

Response Space(y)Class of model $y \in \mathbb{R}$ $\{C_1, C_2, \dots, C_k\}$ $k=2$ $\{C_1, C_2\}$ $\begin{matrix} 0 \\ 1 \end{matrix}$

Regression, Survival

Classification, Probability estimation

Binary Classification, prob. est.

 $\mathcal{Y} = \{0, 1\}$ $y = t(\bar{x})$ $= f(\bar{x}) + \delta, \delta \in \{0, 1, -1\}$ $= h^*(\bar{x}) + \varepsilon, \varepsilon \in \{0, 1, -1\}$ $= g(\bar{x}) + e, e \in \{0, 1, -1\}$

degenerate

 $\Leftrightarrow y \sim \text{Bern}(t(\bar{x}))$

① Assume a function

 $f_{pr}(\bar{x}) : \mathbb{R}^{p+1} \rightarrow (0, 1)$ best guess with \bar{x} of $p(y=1|\bar{x})$ ② Assume $y_i | \bar{x}_i$ independent y_i $\Rightarrow y_i = y_i | \bar{x}_i \sim \text{Bern}(f_{pr}(\bar{x}_i)) = f_{pr}(\bar{x}_i)^{y_i} (1 - f_{pr}(\bar{x}_i))^{1-y_i}$ $\Rightarrow P(\mathcal{D}) = P(y_1 = y_1, y_2 = y_2, \dots, y_n = y_n | \bar{x}_1, \bar{x}_2, \dots, \bar{x}_n)$ $\textcircled{II} = \prod_{i=1}^n f_{pr}(\bar{x}_i)^{y_i} (1 - f_{pr}(\bar{x}_i))^{1-y_i}$ likelihood.Is it possible to $\text{argmax}_{f \in \mathcal{F}} \left\{ \prod_{i=1}^n f_{pr}(\bar{x}_i)^{y_i} (1 - f_{pr}(\bar{x}_i))^{1-y_i} \right\}$?NO, b/c it is impossible. $\mathcal{F} = \text{all functions}$ Instead assume $\mathcal{H}_{pr} = \{\bar{w} \cdot \bar{x} : \bar{w} \in \mathbb{R}^{p+1}\}$. Is this okay?NO, Since $\bar{w} \cdot \bar{x} \notin (0, 1)$ Let's say we want to retain the term $\bar{w} \cdot \bar{x}$ in our hypothesis set. Why? Linear model is very interpretable, monotonic, simple.
generalized linear model (glm) $\Rightarrow \mathcal{H}_{pr} = \{\phi(\bar{w} \cdot \bar{x}) : \bar{w} \in \mathbb{R}^{p+1}\}$ $\phi : \mathbb{R} \rightarrow (0, 1)$

called a "link function"

3 common choices in order of popularity:

① Logistic $\phi(u) = \frac{e^u}{1+e^u} = \frac{1}{1+e^{-u}}$, $1-\phi(u) = \frac{1}{1+e^u}$

② Probit $\phi(u) = \Phi(u) \stackrel{\text{inverse CDF of the standard normal r.v.}}{=} \Phi(u)$

③ Complementary Log-Log $\phi(u) = 1 - e^{-e^u}$, $1-\phi(u) = e^{-e^u}$

Logistic Regression: $\mathcal{H}_{pr} = \left\{ \frac{1}{1+e^{-\vec{w} \cdot \vec{x}}} : \vec{w} \in \mathbb{R}^{p+1} \right\}$

A: $\vec{b} = \underset{\vec{w} \in \mathbb{R}^{p+1}}{\text{argmax}} \left\{ \prod_{i=1}^n \left(\frac{1}{1+e^{-\vec{w} \cdot \vec{x}_i}} \right)^{y_i} \left(\frac{1}{1+e^{\vec{w} \cdot \vec{x}_i}} \right)^{1-y_i} \right\}$ no closed form solution

\downarrow maximum likelihood $P(\mathcal{D})$ i.e. \vec{w}_{pr} is not solvable

need to use numerical optimization to approximate \vec{b} . That is called "fitting a logistic regression".

$g_{pr}(\vec{x}_*) = \frac{1}{1+e^{-\vec{b} \cdot \vec{x}_*}} = \hat{p}(y_*=1 | \vec{x}_*) = \hat{p}_*$
 $\in \mathcal{H}_{pr} \approx P(y_*=1 | \vec{x}_*) = f_{pr}(\vec{x}_*)$

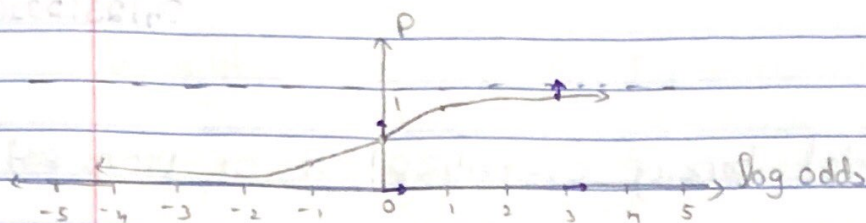
What is the interpretation of b_j 's.

$\hat{p} = \frac{1}{1+e^{-\vec{b} \cdot \vec{x}}} \Rightarrow \frac{1}{\hat{p}} = 1 + e^{-\vec{b} \cdot \vec{x}} \Rightarrow \frac{1-\hat{p}}{\hat{p}} = \frac{1}{\hat{p}} - 1 = e^{-\vec{b} \cdot \vec{x}}$ Odds Against ($y=1 | \vec{x}$)

$\Rightarrow \frac{\hat{p}}{1-\hat{p}} = e^{\vec{b} \cdot \vec{x}} \Rightarrow \vec{b} \cdot \vec{x} = \ln \left(\frac{\hat{p}}{1-\hat{p}} \right) \in \mathbb{R}$

Odds($y=1 | \vec{x}$) Log Odds($y=1 | \vec{x}_*$)

$b_3 = 0.7$, x_3 increases by 1 \Rightarrow Log Odds($y=1 | \vec{x}$) increases by 0.7



log odds	prob
0	0.5
-1	0.27
+1	0.73
-2	0.12
+2	0.88
-∞	0
+∞	1

$$g_{0, p_1} = \bar{y} = \frac{1}{n} \sum \mathbb{1}_{y_i=1}$$

Prob. Est. Honest model validation;
previously $SSE = \sum (y_i - \hat{y}_i)^2$

sum of square prob

Maybe $SSE = \sum (p_i - \hat{p}_i)^2$ $p_i = p(y_i=1 | \vec{x}_i) = f_{\vec{x}}(\vec{x}_i)$

Impossible ...

higher the score, the better the model $\approx R^2$

"Scoring Rule" $S(y, \hat{p})$, A "proper scoring" rule:

$$y_i f_{p_i}(\vec{x}_i) = \text{avgmax}_{\beta \in (0,1)} [S(y_i, \beta)]$$

Two popular proper scoring rule:

① log scoring rule: $S_i = y_i \ln(\hat{p}_i) + (1 - y_i) \ln(1 - \hat{p}_i)$

overall score is average

② Brier score: $S_i = -(y_i - \hat{p}_i)^2$

$$\Rightarrow S = \frac{1}{n} \sum S_i$$