



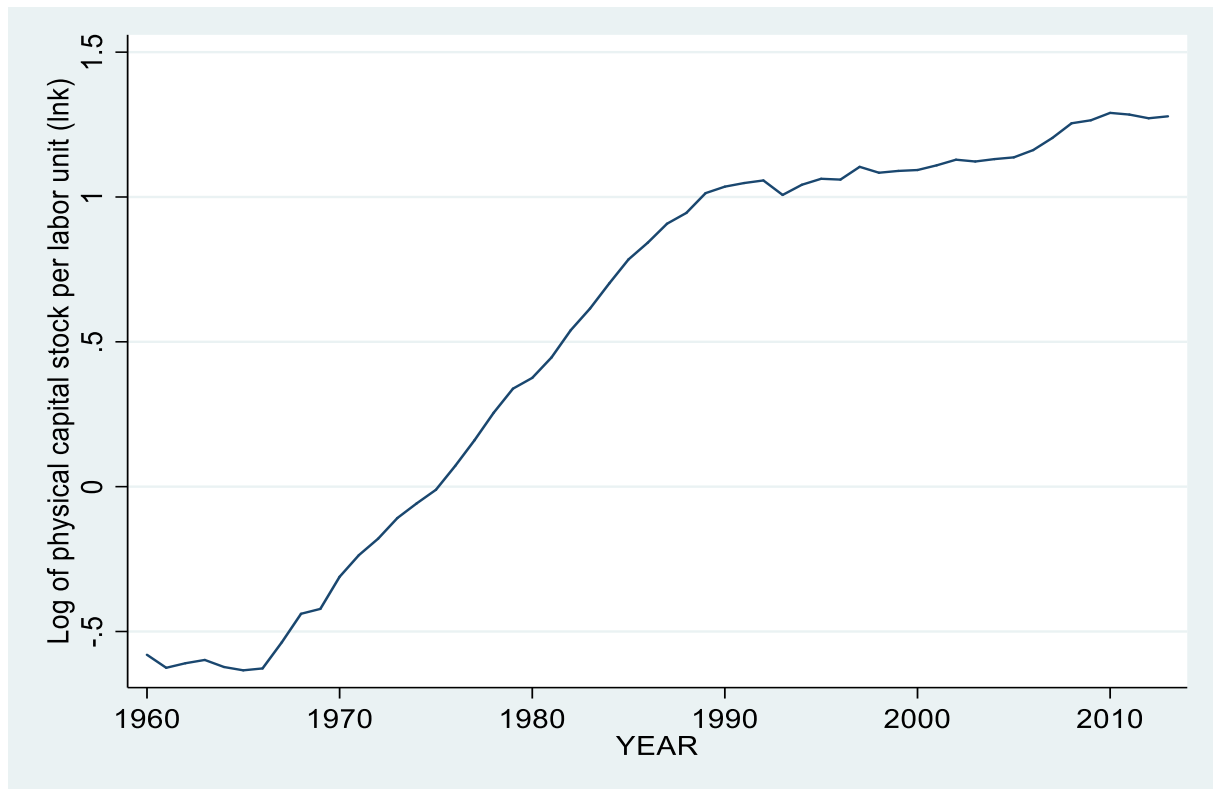
ECON 7350 Applied Econometrics for Macroeconomics and Finance

Research Project 1

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(A)

Figure 1. Time series plot of $\{\ln k_t\}$



Comment:

Apparently, the time series of log of physical capital stock per labor unit $\{\ln k_t\}$ is explosive. On the annual frequency basis, the data especially did not exhibit any signs of mean reversion from year 1970 to 1990. During the 1970~1990 period, it shows a strong deterministic trend. In sum, this time series is not stationary.

