Applied Survival Analysis Using R Chapter 2: Basic Principle of Survival Analysis

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April, 9 2019

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The Hazard and Survival Functions

The Survival Function

Survival analysis is defines the probability of surviving up to a point t. Formally,

$$S(t) = 1 - P(T \le t) = 1 - F(t), \ 0 < t < \infty$$

And it decreases over time

The Hazard Function

It is the instantaneous failure rate. It is the probability that, given that a subject has survived up to time t, he or she fails in the next small interval of time, divided by the length of that interval. Formally,

$$h(t) = \lim_{\delta \to 0} \frac{P(t \le T < t + \delta | T \ge t)}{\delta}.$$

Connections

- Cumulative risk function: $F(t) = P(T \le t), \ 0 < t < \infty$
- Probability density function: $f(t) = \frac{dF(t)}{dt} = -\frac{dS(t)}{dt}$
- So we can deduce the following two formulas:
 - $h(t) = \frac{f(t)}{S(t)}$
 - $H(t) = \int_0^t h(u) du$
 - $S(t) = exp(-\int_0^t h(u)du) = exp(-H(t))$



Mean and Median Survival Time

• The mean survival is the expected value of the survival time

$$\mu = E(T) = \int_0^\infty tf(t)dt$$

or

$$\mu = \int_0^\infty S(t) dt$$

- The mean survival also cannot be computed with the Kaplan-Meier survival curve when the curve does not reach zero, an issue we will discuss in the next chapter.
- The median survival time is defined as the time t such that S(t) = 1/2



Exponential Distribution

- When modeling human or animal survival, it is hard to know what
 parametric family to choose, and often none of the available
 families has sufficient flexibility to model the actual shape of the
 distribution, so commonly we analyze the non-parametric survival
 distribution due to more flexible and applicable.
- The exponential distribution is the simplest but classical survival distribution, has a *constant* hazard, $h(t) = \lambda$
- $H(t) = \lambda t$
- $S(t) = e^{-H(t)} = e^{-\lambda t}$
- $f(t) = S(t)h(t) = \lambda e^{-\lambda t}$
- $E(T) = 1/\lambda$
- $t_{med} = log(2)/\lambda$



Weibull Distribution

- The exponential distribution is easy to work with, but the constant hazard assumption is not often appropriate for describing the lifetimes of humans or animals.
- $h(t) = \alpha \lambda^{\alpha} t^{\alpha 1}$
- $H(t) = (\lambda t)^{\alpha}$
- $S(t) = e^{-(\lambda t)^{\alpha}}$
- $E(T) = \frac{\Gamma(1+1/\alpha)}{\lambda}$
- $t_{med} = \frac{[log(2)]^{1/\alpha}}{\lambda}$



In R

- We may generate random variables from the exponential or Weibull distribution using the functions "rexp" and "rweib".
- The functions "dweibull" and "pweibull" compute the p.d.f. and c.d.f., respectively, of the Weibull distribution.
- These functions use the arguments "shape" and "scale" to represent the parameters α and $1/\lambda$.

```
1 > tt.weib <- rweibull(1000, shape=1.5, scale=1/0.03)
2 > mean(tt.weib)
3 [1] 31.35497
4 > median(tt.weib)
5 [1] 26.84281
```

In R

• Plot the *Weibull survival function* with $\alpha = 1.5$ and $\lambda = 0.03$ by first defining a function "weibSurv" with these parameters and then using the "curve" function to plot the curve as follows:

```
1 > weibSurv <- function(t, shape, scale) pweibull(t, shape=shape,
2 scale=scale, lower.tail=F)
3 > curve(weibSurv(x, shape=1.5, scale=1/0.03), from=0, to=80,
4 ylim=c(0,1), ylab='Survival probability', xlab='Time')
```

Plot the <u>hazard function</u> with p.d.f. divided by the survival function:

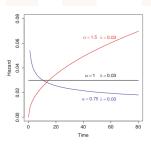
```
1 > weibHaz <- function(x, shape, scale) dweibull(x, shape=shape,
2 scale=scale)/pweibull(x, shape=shape, scale=scale,
3 lower.tail=F)
4 > curve(weibHaz(x, shape=1.5, scale=1/0.03), from=0, to=80,
5 ylab='Hazard', xlab='Time', col="red")
```

In R

• The other two curves obtained using $\alpha=1$ and $\alpha=0.75$ when calling the "curve" function. To place the additional curves on the plot, add "add=T" as an option to the "curve" function.

```
1 > curve(weibHaz(x, shape=1, scale=1/0.03), from=0, to=80,
2 ylab='Hazard', xlab='Time', col="black", add = TRUE)
3 > curve(weibHaz(x, shape=0.75, scale=1/0.03), from=0, to=80,
4 ylab='Hazard', xlab='Time', col="blue", add = TRUE)
```

Fig. 2.4 Weibull hazard functions



M.L.E

No Censoring

$$L(\lambda; t_1, t_2, ..., t_n) = f(t_1, \lambda) \cdot f(t_2, \lambda) \cdot \cdots \cdot f(t_n, \lambda) = \prod_{i=1}^n f(t_i, \lambda)$$

Right-censoring

$$L(\lambda; t_1, t_2, ..., t_n) = \prod_{i=1}^n f(t_i, \lambda)^{\delta_i} \cdot S(t_i, \lambda)^{1-\delta_i} = \prod_{i=1}^n h(t_i, \lambda)^{\delta_i} \cdot S(t_i, \lambda)$$

• This expression means that when t_i is an observed death, the censoring indicator is $\delta_i = 1$, and we enter a p.d.f. factor. When t_i is a censored observation, we have $\delta_i = 0$ we enter a survival factor.

Example

Exponential M.L.E

$$L(\lambda) = \prod_{i=1}^{n} [\lambda e^{-t_i/\lambda}]^{\delta_i} [e^{-\lambda t_i}]^{1-\delta_i}$$

 $d = \sum_{i=1}^{n} \delta_i$, means the total of the death

 $V = \sum_{i=1}^{n} t_i$, means the total of the time of patients on study

- Log-Likelihood $\ell(\lambda) = d \log \lambda \lambda V$
- First Derivative(Score Function) $\ell'(\lambda) = \frac{d}{\lambda} - V \Rightarrow \hat{\lambda} = \frac{d}{V}$



Example

• Second Derivative(*Information*) $\ell''(\lambda) = -\frac{d}{\lambda^2} = -I(\lambda)$

• The inverse of the (*Information*) is approximately the variance of the m.l.e
$$var(\hat{\lambda}) \approx I^{-1}(\lambda) = \lambda^2/d$$

- Substitute the $\hat{\lambda}$ for λ to obtain the observed information $I(\hat{\lambda})$, and get an estimate of the variance $v \hat{a} r(\hat{\lambda}) \approx I^{-1}(\hat{\lambda}) = \hat{\lambda}^2/d = d/V^2$
- Use this formula to carry out hypothesis tests or find a confidence interval for λ .

