Statistical Inference

1 Probability

Theorem 1.1. (The derivation of the pdf of order statistics in a direct way)

$$f_{x(r)}(t) = \frac{n!}{(r-1)!(n-r)!} [F_x(t)]^{r-1} f_x(t) [1 - F_x(t)]^{(n-r)}$$
(1)

Proof. By the definition of $F_{x(r)}(t)$, we have

$$F_{x_{(r)}}(t) = \mathbb{P}(x_{(r)} \le t) = \sum_{i=r}^{n} \mathbb{P}(\sum_{j=1}^{n} I(x_{j} \le t) = i)$$

$$= \sum_{i=r}^{n} \binom{n}{i} [F_{x}(t)]^{i} [1 - F_{x}(t)]^{n-i}. \tag{2}$$

To get the pdf, we can take derivative of both sides of (2)

$$\begin{split} f_{x_{(r)}}(t) &= \sum_{i=r}^{n} \binom{n}{i} \left\{ i [F_x(t)]^{i-1} f_x(t) [1 - F_x(t)]^{n-i} - (n-i) [F_x(t)]^i [1 - F_x(t)]^{n-i-1} f_x(t) \right\} \\ &= \sum_{i=r}^{n} \binom{n}{i} [F_x(t)]^{i-1} [1 - F_x(t)]^{n-i-1} f_x(t) \left\{ i [1 - F_x(t)] - (n-i) F_x(t) \right\} \\ &= \binom{n}{r} r [F_x(t)]^{r-1} [1 - F_x(t)]^{n-r} + \sum_{i=r+1}^{n} \binom{n}{i} i [F_x(t)]^{i-1} [1 - F_x(t)]^{n-i} \\ &- \sum_{i=r}^{n-1} \binom{n}{i} (n-i) [F_x(t)]^{i-1} [1 - F_x(t)]^{n-i} \\ &= \binom{n}{r} r [F_x(t)]^{r-1} [1 - F_x(t)]^{n-r} + \sum_{i=r}^{n-1} \binom{n}{i+1} (i+1) [F_x(t)]^i [1 - F_x(t)]^{n-i-1} \\ &- \sum_{i=r}^{n-1} \binom{n}{i} (n-i) [F_x(t)]^{i-1} [1 - F_x(t)]^{n-i} \\ &= f_{x_{(r)}}(t) = \frac{n!}{(r-1)!(n-r)!} [F_x(t)]^{r-1} f_x(t) [1 - F_x(t)]^{(n-r)} \end{split}$$

Theorem 1.2. (the pdf of noncentral chi-squared distribution)

$$f(x) = \begin{cases} e^{-\delta^2/2} \sum_{i=0}^{\infty} \frac{1}{i!} \left(\frac{\delta^2}{2}\right)^i \frac{x^{i+n/2-1}}{2^{i+n/2}\Gamma(n/2+i)} e^{-x/2}, & x > 0\\ 0, & x \le 0 \end{cases}$$
(3)

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where
$$\delta = \sqrt{\sum_{i=1}^{n} a_i^2}$$
.

Proof. To prove (3), we need to use a orthogonal transformation such that $X_1^2+\cdots+X_n^2=Y_1^2+Z$, where $X_i\sim N(a_i,1), Y_1\sim N(\delta,1)$ and $Z\sim\chi^2_{n-1}$. The orthogonal matrix A can be