



# PROVA 3 -

$$1. \quad P(y_i = 1 | x_i) = S(x_i; \theta) \quad p / \quad S(x_i) = \frac{1}{1 + \exp(-x_i)}$$

$$P(y_i = 0 | x_i) = 1 - S(x_i; \theta)$$

$$P(Y=y | \theta) = \theta^K (1-\theta)^{N-K} = L(\theta)$$

$$L(\theta, y_i, x_i) = [S(x_i; \theta)]^{y_i} [1 - S(x_i; \theta)]^{1-y_i}$$

$$L(\theta, y, X) = \prod_{i=1}^N [S(x_i; \theta)]^{y_i} [1 - S(x_i; \theta)]^{1-y_i}$$

$$\log L(\theta, y, X) = \ln \left[ \prod_{i=1}^N [S(x_i; \theta)]^{y_i} [1 - S(x_i; \theta)]^{1-y_i} \right]$$

$$= \sum_{i=1}^N \left[ y_i \ln \left( \frac{1}{1 + e^{-x_i \theta}} \right) + (1-y_i) \ln \left( \frac{1 + e^{-x_i \theta}}{1 + e^{-x_i \theta}} \right) \right]$$

$$= \sum_{i=1}^N \left[ \ln \left( \frac{\exp(-x_i \theta)}{1 + e^{-x_i \theta}} \right) + y_i \left( \ln \left( \frac{1}{1 + e^{-x_i \theta}} \right) - \ln \left( \frac{e^{-x_i \theta}}{1 + e^{-x_i \theta}} \right) \right) \right]$$

$$= \sum_{i=1}^N \left[ \ln \left( \frac{1}{1 + e^{x_i \theta}} \right) + y_i \left( \ln \left( \frac{1}{e^{-x_i \theta}} \right) \right) \right]$$

$$= \sum_{i=1}^N \left[ \ln(1) - \ln(1 + e^{x_i \theta}) + y_i (\ln(1) - \ln(e^{x_i \theta})) \right]$$

$$= \sum_{i=1}^N \left[ -\ln(1 + \exp(x_i \theta)) + y_i x_i \theta \right]$$





p/ maximizar  $l(\theta) \rightarrow$  derivada

$$0 = \frac{\partial l(\theta)}{\partial \beta_j} = \frac{\partial}{\partial \beta_j} \sum_{i=1}^N [-\ln(1 + e^{x_i \beta_j}) + y_i x_i \beta_j]$$

$$= \sum_{i=1}^N \left( \frac{-e^{x_i \beta_j}}{1 + e^{x_i \beta_j}} x_i + y_i x_i \right)$$

$$= \sum_{i=1}^N \left( y_i - \frac{e^{x_i \beta_j}}{1 + e^{x_i \beta_j}} \right) x_i$$

$$= \sum_{i=1}^N \left( y_i - \frac{1}{1 + e^{-x_i \beta_j}} \right) x_i$$

$$= \sum_{i=1}^N x_i (y_i - p_i)$$

$$= X^t (y - p)$$





2. A primeira derivada, obtida na questão 1 é:

$$\frac{\partial l(\theta)}{\partial \beta_j} = \sum_{i=1}^N x_i (y_i - p_i) \quad (a)$$

$$\hookrightarrow \sum x_{ij} y_i - \sum x_{ij} p_i$$

$$\frac{\partial^2 l(\theta)}{\partial \beta_j \partial \beta_j} = \text{Var } (a)$$

$$= - \sum_{i=1}^N \left( \frac{(1 + e^{x_i \theta}) x_{in} x_{ij} e^{x_i \theta} - x_{in} x_{ij} (e^{x_i \theta})^2}{(1 + e^{x_i \theta})^2} \right)$$

$$= - \sum_{i=1}^N \left( \frac{x_{in} x_{ij} e^{x_i \theta}}{1 + e^{x_i \theta}} \cdot \frac{(1 + e^{x_i \theta})}{1 + e^{x_i \theta}} - x_{in} x_{ij} \frac{(e^{x_i \theta})^2}{(1 + e^{x_i \theta})^2} \right)$$