(B)

Figure 2. ACF of $\{\ln k_t\}$

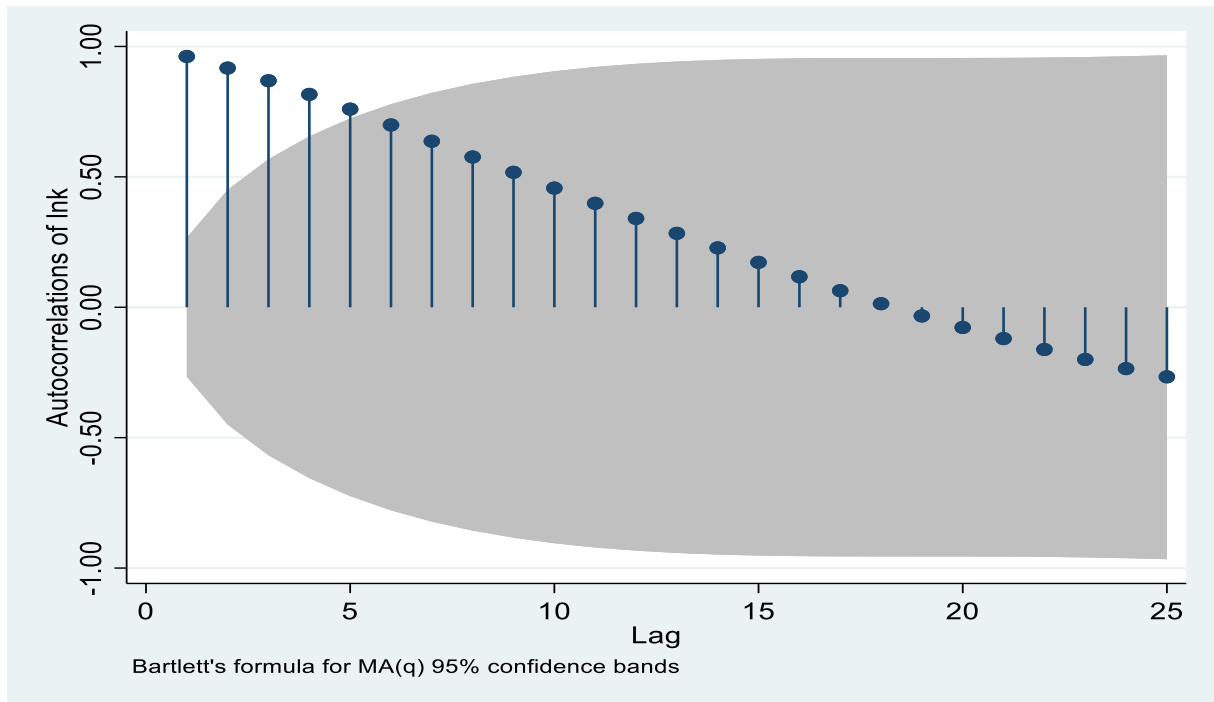
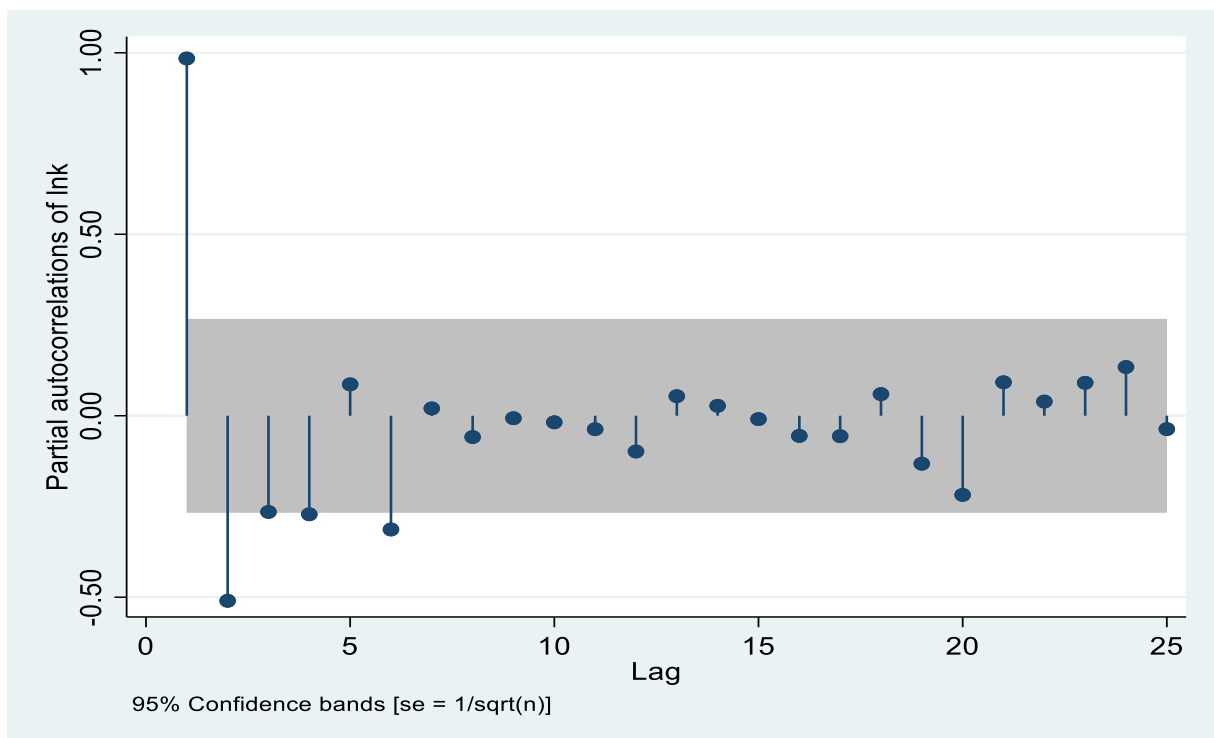


