

REGRESSION METHODS

MULTIPLE COMPARISONS

ANOVA: One-Way Model

What are the groups with differences in means?

MULTIPLE COMPARISONS:

$$\mu_0 = \mu_1?$$

$$\mu_0 = \mu_2?$$
 Pairwise comparisons
$$\mu_1 = \mu_2?$$

$$(\mu_1 + \mu_2)/2 = \mu_0?$$
 — Non-pairwise comparison



Multiple Comparisons: Family-wise error rates

- Illustrating the multiple comparison problem
 - Truth: null hypotheses
 - Tests: pairwise comparisons each at the 5% level.

What is the probability of rejecting at least one?

#groups = K	2	3	4	5	6	7	8	9	10
#pairwise comparisons C = K(K-1)/2	1	3	6	10	15	21	28	36	45
P(at least one sig) =1-(1-0.05) ^C	0.05	0.143	0.265	0.401	0.537	0.659	0.762	0.842	0.901

That is, if you have three groups and make pairwise comparisons, each at the 5% level, your family-wise error rate (probability of making at least one false rejection) is over 14%!

Need to address this issue! Several methods!!!



- Several methods:
 - None (no adjustment)
 - Bonferroni
 - Holm
 - Hochberg
 - Hommel
 - BH
 - BY
 - FDR

- ...

Available in R



- Bonferroni adjustment: for C tests performed, use level a/C (or multiply p-values by C).
 - Simple
 - Conservative
 - Must decide on number of tests beforehand
 - Widely applicable
 - Can be done without software!



- FDR (False Discovery Rate)
 - Less conservative procedure for multiple comparisons
 - Among rejected hypotheses, FDR controls the expected proportion of incorrectly rejected null hypotheses (that is, type I errors).



This option considers all pairwise comparisons

```
> ## call library for multiple comparisons
> library(multcomp)
> ## fit model
> fit1 = lm(chol \sim -1 + factor(rs174548))
>
> ## all pairwise comparisons
> ## -- first, define matrix of contrasts
> M = contrMat(table(rs174548), type="Tukey")
> M
         Multiple Comparisons of Means: Tukey Contrasts
       0 1 2
1 - 0 - 1 \quad 1 \quad 0
2 - 0 - 1 0 1
2 - 1 \quad 0 - 1 \quad 1
> ## -- second, obtain estimates for multiple comparisons
> mc = glht(fit1, linfct =M)
```

Stands for general linear hypothesis testing



```
> ## -- third, adjust the p-values (or not) for multiple comparisons
> summary(mc, test=adjusted("none"))
        Simultaneous Tests for General Linear Hypotheses
Multiple Comparisons of Means: Tukey Contrasts
Fit: lm(formula = chol \sim -1 + factor(rs174548))
Linear Hypotheses:
          Estimate Std. Error t value Pr(>|t|)
1 - 0 == 0 6.802 2.321 2.930 0.00358 **
2 - 0 == 0 5.438 4.540 1.198 0.23167
2 - 1 == 0 -1.364 4.665 -0.292 0.77015
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Adjusted p values reported -- none method)
```



```
> summary(mc, test=adjusted("bonferroni"))
        Simultaneous Tests for General Linear Hypotheses
Multiple Comparisons of Means: Tukey Contrasts
Fit: lm(formula = chol \sim -1 + factor(rs174548))
Linear Hypotheses:
          Estimate Std. Error t value Pr(>|t|)
1 - 0 == 0 6.802 2.321 2.930 0.0107 *
2 - 0 == 0 5.438 4.540 1.198 0.6950
2 - 1 == 0 -1.364 4.665 -0.292 1.0000
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Adjusted p values reported -- bonferroni method)
```

1

```
> summary(mc, test=adjusted("fdr"))
        Simultaneous Tests for General Linear Hypotheses
Multiple Comparisons of Means: Tukey Contrasts
Fit: lm(formula = chol \sim -1 + factor(rs174548))
Linear Hypotheses:
          Estimate Std. Error t value Pr(>|t|)
1 - 0 == 0 6.802 2.321 2.930 0.0107 *
2 - 0 == 0 5.438 4.540 1.198 0.3475
2 - 1 == 0 -1.364 4.665 -0.292 0.7702
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Adjusted p values reported -- fdr method)
```

