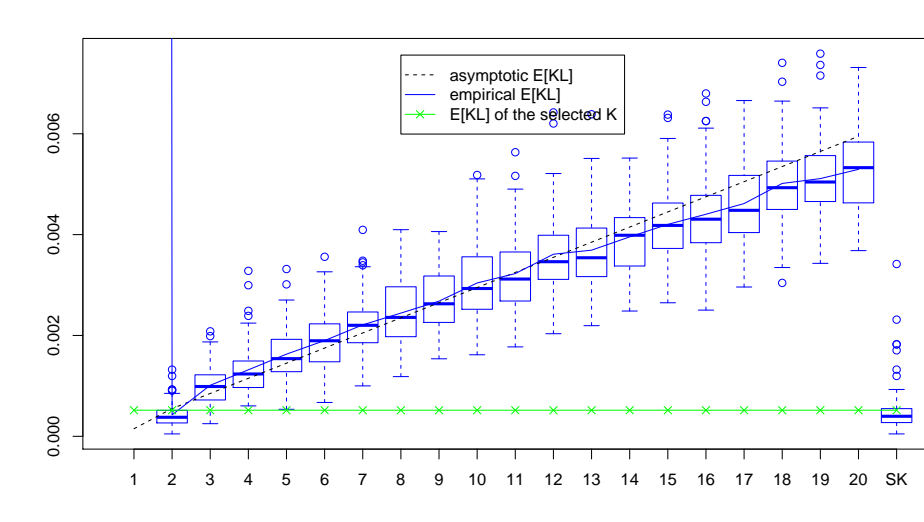
$$\approx -0.6120, t = 20.$$



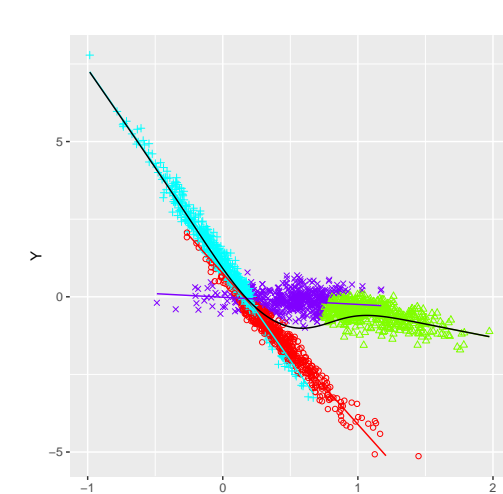
1.3 MS realization



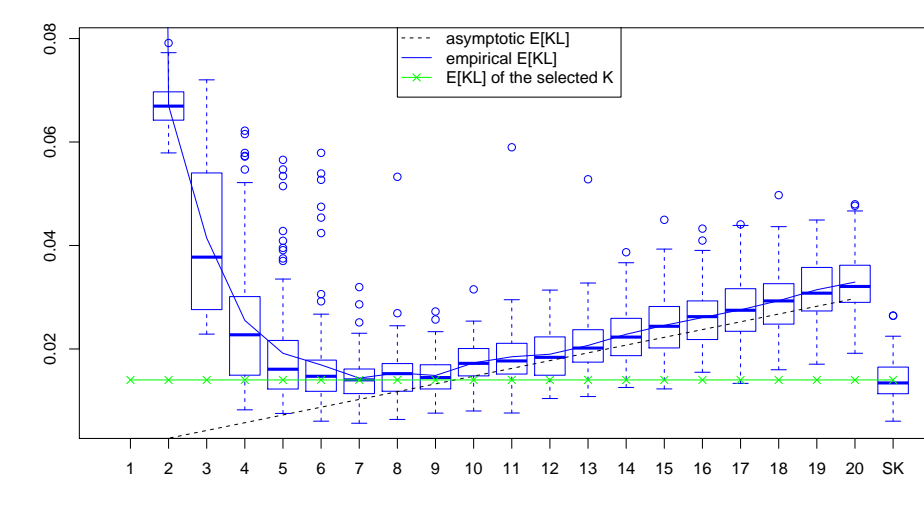
3.1 WS with $n = 2000$



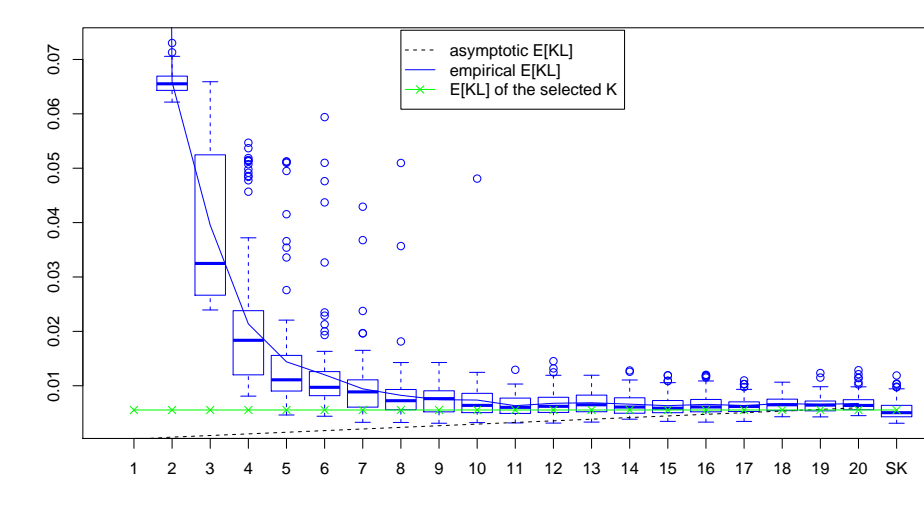
3.2 WS with $n = 10000$



1.4 GLoME clustering



3.3 MS with $n = 2000$



3.4 MS with $n = 10000$