Figure 3. PACF of $\{\ln k_t\}$



Comment:

For the DGP of $\{\ln k_t\}$, its ACF is prevailed by AR(1) process which converges toward zero geometrically from lag 1 to lag 19. However, its ACF turns negative from lag 20. Moreover, its PACF has two non-zero peaks and decays in a oscillatory path from lag 2 as MA(1) process dominates. Altogether, the patterns of ACF and PACF imply that $\{\ln k_t\}$ may contain unit root.

Moreover, it resembles the pattern of ARMA (2,0) because it shows gradual decaying and has two peaks.

(C)

I tested the following AICs and BICs of ARIMA (1,1,0), (1,1,1), (1,1,2), (1,1,3), (1,1,4), (2,1,0), (2,1,1), (2,1,2), (2,1,3), and (2,1,4) models. Among them, ARIMA (2,1,2) is the best model because it has the lowest AIC and BIC.

AIC and BIC tests for ARIMA

Models	AIC	BIC
ARIMA (1,1,0)	-208.9989	-203.088
ARIMA (1,1,1)	-212.5874	-204.7062
ARIMA (1,1,2)	-210.7306	-200.8791
ARIMA (1,1,3)	-210.5381	-198.7164
ARIMA (1,1,4)	-206.8035	-193.0115
ARIMA (2,1,0)	-210.9248	-203.0436
ARIMA (2,1,1)	-210.678	-200.8265
ARIMA (2,1,2)	-216.3281*	-206.4766*
ARIMA (2,1,3)	-214.3194	-202.4976
ARIMA (2,1,4)	-210.4938	-194.7314

Where: * denotes the best model

Below is the estimated model for ARIMA (2,1,2):

$$\{\Delta \ln k_t\} = 0.0366 + 1.9768\Delta \ln k_{t-1} - 0.9948\Delta \ln k_{t-2} - 1.9976\Delta \varepsilon_{t-1} + 1.0000\Delta \varepsilon_{t-2}$$

(0.0039) (0.0091) (0.0115) (0.0206)

