

# Weighted Projection Quantiles Algorithm

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**Algorithm to calculate weighted projection quantile** along the vector  $\mathbf{u} \in \mathcal{B}_p$ , given a set of observations  $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n$ :

1. **Compute  $Q_{proj}(\mathbf{u})$ , the projection quantile along  $\mathbf{u}$**

- Project each  $\mathbf{X}_i$  along  $\mathbf{u}$  to obtain  $X_{\mathbf{u}i} = \frac{\langle \mathbf{X}_i, \mathbf{u} \rangle}{\|\mathbf{u}\|}$ , for  $i = 1, 2, \dots, n$ .
- Find  $\alpha = \frac{1+\|\mathbf{u}\|}{2}$ -th quantile of  $X_{\mathbf{u}1}, \dots, X_{\mathbf{u}n}$ , say  $q_{\mathbf{u}}$ .
- $Q_{proj}(\mathbf{u}) = q_{\mathbf{u}}\mathbf{e}_{\mathbf{u}}$ ,  $\mathbf{e}_{\mathbf{u}} = \mathbf{u}/\|\mathbf{u}\|$  being the unit vector along  $\mathbf{u}$ .

2. **Compute Weights corresponding to this projection quantile  $Q_{proj}(\mathbf{u})$**

- Compute global weights for the direction vector  $\mathbf{u}$  by  $k$ -mean distance:
  - Compute  $k$ -mean distance corresponding to  $Q_{proj}(\mathbf{u})$  using  $\bar{d}_k = \frac{1}{n} \sum_{i=1}^n d_i \mathbb{I}_{\{d_i < d_{(k)}\}}$ , where  $d_i$  is the euclidean distance of  $\mathbf{X}_i$  from  $Q_{proj}(\mathbf{u})$  given by  $\|\mathbf{X}_i - Q_{proj}(\mathbf{u})\|$ .  $k$  is a tuning parameter.
  - Compute the weights corresponding to  $\mathbf{u}$ :

$$w_{\mathbf{u}} = \exp(-a.d_k)$$

where  $a$  is a tuning parameter.

- Compute weights for each sample point  $\mathbf{X}_i; i = 1, 2, \dots, n$ :
  - Compute the orthogonal Norms by  $\|\mathbf{X}_{\mathbf{u}\perp i}\| = \|\mathbf{X}_i - X_{\mathbf{u}i}\mathbf{e}_{\mathbf{u}}\|$ .
  - Compute weight of  $i^{th}$  sample:

$$w_{2i} = \exp \left[ -b \frac{\|\mathbf{X}_{\mathbf{u}\perp i}\|}{\|\mathbf{X}_i\|} \right] \mathbb{I}_{\{\|\mathbf{X}_{\mathbf{u}\perp i}\| \leq \epsilon\}}$$

$b, \epsilon$  being tuning parameters.

3. **Compute the weighted projection quantile**

- Suppose there are  $m$  observations with non-zero weights  $w_{2i}$ , with indices  $i_1, i_2, \dots, i_m$ . Define  $\tilde{X}_{\mathbf{u}i_j} = w_{\mathbf{u}} w_{2i_j} X_{\mathbf{u}i_j}$ .
- Find  $\alpha = \frac{1+\|\mathbf{u}\|}{2}$ -th quantile of  $\tilde{X}_{\mathbf{u}i_1}, \dots, \tilde{X}_{\mathbf{u}i_m}$ . Let it be  $\tilde{q}_{\mathbf{u}}$ .
- Find the weighted projection quantile as  $\tilde{Q}_{proj}(\mathbf{u}) = \tilde{q}_{\mathbf{u}}\mathbf{e}_{\mathbf{u}}$ .

**Definition** Given a random vector  $\mathbf{X} \in \mathbb{R}^p$  that follows a multivariate distribution  $F$ , and a point  $\mathbf{p} \in \mathbb{R}^p$ , find  $\alpha_{\mathbf{p}}$  such that  $\|\mathbf{p}\|$  is the  $\alpha_{\mathbf{p}}$ -th quantile for the projection of  $\mathbf{X}$  on  $\mathbf{p}$ , say  $X_{\mathbf{p}}$ . Then the **Projection Quantile Depth** (PQD) at  $\mathbf{p}$  with respect to  $F$  is defined as

$$D(\mathbf{p}, F) = \exp(-\alpha_{\mathbf{p}})$$

Given data  $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n$ , the PQD at a given  $\mathbf{p}$  can be estimated by finding the two nearest points on either side of  $\|\mathbf{p}\|$  along  $\mathbf{p}$ , say  $\mathbf{p}_1, \mathbf{p}_2$ , obtain their corresponding quantiles, say  $\alpha_1, \alpha_2$  respectively, then estimate  $\alpha_{\mathbf{p}}$  by a linear approximation:

$$\hat{\alpha}_{\mathbf{p}} = \frac{(\alpha_1 - \alpha_2)(\|\mathbf{p}\| - \|\mathbf{p}_1\|)}{\|\mathbf{p}_1\| - \|\mathbf{p}_2\|} + \alpha_1$$

and plugging it in the above definition.

