

HOMEWORK-1

UJVAL BANGALORE UMESH

bangalore039@umn.edu

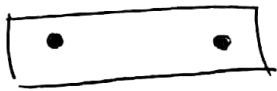
5202540

1)

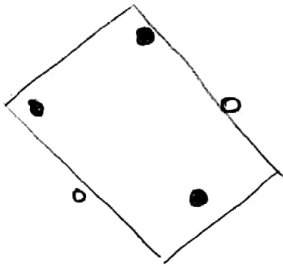
(a) Let H be the space of all rectangles.

The rectangles that are parallel to the axis and the ones that are not parallel to the axis.

The VC dimension of the parallel rectangles is 4.

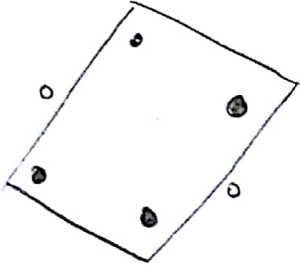


4 points can be shattered by rectangles which are parallel to the axes.



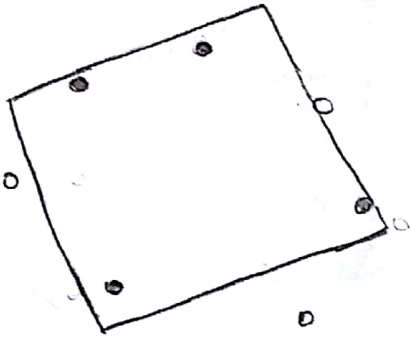
5 points can only be shattered by ~~non~~ non-axis rectangle.

Let us now consider 6 points.



A Rotatable Rectangle can shatter 6 points.

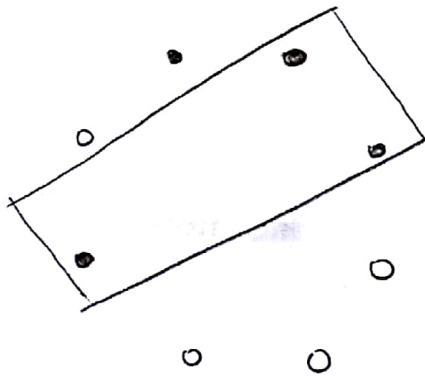
consideration of 7 points.



Rotatable rectangles can shatter 7 points.

To prove it can shatter n points, we need prove that it cannot shatter $n+1$ points.

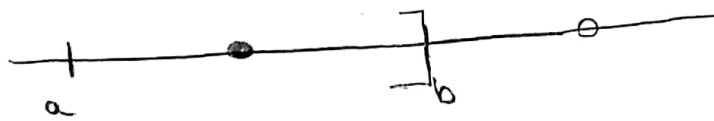
consider 8 points.



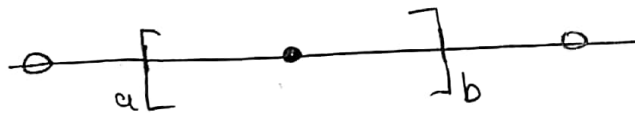
We can see that it cannot shatter 8 points.

Hence the VC dimension of rotatable rectangle is $\boxed{7}$.

b) let d_H be the VC dimension, of intervals in \mathbb{R} . The interval $[a, b]$ consider the interval $[a, b]$ having just one point b in it & ~~one~~ ~~outside~~ one outside



~~one~~ one point can be shattered.
let us consider a positive point in between two negative points.



To shatter this arrangement, ~~one~~ in the interval $[a, b]$.
The VC dimension is two [2].

$$2)(a) b(x|\theta) = \frac{1}{\theta} e^{-x/\theta}, \theta > 0$$

$$l(\theta|x) = p(x|\theta) = \prod_{t=1}^n p(x^t|\theta)$$

$$L(\theta|x) = \log L(\theta|x)$$

$$= \sum_{t=1}^n \log \left(\frac{1}{\theta} e^{-x^t/\theta} \right)$$

$$= \sum_{t=1}^n \log \left(\frac{1}{\theta} \right) + \sum_{t=1}^n \log (e^{-x^t/\theta})$$

$$\frac{\partial}{\partial \theta} \left[\sum_{t=1}^n \log \left(\frac{1}{\theta} \right) + \sum_{t=1}^n \frac{-x^t}{\theta} \right] = 0$$

$$-\frac{N\theta}{\theta^2} + \sum_{t=1}^n \frac{-x^t}{\theta^2} = 0$$

$$\frac{N\theta}{\theta^2} = \sum_{t=1}^n \frac{x^t}{\theta^2}$$

$$\theta = \frac{1}{N} \sum_{t=1}^n x^t, \theta > 0$$

$$(b) b(x|\theta) = \theta x^{\theta-1}, 0 \leq x \leq 1, 0 < \theta < \infty$$

$$L(\theta|x) = \log L(\theta|x)$$

$$= \sum_{t=1}^n \log b(x^t|\theta)$$

$$= \sum_{t=1}^n \log (\theta x^{\theta-1})$$

$$= \sum_{t=1}^n \log \theta + \sum_{t=1}^n \log x^{\theta-1}$$

$$= \frac{\partial}{\partial \theta} \left[\sum_{i=1}^n \log \theta + \sum_{i=1}^n (\theta - 1) \log x_i \right]$$

$$= \frac{N}{\theta} + \sum_{i=1}^n (\log x_i)(\theta - 1) = 0$$

$$\frac{N}{\theta} + \sum_{i=1}^n \log x_i = 0$$

$$\theta = \frac{-N}{\sum_{i=1}^n \log x_i}$$

$$3) f(x|\theta) = \frac{1}{\theta}, \quad 0 \leq x \leq \theta, \quad \theta > 0$$

