

# MATH3160 — Portfolio 6.3

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## 1 Deliverables

### 1.1 Standard Normal Random Variable

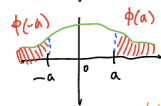
• Density:  $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, x \in \mathbb{R}$



- Area under  $f=1$
- Graph of  $f$  is symmetric
- Peaks at  $x=0$
- $\lim_{x \rightarrow \pm\infty} f = 0$

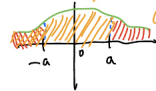
•  $\Phi(-a) = 1 - \Phi(a) \quad \forall a \in \mathbb{R}$  (due to symmetry)

Geometric proof:



$\Phi(-a) = \Phi(a)$   
due to graph symmetry

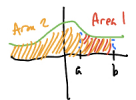
$$\text{Area} = \Phi(-a) + \Phi(a) = 1$$



$$\text{So } \Phi(-a) = 1 - \Phi(a)$$

•  $P(a < Z < b) = \Phi(b) - \Phi(a)$

Geometric Proof:



$$P(a < Z < b) = \text{Area 1}$$

$$P(Z \in (-\infty, b)) = \text{Area 3}$$

$$P(Z \in (-\infty, a)) = \text{Area 2}$$

$$P(a < Z < b) = A_3 - A_2$$

$$\Rightarrow A_1 = \Phi(b) - \Phi(a)$$

$$E[Z] \stackrel{\text{def}}{=} \int_{\mathbb{R}} x \cdot f(x) dx \Rightarrow \int_{\mathbb{R}} x \cdot \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \Rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x \cdot e^{-x^2/2} dx$$

Odd function:  $f(x) = -f(x) \quad \forall x \in \mathbb{R}$

for example,  $f(x) = x^5$

$$\int_{-a}^a f(x) dx = 0$$


Even function:  $f(x) = f(-x)$

$$\int_{-a}^a f(x) dx = 2 \cdot \int_0^a f(x) dx$$

$$f(-x) = x \cdot e^{-x^2/2} = -f(x)$$

(by definition of an odd function)

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x \cdot e^{-x^2/2} dx = 0 \quad \therefore E[Z] = 0$$

$$\begin{aligned}
\bullet \text{Var}(Z) &= \mathbb{E}[Z^2] - (\mathbb{E}[Z])^2 \\
\mathbb{E}[Z^2] &\xrightarrow[\text{from } x]{\text{transform}} \int_{\mathbb{R}} x^2 \cdot f(x) dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^2 \cdot e^{-x^2/2} dx \\
&\quad \uparrow \text{Even function} \\
&\Rightarrow \frac{1}{\sqrt{2\pi}} \cdot 2 \int_0^{\infty} x^2 \cdot e^{-x^2/2} dx \Rightarrow \frac{2}{\sqrt{2\pi}} \int_0^{\infty} x \cdot (-x \cdot e^{-x^2/2}) dx \\
&\xrightarrow[\text{note}]{\text{chain rule}} (e^{-x^2/2})' = -x \cdot e^{-x^2/2} \\
&\quad \downarrow \text{faster} \\
&\quad \text{when } x \rightarrow \infty \\
&\xrightarrow[\text{part}]{\text{int by parts}} \frac{2}{\sqrt{2\pi}} \cdot \left[ \left( x \cdot e^{-x^2/2} \right) \right]_0^{\infty} - \int_0^{\infty} 1 \cdot e^{-x^2/2} dx \\
&\quad \uparrow \text{when } x \rightarrow \infty \\
&\Rightarrow \frac{2}{\sqrt{2\pi}} \cdot \int_0^{\infty} e^{-x^2/2} dx \Rightarrow 2 \cdot \int_0^{\infty} \underbrace{\left( \frac{1}{\sqrt{2\pi}} \right) e^{-x^2/2}}_{f(x)} dx \\
&\quad \uparrow \text{but for half of whole area} \\
&\Rightarrow 2 \cdot \left( \frac{1}{2} \text{ area under } f \right) = 1 \\
&\quad \downarrow \quad \uparrow \\
&\quad 1 \quad 0 \\
\therefore \mathbb{E}[Z^2] = 1, \quad \mathbb{E}[Z^2] - (\mathbb{E}[Z])^2 = 1
\end{aligned}$$