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Multiple Comparisons

- What about using other adjustment methods?
 - For example, we used:

```
> summary(mc, test=adjusted("bonferroni"))
(all pairwise comparisons, with Bonferroni adjustment)
> summary(mc, test=adjusted("fdr"))
(all pairwise comparisons, with FDR adjustment)
```

Other options are:

```
summary(mc, test=adjusted("holm"))
summary(mc, test=adjusted("hochberg"))
summary(mc, test=adjusted("hommel"))
summary(mc, test=adjusted("BH"))
summary(mc, test=adjusted("BY"))
```

Results, in this particular example, are basically the same, but they don't need to be! Different criteria could lead to different results!



GOAL: Comparison of means across K groups

Relationships:

$$\mu_{0} = \beta_{0}$$

$$\mu_{1} = \beta_{0} + \beta_{1}$$

$$\mu_{2} = \beta_{0} + \beta_{2}$$
...
$$\mu_{K-1} = \beta_{0} + \beta_{K-1}$$

One-way ANOVA:

 $H_0: \mu_0 = \mu_1 = ... = \mu_{K-1}$

H₁: not all means are equal

Rejected H₀?

YES

Multiple Regression:

Model: E[Y|groups]= β_0 + β_1 group₂ +...+ β_{k-1} group_k where group₁ is the reference group

$$H_0: \beta_1 = \beta_2 = ... = \beta_{k-1} = 0$$

 H_1 : not all β_i are equal to zero

Multiple Comparisons (control α overall)

REGRESSION METHODS

Two-way ANOVA models



ANOVA: Two-Way Model Motivation:

- Scientific question:
 - Assess the effect of rs174548 and sex on cholesterol levels.

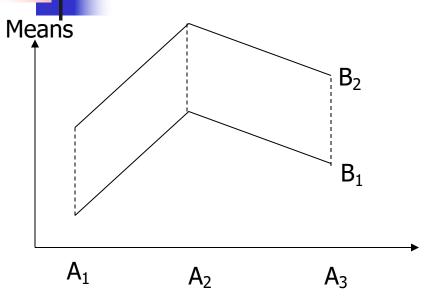


- Factors: A and B
- Goals:
 - Test for main effect of A
 - Test for main effect of B
 - Test for interaction effect of A and B

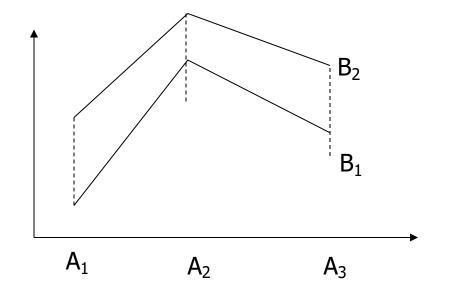
 To simplify discussion, assume that factor A has three levels, while factor B has two levels

Factor	Α
--------	---

		A_1	A_2	A ₃
tor B	B ₁	μ_{11}	μ_{21}	μ_{31}
Fac	B ₂	μ_{12}	μ_{22}	μ_{32}



Parallel lines = No interaction



Lines are not parallel = Interaction



Recall:

- Categorical variables can be represented with "dummy" variables
- Interactions are represented with "cross-products"



Model 1:

$$E[Y|A_2, A_3, B_2] = \beta_0 + \beta_1 A_2 + \beta_2 A_3 + \beta_3 B_2.$$

What are the means in each combination-group?

