

13 -

\bar{t}_m - pour

→ Nota: l'algorithme peut être parallélisé, est simple à coder, mais présente un pb lié à sa nature réursive: ψ de DÉGÉNÉRESCE

la variance des poids augmente avec $k \Rightarrow$ la variance de l'estimateur augmente également

après qq réursions, la plupart des particules sont associées à un poids négligeable

→ Taille efficace du N-échantillon

$$N_{\text{eff}} := \left(\sum_{i=1}^N \frac{1}{(w_h^{(i)})^2} \right)^{-1}$$

$$\in [1; N]$$

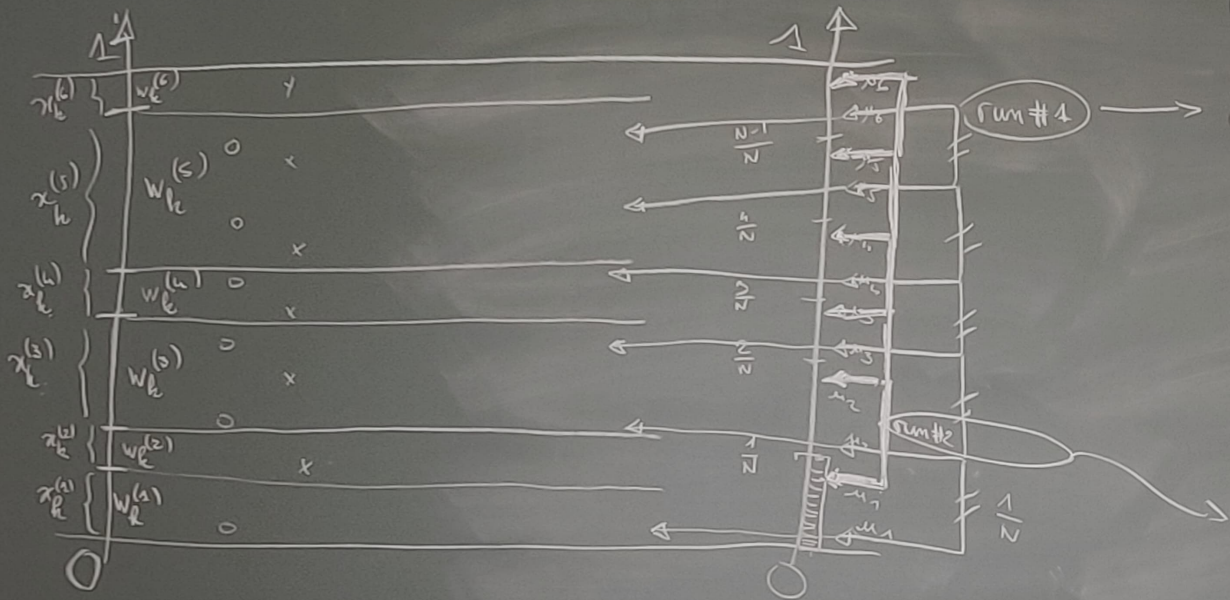
$$= N \text{ si } \forall i, w_h^{(i)} = \frac{1}{N}$$

$$= 1 \text{ si } \forall i \neq i_0, w_h^{(i)} = 0 \\ \text{et } w_h^{(i_0)} = 1$$

* L'algorithme SIR (Sampling Importance Resampling)

$$\left[\left\{ w_k^{(i)}, x_k^{(i)} \right\}_{i=1}^N \right] = \text{SIR} \left(\left\{ w_{k-1}^{(i)}, x_{k-1}^{(i)} \right\}_{i=1}^N, \beta_k \right)$$

1. Si $k=0$, alors % Initialization - Échantillonnage idéal selon la loi initiale
2. Échantillonner $x_0^{(1)} \dots x_0^{(N)} \stackrel{\text{iid}}{\sim} p(x_0)$ et poser $w_0^{(i)} = w_0^N = \frac{1}{N}$
3. Fin - si
4. Si $k \geq 1$ alors
5. Pour $i=1 \dots N$ indépendamment
6. "Propager" $x_{k-1}^{(i)}$ en échantillonnant son successeur selon $x_k^{(i)} \sim q(x_k | x_{k-1}^{(i)}, \beta_k)$