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**Algorithm 1** Algorithm for PQD-based classification

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- 1: **procedure** PQDClassifier(training data  $\mathbf{X}_i \in \mathbb{R}^{n_i \times p}$  with class labels  $i$ ;  $i = 1, 2, \dots, k$ , new data  $\mathbf{x}_{new} \in \mathbb{R}^p$ )
- 2:   Set  $i = 1$ .
- 3:   *top*:
- 4:   Estimate from the sample the PQD of  $\mathbf{p}$  with respect to the  $i^{th}$  population, say  $D(\mathbf{x}_{new}, \mathbf{X}_i)$ .
- 5:   **if**  $i = k$  **then Stop**
- 6:   **else**
- 7:     Set  $i \leftarrow i + 1$ , **goto top**
- 8:
- 9:   Find  $c$  that maximizes the PQD of  $\mathbf{x}_{new}$  w.r.t. all possible classes:

$$D(\mathbf{x}_{new}, \mathbf{X}_c) = \max\{D(\mathbf{x}_{new}, \mathbf{X}_i) : i = 1, 2, \dots, k\}$$

- 10:   Assign class  $c$  to new data  $\mathbf{x}_{new}$ .
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**Note** One can define a weighted version of PQD by replacing  $X_{\mathbf{p}}$  by their weighted version  $\tilde{X}_{\mathbf{p}}$ . A weighted classification scheme follows similarly.

## New notion of data depth

We define a data depth based on projection quantiles (which, of course, can be extended to the weighted versions), on the lines of Zuo's Projection Depth [ref?].

**Definition** Given a random vector  $X \in \mathbb{R}^p$  that follows a multivariate distribution  $F$ , and a point  $x \in \mathbb{R}^p$ , define an outlyingness function

$$O(x, F) = \sup_{\|u\|=1} \left| F_u(x_u) - \frac{1}{2} \right|$$

where  $F_u$  is the distribution function of  $X_u$ , and  $x_u = \langle x, u \rangle / \|u\|$ . Then we define the **Projection Quantile Depth** at  $x$  with respect to  $F$  as

$$PQD(x, F) = \frac{1}{1 + O(x, F)}$$

We now follow the order in Zuo's paper to obtain the properties of the PQD.

**(P1) Affine invariance** Say  $G$  is the cdf of  $AX + b$ . Then we have  $O(Ax + b, G) = O(x, F)$  because for all  $u \in \mathbb{R}^p$  we have

$$F(x) = G(Ax + b) \Rightarrow F_u(x_u) = G_{Au+b}((Ax + b)_{Au+b})$$

where subscripting denotes projection.

**(P2) Quasi-concavity** for  $0 < \lambda < 1, z = \lambda x + (1 - \lambda)y$ ,

$$\begin{aligned} F_u(z_u) &= F_u(\lambda x_u + (1 - \lambda)y_u) \\ \Rightarrow \left| F_u(z_u) - \frac{1}{2} \right| &\leq \begin{cases} \left| \max\{F_u(x_u), F_u(y_u)\} - \frac{1}{2} \right| & \text{if } F_u(z_u) < 1/2 \\ \left| \min\{F_u(x_u), F_u(y_u)\} - \frac{1}{2} \right| & \text{otherwise} \end{cases} \\ &\leq \max \left\{ \left| F_u(x_u) - \frac{1}{2} \right|, \left| F_u(y_u) - \frac{1}{2} \right| \right\} \end{aligned}$$

and we finally get  $PQD(z, F) \geq \max\{PQD(x, F), PQD(y, F)\}$ .

**(P3) Vanishing at infinity** For  $\|x\| \rightarrow \infty$  we have  $F_u(x_u) \rightarrow 1$  hence  $PQD(x, F) \rightarrow 2/3$ . So we can define  $PQD^*(x, F) = PQD(x, F) - 2/3$  to obtain P3.

**(P4) Maximized at center of symmetric  $F$**  Halfspace symmetry of  $F$  w.r.t. center  $\theta$  implies (halfspace) symmetry of  $F_u$  around  $\theta_u$  for any  $u \in \mathbb{R}^p$ , so that  $F_u(\theta_u) = 1/2$  and subsequently  $PQD$  achieves its max value 1 at  $\theta$ .

**Example 1** For  $X \sim \mathcal{N}_2((0, 0)', I_2)$  we have  $X_u \sim N(0, 1)$  for any  $u \in \mathbb{R}^p : \|u\| = 1$ . Due to the symmetry of the distribution function maximizing  $|F_u(x_u) - 1/2|$  is equivalent to maximizing  $F_u(x_u) = \Phi(x_u) = \Phi(\langle x, u \rangle)$ .