	A_1	A ₂	A_3
B ₁	$\mu_{11} = \beta_0$	$\mu_{21} = \beta_0 + \beta_1$	$\mu_{31} = \beta_0 + \beta_2$
B ₂	$\mu_{12} = \beta_0 + \beta_3$	$\mu_{22} = \beta_0 + \beta_1 + \beta_3$	$\mu_{32} = \beta_0 + \beta_2 + \beta_3$

ANO

ANOVA: Two-Way Model

Model 1:

$$E[Y|A_2, A_3, B_2] = \beta_0 + \beta_1 A_2 + \beta_2 A_3 + \beta_3 B_2.$$

	A_1	A ₂	A_3
B ₁	$\mu_{11} = \beta_0$	$\mu_{21} = \beta_0 + \beta_1$	$\mu_{31} = \beta_0 + \beta_2$
B ₂	$\mu_{12} = \beta_0 + \beta_3$	$\mu_{22} = \beta_0 + \beta_1 + \beta_3$	$\mu_{32} = \beta_0 + \beta_2 + \beta_3$

Model with no interaction:

- •Difference in means between groups defined by factor B does not depend on the level of factor A.
- •Difference in means between groups defined by factor A does not depend on the level of factor B.



Model 2:

$$E[Y|A_2, A_3, B_2] = \beta_0 + \beta_1A_2 + \beta_2A_3 + \beta_3B_2 + \beta_4A_2B_2 + \beta_5A_3B_2$$

What are the means in each combination-group?

	A_1	A ₂	A ₃
B ₁	$\mu_{11} = \beta_0$	$\mu_{21} = \beta_0 + \beta_1$	$\mu_{31} = \beta_0 + \beta_2$
B ₂	$\mu_{12} = \beta_0 + \beta_3$	$\mu_{22} = \beta_0 + \beta_1 + \beta_3 + \beta_4$	$\mu_{32} = \beta_0 + \beta_2 + \beta_3 + \beta_5$

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ANOVA: Two-Way Model

- Three (possible) tests
 - Interaction of A and B (may want to start here)
 - Rejection would imply that differences between means of A depends on the level of B (and vice-versa) so stop
 - Main effect of A
 - Test only if no interaction
 - Main effect of B
 - Test only if no interaction

[Note: If you have one observation per cell, you cannot test interaction!]

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ANOVA: Two-Way Model

Model without interaction

$$E[Y|A_2, A_3, B_2] = \beta_0 + \beta_1 A_2 + \beta_2 A_3 + \beta_3 B_2.$$

How do we test for main effect of factor A?

$$H_0$$
: $\beta_1 = \beta_2 = 0$ vs. H_1 : β_1 or β_2 not zero

How do we test for main effect of factor B?

$$H_0$$
: $\beta_3=0$ vs. H_1 : β_3 not zero



Model with interaction:

$$E[Y|A_2, A_3, B_2] = \beta_0 + \beta_1 A_2 + \beta_2 A_3 + \beta_3 B_2 + \beta_4 A_2 B_2 + \beta_5 A_3 B_2$$

How do we test for interactions?

$$H_0$$
: $\beta_4 = \beta_5 = 0$ vs.
 H_1 : β_4 or β_5 not zero

IMPORTANT:

If you reject the null, do not test main effects!!!

ANOVA: Two-Way Model (without interaction)

```
> fit1 = lm(chol \sim factor(sex) + factor(rs174548))
> summary(fit1)
Call:
lm(formula = chol ~ factor(sex) + factor(rs174548))
Residuals:
              10 Median
    Min
                               3Q
                                      Max
-66.6534 -14.4633 -0.6008 15.4450 57.6350
Coefficients:
                   Estimate Std. Error t value Pr(>|t|)
(Intercept)
                   175.365 1.786 98.208 < 2e-16 ***
                    11.053 2.126 5.199 3.22e-07 ***
factor(sex)1
factor(rs174548)1 7.236 2.250 3.215 0.00141 **
factor(rs174548)2
                      5.184 4.398 1.179 0.23928
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.24 on 396 degrees of freedom
Multiple R-squared: 0.08458, Adjusted R-squared: 0.07764
F-statistic: 12.2 on 3 and 396 DF, p-value: 1.196e-07
> fit0 = lm(chol ~ factor(sex))
> anova(fit0,fit1)
Analysis of Variance Table
Model 1: chol ~ factor(sex)
Model 2: chol ~ factor(sex) + factor(rs174548)
           RSS Df Sum of Sq F Pr(>F)
 Res.Df
    398 183480
1
    396 178681 2 4799.1 5.318 0.005259 **
```