$$= - \sum_{i=1}^N \left( \frac{x_{in} x_{ij} e^{x_i \theta}}{1 + e^{x_i \theta}} - x_{in} x_{ij} \frac{(e^{x_i \theta})^2}{(1 + e^{x_i \theta})^2} \right)$$

$$= - \sum_{i=1}^N \left( \frac{x_{in} x_{ij} e^{x_i \theta}}{1 + e^{x_i \theta}} \cdot \frac{e^{-x_i \theta}}{e^{-x_i \theta}} - // \right)$$

$$= - \sum_{i=1}^N \left( x_{in} x_{ij} \frac{1}{1 + e^{x_i \theta}} - x_{in} x_{ij} \frac{e^{x_i \theta}}{1 + e^{x_i \theta}} \cdot \frac{e^{-x_i \theta}}{e^{-x_i \theta}} \right)$$

$$= - \sum_{i=1}^N \left( x_{in} x_{ij} p_i - x_{in} x_{ij} \frac{p_i^2}{(1 + e^{-x_i \theta})^2} \right)$$

$$= - \sum_{i=1}^N x_{in} x_{ij} p_i - x_{in} x_{ij} p^2$$

$$= - \sum_{i=1}^N x_{in} x_{ij} p_i (1 - p_i)$$





Na forma matricial, temos,

$$\frac{d^2 l(\theta)}{d\theta d\theta^t} = - \sum_{i=1}^n x_i x_i^t p_i (1-p_i)$$
$$= - X^t W X$$

$W$  é matriz  $n \times n$  diagonal com  $i$ -ésimo termo dado por  $p_i (1-p_i)$

$$W = \begin{bmatrix} p_1(1-p_1) & 0 & \dots & 0 \\ 0 & p_2(1-p_2) & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & p_n(1-p_n) \end{bmatrix}$$







3.  $N(\mu_0, \sigma^2)$

$$f(y_i, \mu_0, \sigma^2) = \frac{1}{\sqrt{\sigma^2} \sqrt{2\pi}} \exp \left[ -\frac{(y_i - \mu_0)^2}{2\sigma^2} \right]$$

$$\log L(\mu_0, \sigma^2) = \frac{n}{2} \log \sigma^2 - \frac{n}{2} \log(2\pi) - \frac{\sum (y_i - \mu_0)^2}{2\sigma^2}$$

v considerando  $\mu_0$  conhecido

$$\begin{aligned} \frac{\partial \log L(\mu_0, \sigma^2)}{\partial \sigma^2} &= \frac{-n}{2\sigma^2} + \frac{\sum (y_i - \mu_0)^2}{2\sigma^4} \equiv 0 \\ &= -n\sigma^2 + \sum (y_i - \mu_0)^2 = 0 \end{aligned}$$

$$\sigma^2 = \frac{\sum (y_i - \mu_0)^2}{n}$$

v checar pto de máxima

$$\hat{\sigma}^2 = \frac{\sum (y_i - \bar{y})^2}{n}$$

$$\frac{\partial^2 \log L(\mu_0, \sigma^2)}{(\partial \sigma^2)^2} = \frac{-n}{\sigma^4} - \frac{\sum (y_i - \mu_0)^2}{\sigma^6}$$

$$\frac{-n}{\sigma^4} - \frac{\sum (y_i - \mu_0)^2}{\sigma^6} = 0 \rightarrow n\sigma^2 - \sum (y_i - \mu_0)^2 = 0$$
$$\hat{\sigma}^2 = \frac{\sum (y_i - \mu_0)^2}{n}$$

$\rightarrow \sum (y_i - \mu_0)^2$  é sempre  $> 0$  e  $n > 0$   $\therefore \hat{\sigma}^2 > 0$  e, como a 2ª derivada é negativa temos que este é o ponto de máxima.  
Assim,  $\hat{\sigma}^2 = \frac{\sum (y_i - \mu_0)^2}{n}$