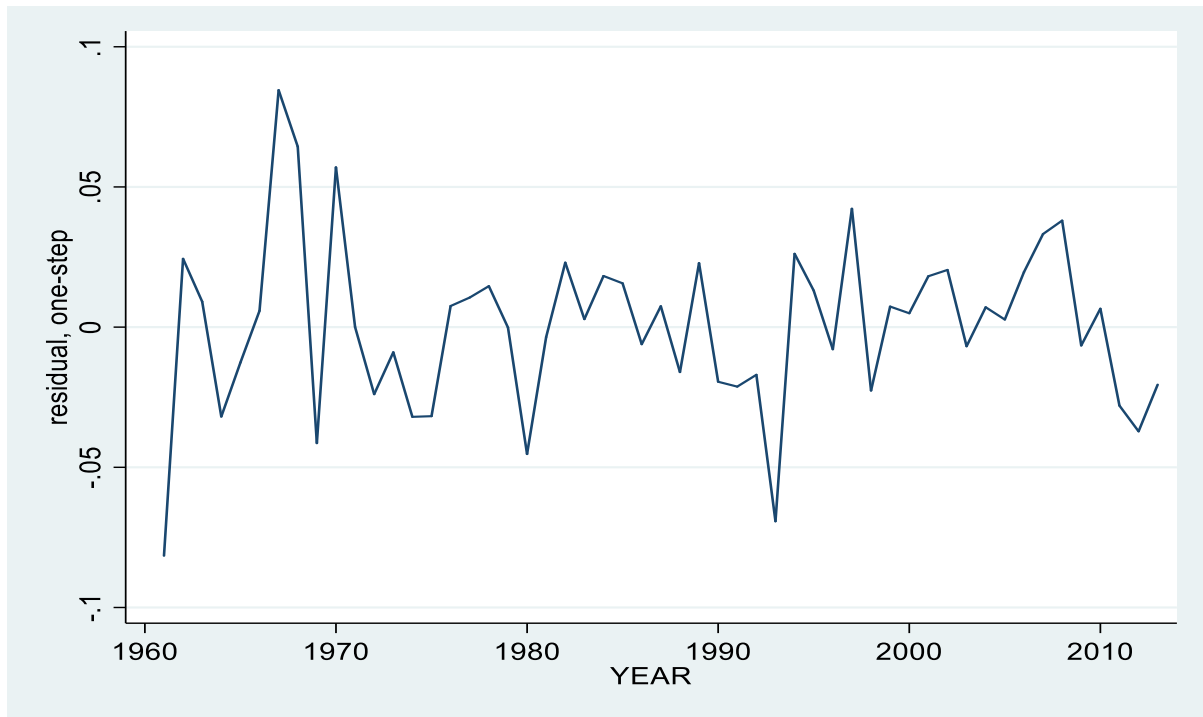
ARIMA regression

Sample: 1961 - 2013 Number of obs = 53
Wald chi2(3) = 438701.83
Log likelihood = -113.164 Prob > chi2 = 0.0000

D.lnk	OPG		z	P> z	[95% Conf. Interval]	
	Coef.	Std. Err.				
lnk						
_cons	.0365961	.0038975	9.39	0.000	.0289572	.0442349
ARMA						
ar						
L1.	1.976776	.0090572	218.26	0.000	1.959024	1.994528
L2.	-.9948172	.0114776	-86.67	0.000	-1.017313	-.9723214
ma						
L1.	-1.997584	.0206332	-96.81	0.000	-2.038025	-1.957144
L2.	1.00001
/sigma	.026253	.0031816	8.25	0.000	.0200173	.0324888

(D)

Figure 4. Residuals of ARIMA (2,1,2) for $\{\ln k_t\}$



The process $\{\hat{\varepsilon}_t\}$ of residuals had been mean reverted around zero. Then I run Ljung-Box test from lag 5 to 9 to check for autocorrelation.

The test of Ljung-Box:

Ho: The variable is independently distributed; so exhibits white noise process.

Ha: The variable is not independently distributed; so exhibits autocorrelation.