ANOVA: Two-Way Model (without interaction)

```
> fit1 = lm(chol \sim factor(sex) + factor(rs174548))
> summary(fit1)
Call:
lm(formula = chol ~ factor(sex) + factor(rs174548))
Residuals:
     Min
               10 Median
                                         Max
-66.6534 -14.4633 -0.6008 15.4450 57.6350
Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
(Intercept)
                      175.365
                                   1.786 98.208 < 2e-16 ***
                       11.053
                                   2.126
                                           5.199 3.22e-07 ***
factor(sex)1
                        7.236
                                   2.250
                                           3.215 0.00141 **
factor(rs174548)1
factor(rs174548)2
                        5.184
                                   4.398
                                           1.179 0.23928
Residual standard error: 21.24 on 396 degrees of freedom
Multiple R-squared: 0.08458,
                                Adjusted R-squared: 0.07764
F-statistic: 12.2 on 3 and 396 DF, p-value: 1.196e-07
> anova(fit0,fit1)
Analysis of Variance Table
Model 1: chol ~ factor(sex)
Model 2: chol ~ factor(sex) + factor(rs174548)
  Res.Df
            RSS Df Sum of Sq
                                     Pr (>F)
     398 183480
                      4799.1 5.318 0.005259 **
     396 178681 2
```

Interpretation of results:

- Estimated mean cholesterol for male C/C group: 175.37 mg/dl
- Estimated difference in mean cholesterol levels between females and males adjusted by genotype: 11.053 mg/dl
- Estimated difference in mean cholesterol levels between C/G and C/C groups adjusted by sex: 7.236 mg/dl
- Estimated difference in mean cholesterol levels between G/G and C/C groups adjusted by sex: 5.184 mg/dl
- There is evidence that cholesterol is associated with sex (p< 0.001).
- There is evidence that cholesterol is associated with genotype (p=0.005)



ANOVA: Two-Way Model (without interaction)

In words:

- Adjusting for sex, the difference in mean cholesterol comparing C/G to C/C is 7.236 and comparing G/G to C/C is 5.184.
 - This difference does not depend on sex
 - (this is because the model does not have an interaction between sex and genotype!)

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ANOVA: Two-Way Model (with interaction)

```
> fit2 = lm(chol ~ factor(sex) * factor(rs174548))
> summary(fit2)
Call:
lm(formula = chol ~ factor(sex) * factor(rs174548))
Residuals:
              10 Median
    Min
                              3Q
                                      Max
-70.5286 -13.6037 -0.9736 14.1709 54.8818
Coefficients:
                                  Estimate Std. Error t value Pr(>|t|)
                                  178.1182 2.0089 88.666 < 2e-16 ***
(Intercept)
                                    5.7109 2.7982 2.041 0.04192 *
factor(sex)1
                                              3.1306 0.307 0.75933
factor(rs174548)1
                                  0.9597
factor(rs174548)2
                                  -0.2015 6.4053 -0.031 0.97492
factor(sex)1:factor(rs174548)1
                                  12.7398 4.4650 2.853 0.00456 **
factor(sex)1:factor(rs174548)2 10.2296 8.7482 1.169 0.24297
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '1
Residual standard error: 21.07 on 394 degrees of freedom
Multiple R-squared: 0.1039, Adjusted R-squared: 0.09257
F-statistic: 9.14 on 5 and 394 DF, p-value: 3.062e-08
```