4.  $E(Y) = \mu$

$$\begin{aligned} E(Y - m)^2 &= E(Y - \mu + \mu - m)^2 \\ &= E[(Y - \mu)^2 + 2(Y - \mu)(\mu - m) + (\mu - m)^2] \\ &= E[(Y - \mu)^2] + 2E[(Y - \mu)(\mu - m)] + E[(\mu - m)^2] \\ &= \text{Var}(Y) + 2(\mu - m)E(Y - \mu) + (\mu - m)^2 \\ E(Y - m)^2 &= \text{Var}(Y) + 2(\mu - m) \cdot 0 + (\mu - m)^2 \\ &= \text{Var}(Y) + 0 + (\mu - m)^2 \\ &= \text{Var}(Y) + (\mu - m)^2 \\ &\quad \rightarrow \text{bias} \end{aligned}$$

$\rightarrow$  derivando em relação a  $m$

$$\rightarrow -2\mu + 2m$$

$$\rightarrow m = \mu \quad p/\text{derivada} = 0$$

$$\rightarrow m = \mu \text{ minimiza}$$







5. Seja  $H_1$  a matriz de projeção ortogonal em  $M(X)$ , ou seja,  $H_1 = X(X'X)^{-1}X'$ . Então temos:

- Para  $Y \in \mathbb{R}^n$ ,  $H_1 Y \in M(X)$  e  $H_1 Y$  é projeção ortogonal de  $Y$  em  $M(X)$ .  $H_1 = I(I + I)^{-1}I$

- A solução de mínimos quadrados  $\hat{\beta} = \arg \min_b \|Y - Xb\|^2$

Assim  $X\hat{\beta}$  é  $H_1 Y = X(X'X)^{-1}X'Y$

A projeção é  $\hat{Y} = X\hat{\beta} \rightarrow \hat{Y} = H_1 Y$

- Em uma boa predição  $Y \approx \hat{Y} = X\hat{\beta}$

Se  $\bar{y} = \sum y_i/n$  e  $\bar{Y} = \bar{y} \mathbf{1}'$  então,

$$Y = \hat{Y} + (Y - \hat{Y}) = \bar{Y} + \hat{Y} - \bar{Y} + (Y - \hat{Y})$$

$$Y - \bar{y} \mathbf{1}' = (\hat{Y} - \bar{y} \mathbf{1}') + (Y - \hat{Y})$$

$$Y = X\hat{\beta} + (Y - X\hat{\beta})$$

$$Y = X(X'X)^{-1}X'Y + (Y - X(X'X)^{-1}X'Y)$$

$$Y = H_1 Y + (Y - H_1 Y)$$

$$Y = H_1 Y + (I - H_1)Y = \bar{Y} + \hat{Y} - \bar{Y} + (Y - \hat{Y})$$

$$Y = H_1 Y + (I - H_1)Y = \bar{y} \mathbf{1}' + H_1 Y - \bar{y} \mathbf{1}' + (I - H_1)Y$$

$$Y = H_1 Y + (I - H_1)Y = \bar{y} \mathbf{1}' + H_1 Y - \bar{y} \mathbf{1}' + (I - H_1)Y$$

~ Aqui devemos notar que  $H_1$  é simétrica e idempotente

$$Y = H_1 Y + (I - H_1)Y = \bar{y} \mathbf{1}' + \overbrace{H_1 Y}^Y - \bar{y} \mathbf{1}' + \overbrace{Y - H_1 Y}^Y$$

$$Y = H_1 Y + (I - H_1)Y = \bar{y} \mathbf{1}' + (Y - \bar{y} \mathbf{1}') + \cancel{Y - Y}$$

~> próxima pg ~>





•  $\langle u, v \rangle$  são ortogonais se o seu produto interno é 0

$$\rightarrow \langle \bar{y}1, (Y - \bar{y}1) \rangle \rightarrow \langle \bar{y}1, (y_i - \bar{y})1 \rangle$$

$$\rightarrow \bar{y}(y_1 - \bar{y}) + \bar{y}(y_2 - \bar{y}) + \dots + \bar{y}(y_n - \bar{y}) = \bar{y} \sum_{i=1}^n (y_i - \bar{y})$$

$$\rightarrow \bar{y} = \sum_{i=1}^n y_i / n \therefore \sum_{i=1}^n \bar{y} = \sum_{i=1}^n y_i \therefore \sum_{i=1}^n (y_i - \bar{y}) = 0 \therefore$$

$$\langle \bar{y}1, (Y - \bar{y}1) \rangle = 0$$