Ljung-Box test results:

Lag	Q-statistics	Critical Value	Decision @ 95 % significance
5	3.8878	3.8415	Reject the null
6	6.3003	5.9915	Reject the null
7	8.8939	7.8147	Reject the null
8	9.1988	9.4877	Do not reject the null
9	12.1670	11.0705	Reject the null

Explanation:

I started the test from lag 5 because Q-statistic requires $(k-2-2)$ degrees of freedom since ARIMA (2,1,2) has $p=2$ and $q=2$. I tested from lag 5 to 9; the result shows that except for lag 8, we can reject the Ho that $\{\hat{\varepsilon}_t\}$ is a white noise process at 95 percent significance level and conclude that $\{\hat{\varepsilon}_t\}$ exhibits autocorrelation. Overall, this ARIMA is inadequate due to presence of autocorrelation which is likely to cause inaccurate computed variance and inaccurate standard error of forecast.

(E)

The ADF test is conducted for $\{\ln k_t\}$, $\{\ln y_t\}$, $\{\ln w_t\}$, and $\{\ln O_t\}$ and each with three different equations.

The three equations are:

Equation (1): the pure random walk model.

Equation (2): adds a drift term to the previous model, Equation (1).

Equation (3): adds a drift and a linear time trend to the random walk model.

The hypothesis test in ADF is described below:

Ho: The variable contains a unit root

Ha: The variable is stationary

(E-1)

For $\{\ln k_t\}$, the test is conducted with lag 2. We failed to reject the null for equations (1) and (3) but can reject the null for equation (2) at 95 percent significance level. Although the results are ambiguous, the ACF and PACF of $\{\ln k_t\}$ show that there exists at least one unit root because $\rho_1 \geq 0.95$. To determine the most suitable specification, we move on to conduct ϕ_3 and ϕ_2 tests subsequently. As a result, we cannot reject the null of ϕ_3 and ϕ_2 tests at 95 percent significance level. Altogether, we can speculate that $\{\ln k_t\}$ contains unit roots and a drift. To be cautious, we should try further with the first differences $\{\Delta \ln k_t\}$. However, we still get the same decision as the previous result. So I try again with second differences $\{\Delta^2 \ln k_t\}$; consequently, we can see that the unit roots of $\{\ln k_t\}$ for equations (1) and (3) will not be eliminated until second differences.

	F-test	Critical Value	Decision@95%
phi 3	2.24	6.73	Do not reject
phi 2	2.63	5.13	Do not reject

Ink	Test Statistics	5% Critical Value	P-Value
Equation (1)	-0.451	-1.950	N.A.
Equation (2)	-2.132*	-1.678	0.0191
Equation (3)	-0.806	-3.499	0.9652

$\Delta \ln k$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-1.130	-1.950	N.A.
Equation (2)	-1.715*	-1.678	0.0466
Equation (3)	-2.503	-3.499	0.3262

$\Delta^2 \ln k$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-4.640*	-1.950	N.A.
Equation (2)	-4.589*	-1.678	0.0000
Equation (3)	-4.749*	-3.499	0.0006

Note: * denotes statistically significant at 5%

(E-2)

For $\{\ln y_t\}$, the test is conducted with lag 2. We failed to reject the null for all three equations at 95 percent significance level. After trying the steps of roadmap, we may conclude that the series is trend stationary and has a quadratic trend. As a result, the next step may require second differences $\{\Delta^2 \ln y_t\}$. Initially, I estimated it with $\{\Delta \ln y_t\}$ and found that we still failed to reject the null for equation (1) at 95 percent significance level. But after second differences, we can certainly eliminate the unit roots for all three equations at 95 percent significance level.

$\ln y$	Test Statistics	5% Critical Value	P-Value
Equation (1)	4.315	-1.950	N.A.
Equation (2)	-1.110	-1.678	0.1363
Equation (3)	-1.704	-3.499	0.7489

$\Delta \ln y$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-1.880	-1.950	N.A.
Equation (2)	-4.386*	-1.678	0.0000
Equation (3)	-4.490*	-3.499	0.0016

$\Delta^2 \ln y$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-7.804*	-1.950	N.A.
Equation (2)	-7.765*	-1.678	0.0000
Equation (3)	-7.677*	-3.499	0.0000

Note: * denotes statistically significant at 5%

(E-3)