ANOVA: Model comparison

ANOVA: Two-Way Model (with interaction)

```
Call:
lm(formula = chol ~ factor(sex) * factor(rs174548))
Residuals:
    Min
               10 Median
                                30
                                        Max
-70.5286 -13.6037 -0.9736 14.1709 54.8818
Coefficients:
                                    Estimate Std. Error t value Pr(>|t|)
(Intercept)
                                    178.1182
                                                 2.0089 88.666 < 2e-16 ***
                                      5.7109
factor(sex)1
                                                 2.7982 2.041 0.04192 *
                                      0.9597
                                                 3.1306 0.307 0.75933
factor(rs174548)1
                                     -0.2015 6.4053 -0.031 0.97492
factor(rs174548)2
                                                 4.4650
factor(sex)1:factor(rs174548)1
                                     12.7398
                                                          2.853 0.00456 **
                                     10.2296
                                                 8.7482 1.169 0.24297
factor (sex) 1: factor (rs174548) 2
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' 1
Residual standard error: 21.07 on 394 degrees of freedom
Multiple R-squared: 0.1039,
                              Adjusted R-squared: 0.09257
F-statistic: 9.14 on 5 and 394 DF, p-value: 3.062e-08
> anova(fit1,fit2)
Analysis of Variance Table
Model 1: chol ~ factor(sex) + factor(rs174548)
Model 2: chol ~ factor(sex) *
                                factor (rs174548)
  Res.Df
           RSS Df Sum of Sq
                                  F Pr(>F)
     396 178681
     394 174902
                        3779 4.2564 0.01483 *
                0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Signif. codes:
```

> fit2 = lm(chol ~ factor(sex) * factor(rs174548))

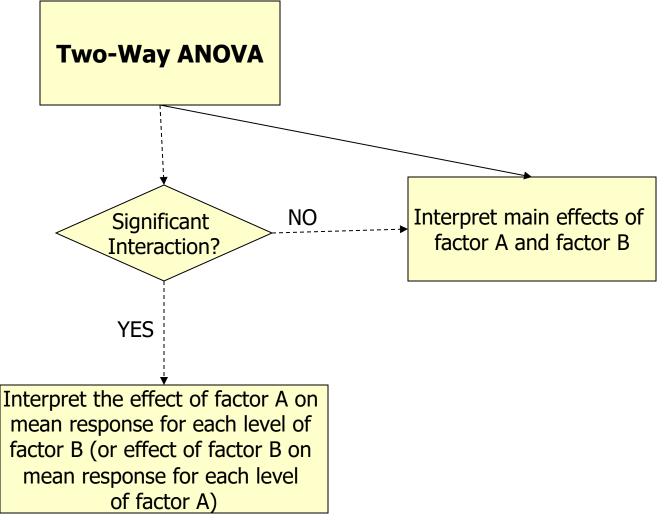
> summary(fit2)

Interpretation of results:

- Estimated mean cholestero for male C/C group:
 178.12 mg/dl
- Estimated mean cholestero for female C/C group? (178.12 + 5.7109) mg/dl
- Estimated mean cholestero for male C/G group: (178.12 +0.9597) mg/dl
- Estimated mean cholestero for female C/G group: (178.12 + 5.7109 + 0.9597 + 12.7398) mg/dl
-

 There is evidence for an interaction between sex and genotype (p= 0.015)







ANalysis of COVAriance Models (ANCOVA) Motivation:

- Scientific question:
 - Assess the effect of rs174548 on cholesterol levels adjusting for age



ANalysis of COVAriance Models (ANCOVA)

- ANOVA with one or more continuous variables
 - Equivalent to regression with "dummy" variables and continuous variables
 - Primary comparison of interest is across k groups defined by a categorical variable, but the k groups may differ on some other potential predictor or confounder variables [also called covariates].



ANalysis of COVAriance Models (ANCOVA)

- To facilitate discussion assume
 - Y: continuous response (e.g. cholesterol)
 - X: continuous variable (e.g. age)
 - Z: dummy variable (e.g. indicator of C/G or G/G versus C/C)

• Model:
$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \beta_3 XZ + \varepsilon$$

Interaction term

Note that:

$$Z = 0 \Rightarrow E[Y \mid X, Z = 0] = \beta_0 + \beta_1 X$$

$$Z = 1 \Rightarrow E[Y \mid X, Z = 1] = (\beta_0 + \beta_2) + (\beta_1 + \beta_3) X$$

This model allows for different intercepts/slopes for each group.

