

Leu-41, DC, 24-25, Sec A

Under each of the hypotheses, you can receive the same $y(t)$.

| | | |
|-------------------------------------|--|-------------------|
| $b_c, b_s = (1, 1) \textcircled{1}$ | $s \textcircled{1}(t) + n_1(t) = y(t)$ | } $p(y(t) (i))$ |
| $(1, -1) \textcircled{2}$ | $s \textcircled{2}(t) + n_2(t) = y(t)$ | |
| $(-1, -1) \textcircled{3}$ | $s \textcircled{3}(t) + n_3(t) = y(t)$ | |
| $(-1, 1) \textcircled{4}$ | $s \textcircled{4}(t) + n_4(t) = y(t)$ | |

The pairs (b_c, b_s) are obtained from the binary waveform (discrete)

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$M=2$

$p_i \rightarrow$ prior prob. of hypothe
ses.

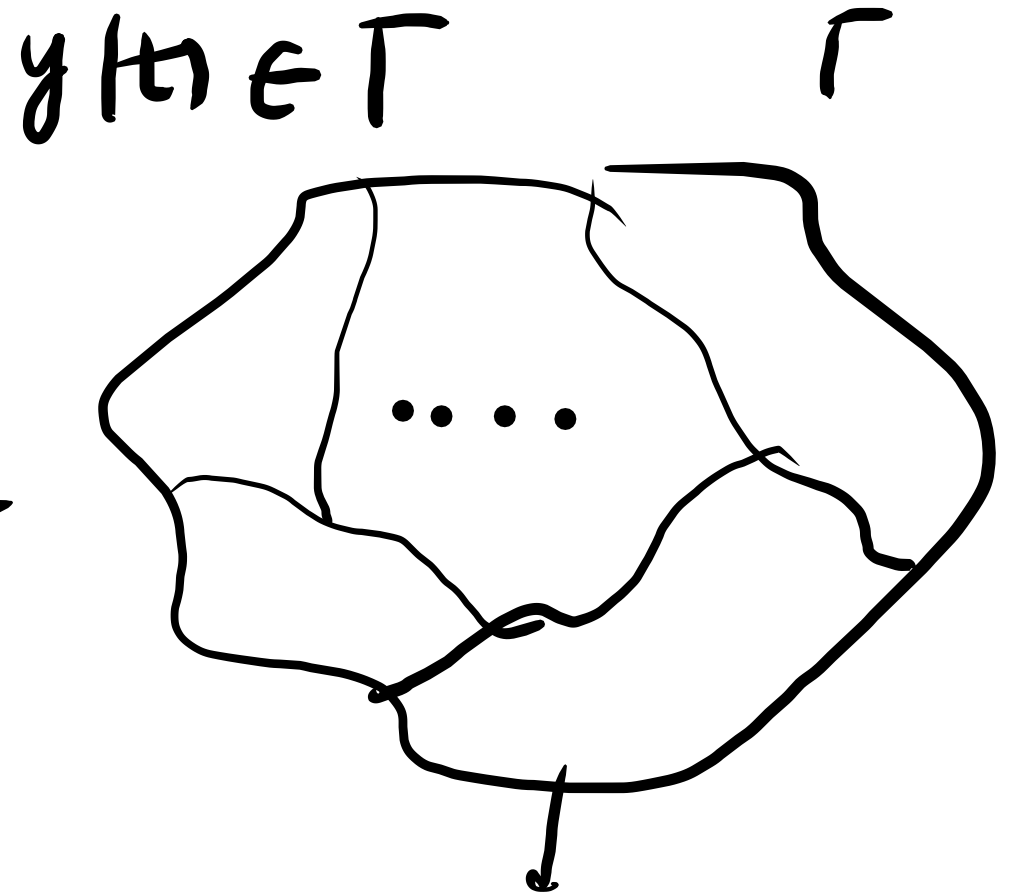
There are 4 prob. \rightarrow (1)

| | | |
|----------------|----|-------|
| | 00 | P_1 |
| (2) \nwarrow | 01 | P_2 |
| (3) \nwarrow | 10 | P_3 |
| (4) \nwarrow | 11 | P_4 |

$$\delta(y(t)) \rightarrow \{1, 2, \dots, M\} \text{ or } \{0, 1, 2, \dots, M-1\}$$