$$\begin{cases} \tilde{x}_k^{(1)} = x_k^{(1)} \\ \tilde{x}_k^{(2)} = x_k^{(2)} \\ \tilde{x}_k^{(3)} = x_k^{(3)} \\ \tilde{x}_k^{(4)} = x_k^{(4)} \\ \tilde{x}_k^{(N-1)} = x_k^{(5)} \\ \tilde{x}_k^{(N)} = x_k^{(5)} \end{cases}$$

$$\begin{cases} \tilde{x}_k^{(1)} = x_k^{(2)} \\ \tilde{x}_k^{(2)} = x_k^{(3)} \\ \tilde{x}_k^{(3)} = x_k^{(4)} \\ \tilde{x}_k^{(4)} = x_k^{(5)} \\ \tilde{x}_k^{(N-1)} = x_k^{(5)} \\ \tilde{x}_k^{(N)} = x_k^{(5)} \end{cases}$$

$$u_1 \sim U[0, \frac{1}{N}]$$

$$u_2 = u_1 + \frac{1}{N}$$

$$u_3 = u_2 + \frac{1}{N} \dots$$

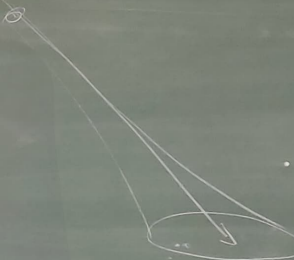
* On obtient le poids associé

$$w_k^{(i)} \propto w_{k-1}^{(i)} \frac{p(z_k | x_k^{(i)}) p(x_k^{(i)} | x_{k-1}^{(i)})}{p(x_k^{(i)} | x_{k-1}^{(i)})}$$



$$w_k^{(i)} \propto w_{k-1}^{(i)} p(z_k | x_k^{(i)})$$

VRAISEMBLANCE de la
particule $x_k^{(i)}$ // mesure z_k



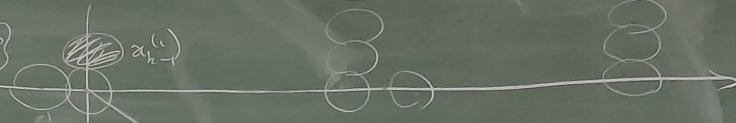
$$\mathcal{P}_{k-1} \left\{ w_{k-1}^{(i)}, x_{k-1}^{(i)} \right\}_{i=1}^N$$

$\downarrow \mathcal{R}$



$$\mathcal{P}_{k-1} \left\{ w_{k-1}^{(i)} = \frac{1}{N}, x_{k-1}^{(i)} \right\}$$

normalized



$\downarrow \text{I.S.}$

$$\mathcal{P}_k \left\{ w_k^{(i)}, x_k^{(i)} \sim p(x_k | x_{k-1}^{(i)}) \right\}$$