ANCOVA

- Testing coincident lines: $H_0: \beta_2 = 0, \beta_3 = 0$
 - Compares overall model with reduced model

$$Y = \beta_0 + \beta_1 X + \varepsilon$$

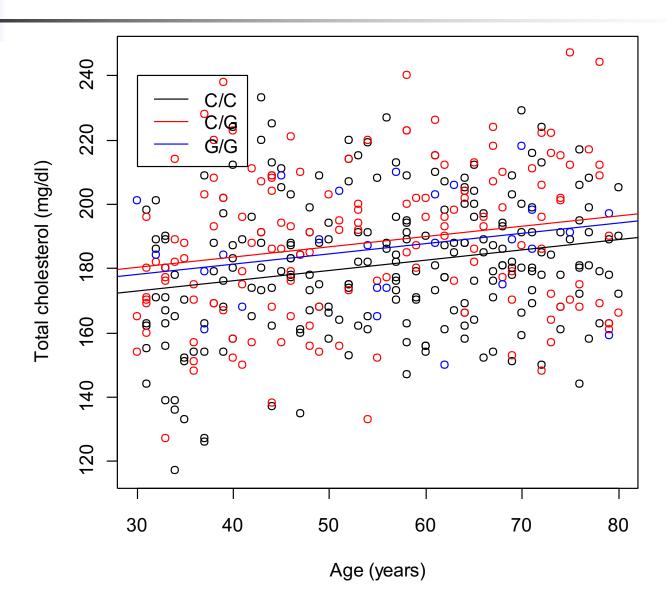
- Testing parallelism: $H_0: \beta_3 = 0$
 - Compares overall model with reduced model

$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \varepsilon$$

ANCOVA

```
> fit0 = lm(chol \sim factor(rs174548))
> summary(fit0)
Call:
lm(formula = chol ~ factor(rs174548))
Residuals:
     Min
                     Median
                10
                                   30
                                           Max
-64.06167 -15.91338 -0.06167 14.93833 59.13605
Coefficients:
                   Estimate Std. Error t value Pr(>|t|)
                   181.062
                                 1.455 124.411 < 2e-16 ***
(Intercept)
                    6.802
factor (rs174548)1
                                2.321 2.930 0.00358 **
factor(rs174548)2 5.438
                                 4.540 1.198 0.23167
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.93 on 397 degrees of freedom
Multiple R-squared: 0.0221, Adjusted R-squared: 0.01718
F-statistic: 4.487 on 2 and 397 DF, p-value: 0.01184
> anova(fit0)
Analysis of Variance Table
Response: chol
                   Df Sum Sq Mean Sq F value Pr(>F)
factor(rs174548)
                                2157 4.4865 0.01184 *
                        4314
                    2
                  397 190875
Residuals
                                 481
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
```

```
> fit1 = lm(chol \sim factor(rs174548) + age)
> summary(fit1)
Call:
lm(formula = chol ~ factor(rs174548) + age)
Residuals:
              1Q Median
    Min
                               30
                                      Max
-57.2089 -14.4293 0.4443 14.2652 55.8985
Coefficients:
                    Estimate Std. Error t value Pr(>|t|)
                   163.28125 4.36422 37.414 < 2e-16 ***
(Intercept)
factor(rs174548)1 7.30137 2.27457 3.210 0.00144 **
factor(rs174548)2 5.08431 4.44331 1.144 0.25321
                     0.32140 0.07457 4.310 2.06e-05 ***
age
Residual standard error: 21.46 on 396 degrees of freedom
Multiple R-squared: 0.06592, Adjusted R-squared: 0.05884
F-statistic: 9.316 on 3 and 396 DF, p-value: 5.778e-06
> anova(fit0,fit1)
Analysis of Variance Table
Model 1: chol ~ factor(rs174548)
Model 2: chol ~ factor(rs174548) + age
 Res.Df
           RSS Df Sum of Sq F Pr(>F)
    397 190875
    396 182322 1 8552.9 18.577 2.062e-05 ***
Signif. codes: 0 \***' 0.001 \**' 0.01 \*' 0.05 \'.' 0.1 \' 1
```



```
> fit2 = lm(chol ~ factor(rs174548) * age)
> summary(fit2)
Call:
lm(formula = chol ~ factor(rs174548) * age)
Residuals:
                 Median
    Min
             1Q
                             3Q
                                    Max
-57.5425 -14.3002 0.7131 14.2138 55.7089
Coefficients:
                       Estimate Std. Error t value Pr(>|t|)
                                 5.79545 28.323 < 2e-16 ***
(Intercept)
                      164.14677
factor(rs174548)1
                        3.42799 8.79946 0.390 0.69707
factor(rs174548)2
                      16.53004 18.28067 0.904 0.36642
                       age
factor(rs174548)1:age
                       0.07159 0.15617 0.458 0.64692
                       -0.20255 0.31488 -0.643 0.52043
factor(rs174548)2:age
Residual standard error: 21.49 on 394 degrees of freedom
Multiple R-squared: 0.06777, Adjusted R-squared: 0.05594
F-statistic: 5.729 on 5 and 394 DF, p-value: 4.065e-05
```