The probability density function $f_n(x|\theta)$ of x_1, \dots, x_n has the form

$$f_n(x|\theta) = \frac{1}{\theta^n}$$

MLE of θ must be a value of θ for which

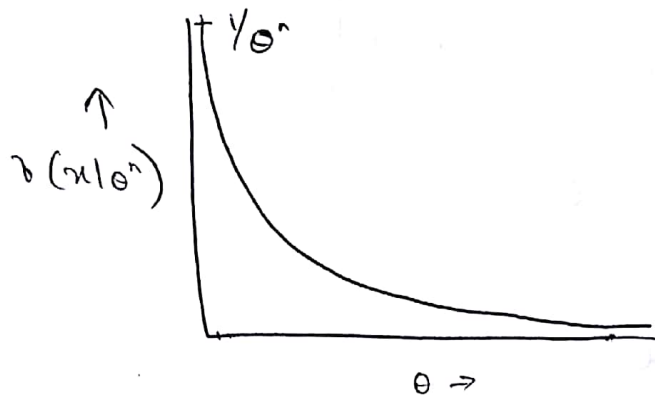
$\theta \geq x_i$ for $i=1, \dots, n$. and which maximises

$\frac{1}{\theta^n}$. Since it is a decreasing function, the

smallest θ such that $\theta \geq x_i$ for $i=1, \dots, n$, the estimate will be the smallest value of θ such that

$\theta \geq x_i$. Since this value is $\theta = \max[x_1, \dots, x_n]$.
Hence the MLE of θ is $\max[x_1, \dots, x_n]$.

The function of $\frac{1}{\theta^n}$ is



~~The value of~~

3 (c)

$$p(x/c_i) = \frac{1}{(2\pi)^{d/2} |\Sigma_i|^{1/2}} \exp \left\{ -\frac{1}{2} (x - \mu)^T \Sigma_i^{-1} (x - \mu) \right\}$$

Estimating the loglikelihood for means.

$$L(x/\mu, \Sigma_i) = \sum_{t=1}^N -\frac{d}{2} \log 2\pi - \frac{1}{2} \log |\Sigma_i| - \frac{1}{2} (x - \mu)^T \Sigma_i^{-1} (x - \mu)$$

$$\frac{\partial L}{\partial \mu} = 0$$

$$\sum_{t=1}^N (x^t - \mu) = 0 \quad N\mu = \sum_{t=1}^N x^t \Rightarrow \mu = \sum_{t=1}^N \frac{x^t}{N}$$

This can be extrapolated to $\mu_i = \sum_{t=1}^N \frac{x^t}{N}$
Means of different classes.

Estimating log likelihood with respect to covariance

$$\frac{\partial L}{\partial \Sigma} = 0$$

$$\frac{1}{2} \Sigma^{-1} = \frac{1}{2} \sum_{t=1}^N (x^t - \mu)^T (x^t - \mu)$$

$$\frac{\partial}{\partial u} [a^T A a]$$

$$= (A^T a)$$

$$\Sigma = \frac{\sum_{t=1}^N (x^t - \mu)^T (x^t - \mu)}{N}$$

$$\Sigma_i = \frac{(x^t - \mu_i)^T (x^t - \mu_i)}{N}$$

$$(E) \quad P(C_i/x) = \frac{P(x/C_i) \times P(C_i)}{P(x)}$$

Bayes formula.

$$96 \quad S_1 = S_2, \quad \text{sigma } S = 0.6 S_1 + 0.4 S_2.$$

Calculation is done based on this.

3)

d) S_1 and S_2 are diagonal.

$$S_1 = \lambda_1 I \quad \text{and} \quad S_2 = \lambda_2 I$$

$$S_1 - \lambda_1 I = 0$$

$$S_2 - \lambda_2 I = 0$$

$$\begin{bmatrix} s_{11} - \lambda_1 & & \\ & s_{22} - \lambda_1 & \\ & & \ddots \\ & & & s_{nn} - \lambda_1 \end{bmatrix} = 0$$

$$\begin{bmatrix} s_{11} - \lambda_2 & & \\ & s_{22} - \lambda_2 & \\ & & \ddots \\ & & & s_{nn} - \lambda_2 \end{bmatrix} = 0$$

Hence λ_1 and λ_2 are eigen values of S_1 and S_2 respectively. We could use the `eig` function in matlab to obtain S_1 and S_2 .

$$[V_1, D_1] = \text{eig}(S_1) \quad \& \quad [V_2, D_2] = \text{eig}(S_2)$$