↑
Input

what is $\Gamma_i \rightarrow$ if $y(t) \in \Gamma_i$, you
say H_i is true



M -disjoint
regions such
that their union
is Γ

$$\Gamma_1 \cup \Gamma_2 \dots \cup \Gamma_M = \Gamma$$

Continuing ex 6.1.1,

$$H_0 = 4/5$$

$$H_1 = 1/5$$

H_1

ans:-

$$y \underset{H_0}{\overset{H_1}{\geq}} \frac{4}{3} \ln 16 = 3.6968$$

average error prob.



$$P(H_0 \rightarrow \text{true})P(e|H_0) + P(H_1 \rightarrow \text{true})P(e|H_1, \text{true})$$

$$P(\text{error}) = P(H_0 \rightarrow \text{true}, \text{error} \cup H_1 \rightarrow \text{true}, \text{error})$$

$$\text{or } P\left(\underbrace{H_0 \rightarrow \text{true} \cap \text{error}}_{E_1} \cup \underbrace{H_1 \rightarrow \text{true} \cap \text{error}}_{E_2}\right)$$

$$P(E_1 \cup E_2) = P(E_1) + P(E_2) - \underbrace{P(E_1 \cap E_2)}_{\emptyset}$$

$$= P(E_1) + P(E_2) = P(H_0 \rightarrow \text{true})P(e|H_0) + P(H_1 \rightarrow \text{true})P(e|H_1)$$

In case of binary hypotheses testing, we can define the likelihood ratio.

$$\begin{aligned} \text{(ML)} \quad P(y|1) \underset{H_0}{\overset{H_1}{\geq}} P(y|0) &\Rightarrow \frac{P(y|1)}{P(y|0)} \underset{H_0}{\overset{H_1}{\geq}} 1 \\ &\triangleq L(y) \end{aligned}$$

LRT (likelihood ratio test) - in general form

$$L(y) \underset{H_0}{\overset{H_1}{\geq}} \gamma$$

$$\begin{aligned} \text{ML: } \gamma &= 1 \\ \text{MAP: } \gamma &= \pi_0 / \pi_1 \end{aligned}$$

In case of MAP,

$$\pi_1 P(y|1) \underset{H_0}{\overset{H_1}{\geq}} \pi_0 P(y|0) \Rightarrow L(y) \underset{H_0}{\overset{H_1}{\geq}} \frac{\pi_0}{\pi_1} = \gamma$$

ex (solved) $H_0: Y \sim \mathcal{N}(0, \sigma^2)$, $H_1: Y \sim \mathcal{N}(m, \sigma^2)$

let π_0 & π_1 be the prior prb.

find ML & MAP rules using LRT

$$\gamma = 1 \quad \gamma = \frac{\pi_0}{\pi_1}$$

$$P(y|0) = \frac{e^{-y^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}} \quad -\infty < y < \infty$$

$$P(y|1) = \frac{e^{-(y-m)^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}} \quad "$$

n.w. derive
for $m < 0$.

MAP:

$$\gamma \underset{H_0}{\overset{H_1}{\geq}} \frac{m}{2} + \frac{\sigma^2}{m} \log\left(\frac{\pi_0}{\pi_1}\right) \quad m > 0$$

$$L(y) \underset{H_0}{\overset{H_1}{\geq}} \gamma$$

$$\Rightarrow ?$$

$$\log_e L(y) \underset{H_0}{\overset{H_1}{\geq}} \log_e \gamma$$

$$\gamma \underset{H_0}{\overset{H_1}{\geq}} \frac{m}{2} \quad \text{ML, } (m > 0)$$

Practice problem :- ex - 6.1.2

So, we have seen hypotheses testing when Y is a Random Variable. What if it is a random process?