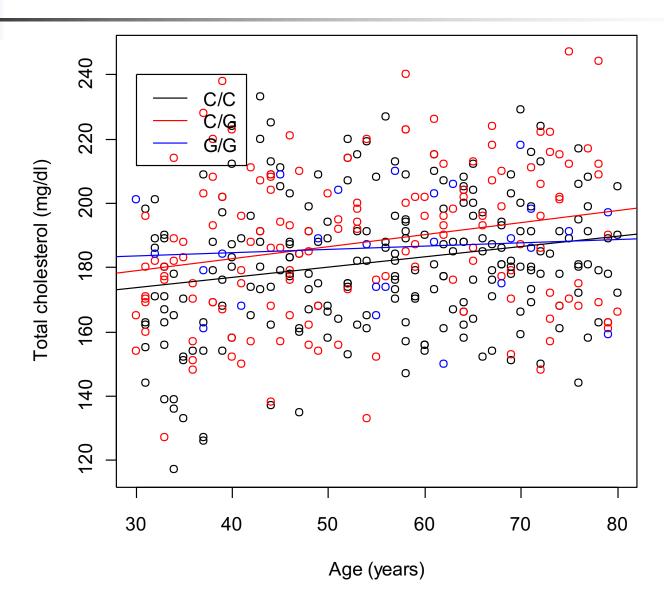
```
> fit0 = lm(chol \sim age)
> summary(fit0)
Call:
lm(formula = chol ~ age)
Residuals:
         1Q Median 3Q
   Min
                                Max
-60.453 -14.643 -0.022 14.659 58.995
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 166.90168  4.26488  39.134  < 2e-16 ***
          age
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 21.69 on 398 degrees of freedom
Multiple R-squared: 0.04099, Adjusted R-squared: 0.03858
F-statistic: 17.01 on 1 and 398 DF, p-value: 4.522e-05
> anova(fit0,fit2)
Analysis of Variance Table
Model 1: chol ~ age
Model 2: chol ~ factor(rs174548) * age
 Res.Df RSS Df Sum of Sq F Pr(>F)
   398 187187
   394 181961 4 5226.6 2.8293 0.02455 *
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
```

Test of coincident lines

Test of parallel lines

```
> anova(fit1,fit2)
Analysis of Variance Table

Model 1: chol ~ factor(rs174548) + age
Model 2: chol ~ factor(rs174548) * age
   Res.Df   RSS Df Sum of Sq   F Pr(>F)
1    396 182322
2   394 181961 2   361.11 0.391 0.6767
```



- In summary:
 - If the slopes are not equal, then age is an effect modifier

$$E[Y | x, z] = \beta_0 + \beta_1 x + \beta_2 (CG) + \beta_3 (GG) + \beta_4 (x * CG) + \beta_5 (x * GG)$$