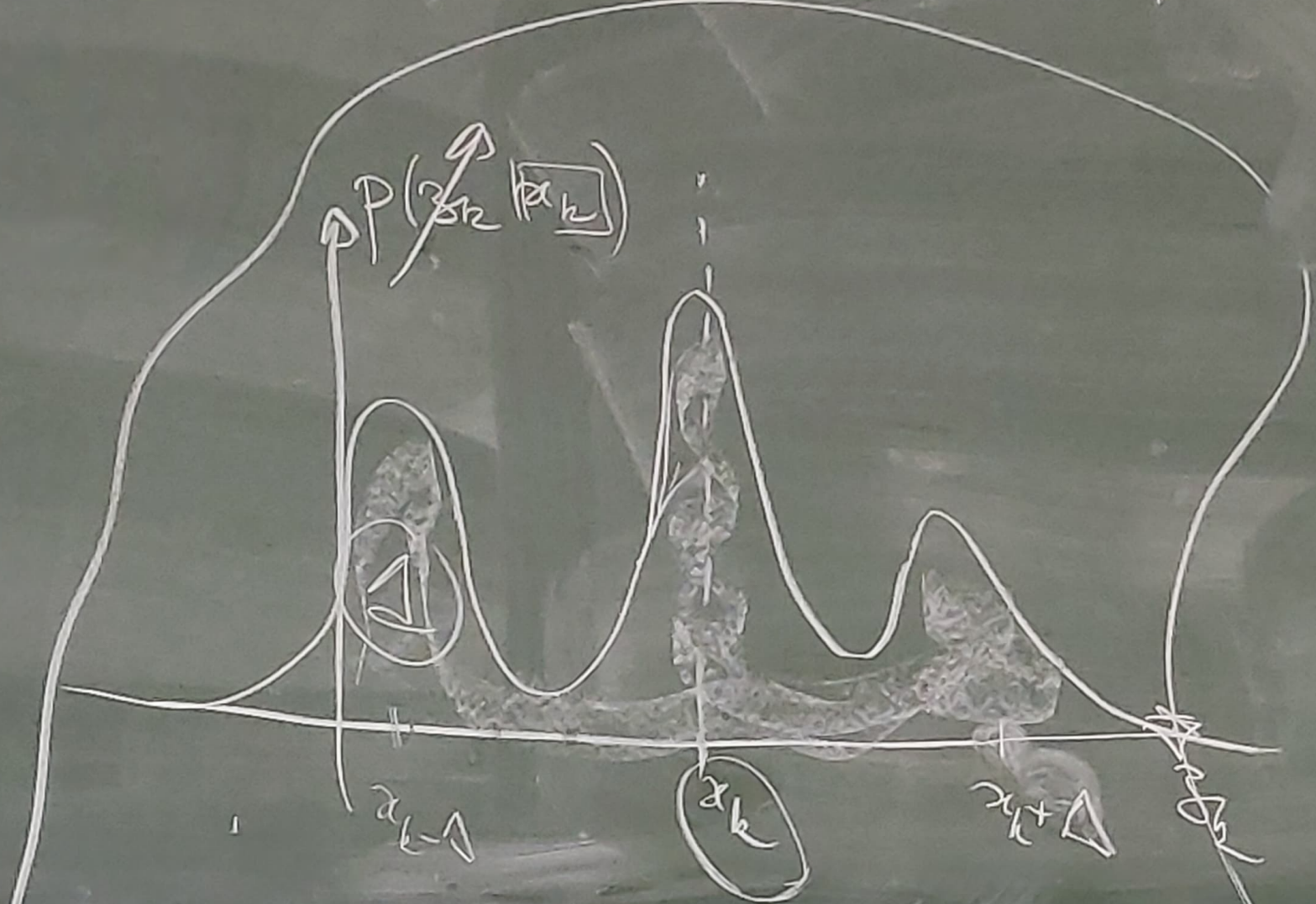
$\downarrow w$

$$\mathcal{P}_{k-1} \left\{ w_k^{(i)}, x_k^{(i)} \right\}$$

Note:

$$\sum_{i=1}^N w_{k-1}^{(i)} \delta(x_k - x_k^{(i)}) \approx p(x_k | \mathcal{D}_{k-1})$$

$P(z_k | x_k)$



The diagram illustrates a Markov chain Monte Carlo process. A wavy line represents the target distribution. A horizontal line represents the current state. Points x_{k-1} , x_k , and x_{k+1} are marked on the horizontal line. A vertical line at x_k is labeled $P(z_k | x_k)$. A shaded region is shown above x_k , and a vertical dashed line extends upwards from x_k .