## 1.2 General Normal $\rightarrow$ Standard Normal

Let  $X \sim \mathcal{N}(\mu, \sigma^2)$ , then  $\frac{x-\mu}{\sigma} \sim \mathcal{N}(0, 1)$ . This is called the reduction to a standard normal random variable, and can be accomplished by transforming the random variable  $X$  into a standard normal random variable  $Z$  by the above reduction.

The opposite holds, given  $Z \sim \mathcal{N}(0, 1)$ , we can transform  $Z$  into a random variable  $X$  with a general normal distribution by the following transformation to revert back to a general normal random variable:

$$X = \mu + \sigma Z$$

### 1.3 General Normal Random Variable

•  $\text{Dom } X = \mathbb{R}$  (As  $f \neq 0 \forall x \in \mathbb{R}$ )

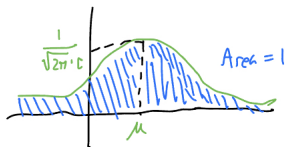
• Density:

$$\begin{aligned} \textcircled{1} \text{ Get } F_X &\longrightarrow F_X(x) \stackrel{\text{def}}{=} P(X \leq x) = P(c \cdot Z + \mu \leq x) \\ &\Rightarrow P\left(Z \leq \frac{x - \mu}{c}\right) = \Phi\left(\frac{x - \mu}{c}\right) \end{aligned}$$

$\underbrace{\qquad\qquad\qquad}_{\Phi\left(\frac{x - \mu}{c}\right)}$

$$\begin{aligned} \textcircled{2} \text{ Get } f_X = F'_X &\longrightarrow f_X = \left(\Phi\left(\frac{x - \mu}{c}\right)\right)' \\ &\xRightarrow[\text{rule}]{\text{chain}} \Phi'\left(\frac{x - \mu}{c}\right) \cdot \left(\frac{x - \mu}{c}\right)' \quad \left[\frac{x}{c} - \frac{\mu}{c}\right] \\ &\Rightarrow f_Z\left(\frac{x - \mu}{c}\right) \cdot \frac{1}{c} \\ &\quad \left[f_Z(x) \stackrel{\text{def}}{=} \frac{1}{\sqrt{2\pi}} e^{-x^2/2}\right] \\ &\Rightarrow \frac{1}{\sqrt{2\pi} \cdot c} e^{-\frac{(x - \mu)^2}{2c^2}}, \quad x \in \mathbb{R} \end{aligned}$$

plot:



•  $E[X]$ :

$$\begin{aligned} E[X] &= E[cZ + \mu] \xrightarrow{\text{linearity}} c \cdot E[Z] + \mu \\ &\quad \begin{array}{c} \vdots \\ \text{std normal} \\ \text{r.v.} \end{array} \quad \begin{array}{c} \uparrow \\ 0 \end{array} \\ &= c \cdot 0 + \mu = \mu \end{aligned}$$

•  $\text{Var}(X)$ :

$$\begin{aligned} \text{Var}(X) &= \text{Var}(c \cdot Z + \mu) \\ &\quad \begin{array}{c} \uparrow \\ \text{scaling} \end{array} \end{aligned}$$

$$\boxed{\text{Var}(aX + b) = a^2 \cdot \text{Var}(X)}$$

$$\begin{aligned} \Rightarrow c^2 \cdot \text{Var}(Z) &= c^2 \\ &\quad \uparrow \\ &\quad 1 \end{aligned}$$

(So, if  $X \sim \mathcal{N}(\mu, c^2)$ , then  $\mu = E[X]$   
and  $c^2 = \text{Var}(X)$ )