If the slopes are the same,

$$E[Y | x, z] = \beta_0 + \beta_1 x + \beta_2 (CG) + \beta_3 (GG)$$

If the slopes are the same,

$$E[Y | x, z] = \beta_0 + \beta_1 x + \beta_2 (CG) + \beta_3 (GG)$$

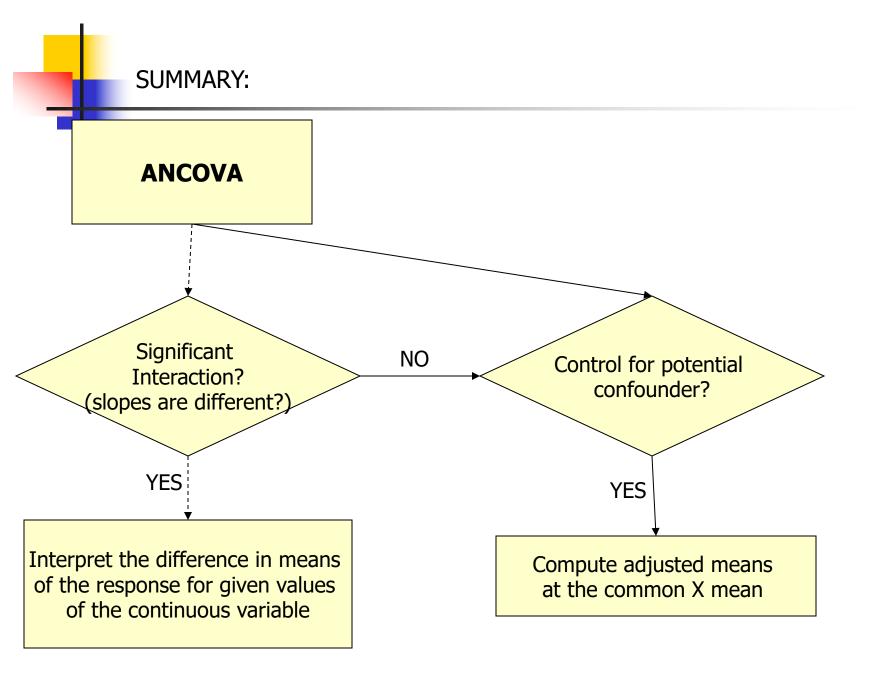
- then one can obtain adjusted means for the three genotypes using the mean age over all groups
 - For example, the adjusted means for the three groups would be

$$\overline{Y}_{1}(adj) = \hat{\beta}_{0} + \overline{x} \hat{\beta}_{1}$$

$$\overline{Y}_{2}(adj) = (\hat{\beta}_{0} + \hat{\beta}_{2}) + \overline{x} \hat{\beta}_{1}$$

$$\overline{Y}_{3}(adj) = (\hat{\beta}_{0} + \hat{\beta}_{3}) + \overline{x} \hat{\beta}_{1}$$

```
> ## mean cholesterol for different genotypes adjusted by age
> predict(fit1, new=data.frame(age=mean(age),rs174548=0))
180.9013
> predict(fit1, new=data.frame(age=mean(age),rs174548=1))
188, 2026
> predict(fit1, new=data.frame(age=mean(age),rs174548=2))
185.9856
> ## mean cholesterol for different genotypes adjusted by age
> mean(predict(fit1, new=data.frame(age=age,rs174548=0)))
180.9013
> mean(predict(fit1, new=data.frame(age=age,rs174548=1)))
188.2026
> mean(predict(fit1, new=data.frame(age=age,rs174548=2)))
185.9856
```





We have considered:

- ANOVA and ANCOVA
 - Interpretation
 - Estimation
 - Interaction

Multiple comparisons

Exercise

- Work on Exercise 9-12
 - Try each exercise on your own
 - Make note of any questions or difficulties you have
 - At 3:30ET?? we will meet as a group to go over the solutions and discuss your questions