x_{k-1}

x_k

x_{k+1}

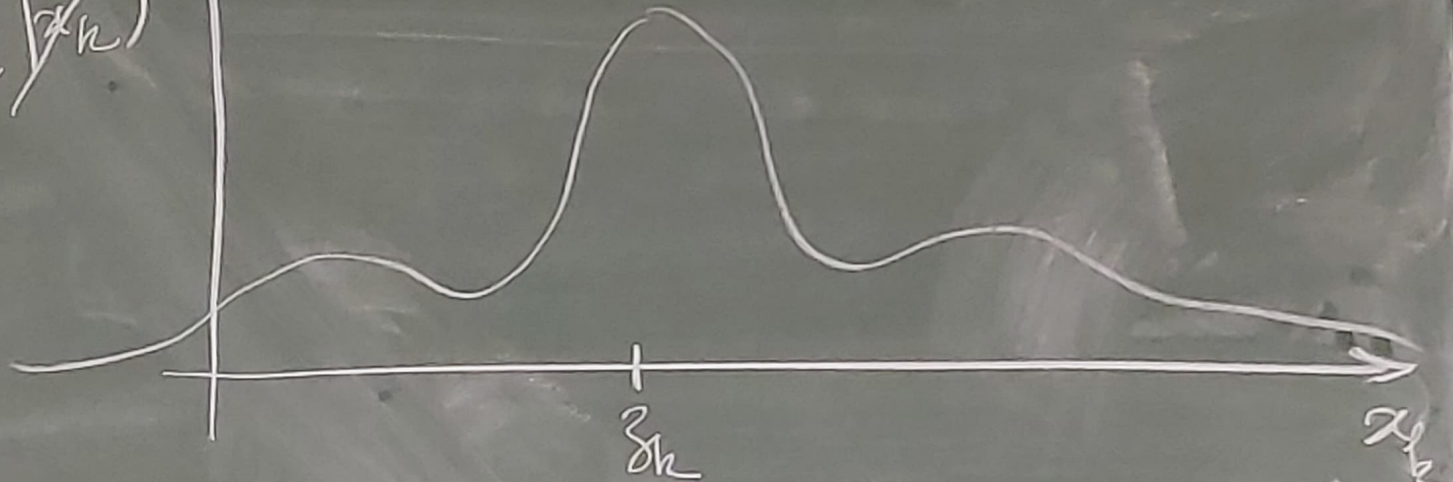
x_{k+2}

$$L(x_k; z_k)$$

$=$ \nearrow

$$p(z_k | x_k)$$

soit z_k la mesure

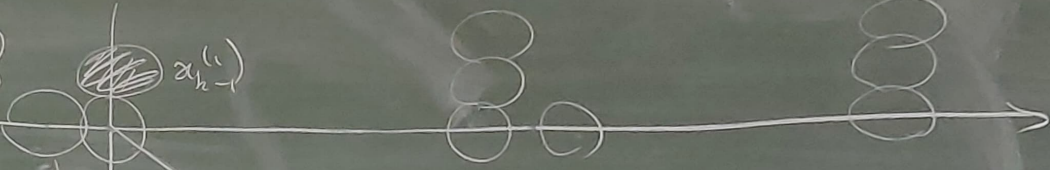


$$\{p_{k-1}^{(i)}\} \left\{ w_{k-1}^{(i)}, x_{k-1}^{(i)} \right\}_{i=1}^N$$

$\downarrow R$



$$p_{k-1} \left\{ w_{k-1}^{(i)} = \frac{1}{N}, x_{k-1}^{(i)} \right\}$$



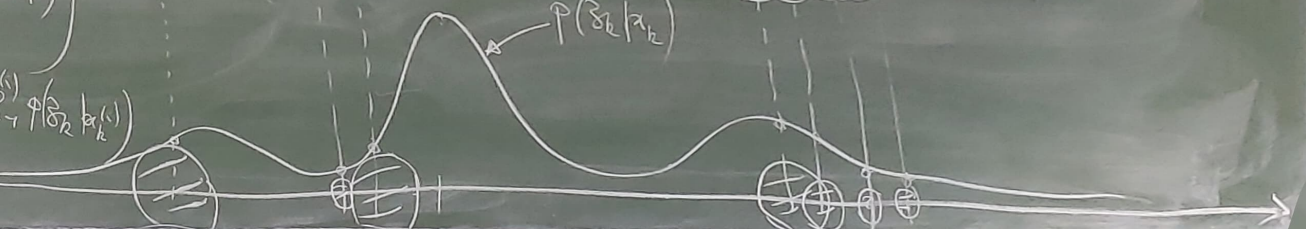
$\downarrow I.S$

$$p_k \left\{ w_k^{(i)}, x_k^{(i)} \sim p(x_k | x_{k-1}^{(i)}) \right\}$$



$p(x_k | x_{k-1})$

$$\left\{ w_k^{(i)}, x_k^{(i)} \right\}$$



$$\left\{ w_{h-1}^{(i)}, x_{h-1}^{(i)} \right\}_{i=1}^N$$

↓ ± s

$$x_h^{(i)} \sim q(x_h | z_h)$$

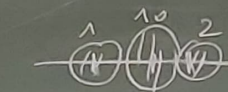
ou

$$q(x_h | z_h) \text{ est}$$

$\sim q(z_h | x_h)$ modulo
une constante de normalisation

$$w_h^{(i)} \propto w_{h-1}^{(i)} \frac{p(z_h | x_h^{(i)}) q(x_h^{(i)} | z_{h-1}^{(i)})}{q(x_h^{(i)} | z_h)}$$

$$p(x_{h-1} | z_{h-1})$$



$$q(x_h | z_h)$$

$$p(z_h | x_h)$$