For $\{\ln w_t\}$, the test is conducted with lag 2. We failed to reject the null for all three equations at 95 percent significance level. Similar to $\{\ln y_t\}$, we may conclude that the series is trend stationary and has a quadratic trend. Thus, we continue to try the first differences but still failed to reject the null for equation (3). Then we should try second differences $\{\Delta^2 \ln w_t\}$; eventually, we can reject the null for all three equations at 95 percent significance level.

lnw	Test Statistics	5% Critical Value	P-Value
Equation (1)	-1.033	-1.950	N.A.
Equation (2)	-1.213	-1.678	0.1156
Equation (3)	-1.912	-3.499	0.6483

Δlnw	Test Statistics	5% Critical Value	P-Value
Equation (1)	-2.190*	-1.950	N.A.
Equation (2)	-2.383*	-1.678	0.0107
Equation (3)	-2.294	-3.499	0.4374

Δ²lnw	Test Statistics	5% Critical Value	P-Value
Equation (1)	-5.079*	-1.950	N.A.
Equation (2)	-5.023*	-1.678	0.0000
Equation (3)	-5.061*	-3.499	0.0002

Note: * denotes statistically significant at 5%

(E-4)

For $\{ln O_t\}$, the test is conducted with lag 2. We failed to reject the null for equations (1) and (3) but can reject the null for equation (2) at 95 percent significance level. Although the results are ambiguous, the ACF and PACF of $\{ln O_t\}$ show that there exists at least one unit root because $\rho_1 \geq 0.95$. Moreover, to determine the most suitable specification, we move on to conduct ϕ_3 and ϕ_2 subsequently. As a result, we cannot reject the null of ϕ_3 and ϕ_2 tests at 95 percent significance level. To be cautious, we should try further with first differences $\{\Delta ln O_t\}$; eventually, we can reject the null for all three equations at 95 percent significance level.

	F-test	Critical Value	Decision@95%
phi 3	2.53	6.73	Do not reject
phi 2	1.70	5.13	Do not reject

lnO	Test Statistics	5% Critical Value	P-Value
Equation (1)	-0.830	-1.950	N.A.
Equation (2)	-1.926*	-1.678	0.0301
Equation (3)	-2.231	-3.499	0.4723

$\Delta \ln o$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-4.179*	-1.950	N.A.
Equation (2)	-4.136*	-1.678	0.0001
Equation (3)	-4.087*	-3.499	0.0066

Note: * denotes statistically significant at 5%

(F)

Based on BIC, we choose the **ARDL (3,0,1,0)** regression model. **The estimated model** is shown below:

$$\begin{aligned}
 \{\widehat{\ln y_t}\} = & 2.7321 + 0.0048t - 0.0313D_{01} - 0.0289D_0 + 0.0023D_{77} + 1.2466\ln y_{t-1} \\
 & (1.0844) \quad (0.0046) \quad (0.0223) \quad (0.0166) \quad (0.0267) \quad (0.1475) \\
 & - 0.7763\ln y_{t-2} + 0.3286\ln y_{t-3} + 0.0493 \ln k_t - 0.1471\ln w_t \\
 & (0.2126) \quad (0.1421) \quad (0.0387) \quad (0.0730) \\
 & + 0.1880 \ln w_{t-1} + 0.0101\ln O_t \\
 & (0.0751) \quad (0.0220)
 \end{aligned}$$

$$R^2 = 0.9993$$

Sample: 1964 - 2013	Number of obs	=	50
	F(11, 38)	=	4742.94
	Prob > F	=	0.0000
	R-squared	=	0.9993
	Adj R-squared	=	0.9991
Log likelihood = 129.45628	Root MSE	=	0.0208

lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lny						
L1.	1.246609	.1474874	8.45	0.000	.9480363	1.545182
L2.	-.7762955	.2126313	-3.65	0.001	-1.206745	-.3458458
L3.	.3285665	.1420503	2.31	0.026	.0410006	.6161323
lnk	.0492925	.0386811	1.27	0.210	-.0290134	.1275983
lnw						
--.	-.1470679	.0729489	-2.02	0.051	-.2947453	.0006095
L1.	.1879981	.075057	2.50	0.017	.0360532	.339943
lno	.0101193	.0220055	0.46	0.648	-.0344284	.0546671
d01	-.0312674	.0222532	-1.41	0.168	-.0763166	.0137818
d0	-.0289431	.0166422	-1.74	0.090	-.0626335	.0047473
d77	.0022812	.0266803	0.09	0.932	-.0517303	.0562927
t	.0047761	.004608	1.04	0.307	-.0045523	.0141046
_cons	2.732141	1.084377	2.52	0.016	.5369344	4.927347

(G)

The estimated ECM can be represented as:

$$\begin{aligned}\{\widehat{\Delta \ln y_t}\} = & 2.4168 + 0.0037t - 0.1768 [\ln y_{t-1} - 0.3138 \ln k_{t-1} - 0.2074 \ln w_{t-1} \\ & (1.0580) \quad (0.0046) \quad (0.0884) \quad (0.3103) \quad (0.0975) \\ & - 0.0419 \ln O_{t-1}] + 0.4107 \Delta \ln y_{t-1} - 0.2959 \Delta \ln y_{t-2} + 0.0555 \Delta \ln k_t \\ & (0.1269) \quad (0.1236) \quad (0.1402) \quad (0.0386) \\ & - 0.1492 \Delta \ln w_t + 0.0074 \Delta \ln O_t - 0.0306 \Delta D_{01t-1} \\ & (0.0733) \quad (0.0220) \quad (0.0224) \\ & - 0.0286 \Delta D_{07t-1} + 0.0001 \Delta D_{77t-1} \\ & (0.0167) \quad (0.0268)\end{aligned}$$

$$R^2 = 0.5889$$

Comment:

After estimating the ECM, I find that the speed of adjustment is slow based on the magnitude of coefficient and if we incorporate the result from part (E), we learn that $\ln y_t$ is possibly a trend stationary time series. Thus, $\ln y_t$ may take quite some time to return to the long run equilibrium or steady state.

ARDL(3,0,1,0) regression

Sample: 1963 - 2013

Number of obs = 51
R-squared = 0.5889
Adj R-squared = 0.4730
Root MSE = 0.0210

Log likelihood = 131.60073

	D.lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ADJ						
	lny L1.	-.1767611	.0884139	-2.00	0.053	-.3555952 .0020729
LR						
	lnk L1.	.3138241	.3103313	1.01	0.318	-.3138803 .9415285
	lnw L1.	.2073692	.0974537	2.13	0.040	.0102504 .4044879
	lno L1.	.0418554	.1268715	0.33	0.743	-.2147663 .2984772
SR						
	lny LD.	.4106602	.1235888	3.32	0.002	.1606783 .6606422
	L2D.	-.2959411	.1402152	-2.11	0.041	-.579553 -.0123292
	lnk D1.	.0554719	.0385538	1.44	0.158	-.0225106 .1334544
	lnw D1.	-.1491997	.0733392	-2.03	0.049	-.2975423 -.0008571
	lno D1.	.0073984	.0220121	0.34	0.739	-.0371253 .0519221
	d01	-.0305867	.0223716	-1.37	0.179	-.0758376 .0146641
	d0	-.0286444	.0167344	-1.71	0.095	-.0624929 .005204
	d77	.0001318	.0267706	0.00	0.996	-.0540168 .0542804
	t	.0037351	.0045513	0.82	0.417	-.0054707 .012941
	_cons	2.416759	1.05803	2.28	0.028	.2766915 4.556827

(H)

I conducted the Engle-Granger test with lag 2 to see whether ($\{ln k_t\}, \{ln y_t\}, \{ln w_t\}$, and $\{ln O_t\}$) are cointegrated. Hypothesis test and test results are represented below.