Converting  $u$  to polar coordinates we can maximize over the angle  $\theta$  that  $u$  makes with the X-axis, and can easily find that the above function maximizes at  $\theta = \tan^{-1}(x_2/x_1)$ , which is the direction for  $x$  itself. Thus for given  $x$  the sup is obtained at  $u = x$ , and we have

$$PQD(x, F) = \frac{1}{1 + \Phi(\|x\|) - 1/2} = \frac{2}{1 + 2\Phi(\|x\|)} \quad (1)$$

Fig. 1 shows a comparison of PQD with Zuo's PD for standard bivariate normal.

**Note** In place of standard bivariate normal, one can use any circularly symmetric  $F$ , i.e. a distribution that has the same marginal distribution  $F_0$  along all one-dimensional projections to come up with an exact formula for  $PQD(x, F)$ . In that general scenario,  $\Phi(\cdot)$  will be replaced by  $F_0(\cdot)$  in (1).

## A general class of projection based depth functions

**Definition** We consider a random variable  $X \in \mathbb{R}^p$  following the distribution  $F$ . Define  $F_u$  as the projection of  $F$  towards any  $u \in \mathbb{R}^p$ . In this setup, we define a general class of outlyingness (or htpe) function:

$$O(x, F) = \sup_{\|u\|=1} f(x_u, \theta(F_u)) \quad (2)$$

where  $f(\cdot)$  is a function satisfying properties mentioned below.  $x_u = \langle x, u \rangle / \|u\|$ , and  $\theta : \mathcal{F} \mapsto \Theta \subseteq \mathbb{R}^k$  is a parameter functional (?) characterizing the distribution function  $F_u \in \mathcal{F}$ , the set of all univariate cdf's. Then we define the depth at  $x$  with respect to  $F$  as  $D(x, F) = 1/(1 + O(x, F))$ .

The necessary properties of  $f(\cdot)$  are formulated keeping necessary properties of a depth functions in mind [ref zuo-serf]:

**(F1) Affine invariant**  $f(x_u, \theta_u) = f(Ax_u + b, \theta_u^{AX+b})$  for non-singular matrix  $A \in \mathbb{R}^{p \times p}$ . where subscripting denotes projection.

**(F2) Quasi-convex in  $x_u$**  for  $0 < \lambda < 1, z = \lambda x + (1 - \lambda)y$ ,

$$f(z_u) = f(\lambda x_u + (1 - \lambda)y_u) \leq \max\{f(x_u, \theta_u), f(y_u, \theta_u)\}$$

from which quasi-concavity of  $D(x, F)$  follows.

**(F3) Monotonically increasing in  $x_u$**  ???

**(F4) Maximized at center of symmetric  $F$**  ??

**(F5) Uniformly continuous in both  $x_u$  and  $\theta(F_u)$**

**Theorem 0** Given an outlyingness function like (1) and an  $f(\cdot)$  satisfying properties (F1)-(F5), the resulting depth function is:

- (i) Affine invariant,
- (ii) Quasi-concave, and as a result  $D(x, F)$  decreases monotonically along any ray starting from the deepest point, [ref]
- (iii) Maximized at the center of a symmetric  $F$ .
- (iv) Uniformly continuous.

The proofs easily follow from the corresponding properties of  $f(\cdot)$ .

Now suppose that  $\theta_{nu}$  is the estimate of  $\theta_u$  from a size- $n$  iid sample. Following Zuo, we state some properties regarding the convergence  $\theta_{nu} \rightarrow \theta_u$ .

- (C0)  $\sup_{\|u\|=1} |\theta_i(F_u)| < \infty$  for  $i = 1, 2, \dots, k$ .
- (C1)  $\sup_{\|u\|=1} |\theta_i(F_{nu}) - \theta_i(F_u)| = o_P(1)$  for  $i = 1, 2, \dots, k$ .
- (C2)  $\sup_{\|u\|=1} |\theta_i(F_{nu}) - \theta_i(F_u)| = o(1)$  a.s. for  $i = 1, 2, \dots, k$ .
- (C3)  $\sup_{\|u\|=1} |\theta_i(F_{nu}) - \theta_i(F_u)| = O_P(1/\sqrt{n})$  for  $i = 1, 2, \dots, k$ .

Again following the lines of Zuo's paper, we say that:

**Theorem 1** Given the condition (C1) above (?):

- (i)  $\sup_{x \in \mathbb{R}^d} |D(x, F_n) - D(x, F)| = o_P(1)$  given (C2) holds.
- (ii)  $\sup_{x \in \mathbb{R}^d} |D(x, F_n) - D(x, F)| = O_P(1)$  given (C3) holds.
- (iii)  $\sup_{x \in \mathbb{R}^d} |D(x, F_n) - D(x, F)| = O(1/\sqrt{n})$  given (C4) holds.