Engle-Granger test

Ho: No cointegration among the variables

Ha: Cointegration exist among the variables

Explanation:

The pre-test shows that none of $\{ln k_t\}$, $\{ln y_t\}$, and $\{ln w_t\}$ can be certainly stationary even after taking the first difference. We then conduct cointegration test with Engle-Granger to see whether the system of multivariate equation can be stationary. The Engle-Granger shows that we failed to reject the null hypothesis and; thus there is no cointegration among them at 95 percent significance level. For the purpose of robustness check, I also estimated the Engle-Granger test with eight different orders among ($\{ln k_t\}, \{ln y_t\}, \{ln w_t\}$, and $\{ln O_t\}$). As a result, they return the same decision which concludes that there is no cointegration among them at 95 percent significance level; in other words, there is no existence of long run stable relationship among them. Since there is no evidence of cointegration, we can then safely move on to do multivariate time series model like the vector autoregression model.

Pre-Test

$\Delta ln k$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-1.130	-1.950	N.A.
Equation (2)	-1.715*	-1.678	0.0466
Equation (3)	-2.503	-3.499	0.3262
$\Delta ln y$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-1.880	-1.950	N.A.
Equation (2)	-4.386*	-1.678	0.0000
Equation (3)	-4.490*	-3.499	0.0016
$\Delta ln w$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-2.190*	-1.950	N.A.
Equation (2)	-2.383*	-1.678	0.0107
Equation (3)	-2.294	-3.499	0.4374
$\Delta ln O$	Test Statistics	5% Critical Value	P-Value
Equation (1)	-4.179*	-1.950	N.A.
Equation (2)	-4.136*	-1.678	0.0001
Equation (3)	-4.087*	-3.499	0.0066

Where: * denotes statistically significant at 5%

Equation (1): the pure random walk model.

Equation (2): adds a drift term to the previous model, Equation (1).

Equation (3): adds a drift and a linear time trend to the random walk model.

Engle-Granger Cointegration test table

Orders	Test Statistics	5% Critical Value	Decision
lnk lny lnw lno	-2.3260	-4.316	Do not reject the null
lny lnk lnw lno	-1.7240	-4.316	Do not reject the null
lnw lnk lny lno	-1.8030	-4.316	Do not reject the null
lno lnk lny lnw	-3.3500	-4.316	Do not reject the null
lny lno lnw lnk	-1.7240	-4.316	Do not reject the null
lnw lno lny lnk	-1.8030	-4.316	Do not reject the null
lno lnw lny lnk	-3.3500	-4.316	Do not reject the null
lno lnw lnk lny	-3.3500	-4.316	Do not reject the null

(I)**The first differences model of the form:**

$$\begin{aligned}
\{\Delta \widehat{\ln y_t}\} = & 0.0660 - 0.0220D_{01} - 0.0323D_0 - 0.0148D_{77} \\
& (0.0102) \quad (0.0300) \quad (0.0163) \quad (0.0108) \\
& - 0.1384 \Delta \ln k_t - 0.0251 \Delta \ln w_t + 0.0243 \Delta \ln O_t \\
& (0.1169) \quad (0.0622) \quad (0.0339)
\end{aligned}$$

$$R^2 = 0.1394$$

Source	SS	df	MS	Number of obs	=	52
Model	.006251543	6	.001041924	F(6, 45)	=	1.22
Residual	.038585289	45	.000857451	Prob > F	=	0.3163
Total	.044836831	51	.000879154	R-squared	=	0.1394
				Adj R-squared	=	0.0247
				Root MSE	=	.02928

dlny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
d01	-.0220242	.0300027	-0.73	0.467	-.0824526 .0384042
d0	-.0322923	.0162541	-1.99	0.053	-.0650297 .000445
d77	-.0148066	.0108343	-1.37	0.179	-.0366279 .0070147
dlnk	-.1383957	.116944	-1.18	0.243	-.3739329 .0971415
dlnw	-.0250926	.0622368	-0.40	0.689	-.1504438 .1002587
dlno	.0242653	.0339092	0.72	0.478	-.0440314 .0925621
_cons	.065984	.0102216	6.46	0.000	.0453967 .0865714